Analysis of Residuals¹ STA442/2101 Fall 2016

 $^{^{1}\}mathrm{See}$ last slide for copyright information.

Residual means left over: $e_i = Y_i - \hat{Y}_i$

- Vertical distance of Y_i from the regression hyper-plane
- An error of "prediction."
- Big residuals merit further investigation.
- Big compared to what?
- They are normally distributed
- Consider standardizing
- Maybe detect outliers
- Plots can also be informative.

Residuals are like estimated error terms

$$e_i = Y_i - \widehat{Y}_i \quad \Leftrightarrow \quad Y_i = \widehat{Y}_i + e_i$$

$$Y_{i} = \widehat{Y}_{i} + e_{i}$$

$$= \widehat{\beta}_{0} + \widehat{\beta}_{1}x_{i,1} + \dots + \widehat{\beta}_{p-1}x_{i,p-1} + e_{i}$$

$$= \beta_{0} + \beta_{1}x_{i,1} + \dots + \beta_{p-1}x_{i,p-1} + \epsilon_{i}$$

Normal distribution of e_i implies normal distribution of e_i , but the e_i are not independent, and they do not have equal variance.

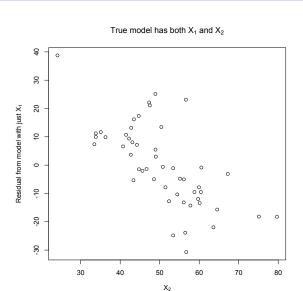
Data = Fit + Residual

$$Y_i = \widehat{Y}_i + e_i$$

Plotting residuals can be helpful

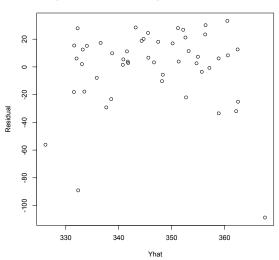
- \bullet Against predicted Y.
- Against explanatory variables not in the equation.
- Against explanatory variables in the equation.
- Look for serious departures from normality.

Plot Residuals Against Explanatory Variables Not in the Equation



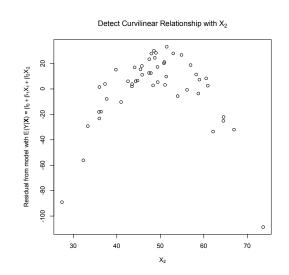
Plot Residuals Against \widehat{Y}

Suspect Curvilinear Relationship with one or more X variables



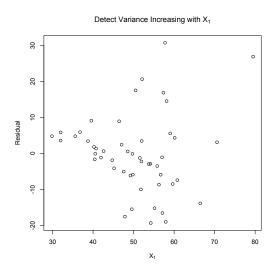
Plot Residuals Against Explanatory Variables in the Equation

Plot versus X_1 showed nothing



Plot Residuals Against Explanatory Variables in the Equation

Can show non-constant variance



Outlier detection

- Big residuals may be outliers. What's "big?"
- Consider standardizing.
- Could divide by square root of sample variance of e_1, \ldots, e_n .
- Semi-Studentized: Estimate $Var(e_i)$ and divide by square root of that. $e_i^* = \frac{e_i}{\sqrt{MSE(1-h_{i,i})}}$
- In R, this is produced with rstandard.

Studentized deleted residuals The idea

- \bullet An outlier will make MSE big.
- In that case, the standardized (Semi-Studentized) residual will be too small less noticeable.
- So calculate \widehat{Y} for each observation based on all the other observations, but not that one.
- Predict each observed Y based on all the others, and assess error of prediction (divided by standard error).

Apply prediction interval technology

$$T = \frac{Y_{n+1} - \mathbf{x}_{n+1}^{\top} \widehat{\boldsymbol{\beta}}}{\sqrt{MSE(1 + \mathbf{x}_{n+1}^{\top} (\mathbf{X}^{\top} \mathbf{X})^{-1} \mathbf{x}_{n+1})}} \sim t(n - p)$$

- Note that Y_i is now being called Y_{n+1} .
- If the "prediction" is too far off there is trouble.
- Use T as a test statistic.
- Need to change the notation.

Studentized deleted residual

$$T_i = \frac{Y_i - \mathbf{x}_i^{\top} \widehat{\boldsymbol{\beta}}_{(i)}}{\sqrt{MSE_{(i)}(1 + \mathbf{x}_i^{\top}(\mathbf{X}_{(i)}^{\top}\mathbf{X}_{(i)})^{-1}\mathbf{x}_i)}} \sim t(n - 1 - p)$$

- In R, this is produced with rstudent.
- There is a more efficient formula.
- Use T_i as a test statistic of $H_0: E(Y_i) = \mathbf{x}_i^{\top} \boldsymbol{\beta}$.
- If H_0 is rejected, investigate.
- We are doing n tests.
- Type I errors are very time consuming and disturbing.
- If independent, probability of no false positives would be $(1-\alpha)^n \to 0$.
- But they are not independent.
- How about a Bonferroni correction?

Copyright Information

This slide show was prepared by Jerry Brunner, Department of Statistics, University of Toronto. It is licensed under a Creative Commons Attribution - ShareAlike 3.0 Unported License. Use any part of it as you like and share the result freely. The LATEX source code is available from the course website: http://www.utstat.toronto.edu/~brunner/oldclass/appliedf16