Factorial ANOVA

More than one categorical explanatory variable

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Factorial ANOVA

- Categorical explanatory variables are called factors
- More than one at a time
- Designed for true experiments, but also useful with observational data
- If there are observations at all combinations of explanatory variable values, it's called a *complete* factorial design (as opposed to a fractional factorial).

The potato study

- Cases are potatoes
- Inoculate with bacteria, store for a fixed time period.
- Response variable is diameter of rotten spot in millimeters
- Two explanatory variables, randomly assigned
 - Bacteria Type (1, 2, 3)
 - Temperature (1=Cool, 2=Warm)

Two-factor design

	Bacteria Type			
Temp	1	2	3	
1=Cool				
2=Warm				

Six treatment conditions

Factorial experiments

- Allow more than one factor to be investigated in the same study: Efficiency!
- Allow the scientist to see whether the effect of an explanatory variable *depends* on the value of another explanatory variable: Interactions
- Thank you again, Mr. Fisher.

Model: Data are normal with equal variance and conditional (cell) means $\mu_{i,j}$

	Bacteria Type						
Temp	1	2	3				
1=Cool	$\mu_{1,1}$	$\mu_{1,2}$	$\mu_{1,3}$	$\frac{\mu_{1,1} + \mu_{1,2} + \mu_{1,3}}{3}$			
2=Warm	$\mu_{2,1}$	$\mu_{2,2}$	$\mu_{2,3}$	$\frac{\mu_{2,1} + \mu_{2,2} + \mu_{2,3}}{3}$			
	$\frac{\mu_{1,1} + \mu_{2,1}}{2}$	$\frac{\mu_{1,2} + \mu_{2,2}}{2}$	$\frac{\mu_{1,3} + \mu_{2,3}}{2}$	μ			

Tests

- Main effects: Differences among marginal means
- Interactions: Differences between differences (What is the effect of Factor A? It depends on level of Factor B.)

To understand the interaction, plot the means



Either Way



Non-parallel profiles = Interaction



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Main effects for both variables, no interaction



Main effect for Bacteria only



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Main Effect for Temperature Only



Both Main Effects, and the Interaction



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Should you interpret the main effects?



Testing Contrasts

	Bacteria Type						
Temp	1	2	3				
1=Cool	$\mu_{1,1}$	$\mu_{1,2}$	$\mu_{1,3}$	$\frac{\mu_{1,1} + \mu_{1,2} + \mu_{1,3}}{3}$			
2=Warm	$\mu_{2,1}$	$\mu_{2,2}$	$\mu_{2,3}$	$\frac{\mu_{2,1} + \mu_{2,2} + \mu_{2,3}}{3}$			
	$\frac{\mu_{1,1} + \mu_{2,1}}{2}$	$\frac{\mu_{1,2} + \mu_{2,2}}{2}$	$\frac{\mu_{1,3} + \mu_{2,3}}{2}$	μ			

- Differences between marginal means are definitely contrasts
- Interactions are also sets of contrasts

Interactions are sets of Contrasts

	Bacteria Type						
Temp	1	2	3				
1=Cool	$\mu_{1,1}$	$\mu_{1,2}$	$\mu_{1,3}$	$\frac{\mu_{1,1} + \mu_{1,2} + \mu_{1,3}}{3}$			
2=Warm	$\mu_{2,1}$	$\mu_{2,2}$	$\mu_{2,3}$	$\frac{\mu_{2,1} + \mu_{2,2} + \mu_{2,3}}{3}$			
	$\frac{\mu_{1,1} + \mu_{2,1}}{2}$	$\frac{\mu_{1,2} + \mu_{2,2}}{2}$	$\frac{\mu_{1,3} + \mu_{2,3}}{2}$	μ			

• $H_0: \mu_{1,1} - \mu_{2,1} = \mu_{1,2} - \mu_{2,2} = \mu_{1,3} - \mu_{2,3}$

•
$$H_0: \mu_{1,2} - \mu_{1,1} = \mu_{2,2} - \mu_{2,1}$$
 and
 $\mu_{1,3} - \mu_{1,2} = \mu_{2,3} - \mu_{2,2}$

Interactions are sets of Contrasts



- $H_0: \mu_{1,1} \mu_{2,1} = \mu_{1,2} \mu_{2,2} = \mu_{1,3} \mu_{2,3}$
- $H_0: \mu_{1,2} \mu_{1,1} = \mu_{2,2} \mu_{2,1}$ and $\mu_{1,3} - \mu_{1,2} = \mu_{2,3} - \mu_{2,2}$ ¹⁸

Equivalent statements

- The effect of A depends upon B
- The effect of B depends on A

$$H_0: \mu_{1,1} - \mu_{2,1} = \mu_{1,2} - \mu_{2,2} = \mu_{1,3} - \mu_{2,3}$$

$$H_0: \mu_{1,2} - \mu_{1,1} = \mu_{2,2} - \mu_{2,1}$$
 and
 $\mu_{1,3} - \mu_{1,2} = \mu_{2,3} - \mu_{2,2}$

Three factors: A, B and C

- There are three (sets of) main effects: One each for A, B, C
- There are three two-factor interactions
 - A by B (Averaging over C)
 - A by C (Averaging over B)
 - B by C (Averaging over A)
- There is one three-factor interaction: AxBxC

Meaning of the 3-factor interaction

- The form of the A x B interaction depends on the value of C
- The form of the A x C interaction depends on the value of B
- The form of the B x C interaction depends on the value of A
- These statements are equivalent. Use the one that is easiest to understand.

To graph a three-factor interaction

- Make a separate mean plot (showing a 2-factor interaction) for each value of the third variable.
- In the potato study, a graph for each oxygen level.

Four-factor design

- Four sets of main effects
- Six two-factor interactions
- Four three-factor interactions
- One four-factor interaction: The nature of the three-factor interaction depends on the value of the 4th factor
- There is an F test for each one
- And so on ...

As the number of factors increases

- The higher-way interactions get harder and harder to understand
- All the tests are still tests of sets of contrasts (differences between differences of differences ...)
- But it gets harder and harder to write down the contrasts
- Effect coding becomes easier

Effect coding

Bact	B ₁	B ₂		
1	1	0	Temperature	Т
	0	1	1=Cool	1
	0		2=Warm	-1
3	-1	-1		

 $E(Y|\mathbf{X} = \mathbf{x}) = \beta_0 + \beta_1 B_1 + \beta_2 B_2 + \beta_3 T + \beta_4 B_1 T + \beta_5 B_2 T$

Interaction effects correspond to products of dummy variables

 $E(Y|\mathbf{X} = \mathbf{x}) = \beta_0 + \beta_1 B_1 + \beta_2 B_2 + \beta_3 T + \beta_4 B_1 T + \beta_5 B_2 T$

- The A x B interaction: Multiply each dummy variable for A by each dummy variable for B
- Use these products as additional explanatory variables in the multiple regression
- The A x B x C interaction: Multiply each dummy variable for C by each product term from the A x B interaction
- Test the sets of product terms simultaneously 26

Make a table

 $E(Y|\mathbf{X} = \mathbf{x}) = \beta_0 + \beta_1 B_1 + \beta_2 B_2 + \beta_3 T + \beta_4 B_1 T + \beta_5 B_2 T$

Bact	Temp	B ₁	B ₂	Т	B ₁ T	B ₂ T	$E(Y \mathbf{X} = \mathbf{x})$
1	1	1	0	1	1	0	$\beta_0 + \beta_1 + \beta_3 + \beta_4$
1	2	1	0	-1	-1	0	$\beta_0 + \beta_1 - \beta_3 - \beta_4$
2	1	0	1	1	0	1	$\beta_0 + \beta_2 + \beta_3 + \beta_5$
2	2	0	1	-1	0	-1	$\beta_0 + \beta_2 - \beta_3 - \beta_5$
3	1	-1	-1	1	-1	-1	$\beta_0 - \beta_1 - \beta_2 + \beta_3 - \beta_4 - \beta_5$
3	2	-1	-1	-1	1	1	$\beta_0 - \beta_1 - \beta_2 - \beta_3 + \beta_4 + \beta_5$

Cell and Marginal Means

	Bacteria Type								
Tmp	1	2	3						
1=C	$\beta_0 + \beta_1 + \beta_3 + \beta_4$	$\beta_0 + \beta_2 + \beta_3 + \beta_5$	$\begin{array}{c} \beta_0-\beta_1-\beta_2\\ +\beta_3-\beta_4-\beta_5 \end{array}$	$\begin{array}{c} \beta_0 \\ +\beta_3 \end{array}$					
2=W	$\beta_0 + \beta_1 - \beta_3 - \beta_4$	$\beta_0 + \beta_2 - \beta_3 - \beta_5$	$\beta_0 - \beta_1 - \beta_2 \\ -\beta_3 + \beta_4 + \beta_5$	$egin{array}{c} eta_0 \ -eta_3 \end{array}$					
	$\beta_0 + \beta_1$	$\beta_0 + \beta_2$	$\beta_0 - \beta_1 - \beta_2$	eta_0					

We see

- Intercept is the grand mean
- Regression coefficients for the dummy variables are deviations of the marginal means from the grand mean
- What about the interactions?

 $E(Y|\mathbf{X} = \mathbf{x}) = \beta_0 + \beta_1 B_1 + \beta_2 B_2 + \beta_3 T + \beta_4 B_1 T + \beta_5 B_2 T$

A bit of algebra shows

 $\mu_{1,1} - \mu_{2,1} = \mu_{1,2} - \mu_{2,2}$ is equivalent to $\beta_4 = \beta_5$

 $\mu_{1,2} - \mu_{2,2} = \mu_{1,3} - \mu_{2,3}$ is equivalent to $2\beta_4 = -\beta_5$

So
$$\beta_4 = \beta_5 = 0$$

Factorial ANOVA with effect coding is pretty automatic

- You don't have to make a table unless asked
- It always works as you expect it will
- Significance tests are the same as testing sets of contrasts
- Covariates present no problem. Main effects and interactions have their usual meanings, "controlling" for the covariates.
- Could plot the least squares means

Again

- Intercept is the grand mean
- Regression coefficients for the dummy variables are deviations of the marginal means from the grand mean
- Test of main effect(s) is test of the dummy variables for a factor.
- Interaction effects are regression coefficients corresponding to products of dummy variables.

Balanced vs. Unbalanced Experimental Designs

- Balanced design: Cell sample sizes are proportional (usually equal)
- Explanatory variables have zero relationship to one another
- Numerator SS in ANOVA are independent.
- Everything is nice and simple
- Most experimental studies are designed this way.
- As soon as somebody drops a test tube, it's no longer true

Analysis of unbalanced data

- When explanatory variables are related, there is potential ambiguity.
- A is related to Y, B is related to Y, and A is related to B.
- Who gets credit for the portion of variation in Y that could be explained by either A or B?
- With a regression approach, whether you use contrasts or dummy variables (equivalent), the answer is **nobody**.
- Think of full, restricted models.
- Equivalently, general linear test.

Some software is designed for balanced data

- The special purpose formulas are much simpler.
- They were very useful *in the past*.
- Since most real data are at least a little unbalanced, these formulas are a recipe for trouble.
- Most textbook data are balanced, so they cannot tell you what your software is really doing.
- R's anova and aov functions are designed for balanced data, though anova applied to Im objects can give you what you want if you use it with care.
- SAS proc glm is much more convenient. SAS proc anova is for balanced data. Avoid it.

Type I and Type III Tests

- proc glm displays both by default.
- Type III is the regression approach we know and love.
- We will use Type III and ignore Type I.
- But just for the record ...

Tests based on Type III SS

- Type III Sums of squares are sequential.
- In order of the effects in the model statement.
- Numerator of F-ratio is (SSR_F-SSR_R)/s, where the "restricted" model has all preceding terms, and the "full" model has those and also the effect being tested.
 - First term is tested not controlling for anything.
 - Second term is tested controlling for the first.
 - Third term is tested controlling for the first two.
 - And so on.
- But the denominator is MSE from the model with *all* effects.

Type I vs. Type III

- Type I test controls for all preceding effects.
- Type III test controls for *all* other effects.
- Type I and Type III tests of the *last* effect in the model are identical.
- For balanced data, Type I and Type III tests of all effects are the same.
- I can't remember what the Type II tests are.

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