

STA 431 Assignment 8

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		Endogenous	Exogenous
(a)	Observable	Y_1, Y_2	X
	Latent		$\varepsilon_1, \varepsilon_2$

(b) Independently for $i = 1, \dots, n$ let

$$Y_{i1} = \gamma_0 + \gamma_1 X_i + \varepsilon_{i1}$$

$$Y_{i2} = \beta_0 + \beta_1 Y_{i1} + \varepsilon_{i2} \quad \text{where}$$

$$X_i \sim N(\mu_x, \sigma^2), \quad \varepsilon_{i1} \sim N(0, \psi_1), \quad \varepsilon_{i2} \sim N(0, \psi_2)$$

and $X_i, \varepsilon_{i1}, \varepsilon_{i2}$ are independent.

$$(c) \quad \Theta = (\gamma_0, \gamma_1, \beta_0, \beta_1, \mu_x, \sigma, \psi_1, \psi_2)$$

~~(c)~~ $E(X) = \mu_x$

$$(d) \quad E(Y_1) = E(\gamma_0 + \gamma_1 X + \varepsilon_1)$$

$$= \gamma_0 + \gamma_1 \mu_x$$

$$E(Y_2) = E(\beta_0 + \beta_1 Y_1 + \varepsilon_2)$$

$$= \beta_0 + \beta_1 E(Y_1) + 0$$

$$= \beta_0 + \beta_1 (\gamma_0 + \gamma_1 \mu_x)$$

$$= \beta_0 + \beta_1 \gamma_0 + \beta_1 \gamma_1 \mu_x$$

(1d continued)

$$Y_1 = \gamma_0 + \gamma_1 X + \varepsilon_1$$

$$Y_2 = \beta_0 + \beta_1 Y_1 + \varepsilon_2$$

$$= \beta_0 + \beta_1 \gamma_0 + \gamma_1 \beta_1 X + \beta_1 \varepsilon_1 + \varepsilon_2$$

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	X	Y_1	Y_2
Σ	σ^2	$\gamma_1 \sigma^2$	$\gamma_1 \beta_1 \sigma^2$
=		$\gamma_1^2 \sigma^2 + \psi_1$	$\beta_1 (\gamma_1^2 \sigma^2 + \psi_1)$
			$\gamma_1^2 \beta_1^2 \sigma^2 + \beta_1^2 \psi_1$ + ψ_2

$$\text{cov}(Y_1, Y_2) = \text{cov}(Y_1, \beta_0 + \beta_1 Y_1 + \varepsilon_2)$$

$$= \beta_1 \text{var}(Y_1) = \beta_1 (\gamma_1^2 \sigma^2 + \psi_1)$$

$$\mu = \begin{pmatrix} \mu_x \\ \gamma_0 + \gamma_1 \mu_x \\ \beta_0 + \beta_1 (\gamma_0 + \gamma_1 \mu_x) \end{pmatrix}$$

(1e) Again, $\Theta = (\gamma_0, \gamma_1, \beta_0, \beta_1, \mu_x, \phi, \psi_1, \psi_2)$

$$\phi = \sigma_{11}$$

$$\gamma_1 = \sigma_{12} / \phi$$

$$\beta_1 = \sigma_{13} / (\gamma_1 \phi) \text{ That's } \sigma_{13} / (\gamma_1 \phi)$$

$$\psi_1 = \sigma_{22} - \gamma_1^2 \phi$$

$$\psi_2 = \sigma_{33} - \gamma_1^2 \beta_1^2 \phi - \beta_1^2 \psi_1$$

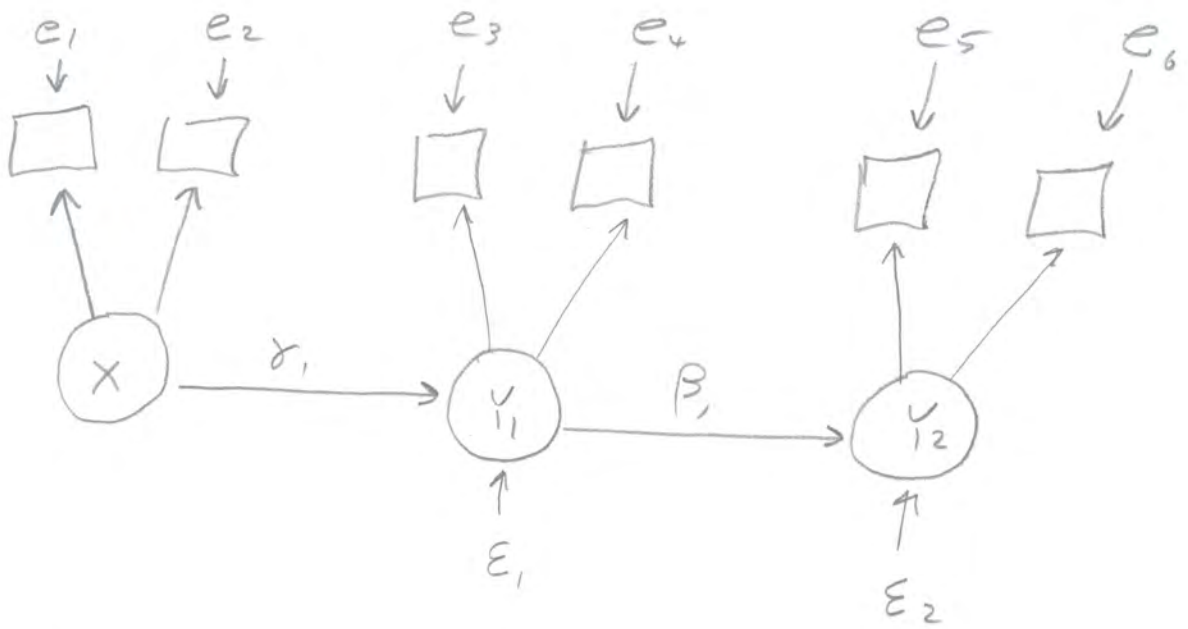
$$\mu_x = \mu_1 \text{ That's } \mu_1$$

$$\gamma_0 = \mu_2 - \gamma_1 \mu_x$$

$$\beta_0 = \mu_3 - \beta_1 (\gamma_0 + \gamma_1 \mu_x)$$

(f) No. There are 8 parameters and 9 moment structure equations.

(1g)

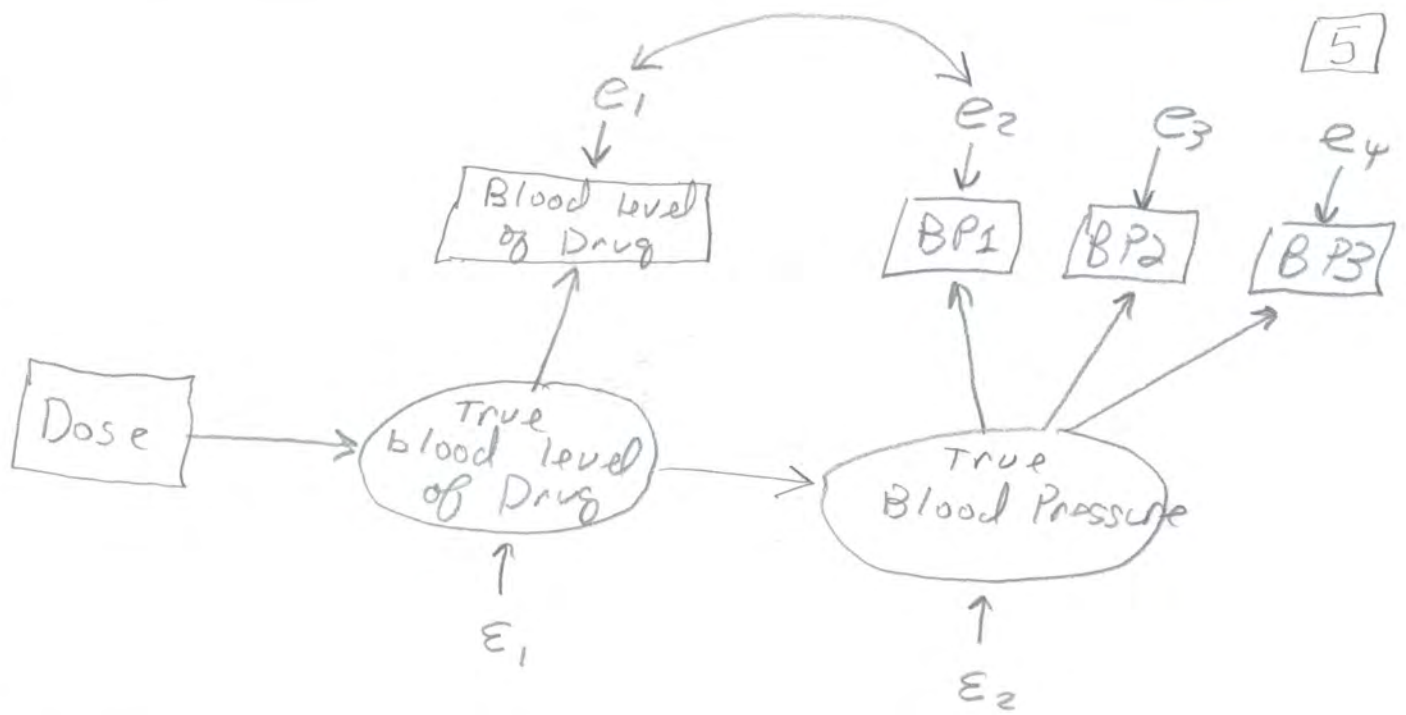


(i)

(ii)

Yes, all parameters are identifiable. Identifiability of two parameters in two latent model was established in part (e) of this question - and anyway it's established by the Acyclic Rule. Parameters of two measurement model are identifiable by the Double Measurement Rule, and parameters of two combined model are identifiable by the Two-Stage Rule.

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(a) Model Equations in scalar form

$$Y_1 = \gamma_1 X_1 + \beta_1 Y_3 + \epsilon_1$$

$$Y_2 = \gamma_2 X_2 + \beta_2 Y_1 + \epsilon_2$$

$$Y_3 = \gamma_3 X_1 + \gamma_3 X_2 + \beta_3 Y_2 + \epsilon_3$$

$$D_1 = \lambda_1 X_1 + e_1$$

$$D_2 = \lambda_1 X_1 + e_2$$

$$D_3 = \lambda_2 Y_1 + e_3$$

$$D_4 = \lambda_3 Y_2 + e_4$$

$$D_5 = \lambda_4 X_2 + e_5$$

$$D_6 = \lambda_4 X_2 + e_6$$

(3b.)

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$$\begin{pmatrix} Y_1 \\ Y_2 \\ Y_3 \end{pmatrix} = \begin{pmatrix} 0 & 0 & \beta_1 \\ \beta_2 & 0 & 0 \\ 0 & \beta_3 & 0 \end{pmatrix} \begin{pmatrix} Y_1 \\ Y_2 \\ Y_3 \end{pmatrix} + \begin{pmatrix} \gamma_1 & 0 \\ 0 & \gamma_2 \\ \gamma_3 & \gamma_3 \end{pmatrix} \begin{pmatrix} X_1 \\ X_2 \end{pmatrix} + \begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \varepsilon_3 \end{pmatrix}$$

(c)

$$\begin{pmatrix} D_1 \\ D_2 \\ D_3 \\ D_4 \\ D_5 \\ D_6 \end{pmatrix} = \begin{pmatrix} \lambda_1 & 0 & 0 & 0 & 0 \\ \lambda_1 & 0 & 0 & 0 & 0 \\ 0 & 0 & \lambda_2 & 0 & 0 \\ 0 & 0 & 0 & \lambda_3 & 0 \\ 0 & \lambda_4 & 0 & 0 & 0 \\ 0 & \lambda_4 & 0 & 0 & 0 \end{pmatrix} \begin{pmatrix} X_1 \\ X_2 \\ Y_1 \\ Y_2 \\ Y_3 \end{pmatrix} + \begin{pmatrix} e_1 \\ e_2 \\ e_3 \\ e_4 \\ e_5 \\ e_6 \end{pmatrix}$$

$$d = \Lambda F + e$$

(d)

$$\Phi = (\varphi)$$

$$\varphi = \begin{pmatrix} \varphi_1 & 0 & 0 \\ 0 & \varphi_2 & 0 \\ 0 & 0 & \varphi_3 \end{pmatrix}$$

(3d continued)

$$\Omega =$$

	e_1	e_2	e_3	e_4	e_5	e_6
e_1	w_{11}	w_{12}	0	0	0	0
e_2		w_{22}	0	0	0	0
e_3			w_{33}	0	0	0
e_4				w_{44}	w_{45}	0
e_5					w_{55}	w_{56}
e_6						w_{66}

It's symmetric of course.

④ (a)

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$$(4) y_i = \beta y_i + \Gamma' x_i + \varepsilon_i$$

$$(a) y_i - \beta y_i = \Gamma' x_i + \varepsilon_i$$

$$\Rightarrow I y_i - \beta y_i = \Gamma' x_i + \varepsilon_i$$

$$\Rightarrow (I - \beta) y_i = \Gamma' x_i + \varepsilon_i$$

$$\Rightarrow y_i = (I - \beta)^{-1} \Gamma' x_i + (I - \beta)^{-1} \varepsilon_i$$

$$\text{So } \text{cov}(y_i) = (I - \beta)^{-1} \Gamma \text{cov}(x_i) (\Gamma' (I - \beta)^{-1})' + (I - \beta)^{-1} \Psi (I - \beta)^{-1'}'$$

$$= (I - \beta)^{-1} \Gamma \Phi_x \Gamma' ((I - \beta)^{-1})' + (I - \beta)^{-1} \Psi (I - \beta)^{-1'}$$

$$= (I - \beta)^{-1} \Gamma \Phi_x \Gamma' (I - \beta^T)^{-1} + (I - \beta)^{-1} \Psi (I - \beta^T)^{-1}$$

$$(b) \text{cov}(x_i, y_i) = \text{cov}(x_i, (I - \beta)^{-1} \Gamma' x_i)$$

$$= \text{cov}(x_i, x_i) (\Gamma' (I - \beta)^{-1})'$$

$$= \Phi_x \Gamma' (I - \beta^T)^{-1}$$

(5) (a)

$$(I - \beta) y_i = \Gamma' x_i + \varepsilon_i$$

$$\Rightarrow v^T (I - \beta) y_i = v^T \Gamma' x_i + v^T \varepsilon_i$$

$$\Rightarrow 0 = v^T \Gamma' x_i + v^T \varepsilon_i$$

$$\Rightarrow v^T \varepsilon_i = -v^T \Gamma' x_i$$

$$(b) \Rightarrow \text{cov}(v^T \varepsilon_i) = -\text{cov}(v^T \Gamma' x_i)$$

$$\Rightarrow v^T \Psi v = -v^T \Gamma' \Phi_x \Gamma' v \quad (*)$$

$v^T \Psi v > 0$ because Ψ is positive definite and $v^T \Gamma' \Phi_x \Gamma' v \geq 0$ because $\text{cov}(\Gamma' x_i) = \Gamma' \Phi_x \Gamma'$ is a covariance matrix, and all covariance matrices are non-negative definite.

Thus, $0 < v^T \Psi v = -v^T \Gamma' \Phi_x \Gamma' v \leq 0$, a contradiction. The existence of such a v is impossible. Thus the rows of $(I - \beta)$ are linearly independent, and $(I - \beta)^{-1}$ exists. \square

Another way to produce a contradiction is from (*).

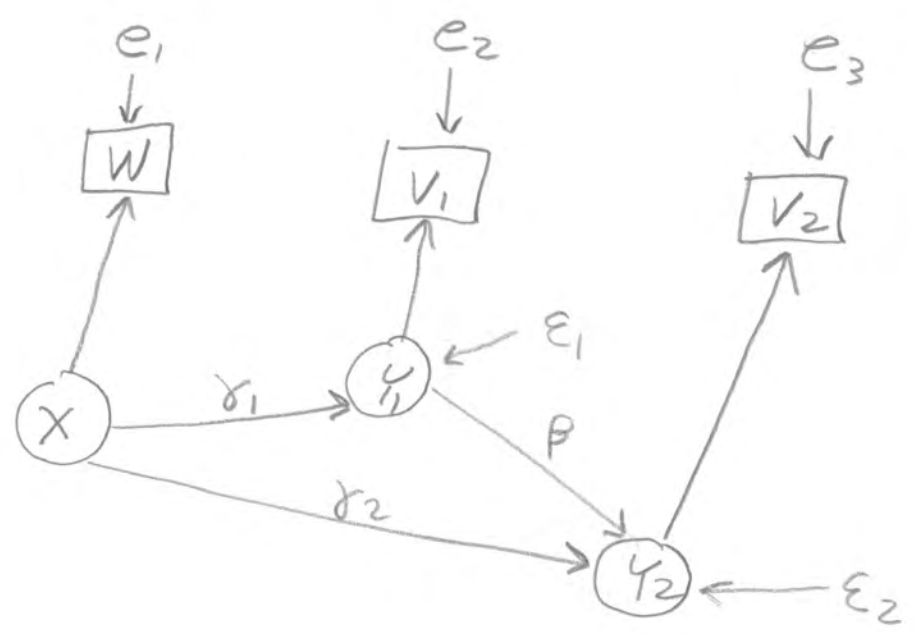
$$\text{cov}(v^T \varepsilon_i, v^T \varepsilon_i) = \text{cov}(v^T \varepsilon_i, -v^T \Gamma' x_i)$$

$$\begin{array}{ccc} \parallel & & \parallel \\ v^T \Psi v > 0 & & 0 \end{array}$$

(6)

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(a)



(b) No. $\theta = (\varnothing, \gamma_1, \gamma_2, \beta, \psi_1, \psi_2, w_1, w_2, w_3)$
 9 parameters and 6 covariance structure equations. Fails the parameter count rule

(7) (a) Independently for $i=1, \dots, n$, let

$$D_{i,1} = F_i + e_{i,1}$$

$$D_{i,2} = \lambda F_i + e_{i,2}$$

$$D_{i,3} = \lambda F_i + e_{i,3}$$

where

$$\text{Var}(F_i) = \sigma$$

$$\text{Var}(e_{i,1}) = w_1$$

$$\text{Var}(e_{i,2}) = w_2$$

$$\text{Var}(e_{i,3}) = w_3$$

(b) $\theta = (\lambda, \sigma, w_1, w_2, w_3)$

(c) Yes. 6 equations in 5 unknown parameters.

	D_1	D_2	D_3
D_1	$\sigma + w_1$	$\lambda \sigma$	$\lambda \sigma$
D_2		$\lambda^2 \sigma + w_2$	$\lambda^2 \sigma$
D_3			$\lambda^2 \sigma + w_3$

If $\lambda = 0$, then λ, w_2 and w_3 are identifiable but σ and w_1 are not. If $\lambda \neq 0$,

$$\sigma = \sigma_{12} \sigma_{13} / \sigma_{23}, \quad \lambda = \sigma_{13} / \sigma_{12}, \quad w_1 = \sigma_{11} - \sigma$$

$$w_2 = \sigma_{22} - \lambda^2 \sigma, \quad w_3 = \sigma_{33} - \lambda^2 \sigma$$

And all parameters are identifiable.

(e) 6 equations in 5 unknowns: one df

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(a)

$$Y_1 = \gamma X + \epsilon_1$$

$$Y_2 = \beta_1 Y_1 + \epsilon_2$$

$$Y_3 = \beta_2 Y_1 + \epsilon_3$$

$$V = Y_1 + \epsilon$$

(b) $\Theta = (\gamma \beta_1 \beta_2 \phi \omega \psi_1 \psi_2 \psi_3 \psi_{23})$

(c) Yes $4(4+1)/2 = 10$ equations in 9 unknown parameters

(d) Just part of the covariance matrix:

	X	V	Y_2	Y_3
X	ϕ	$\gamma\phi$	$\gamma\beta_1\phi$	$\gamma\beta_2\phi$

$$\phi = \sigma_{11}$$

$$\gamma = \sigma_{12} / \sigma_{11}$$

$$\beta_1 = \sigma_{13} / \sigma_{12}$$

$$\beta_2 = \sigma_{14} / \sigma_{12}$$

Showing identifiability provided $\beta_1, \beta_2 \neq 0$

$$Y_1 = \gamma X + \varepsilon_1$$

$$Y_2 = \beta_1 Y_1 + \varepsilon_2 = \beta_1 (\gamma X + \varepsilon_1) + \varepsilon_2$$

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$$= \gamma \beta_1 X + \beta_1 \varepsilon_1 + \varepsilon_2$$

$$Y_3 = \gamma \beta_2 X + \beta_2 \varepsilon_1 + \varepsilon_3$$

$$V = Y_1 + e = \gamma X + \varepsilon_1 + e$$

$\Sigma =$

	X	V	Y_{12}	Y_{13}
X	σ^2	$\gamma \sigma^2$	$\gamma \beta_1 \sigma^2$	$\gamma \beta_2 \sigma^2$
V		$\gamma^2 \sigma^2 + \psi_1 + \omega$	$\gamma^2 \beta_1 \sigma^2 + \beta_1 \psi_1$	$\gamma^2 \beta_2 \sigma^2 + \beta_2 \psi_1$
Y_{12}			$\gamma^2 \beta_1^2 \sigma^2 + \beta_1^2 \psi_1 + \psi_2$	$\gamma^2 \beta_1 \beta_2 \sigma^2 + \beta_1 \beta_2 \psi_1 + \psi_{23}$
Y_{13}				$\gamma^2 \beta_2^2 \sigma^2 + \beta_2^2 \psi_1 + \psi_3$

$$\Theta = (\gamma \quad \sigma^2 \quad \psi_1 \quad \beta_1 \quad \psi_2 \quad \beta_2 \quad \psi_3 \quad \omega \quad \psi_{23})$$

Assignment 8, Question 9

```
> # Simulate manipulation check design
> # Assignment 8, Question 9
>
> rm(list=ls())
> # source("rmvn.txt")
> source("https://www.utstat.toronto.edu/~brunner/openSEM/fun/rmvn.txt")
>
> # install.packages("lavaan", dependencies = TRUE) # Only need to do this once
> library(lavaan)
```

This is lavaan 0.6-11

lavaan is FREE software! Please report any bugs.

```
>
> # Set parameter values and sample size
> gamma = 1; beta1 = 0.75; beta2 = 1.25; phi = 2
> omega = 1; psi1 = 3; psi2 = 4; psi3 = 5; psi23 = 2
> n = 10000
>
> ##### Generate data #####
> set.seed(888999)
> X = rnorm(n,0,sqrt(phi))
> epsilon1 = rnorm(n,0,sqrt(psi1))
> e = rnorm(n,0,sqrt(omega))
> V23 = rbind(c(psi2,psi23),
+           c(psi23,psi3))
> er = rmvn(n,c(0,0),V23)
> epsilon2 = er[,1]
> epsilon3 = er[,2]
>
> Y1 = gamma*X + epsilon1
> Y2 = beta1*Y1 + epsilon2
> Y3 = beta2*Y1 + epsilon3
> V = Y1 + e
> Mcheck = cbind(X,V,Y2,Y3)
> #####
>
> cor(Mcheck)
      X          V          Y2          Y3
X  1.0000000  0.5726864  0.3931787  0.4845121
V  0.5726864  1.0000000  0.5790716  0.7084621
Y2 0.3931787  0.5790716  1.0000000  0.7146340
Y3 0.4845121  0.7084621  0.7146340  1.0000000
```

```

> ##### Fit model with lavaan
>
> mod1 = 'Y1 ~ gamma*X
+       Y1 =~ 1.0*V + beta1*Y2 + beta2*Y3
+       # Variances and covariances
+       X ~~ phi*X      # Var(X) = phi
+       V ~~ omega*V    # Var(e) = omega
+       Y1 ~~ psi1*Y1   # Var(epsilon1) = psi1
+       Y2 ~~ psi2*Y2   # Var(epsilon2) = psi2
+       Y3 ~~ psi3*Y3   # Var(epsilon3) = psi3
+       Y2 ~~ psi23*Y3  # Cov(epsilon2,epsilon3) = psi23
+       ' # End of mod1
>
> fit1 = lavaan(mod1, data=Mcheck)
> # summary(fit1)
> parameterEstimates(fit1)
  lhs op rhs label  est    se      z pvalue ci.lower ci.upper
1  Y1 ~  X gamma 0.983 0.014 69.859    0    0.955  1.010
2  Y1 =~ V      1.000 0.000    NA     NA    1.000  1.000
3  Y1 =~ Y2 beta1 0.732 0.013 56.490    0    0.707  0.758
4  Y1 =~ Y3 beta2 1.247 0.018 68.087    0    1.211  1.283
5  X  ~~ X  phi  2.010 0.028 70.711    0    1.954  2.066
6  V  ~~ V  omega 0.956 0.060 15.817    0    0.838  1.075
7  Y1 ~~ Y1 psi1  3.020 0.073 41.500    0    2.877  3.162
8  Y2 ~~ Y2 psi2  3.997 0.068 58.848    0    3.864  4.130
9  Y3 ~~ Y3 psi3  5.164 0.118 43.867    0    4.933  5.395
10 Y2 ~~ Y3 psi23 2.088 0.075 27.934    0    1.941  2.234

>
> # Try mod2 with just one measurement of Y1
> # and regressions connecting Y1 to Y2 and Y3
>
> mod2 = 'Y1 ~ gamma*X
+       Y2 ~ beta1*Y1
+       Y3 ~ beta2*Y1
+       Y1 =~ 1.0*V
+       # Variances and covariances
+       X ~~ phi*X      # Var(X) = phi
+       V ~~ omega*V    # Var(e) = omega
+       Y1 ~~ psi1*Y1   # Var(epsilon1) = psi1
+       Y2 ~~ psi2*Y2   # Var(epsilon2) = psi2
+       Y3 ~~ psi3*Y3   # Var(epsilon3) = psi3
+       Y2 ~~ psi23*Y3  # Cov(epsilon2,epsilon3) = psi23
+       ' # End of mod2
>
> fit2 = lavaan(mod2, data=Mcheck)
> parameterEstimates(fit2)
  lhs op rhs label  est    se      z pvalue ci.lower ci.upper
1  Y1 ~  X gamma 0.983 0.014 69.859    0    0.955  1.010
2  Y2 ~  Y1 beta1 0.732 0.013 56.490    0    0.707  0.758
3  Y3 ~  Y1 beta2 1.247 0.018 68.087    0    1.211  1.283
4  Y1 =~ V      1.000 0.000    NA     NA    1.000  1.000
5  X  ~~ X  phi  2.010 0.028 70.711    0    1.954  2.066
6  V  ~~ V  omega 0.956 0.060 15.817    0    0.838  1.075
7  Y1 ~~ Y1 psi1  3.020 0.073 41.500    0    2.877  3.162
8  Y2 ~~ Y2 psi2  3.997 0.068 58.848    0    3.864  4.130
9  Y3 ~~ Y3 psi3  5.164 0.118 43.867    0    4.933  5.395
10 Y2 ~~ Y3 psi23 2.088 0.075 27.934    0    1.941  2.234

```