Background: Matrices and Random Vectors¹ STA431 Spring 2017

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3 Multivariate Normal

Matrices

- $\mathbf{A} = [a_{ij}]$
- Transpose: $\mathbf{A}^{\top} = [a_{ji}]$
- Multiplication: $\mathbf{AB} \neq \mathbf{BA}$
- $(\mathbf{A}\mathbf{B})^{\top} = \mathbf{B}^{\top}\mathbf{A}^{\top}$
- Inverse of a square matrix: $\mathbf{A}^{-1}\mathbf{A} = \mathbf{A}\mathbf{A}^{-1} = \mathbf{I}$. (Only need to show it in one direction.)

•
$$(\mathbf{A}^{-1})^{\top} = (\mathbf{A}^{\top})^{-1}$$

Trace of a square matrix: Sum of the diagonal elements

$$tr(\mathbf{A}) = \sum_{i=1}^{n} a_{i,i}$$

• Of course $tr(\mathbf{A} + \mathbf{B}) = tr(\mathbf{A}) + tr(\mathbf{B})$,

•
$$tr(\mathbf{A}) = tr(\mathbf{A}^{\top})$$
, etc.

- But less obviously, even though $AB \neq BA$,
- $tr(\mathbf{AB}) = tr(\mathbf{BA})$

Proof of $tr(\mathbf{AB}) = tr(\mathbf{BA})$ Using $\mathbf{AB} = \mathbf{C} = [c_{i,j}] = \sum_k a_{i,k} b_{k,j}$

Let **A** be an $r \times p$ matrix and **B** be a $p \times r$ matrix, so that the product matrices **AB** and **BA** are both defined.

$$tr(\mathbf{AB}) = \sum_{i=1}^{r} \left(\sum_{k=1}^{p} a_{i,k} b_{k,i} \right)$$
$$= \sum_{k=1}^{p} \left(\sum_{i=1}^{r} b_{k,i} a_{i,k} \right)$$
$$= tr(\mathbf{BA})$$

Random vectors Expected values and variance-covariance matrices

•
$$E(\mathbf{X}) = [E(X_{i,j})]$$

• $E(\mathbf{X} + \mathbf{Y}) = E(\mathbf{X}) + E(\mathbf{Y})$

•
$$E(\mathbf{AXB}) = \mathbf{A}E(\mathbf{X})\mathbf{B}$$

•
$$cov(\mathbf{X}) = E\left\{ (\mathbf{X} - \boldsymbol{\mu})(\mathbf{X} - \boldsymbol{\mu})^{\top} \right\}$$

• $cov(\mathbf{X}) = A cov(\mathbf{X}) A^{\top}$

•
$$cov(\mathbf{A}\mathbf{X}) = \mathbf{A}cov(\mathbf{X})\mathbf{A}$$

•
$$cov(\mathbf{X}, \mathbf{Y}) = E\left\{ (\mathbf{X} - \boldsymbol{\mu}_x)(\mathbf{Y} - \boldsymbol{\mu}_y)^\top \right\}$$

•
$$cov(\mathbf{X} + \mathbf{a}) = cov(\mathbf{X})$$

•
$$cov(\mathbf{X} + \mathbf{a}, \mathbf{Y} + \mathbf{b}) = cov(\mathbf{X}, \mathbf{Y})$$

The Centering Rule Based on $cov(\mathbf{X} + \mathbf{a}) = cov(\mathbf{X})$

Often, variance and covariance calculations can be simplified by subtracting off constants first.

Denote the *centered* version of **X** by $\overset{c}{\mathbf{X}} = \mathbf{X} - E(\mathbf{X})$, so that

•
$$E(\mathbf{X}^{c}) = \mathbf{0}$$
 and

•
$$cov(\mathbf{X}) = E(\mathbf{X}\mathbf{X}^{\top}) = cov(\mathbf{X})$$

Linear combinations These are matrices, but they could be scalars

$$\mathbf{L} = \mathbf{A}_1 \mathbf{X}_1 + \dots + \mathbf{A}_m \mathbf{X}_m + \mathbf{b}$$

$$\mathbf{\tilde{L}} = \mathbf{A}_1 \mathbf{\tilde{X}}_1 + \dots + \mathbf{A}_m \mathbf{\tilde{X}}_m, \text{ where}$$

$$\mathbf{\tilde{X}}_j = \mathbf{X}_j - E(\mathbf{X}_j) \text{ for } j = 1, \dots, m.$$

The centering rule says

$$cov(\mathbf{L}) = E(\overset{c}{\mathbf{L}}\overset{c}{\mathbf{L}}^{\top})$$
$$cov(\mathbf{L}_1, \mathbf{L}_2) = E(\overset{c}{\mathbf{L}}_1 \overset{c}{\mathbf{L}}_2^{\top})$$

In words: To calculate variances and covariances of linear combinations, one may simply discard added constants, center all the random vectors, and take expected values of products.

Example: $cov(\mathbf{X} + \mathbf{Y})$ Using the centering rule

$$\begin{aligned} cov(\mathbf{X} + \mathbf{Y}) &= E(\overset{c}{\mathbf{X}} + \overset{c}{\mathbf{Y}})(\overset{c}{\mathbf{X}} + \overset{c}{\mathbf{Y}})^{\top} \\ &= E(\overset{c}{\mathbf{X}} + \overset{c}{\mathbf{Y}})(\overset{c}{\mathbf{X}}^{\top} + \overset{c}{\mathbf{Y}}^{\top}) \\ &= E(\overset{c}{\mathbf{X}}\overset{c}{\mathbf{X}}^{\top}) + E(\overset{c}{\mathbf{Y}}\overset{c}{\mathbf{Y}}^{\top}) + E(\overset{c}{\mathbf{X}}\overset{c}{\mathbf{Y}}^{\top}) + E(\overset{c}{\mathbf{Y}}\overset{c}{\mathbf{X}}^{\top}) \\ &= cov(\mathbf{X}) + cov(\mathbf{Y}) + cov(\mathbf{X}, \mathbf{Y}) + cov(\mathbf{Y}, \mathbf{X}) \end{aligned}$$

• Does
$$cov(\mathbf{Y}, \mathbf{X}) = cov(\mathbf{X}, \mathbf{Y})$$
?
• Does $cov(\mathbf{Y}, \mathbf{X}) = cov(\mathbf{X}, \mathbf{Y})^{\top}$?

Use $cov(\mathbf{X}, \mathbf{Y}) = E\left\{ (\mathbf{X} - \boldsymbol{\mu}_x)(\mathbf{Y} - \boldsymbol{\mu}_y)^\top \right\}$

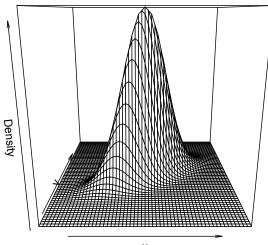
The Multivariate Normal Distribution

The $p \times 1$ random vector **X** is said to have a *multivariate normal* distribution, and we write $\mathbf{X} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, if **X** has (joint) density

$$f(\mathbf{x}) = \frac{1}{|\mathbf{\Sigma}|^{\frac{1}{2}} (2\pi)^{\frac{p}{2}}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^{\top} \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right\},\$$

where $\boldsymbol{\mu}$ is $p \times 1$ and $\boldsymbol{\Sigma}$ is $p \times p$ symmetric and positive definite.

The Bivariate Normal Density Multivariate normal with p = 2 variables



Analogies

Multivariate normal reduces to the univariate normal when $p=1\,$

• Univariate Normal

•
$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left\{-\frac{1}{2}\frac{(x-\mu)^2}{\sigma^2}\right\}$$

• $E(X) = \mu, Var(X) = \sigma^2$

•
$$\frac{(X-\mu)^2}{\sigma^2} \sim \chi^2(1)$$

• Multivariate Normal

•
$$f(\mathbf{x}) = \frac{1}{|\mathbf{\Sigma}|^{\frac{1}{2}} (2\pi)^{\frac{p}{2}}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^{\top} \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right\}$$

• $E(\mathbf{X}) = \boldsymbol{\mu}, cov(\mathbf{X}) = \boldsymbol{\Sigma}$

•
$$(\mathbf{X} - \boldsymbol{\mu})^{\top} \boldsymbol{\Sigma}^{-1} (\mathbf{X} - \boldsymbol{\mu}) \sim \chi^2(p)$$

More properties of the multivariate normal

- If **c** is a vector of constants, $\mathbf{X} + \mathbf{c} \sim N(\mathbf{c} + \boldsymbol{\mu}, \boldsymbol{\Sigma})$
- If **A** is a matrix of constants, $\mathbf{A}\mathbf{X} \sim N(\mathbf{A}\boldsymbol{\mu}, \mathbf{A}\boldsymbol{\Sigma}\mathbf{A}^{\top})$
- Linear combinations of multivariate normals are multivariate normal.
- All the marginals (dimension less than p) of **X** are (multivariate) normal.
- For the multivariate normal, zero covariance implies independence. The multivariate normal is the only continuous distribution with this property.

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http://www.utstat.toronto.edu/~brunner/oldclass/431s17