NAME (PRINT): Last/Surname

First /Given Name

STUDENT #:

SIGNATURE:

# UNIVERSITY OF TORONTO MISSISSAUGA APRIL 2011 FINAL EXAMINATION STA431H5S Structural Equation Models Jerry Brunner Duration - 3 hours Aids: Statistical Calculators; Reference Sheet (Supplied)

The University of Toronto Mississauga and you, as a student, share a commitment to academic integrity. You are reminded that you may be charged with an academic offence for possessing any unauthorized aids during the writing of an exam, including but not limited to any electronic devices with storage, such as cell phones, pagers, personal digital assistants (PDAs), iPods, and MP3 players. Unauthorized calculators and notes are also not permitted. Do not have any of these items in your possession in the area of your desk. Please turn the electronics off and put all unauthorized aids with your belongings at the front of the room before the examination begins. If any of these items are kept with you during the writing of your exam, you may be charged with an academic offence. A typical penalty may cause you to fail the course.

Please note, you **CANNOT** petition to **re-write** an examination once the exam has begun.

Qn. #	Value	Score	
1	10		
2	27		
3	23		
4	12		
5	10		
6	18		
Total = 100 Points			

1. Let F be a latent variable. Consider two equivalent measurements with uncorrelated measurement error:

$$D_1 = \nu + F + e_1$$
$$D_2 = \nu + F + e_2,$$

where  $E(F) = \mu$ ,  $Var(F) = \phi$ ,  $E(e_1) = E(e_2) = 0$ ,  $Var(e_1) = Var(e_2) = \omega$ , and F,  $e_1$  and  $e_2$  are all independent. The observable variables in this model are  $D_1$  and  $D_2$ . The *reliability* of an observable measurement is defined as the squared correlation between the true (latent) measurement and the observable measurement. Show that the reliability of  $D_1$  equals the correlation between  $D_1$  and  $D_2$ . Hint: both quantities equal a familiar formula.

27 points 2. Consider the following path diagram.



(a) Write down the model equations in scalar form.

(b) Refer to the general two-stage structural equation model on the formula sheet. Write down the following, using symbols from the path diagram. This part is intended to help you use matrices with the correct dimensions.

i.  $\mathbf{X}$ 

ii.  $\mathbf Y$ 

iii.  $\mathbf{F}$ 

iv.  $\mathbf{D}$ 

(c) Give the matrix  $\beta$ , using symbols from the path diagram. Make sure it has the correct dimensions.

(d) Give the matrix  $\Gamma$ , using symbols from the path diagram. Make sure it has the correct dimensions.

(e) Give the matrix  $\Lambda$ , using symbols from the path diagram. Make sure it has the correct dimensions.

(f) Give the matrix  $\Phi_{11}$ , using symbols from the path diagram. Make sure it has the correct dimensions. Not all symbols in  $\Phi_{11}$  appear on the path diagram.

(g) Give the matrix  $\Psi$ . Make sure it has the correct dimensions. The symbols in  $\Psi$  do not appear on the path diagram. Use standard notation.

(h) Give the matrix  $\Omega$ , using symbols from the path diagram. Make sure it has the correct dimensions. Not all symbols in  $\Omega$  appear on the path diagram.

- 3. For the measurement part of the general two-stage structural equation model (see formula sheet), suppose  $\Phi$  is the identity and  $\Omega$  is diagonal. In addition, suppose that the observable variables have been standardized, and we write **Z** in place of **D**. This yields the classical exploratory factor analysis model. In your calculations, assume **F** is  $p \times 1$  and **Z** is  $k \times 1$ .
  - (a) Show that factor loadings are actually correlations by calculating  $Corr(Z_i, F_j)$ . This is a scalar (not matrix) calculation. Show your work. Circle your final answer.

(b) Show that  $\omega_{i,i} = Var(e_i)$  is a function of the  $\lambda_{i,j}$  quantities. This is a scalar (not matrix) calculation. Show your work. Circle your final answer.

(c) What are the *unknown* parameters in this classical factor analysis model? The answer is a list of one or more matrices.

(d) Prove that the parameters of the the classical exploratory factor analysis model (the model you are working with in this question) are not identifiable. This is the case even if there are enough observable variables so that the counting rule is satisfied, so don't try to use the counting rule. Show your work.

4. let

 $\begin{array}{rcl} D_1 &=& \lambda_1 F + e_1 \\ D_2 &=& \lambda_2 F + e_2 \\ D_3 &=& \lambda_3 F + e_3, \end{array}$ 

where all expected values are zero,  $V(e_i) = \omega_i$  for i = 1, 2, 3,  $V(F) = \phi$ , the error terms are independent of one another and of the factor,  $\lambda_1 > 0$ ,  $\lambda_2 \neq 0$  and  $\lambda_3 \neq 0$ . The parameters of the model are not identifiable (you don't have to show this), but you can work with it anyway; do not re-parameterize!

(a) Give the covariance matrix  $\Sigma$ . This is a scalar calculation.

(b) Show that the ratio of factor loadings  $\lambda_1/\lambda_2$  is identifiable.

(c) Show that the parameter  $\omega_1$  is identifiable. You have more room on this page than you need.

10 points 5. The parameters of the model represented by the following path diagram are identifiable. Briefly explain why. Make specific reference to the identifiability rules on the reference sheet.



6. The *Poverty Data* contain information from a sample of 97 countries. The variables include Live birth rate per 1,000 of population, Death rate per 1,000 of population, Infant deaths per 1,000 of population under 1 year old, Life expectancy at birth for males, Life expectancy at birth for females, and Gross National Product per capita in U.S. dollars. Please answer these questions based on the SAS program and list file starting on Page 12.

Note: The printout shows two runs of proc calis. Each one fits a different model.

- (a) What is the parameter  $\theta$  for the factor analysis model? That's the first run of proc calis. Give your answer in the form of a list of names from the SAS job.
- (b) Does the model fit the data adequately? Answer Yes or No, and back up your answer with two numbers from the printout: The value of a test statistic, and a *p*-value.
- (c) What is the maximum likelihood estimate of the correlation between factors? The answer is a single number from the printout.
- (d) Clearly, the factor loadings are not identifiable for points in the parameter space where  $\phi_{12} = 0$  In your view, is this likely to be a problem in the present case? Answer "Yes, a problem," or "No problem" and briefly explain. Does the printout support your position?
- (e) Now consider the second run of **proc calis** the latent regression model. Does this model fit the data adequately? Answer Yes or No. Is the fit better than the factor analysis model, worse than the factor analysis model, or is it the same? Write one of these alternatives and *circle it*.
- (f) According to the regression model, wealth causes health if there is any connection between the two. How do you know *from the printout* that more money translates into better health (and not worse health)? Cite a particular number from the printout.

That's the end of the questions. The rest of this exam script is computer printout.

```
/************************* finalpoverty2011B.sas *******************************/
options linesize=79 noovp formdlim='-';
title 'UN Poverty Data: STA431 S 2011 Final Exam';
data misery;
     infile 'poverty.data';
    input birthrate deathrate infmort lifexM lifexF gnp group country $;
    gnp1000 = gnp/1000; /* In thousands of dollars */
    lifex = (lifexM+lifexF)/2;
proc calis cov;
    title2 'Factor analysis model standardizing the factors';
    var lifex infmort gnp1000 birthrate;
    lineqs
                  = lambda1 Fhealth + e1,
         lifex
         infmort = lambda2 Fhealth + e2,
         gnp1000 = lambda3 Fwealth + e3,
         birthrate = lambda4 Fwealth + e4;
     std
         Fhealth = 1, Fwealth=1,
         e1-e4 = 4 * omega: ;
     cov Fhealth Fwealth = phi12;
    bounds 0.0 < omega1-omega4;</pre>
proc calis cov;
    title2 'Latent regression model setting loadings to one';
    var lifex infmort gnp1000 birthrate;
    lineqs
         Fhealth = gamma Fwealth + epsilon,
         lifex
                             Fhealth + e1,
                  =
          infmort = lambda2 Fhealth + e2,
         gnp1000 =
                            Fwealth + e3,
         birthrate = lambda4 Fwealth + e4;
     std
         Fwealth=phi, epsilon = psi,
          e1-e4 = 4 * omega: ;
     bounds 0.0 < phi psi omega1-omega4;</pre>
```

Parts of finalpoverty2011B.lst \_\_\_\_\_ UN Poverty Data: STA431 S 2011 Final Exam 1 Factor analysis model standardizing the factors The 4 Endogenous Variables Manifest lifex infmort gnp1000 birthrate Latent The 6 Exogenous Variables Manifest Latent Fhealth Fwealth Error e1 e2 e3 e4 \_\_\_\_\_ UN Poverty Data: STA431 S 2011 Final Exam 5 Factor analysis model standardizing the factors Actual Over Max Abs Rest Func Act Objective Obj Fun Gradient Pred Iter arts Calls Con Function Change Element Lambda Change 0 2 0 0.01215 0.00548 0.0285 0 1.099 1 3 4 5 0.01207 0.000077 0.000084 2 0 0 0 1.024 0 0 0.01207 1.067E-6 0.000012 3 0 0.01207 1.067E-6 0.000012 0 1.024 0.01207 1.6E-8 1.267E-6 0 1.023 1.024 0 4 0 Optimization Results Iterations 4 Function Calls 6 Jacobian Calls 5 Active Constraints 0 Objective Function 0.0120694315 Max Abs Gradient Element 1.2666839E-6 Lambda 0 Actual Over Pred Change 1.0233020931 Radius 0.0010158686

ABSGCONV convergence criterion satisfied.

\_\_\_\_\_ UN Poverty Data: STA431 S 2011 Final Exam 6 Factor analysis model standardizing the factors The CALIS Procedure Covariance Structure Analysis: Maximum Likelihood Estimation Fit Function 0.0121 Goodness of Fit Index (GFI) 0.9940 GFI Adjusted for Degrees of Freedom (AGFI) 0.9404 Root Mean Square Residual (RMR) 2.5734 Standardized Root Mean Square Residual (SRMR) 0.0067 Parsimonious GFI (Mulaik, 1989) 0.1657 Chi-Square 1.0862 Chi-Square DF 1 Pr > Chi-Square 0.2973 Independence Model Chi-Square 399.83 Independence Model Chi-Square DF 6 -----\_\_\_\_ 7 UN Poverty Data: STA431 S 2011 Final Exam Factor analysis model standardizing the factors The CALIS Procedure

Covariance Structure Analysis: Maximum Likelihood Estimation

Manifest Variable Equations with Estimates

lifex	=	10.3124*Fhealth	+	1.0000	e1
Std Err		0.7841 lambda1			
t Value		13.1522			
infmort	=	-44.2690*Fhealth	+	1.0000	e2
Std Err		3.6165 lambda2			
t Value		-12.2408			
gnp1000	=	5.4842*Fwealth	+	1.0000	e3
Std Err		0.7664 lambda3			
t Value		7.1560			
birthrate	=	-12.7179*Fwealth	+	1.0000	e4
Std Err		1.1530 lambda4			
t Value		-11.0302			

#### Variances of Exogenous Variables

			Standard	
Variable	Parameter	Estimate	Error	t Value
Fhealth		1.00000		
Fwealth		1.00000		
e1	omega1	1.44765	1.84498	0.78
e2	omega2	184.16184	43.51831	4.23
e3	omega3	35.43109	5.57604	6.35
e4	omega4	25.92052	10.36055	2.50

Var1	Var2	Parameter	Estimate	Standard Error	t Value
Fhealth	Fwealth	phi12	0.96065	0.03051	31.49

# Covariances Among Exogenous Variables

UN Poverty Data: STA431 S 2011 Final Exam Latent regression model setting loadings to one					9			
			The 5	Endogenous	Variables	3		
M L	lanifest .atent	life: Fhea	x i lth	nfmort į	gnp1000	birthrate		
			The 6	Exogenous V	Variables			
M L E	lanifest .atent Crror	Fwea	lth lon e	1 .	e2	e3	e4	
UN Poverty Data: STA431 S 2011 Final Exam 14 Latent regression model setting loadings to one								
	Optimization Start							
Active Constraints0Objective Function0.0176241672Max Abs Gradient Element0.0096430636Radius1								
						Max Abs		Actual Over
τ.	Rest	Func	Act	Objective	e ObjFur	n Gradient	Iomhdo	Pred

0.01216 0.00547 0.00296

0.01207 0.000086 0.000199

0.01207 1.064E-6 0.000012

0.01207 1.596E-8 3.008E-6

continued	on	page	16
continucu	010	puge	10

1.097

1.019

1.018

1.017

## Optimization Results

Iterations	4	Function Calls	6
Jacobian Calls	5	Active Constraints	0
Objective Function	0.0120694315	Max Abs Gradient Element	3.0080773E-6
Lambda	0	Actual Over Pred Change	1.0166650766
Radius	0.0010397183	-	
ABSGCONV convergence c	riterion satisfie	d.	
UN 1	Poverty Data: STA	431 S 2011 Final Exam	15
Latent	regression model	setting loadings to one	
	The CALIS	Procedure	
Covariance St	ructure Analysis:	Maximum Likelihood Estim	ation
Fit Function		0	.0121
Goodness of F:	it Index (GFI)	0	.9940
GFI Adjusted :	for Degrees of Fr	eedom (AGFI) 0	.9404
Root Mean Squa	are Residual (RMR	.) 2	.5734
Standardized 1	Root Mean Square	Residual (SRMR) 0	.0067
Parsimonious (	GFI (Mulaik, 1989	) 0	.1657
Chi-Square		1	.0862
Chi-Square DF			1
Pr > Chi-Squar	re	0	.2973
Independence l	Model Chi-Square	3	99.83
Independence 1	Model Chi-Square	DF	6

UN Poverty Data: STA431 S 2011 Final Exam Latent regression model setting loadings to one 16

The CALIS Procedure Covariance Structure Analysis: Maximum Likelihood Estimation

Manifest Variable Equations with Estimates

lifex = 1.0000 Fhealth + 1.0000 e1
infmort = -4.2928\*Fhealth + 1.0000 e2
Std Err 0.1650 lambda2
t Value -26.0109
gnp1000 = 1.0000 Fwealth + 1.0000 e3
birthrate = -2.3190\*Fwealth + 1.0000 e4
Std Err 0.2961 lambda4
t Value -7.8315

UN Poverty Data: STA431 S 2011 Final Exam Latent regression model setting loadings to one

The CALIS Procedure Covariance Structure Analysis: Maximum Likelihood Estimation

Latent Variable Equations with Estimates

Fhealth = 1.8064\*Fwealth + 1.0000 epsilon Std Err 0.2317 gamma t Value 7.7958

# Variances of Exogenous Variables

			Standard	
Variable	Parameter	Estimate	Error	t Value
Fwealth	phi	30.07656	8.40592	3.58
epsilon	psi	8.20441	6.16622	1.33
e1	omega1	1.44764	1.84498	0.78
e2	omega2	184.16198	43.51832	4.23
e3	omega3	35.43109	5.57604	6.35
e4	omega4	25.92052	10.36055	2.50

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