

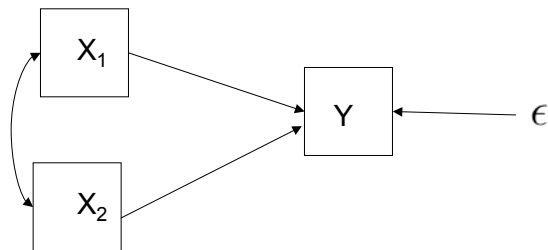
# Structural equation models are extensions of multiple regression

## Introduction to Structural Equation Models

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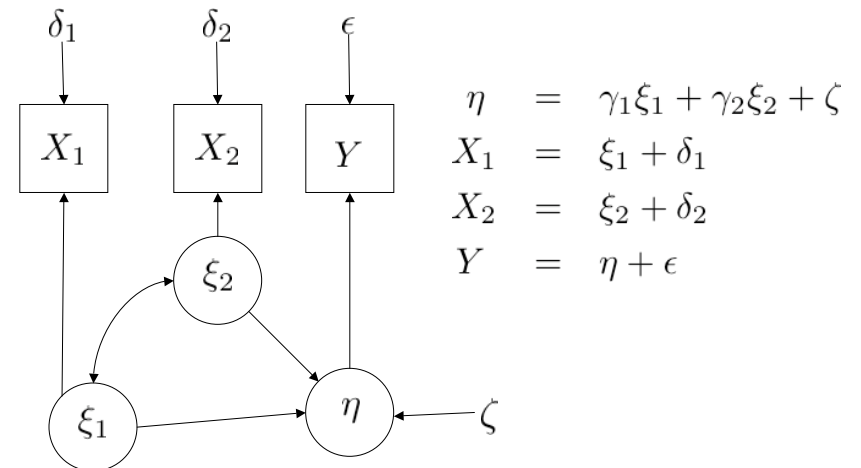
- Simultaneous equations
- Include latent variables
- DV in one equation can be IV in another

### Multiple Regression

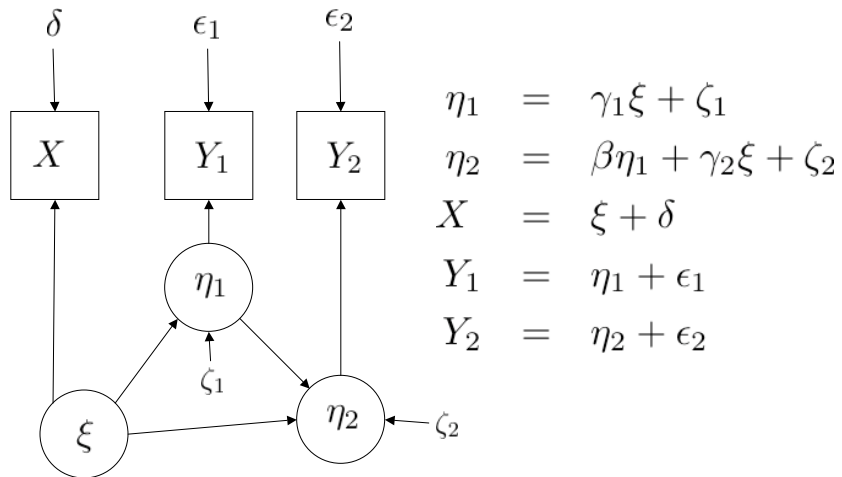


$$Y = \beta_1 X_1 + \beta_2 X_2 + \epsilon$$

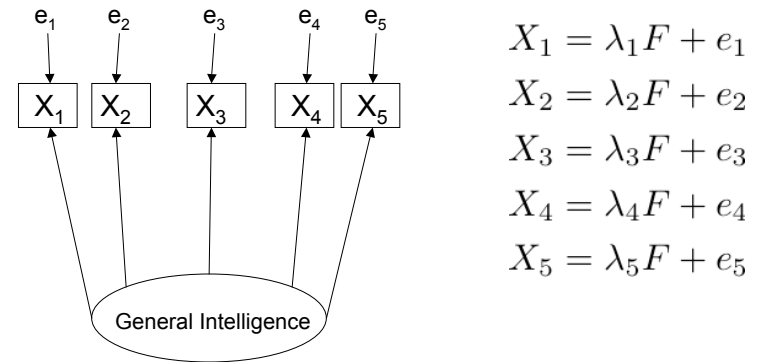
### An Extension



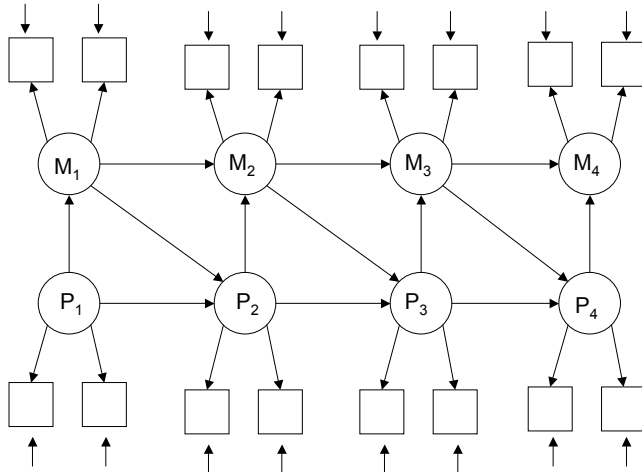
### A Further Extension



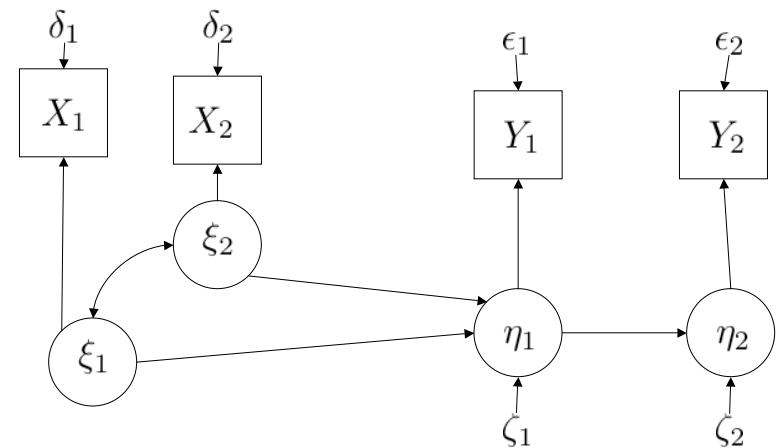
### A Factor Analysis Model



### A Longitudinal Model



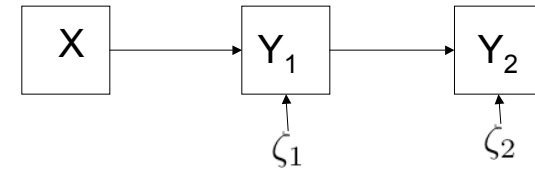
### Vocabulary: Latent v.s. Manifest, Endogenous v.s. Exogenous



## Notation

- LISREL (Bollen, 1989)
- Classical Factor Analysis (Lawley, 1971)
- LINEQS (Bentler and Weeks, 1980)
- RAM (McArdale, 1980)
- COSAN (McDonald 1978, 1980)

## Estimation and Testing



$$\begin{aligned}
 Y_1 &= \gamma X + \zeta_1 & E(X) &= E(\zeta_1) = E(\zeta_2) = 0 \\
 Y_2 &= \beta Y_1 + \zeta_2 & V(X) &= \phi, V(\zeta_1) = \psi_1, V(\zeta_2) = \psi_2 \\
 & & X, \zeta_1, \zeta_2 & \text{are independent} \\
 & & & \text{Everything is normal}
 \end{aligned}$$

## Distribution of the data

$$\begin{bmatrix} X_1 \\ Y_{1,1} \\ Y_{1,2} \end{bmatrix} \dots \begin{bmatrix} X_n \\ Y_{n,1} \\ Y_{n,2} \end{bmatrix} \text{ are independent normal with mean zero}$$

and covariance matrix

$$\Sigma = \begin{bmatrix} \phi & \gamma\phi & \beta\gamma\phi \\ \gamma\phi & \gamma^2\phi + \psi_1 & \beta(\gamma^2\phi + \psi_1) \\ \beta\gamma\phi & \beta(\gamma^2\phi + \psi_1) & \beta^2(\gamma^2\phi + \psi_1) + \psi_2 \end{bmatrix}$$

$$\boldsymbol{\theta} = (\gamma, \beta, \phi, \psi_1, \psi_2)$$

## Maximum Likelihood

$$L(\boldsymbol{\mu}, \Sigma) = |\Sigma|^{-n/2} (2\pi)^{-nk/2} \exp -\frac{n}{2} \left[ \text{tr}(\widehat{\Sigma}\Sigma^{-1}) + (\bar{\mathbf{x}} - \boldsymbol{\mu})' \Sigma^{-1} (\bar{\mathbf{x}} - \boldsymbol{\mu}) \right]$$

Minimize  $-2\ell(\boldsymbol{\mu}(\boldsymbol{\theta}), \Sigma(\boldsymbol{\theta}))$

$$\begin{aligned}
 &= n \left[ \log |\Sigma(\boldsymbol{\theta})| + k \log(2\pi) + \text{tr}(\widehat{\Sigma}\Sigma(\boldsymbol{\theta})^{-1}) \right. \\
 &\quad \left. + (\bar{\mathbf{x}} - \boldsymbol{\mu}(\boldsymbol{\theta}))' \Sigma(\boldsymbol{\theta})^{-1} (\bar{\mathbf{x}} - \boldsymbol{\mu}(\boldsymbol{\theta})) \right]
 \end{aligned}$$

## Likelihood Ratio Test for Goodness of Fit

$$\begin{aligned}
 G &= -2 \log \frac{L(\bar{\mathbf{X}}, \boldsymbol{\Sigma}(\hat{\boldsymbol{\theta}}))}{L(\bar{\mathbf{X}}, \hat{\boldsymbol{\Sigma}})} \\
 &= n[\log |\boldsymbol{\Sigma}(\hat{\boldsymbol{\theta}})| + k \log(2\pi) + \text{tr}(\hat{\boldsymbol{\Sigma}} \boldsymbol{\Sigma}(\hat{\boldsymbol{\theta}})^{-1})] \\
 &\quad - n[\log |\hat{\boldsymbol{\Sigma}}| + k \log(2\pi) + k] \\
 &= n[\log |\boldsymbol{\Sigma}(\hat{\boldsymbol{\theta}})| - \log |\hat{\boldsymbol{\Sigma}}| + \text{tr}(\hat{\boldsymbol{\Sigma}} \boldsymbol{\Sigma}(\hat{\boldsymbol{\theta}})^{-1}) - k]
 \end{aligned}$$

Do it all at once: Minimize

$$G(\boldsymbol{\theta}) = n[\log |\boldsymbol{\Sigma}(\boldsymbol{\theta})| - \log |\hat{\boldsymbol{\Sigma}}| + \text{tr}(\hat{\boldsymbol{\Sigma}} \boldsymbol{\Sigma}(\boldsymbol{\theta})^{-1}) - k]$$

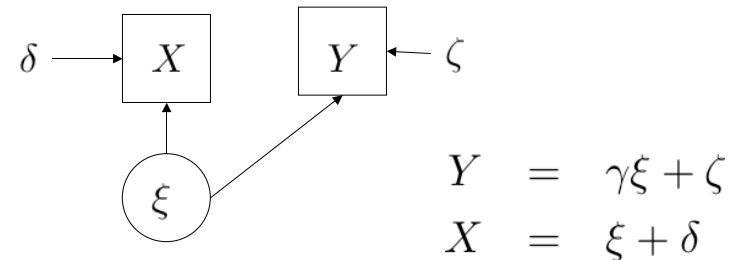
Actually, SAS minimizes the “Objective Function”

$$\log |\boldsymbol{\Sigma}(\boldsymbol{\theta})| - \log |\hat{\boldsymbol{\Sigma}}| + \text{tr}(\hat{\boldsymbol{\Sigma}} \boldsymbol{\Sigma}(\boldsymbol{\theta})^{-1}) - k$$

## Chi-square and Z Tests

- “Chisquare” is (n-1) times minimum objective function.
- Test nested models by difference between chi-square values
- Z tests are produced by default; Asymptotic Covariance matrix is available
- Likelihood ratio tests perform better

## Simple Regression with measurement Error



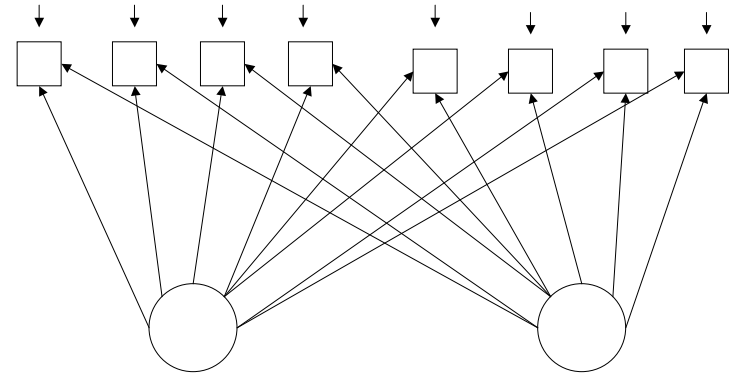
$$V(\xi) = \phi, V(\zeta) = \psi, V(\delta) = \theta_\delta$$

## The Model is not Identified

$$\theta = (\gamma, \phi, \theta_\delta, \psi)$$

$$\Sigma = \begin{bmatrix} \sigma_{1,1} & \sigma_{1,2} \\ \sigma_{2,2} & \end{bmatrix} = \begin{bmatrix} \phi + \theta_\delta & \gamma\phi \\ \gamma^2\phi + \psi & \end{bmatrix}$$

## Unconstrained (Exploratory) Factor Analysis



## Model Identification

$$\theta \in \Theta$$

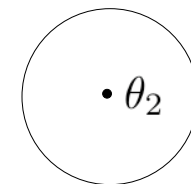
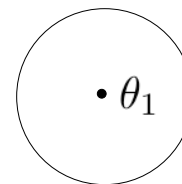
$$\mathbf{D} \sim P_\theta = g(\theta) \in \mathcal{P}$$

$$g : \Theta \rightarrow \mathcal{P}$$

If the function  $g$  is one to one,  
then the model is identified.

## Consistent Estimation is Impossible

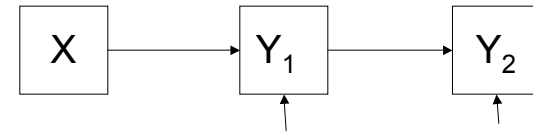
Suppose  $\theta_1 \neq \theta_2$  with  $P_{\theta_1} = P_{\theta_2}$



## To prove model identification

- Show that the parameter can be recovered from the distribution of the observed data.
- In practice, recover it from the moments (usually, the covariance matrix).
- Sometimes, only a function of the parameter is identified.

## Remember the example



$$\theta = (\gamma, \beta, \phi, \psi_1, \psi_2)$$

$$\Sigma = \begin{bmatrix} \phi & \gamma\phi & \beta\gamma\phi \\ \gamma\phi & \gamma^2\phi + \psi_1 & \beta(\gamma^2\phi + \psi_1) \\ \beta\gamma\phi & \beta(\gamma^2\phi + \psi_1) & \beta^2(\gamma^2\phi + \psi_1) + \psi_2 \end{bmatrix}$$

## Solve the identifying equations

$$\Sigma = \begin{bmatrix} \sigma_{1,1} & \sigma_{1,2} & \sigma_{1,3} \\ & \sigma_{2,2} & \sigma_{2,3} \\ & & \sigma_{3,3} \end{bmatrix} = \begin{bmatrix} \phi & \gamma\phi & \beta\gamma\phi \\ & \gamma^2\phi + \psi_1 & \beta(\gamma^2\phi + \psi_1) \\ & & \beta^2(\gamma^2\phi + \psi_1) + \psi_2 \end{bmatrix}$$

For  $\theta = (\gamma, \beta, \phi, \psi_1, \psi_2)$

## An Identified model can be

- Just-Identified (saturated): Same number of parameters and identifying equations
- Over-identified: More equations than unknowns

## Identification Rules

- A necessary condition for all Models
- Sufficient conditions for models with just observed variables
- Sufficient conditions for measurement models (factor analysis)
- Sufficient conditions for combined models

## Parameter Count

- The number of parameters must be no more than the number of unique elements in the covariance matrix of the observed variables.
- Necessary for all models

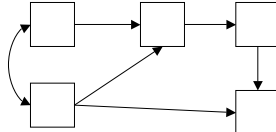
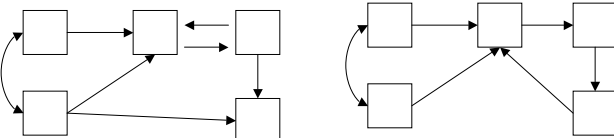
## Observed variable models

$$\mathbf{Y} = \beta\mathbf{Y} + \mathbf{\Gamma}\mathbf{X} + \zeta$$

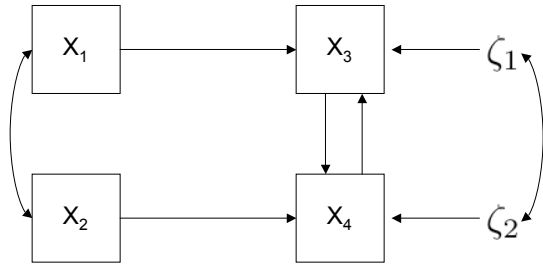
Identified if  $Cov(\mathbf{X}, \zeta) = \mathbf{0}$  and

- $\beta = \mathbf{0}$  (Regression model), or
- Model is Recursive, and  $Var(\zeta)$  is diagonal

## Recursive means no Feedback Loops

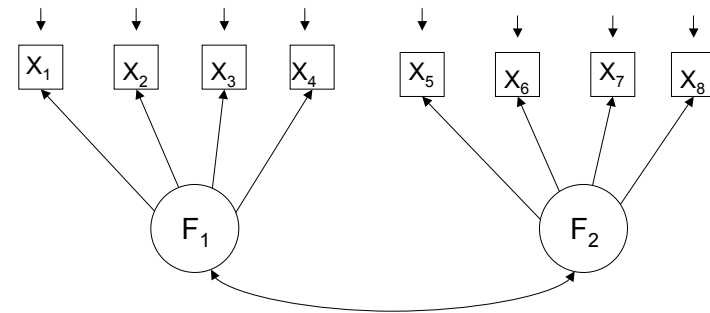
- Like this 
- Not this 

## A just-identified, Non-recursive Model



## Measurement Models (Confirmatory Factor Analysis)

$$\mathbf{X} = \mathbf{\Lambda F} + \mathbf{e}$$



## Rules for Confirmatory Factor Analysis

- Three-indicator rules
- Two-indicator rules

## “Pick a Scale” for each factor

- F is obesity
- $X_1$  is percent body fat from immersion
- $X_2$  is triceps skin fold
- $X_3$  is Body Mass Index

$$\begin{array}{l} X_1 = \lambda_1 F + e_1 \\ X_2 = \lambda_2 F + e_2 \\ X_3 = \lambda_3 F + e_3 \end{array} \longrightarrow \begin{array}{l} X_1 = F + e_1 \\ X_2 = \lambda_2 F + e_2 \\ X_3 = \lambda_3 F + e_3 \end{array}$$



## Three-indicator rules

- At least three variables per factor
- Each variable caused by only one factor
- Errors uncorrelated with factors and with each other
- Pick a scale for each factor
- No restrictions on  $\text{Var}(\mathbf{F})$

## Two-Indicator Rules

- At least two variables per factor
- Each variable caused by only one factor
- Errors uncorrelated with factors and with each other
- Pick a scale for each factor
- At least one non-zero off-diagonal element in each row of  $\text{Var}(\mathbf{F})$

## Combined Models

- Consider the latent part of the model as a model for observed variables. Verify identification.
- Consider the measurement part of the model as a confirmatory factor analysis, ignoring structure in  $\text{Var}(\mathbf{F})$ . Verify identification.

## Fixing up non-identified models

- Negotiation
- Deeper study may be rewarded
- “Model” is not identified.
- Consider identification before collecting data!

## Further Issues

- Normality
- Numerical problems
- Sample size
- Categorical data

## Software

- LISREL, EQS, RAM
- Mplus
- Stata
- R
- AMOS (Graphical interface)
- SAS proc calis