

Family (Last) Name _____

Given (First) Name Jerry

Student Number _____

STA 302s13 Quiz 6A

For reference, the general linear model with normal error terms is $\mathbf{Y} = \mathbf{X}\beta + \epsilon$, the columns of \mathbf{X} are linearly independent, and $\epsilon \sim N_n(\mathbf{0}, \sigma^2 \mathbf{I}_n)$. In this quiz, you may use the following facts without proving them here.

1. $\mathbf{Y} \sim N_n(\mathbf{X}\beta, \sigma^2 \mathbf{I}_n)$
2. $\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} \sim N_{k+1}(\beta, \sigma^2(\mathbf{X}'\mathbf{X})^{-1})$
3. $\hat{\beta}$ and $SSE = (\mathbf{Y} - \mathbf{X}\hat{\beta})'(\mathbf{Y} - \mathbf{X}\hat{\beta})$ are independent.
4. If $\mathbf{T} \sim N_p(\mu, \Sigma)$ where the covariance matrix Σ is strictly positive definite, $(\mathbf{T} - \mu)' \Sigma^{-1} (\mathbf{T} - \mu) \sim \chi^2(p)$.
5. Let $W = W_1 + W_2$, where W_1 and W_2 are independent, $W_2 \sim \chi^2(\nu_2)$ and $W \sim \chi^2(\nu_1 + \nu_2)$, where ν_1 and ν_2 are both positive. Then $W_1 \sim \chi^2(\nu_1)$.
6. $(\mathbf{Y} - \mathbf{X}\beta)'(\mathbf{Y} - \mathbf{X}\beta) = (\mathbf{Y} - \mathbf{X}\hat{\beta})'(\mathbf{Y} - \mathbf{X}\hat{\beta}) + (\hat{\beta} - \beta)'(\mathbf{X}'\mathbf{X})(\hat{\beta} - \beta)$

Show $SSE/\sigma^2 \sim \chi^2(n - k - 1)$. Refer to the facts above by number when you use them.

- By (1) and (4), $\frac{1}{\sigma^2} (\mathbf{Y} - \mathbf{X}\beta)'(\mathbf{Y} - \mathbf{X}\beta) = (\mathbf{Y} - \mathbf{X}\beta)'(\sigma^{-2}\mathbf{I}_n)(\mathbf{Y} - \mathbf{X}\beta) \sim \chi^2(n)$
- By (2) and (4), $\frac{1}{\sigma^2} (\hat{\beta} - \beta)'(\mathbf{X}'\mathbf{X})(\hat{\beta} - \beta) = (\hat{\beta} - \beta)'(\sigma^{-2}(\mathbf{X}'\mathbf{X})^{-1})^{-1}(\hat{\beta} - \beta) \sim \chi^2(k+1)$

Let $W = \frac{1}{\sigma^2} (\mathbf{Y} - \mathbf{X}\beta)'(\mathbf{Y} - \mathbf{X}\beta)$
 $W_1 = \frac{1}{\sigma^2} (\mathbf{Y} - \mathbf{X}\hat{\beta})'(\mathbf{Y} - \mathbf{X}\hat{\beta})$
 $W_2 = \frac{1}{\sigma^2} (\hat{\beta} - \beta)'(\mathbf{X}'\mathbf{X})(\hat{\beta} - \beta)$

- By (6), $W = W_1 + W_2$
- By (3), W_1 & W_2 are independent
- By (5), $W_1 = \frac{SSE}{\sigma^2}$ has a chi-squared distribution, with df $n - (k+1) = n - k - 1$

[Signature]

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STA 302s13 Quiz 6B

For reference, the general linear model with normal error terms is $\mathbf{Y} = \mathbf{X}\beta + \epsilon$, the columns of \mathbf{X} are linearly independent, and $\epsilon \sim N_n(\mathbf{0}, \sigma^2 \mathbf{I}_n)$. Recall that

- $\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$
 - $\hat{\mathbf{Y}} = \mathbf{X}\hat{\beta}$
 - $\hat{\epsilon} = \mathbf{Y} - \hat{\mathbf{Y}}$
- $\sum_{i=1}^n \hat{y}_i = \sum_{i=1}^n y_i$

Letting $SST = \sum_{i=1}^n (Y_i - \bar{Y})^2$, $SSE = \sum_{i=1}^n (Y_i - \hat{Y}_i)^2$ and $SSR = \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2$, show $SST = SSR + SSE$.

$$\begin{aligned} SST &= \sum_{i=1}^n (Y_i - \bar{Y})^2 = \sum_{i=1}^n (Y_i - \hat{Y}_i + \hat{Y}_i - \bar{Y})^2 \\ &= \sum_{i=1}^n \left((Y_i - \hat{Y}_i)^2 + 2(Y_i - \hat{Y}_i)(\hat{Y}_i - \bar{Y}) + (\hat{Y}_i - \bar{Y})^2 \right) \\ &= \sum_{i=1}^n (Y_i - \hat{Y}_i)^2 + 2 \sum_{i=1}^n (Y_i - \hat{Y}_i)(\hat{Y}_i - \bar{Y}) + \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2 \\ &= SSE + 2 \sum_{i=1}^n (Y_i - \hat{Y}_i)(\hat{Y}_i - \bar{Y}) + SSR \end{aligned}$$

$$\begin{aligned} \text{Now, } \sum_{i=1}^n (Y_i - \hat{Y}_i)(\hat{Y}_i - \bar{Y}) &= \sum_{i=1}^n (Y_i - \hat{Y}_i)\hat{Y}_i - \bar{Y} \sum_{i=1}^n (Y_i - \hat{Y}_i) \\ &= \sum_{i=1}^n (Y_i - \hat{Y}_i)\hat{Y}_i - 0 = (\mathbf{Y} - \hat{\mathbf{Y}})' \hat{\mathbf{Y}} = \mathbf{Y}'\hat{\mathbf{Y}} - \hat{\mathbf{Y}}'\hat{\mathbf{Y}} \\ &= \mathbf{Y}'\hat{\mathbf{X}}\hat{\beta} - (\hat{\mathbf{X}}\hat{\beta})'\hat{\mathbf{X}}\hat{\beta} \\ &= \mathbf{Y}'\hat{\mathbf{X}}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} - (\hat{\mathbf{X}}\hat{\beta})'\hat{\mathbf{X}}\hat{\beta} \\ &= \mathbf{Y}'\hat{\mathbf{X}}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} - \mathbf{Y}'\hat{\mathbf{X}}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} \\ &= \mathbf{Y}'\hat{\mathbf{X}}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} - \mathbf{Y}'\hat{\mathbf{X}}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} = 0 \end{aligned}$$

$$\text{SO } SST = SSE + SSR$$

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