

Old Test Questions, mostly from STA257

1. Let X be a *continuous* random variable and let a and b be constants. Prove $\text{Var}[aX + b] = a^2\text{Var}(X)$. You may use the theorem $\text{Var}(Y) = E[Y^2] - (E[Y])^2$ if you wish. If you know a result about $E[aX + b]$, you may use it without proof.
2. Let X and Y be continuous random variables that are *independent*. Prove that $E[XY] = E[X]E[Y]$. Be very clear about where you are using the assumption of independence.
3. Let X have a binomial distribution with parameters n and θ ; that is, $f_X(x) = \binom{n}{x}\theta^x(1-\theta)^{n-x}I(x=0, \dots, n)$.
 - a. (10 points) Derive the moment-generating function $M_X(t)$.
 - b. (10 points) Use the preceding answer to find $E[X]$.
4. Derive the moment-generating function of a random variable with a Gamma density (see formula sheet) with parameters $\alpha > 0$ and $\beta > 0$. **Don't** bother to specify the range of values for which the function exists. *Note that on this question, if you make crazy mistakes to force your answer to match $M(t)$ on the formula sheet, you will get a zero.*
5. Chebyshev's inequality states that for any random variable X with $E(X) = \mu$ and $\text{Var}(X) = \sigma^2$ and for any $k > 0$, $P(|X - \mu| \geq k\sigma) \leq \frac{1}{k^2}$. Use this result to prove the following. Let X_1, \dots, X_n be a random sample from a population with expected value μ and variance σ^2 . Then for all $\epsilon > 0$,
$$\lim_{n \rightarrow \infty} P(|\bar{X}_n - \mu| \geq \epsilon) = 0.$$
6. The random variable X has density $f_X(x) = 4e^{x-4e^x}I(-\infty < x < \infty)$. (The indicator is not really necessary but it may be helpful to you.) Find the density of $Y = e^X$. Make sure $f_Y(y)$ is correct for all real y . **Circle your answer.**
7. Let X_1, \dots, X_n be a random sample from a *normal* population with $\mu = 0$ and $\sigma^2 = 1$, and let $Y = \sum_{i=1}^n X_i^2$. Find $f_Y(y)$. Make sure it is correct for all real y . **Circle your answer.**
8. The joint density of X_1 and X_2 is $f_{X_1, X_2}(x_1, x_2) = e^{-x_1-x_2}I(x_1 > 0)I(x_2 > 0)$. Find the density of $Y = X_1 + X_2$ any way you wish (more than one way will work). Make sure $f_Y(y)$ is correct for all real y . **Circle your answer.**

9. Let X_1, \dots, X_n be a random sample from a distribution for which the moment-generating function exists. Show $\bar{X}_n \xrightarrow{d} Y$, where Y is a “degenerate” random variable with $Pr\{Y = \mu\} = 1$. Show your work.

Even Older Problems, mostly STA257

1. Let X have a geometric distribution; that is, $f(x) = \theta (1-\theta)^{x-1} I_{\{x = 1, 2, \dots\}}$. Find the moment-generating function $M_X(t)$. You do *NOT* have to say anything about the values of t for which this function exists.
2. Let X and Y be discrete random variables. Starting from an expression for $E[g(X,Y)]$, prove that $E[X+Y] = E[X] + E[Y]$. **Do NOT assume X and Y are independent!** If you assume independence you will get zero marks on this question.
3. Let X have a binomial distribution with parameters n and θ ; that is, $X \sim B(n,\theta)$. Starting with the definition of a moment-generating function, find $M_X(t)$; show all your work and **circle your answer**.
4. Let X have a normal distribution with mean μ and variance σ^2 ; that is, $X \sim N(\mu,\sigma^2)$. Let $Y = e^X$. (This means $X = \ln(Y)$, so Y has a log-normal distribution) Use the distribution function technique to find the density $f_Y(y)$. Don't forget the support! **Circle your answer**.
5. Let X_1, \dots, X_n be independent random variables with moment-generating functions $M_{X_i}(t)$, $i = 1, \dots, n$. Let $Y = \sum_{i=1}^n X_i$. Starting from the definition of a moment-generating function and then using a convenient expression for $E[g(X_1, \dots, X_n)]$, find $M_Y(t)$. Assume X_1, \dots, X_n are continuous, so you'll integrate. Show all your work.
6. Let X_1, \dots, X_n be independent Poisson random variables, all with the same parameter $\lambda > 0$. Let $Y = \sum_{i=1}^n X_i$. Give the probability distribution $f_Y(y)$. Show your work. Remember, the support counts for half marks. You have more room than you need. **Circle your answer**.
7. Let the continuous random variable X have density $f_X(x) = \frac{1}{(n-1)!} e^{-x} x^{n-1} I_{\{x>0\}}$. Let $Y = \ln(X)$. **Find the density $f_Y(y)$** . Clearly indicate where this density is greater than zero or your answer is wrong.
8. Let X have a Poisson distribution with parameter $\lambda > 0$. That is, $f_X(x) = \frac{e^{-\lambda} \lambda^x}{x!} I_{\{x=0,1, \dots\}}$. Find the moment-generating function $M_X(t)$.

9. Let the random variable X have an exponential distribution with mean θ . Let $Y = \frac{2X}{\theta}$. Derive the probability density function $f_Y(y)$. Show your work.

10. Let X_1, \dots, X_n be a random sample from a Normal $(0, \sigma^2)$ distribution.

a) Let $Y_i = \frac{X_i^2}{\sigma^2}$. Find the density of Y_i . You may use the symmetry of the normal distribution if necessary, but don't use any theorems about the normal distribution. Derive the result or get no marks. **Circle your answer.**

b) Let $W = \sum_{i=1}^n Y_i = \sum_{i=1}^n \frac{X_i^2}{\sigma^2}$. Find the distribution of W . (That is, the distribution of W has a name. Name the distribution and give the value of the parameter.) Show your work.

11. Let X have a Gamma distribution with parameters α and β , i.e. $f_X(x) = \frac{1}{\beta^\alpha \Gamma(\alpha)} e^{-x/\beta} x^{\alpha-1}$ for $x > 0$ and zero otherwise. Find the density of $Y = 1/X$.

Be sure to indicate the support of Y !

12. Let X_1, X_2, X_3 be a random sample from a distribution that is $N(6, 4)$.

Let Y be the largest sample value. Find $P(Y < 8)$.

a) .405

b) .595 *

13. Let $X_1 \sim N(\mu_1, \sigma_1^2)$ and $X_2 \sim N(\mu_2, \sigma_2^2)$ be independent. What is the variance of $Y = X_1 - X_2$?

14. Let X_1, \dots, X_n be a random sample from a $N(\mu, \sigma^2)$ population. What is the distribution of \bar{X} ?

15. Let X_1, \dots, X_n be a random sample from a Poisson(μ) population. What is

the distribution of $Y = \sum_{i=1}^n X_i$?

16. Let $F_n(x) = \frac{x}{n} I_{\{0 < x < 1\}} - \frac{n-1}{n} e^{-x} I_{\{x > 0\}} + \frac{1}{n} I_{\{x \geq 1\}} + \frac{n-1}{n} I_{\{x > 0\}}$. The

limiting distribution of F_n is

a) Exponential with $\theta=1$ *

17. Let X_1, \dots, X_n be a random sample from a $N(0, \sigma^2)$ population. What is the limiting distribution of \bar{X}_n ? Hint: Disregard the behavior of the limiting distribution at any points of discontinuity.

a) Exponential with $\theta=1$

b) Standard Normal

c) Degenerate at zero *

18. Let X_1, \dots, X_n be a random sample from a $\chi^2(1)$ distribution. Show

that the distribution of $Y = \sum_{i=1}^n X_i$ is also chi-squared, and find its

parameter (degrees of freedom).

19. Let $X \sim N(\mu, \sigma^2)$. Show that $Z = \frac{X-\mu}{\sigma} \sim N(0,1)$