

Weibull Regression¹

STA312 Fall 2023

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Background Reading

Section 10.6 in the text, but it refers to a lot of things we have not covered yet.

Overview

① Exponential

② Weibull

A multiplicative regression model

Exponential model, just one explanatory variable

Independently for $i = 1, \dots, n$,

$$t_i^* = e^{\beta_0 + \beta_1 x_i} \times \epsilon_i$$

where

t_i^* are failure times, possibly censored.

β_0 and β_1 are unknown constants (parameters).

x_1, \dots, x_n are known, observed constants.

$\epsilon_1, \dots, \epsilon_n$ are independent Exponential(1) random variables.

t_1, \dots, t_n are observed event times, either failure or censoring.

$\delta_1, \dots, \delta_n$ are observed indicators for uncensored.

- These are sometimes called *accelerated failure time* models.
- Because the effect of $x_i \neq 0$ is to *multiply* the failure time by a constant.

Distribution of $t_i^* = e^{\beta_0 + \beta_1 x_i} \times \epsilon_i$, with $\epsilon_i \sim \exp(1)$

- If $\epsilon \sim \exp(1)$ and $a > 0$, $x = a\epsilon$ is also exponential.
- Expected value a (or $\lambda = 1/a$).
- Thus, $E(t_i^*) = e^{\beta_0 + \beta_1 x_i} \Leftrightarrow \log E(t_i^*) = \beta_0 + \beta_1 x_i$.
- We are adopting a linear model for the log of the expected value.
- Or, we can transform the failure times by taking logs.

$$\begin{aligned}\log t_i^* &= \beta_0 + \beta_1 x_i + \log \epsilon_i \\ &= \beta_0 + \beta_1 x_i + \epsilon_i^*\end{aligned}$$

where $\epsilon_i^* = \log \epsilon_i \sim G(0, 1)$.

Meaning of β_1

With $E(t_i^*) = e^{\beta_0 + \beta_1 x_i}$

- Increase x_i by one unit.
- The effect is to multiply $E(t_i^*)$ by a constant.

$$\begin{aligned}
 e^{\beta_0 + \beta_1(x_i+1)} &= c e^{\beta_0 + \beta_1 x_i} \\
 \Leftrightarrow c &= \frac{e^{\beta_0 + \beta_1(x_i+1)}}{e^{\beta_0 + \beta_1 x_i}} \\
 &= \frac{e^{\beta_0 + \beta_1 x_i + \beta_1}}{e^{\beta_0 + \beta_1 x_i}} \\
 &= e^{\beta_1}
 \end{aligned}$$

- So when x_i is increased by one unit, $E(t_i^*)$ is multiplied by e^{β_1} .
- If $\beta_1 > 0$, $E(t_i^*)$ goes up.
- If $\beta_1 < 0$, $E(t_i^*)$ goes down.

Natural extensions

- More than one explanatory variable.
- Centering the quantitative explanatory variables.

$$t_i^* = \exp\{\beta_0 + \beta_1(x_{i,1} - \bar{x}_1) + \dots + \beta_{p-1}(x_{i,p-1} - \bar{x}_{p-1})\} \cdot \epsilon_i$$

- In this case, e^{β_0} is the expected failure time for average values of all the explanatory variables.
- If there are dummy variables, center only the quantitative variables (covariates).

Equivalent model on the log scale

Starting with $t_i^* = \exp\{\beta_0 + \beta_1 x_{i,1} + \dots + \beta_{p-1} x_{i,p-1}\} \cdot \epsilon_i$

$$\begin{aligned} \log t_i^* &= \beta_0 + \beta_1 x_{i,1} + \dots + \beta_{p-1} x_{i,p-1} + \log \epsilon_i \\ &= \beta_0 + \beta_1 x_{i,1} + \dots + \beta_{p-1} x_{i,p-1} + \epsilon_i^* \\ &= \mathbf{x}_i^\top \boldsymbol{\beta} + \epsilon_i^*, \end{aligned}$$

where $\epsilon_i^* \sim G(0, 1)$.

- Recall, if $Z \sim G(0, 1)$, then $\sigma Z + \mu \sim G(\mu, \sigma)$.
- So the model says $\log t_i^* \sim G(\mathbf{x}_i^\top \boldsymbol{\beta}, 1)$
- Why should the variance of log survival time be $\frac{\pi^2}{6}$?
- Much more reasonable is

$$\log t_i^* = \beta_0 + \beta_1 x_{i,1} + \dots + \beta_{p-1} x_{i,p-1} + \sigma \epsilon_i^*$$
- In this case, $\log t_i^* \sim G(\mathbf{x}_i^\top \boldsymbol{\beta}, \sigma)$.

Switching back to the time scale

From the log time scale

$$\begin{aligned}\log t_i^* &= \beta_0 + \beta_1 x_{i,1} + \dots + \beta_{p-1} x_{i,p-1} + \sigma \epsilon_i^* \\ \Leftrightarrow t_i^* &= e^{\mathbf{x}_i^\top \boldsymbol{\beta}} e^{\sigma \epsilon_i^*} = e^{\mathbf{x}_i^\top \boldsymbol{\beta}} e^{\sigma \log \epsilon_i} = e^{\mathbf{x}_i^\top \boldsymbol{\beta}} e^{\log(\epsilon_i^\sigma)} \\ \Leftrightarrow t_i^* &= e^{\mathbf{x}_i^\top \boldsymbol{\beta}} \epsilon_i^\sigma\end{aligned}$$

We have arrived at the multiplicative regression model:

$$t_i^* = \exp\{\beta_0 + \beta_1 x_{i,1} + \dots + \beta_{p-1} x_{i,p-1}\} \cdot \epsilon_i^\sigma$$

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- It's an accelerated failure time model. Changing one of the x values multiplies t_i^* by something.
- In particular, if $x_{i,k}$ is increased by one unit while holding all other $x_{i,j}$ values constant,
- Then t_i^* is multiplied by e^{β_k} .
- Holding $x_{i,j}$ values constant is the meaning of “controlling” for explanatory variables in Weibull regression.
- Note that if β_k is negative, $e^{\beta_k} < 1$ and t_i^* goes down.
- Call it a “negative relationship” (controlling for the other variables).
- If β_k is positive, $e^{\beta_k} > 1$ and t_i^* goes up.
- Call this a “positive relationship” (controlling for the other variables).

Distribution of t_i^*

Recall

- We have established that $\log t_i^* \sim G(\mathbf{x}_i^\top \boldsymbol{\beta}, \sigma)$.
- Exponential function of Gumbel(μ, σ) is Weibull(α, λ) with $\lambda = e^{-\mu}$ and $\alpha = 1/\sigma$.
- Note that here, $\mu_i = \mathbf{x}_i^\top \boldsymbol{\beta}$.
- So, t_i^* is Weibull, with $\lambda_i = e^{-\mathbf{x}_i^\top \boldsymbol{\beta}}$ and $\alpha = 1/\sigma$.
- This means

$$\begin{aligned}
 E(t_i^*) &= \frac{\Gamma(1 + \frac{1}{\alpha})}{\lambda} = e^{\mathbf{x}_i^\top \boldsymbol{\beta}} \Gamma(1 + \sigma) \\
 \text{Median}(t_i^*) &= \frac{[\log(2)]^{1/\alpha}}{\lambda} = e^{\mathbf{x}_i^\top \boldsymbol{\beta}} \log(2)^\sigma \\
 h(t) &= \alpha \lambda^\alpha t^{\alpha-1} = \frac{1}{\sigma} \exp\left\{-\frac{1}{\sigma} \mathbf{x}^\top \boldsymbol{\beta}\right\} t^{\frac{1}{\sigma}-1}
 \end{aligned}$$

Conclusions

Following from $\log t_i^* \sim G(\mathbf{x}_i^\top \boldsymbol{\beta}, \sigma)$

$$\begin{aligned} E(t_i^*) &= e^{\mathbf{x}_i^\top \boldsymbol{\beta}} \Gamma(1 + \sigma) \\ \text{Median}(t_i^*) &= e^{\mathbf{x}_i^\top \boldsymbol{\beta}} \log(2)^\sigma \\ h(t) &= \frac{1}{\sigma} \exp\left\{-\frac{1}{\sigma} \mathbf{x}^\top \boldsymbol{\beta}\right\} t^{\frac{1}{\sigma}-1} \end{aligned}$$

- Increasing value of x_j by c units multiplies the mean and median by $e^{c\beta_j}$.
- Increasing value of x_j by c units multiplies the hazard function by $e^{-\frac{c}{\sigma}\beta_j}$. If β_j is positive, the hazard goes *down*.
- Remarkable because the hazard function is a function of time t .
- And the effect is *identical* for every value of t .

Proportional Hazards

$$h(t) = \frac{1}{\sigma} \exp\left\{-\frac{1}{\sigma} \mathbf{x}^\top \boldsymbol{\beta}\right\} t^{\frac{1}{\sigma}-1}$$

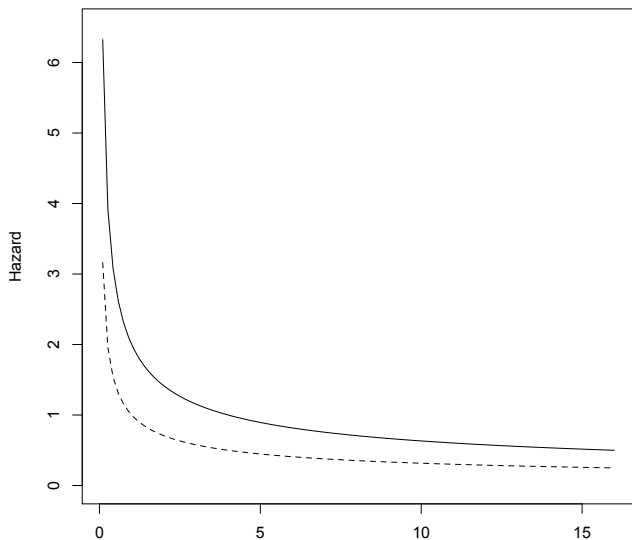
- Suppose two individuals have different \mathbf{x} vectors of explanatory variable values.
- They have different hazard functions because their $\lambda_i = e^{-\mathbf{x}_i^\top \boldsymbol{\beta}}$ values are different.
- Look at the *ratio*:

$$\begin{aligned} \frac{h_1(t)}{h_2(t)} &= \frac{\frac{1}{\sigma} \exp\left\{-\frac{1}{\sigma} \mathbf{x}_1^\top \boldsymbol{\beta}\right\} t^{\frac{1}{\sigma}-1}}{\frac{1}{\sigma} \exp\left\{-\frac{1}{\sigma} \mathbf{x}_2^\top \boldsymbol{\beta}\right\} t^{\frac{1}{\sigma}-1}} \\ &= \frac{\exp\left\{-\frac{1}{\sigma} \mathbf{x}_1^\top \boldsymbol{\beta}\right\}}{\exp\left\{-\frac{1}{\sigma} \mathbf{x}_2^\top \boldsymbol{\beta}\right\}} \\ &= \exp\left\{\frac{1}{\sigma} (\mathbf{x}_2 - \mathbf{x}_1)^\top \boldsymbol{\beta}\right\} \end{aligned}$$

The point is that $h_1(t)$ and $h_2(t)$ are *always in the same proportion for every value of t* .

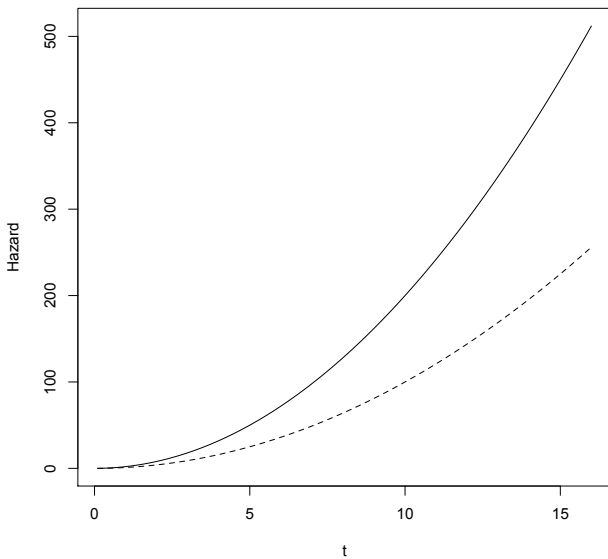
Proportional Hazards

$$h_1(t) = 2h_2(t) \text{ with } \sigma = 2$$



Proportional Hazards

$$h_1(t) = 2h_2(t) \text{ with } \sigma = 1/3$$



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