

Assignment 11

1

① (a) $h(x) = h_0(x) e^{\beta_1 a + \beta_2 d_1 + \beta_3 d_2 + \beta_4 d_3 + \beta_5 d_4}$

(b)

	d_1	d_2	d_3	d_4	$h(x)$
Obs only	1	0	0	0	$h_0(x) e^{\beta_1 a} e^{\beta_2}$
Placebo	0	0	0	0	$h_0(x) e^{\beta_1 a}$
100mg	0	1	0	0	$h_0(x) e^{\beta_1 a} e^{\beta_3}$
500mg	0	0	1	0	$h_0(x) e^{\beta_1 a} e^{\beta_4}$
2,000mg	0	0	0	1	$h_0(x) e^{\beta_1 a} e^{\beta_5}$

(c) (i) $H_0: \beta_2 = \beta_3 = \beta_4 = \beta_5 = 0$

(ii)

$$\begin{pmatrix} 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \\ \beta_4 \\ \beta_5 \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \\ 0 \\ 0 \end{pmatrix}$$

$L \quad \beta = 0$

(1d) (i) $H_0: \beta_2 = 0$

(ii)

$$(0 \ 1 \ 0 \ 0 \ 0) \begin{pmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \\ \beta_4 \\ \beta_5 \end{pmatrix} = 0$$

(e) (i) $H_0: \beta_3 = \beta_4 = \beta_5 = 0$

$$\begin{pmatrix} 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \\ \beta_4 \\ \beta_5 \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \\ 0 \end{pmatrix}$$

L

$$B = 0$$

(f) $\binom{4}{2} = 6$

(1g) (i) $H_0: \beta_5 = 0$
(ii)

$$(0 \ 0 \ 0 \ 0 \ 1) \begin{pmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \\ \beta_4 \\ \beta_5 \end{pmatrix} = 0$$

(1h) (i) $H_0: \beta_3 = \beta_4$

$$(0 \ 0 \ 1 \ -1 \ 0) \begin{pmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \\ \beta_4 \\ \beta_5 \end{pmatrix} = 0$$

$$L \quad \beta = 0$$

Q2

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[R.app GUI 1.79 (8198) x86_64-apple-darwin17.0]

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```
> # Q2: channing
> rm(list=ls()); # options(scipen=999)
> # install.packages("KMsurv",dependencies=TRUE) # Only need to do this once
> library(KMsurv); library(survival)
> data(channing) # For some reason this is necessary
>
> # We want age at entry. Age at censoring or death is the response variable
>
> retired = within(channing,
+ {
+ ageentry = ageentry/12 # Age in years
+ cageentry = (ageentry-mean(ageentry)) # Centered ageentry in years
+ gender = gender-1 # 1=F, 0=M
+ })
> summary(retired) # Q2a
```

obs	death	ageentry	age	time
Min. : 1.0	Min. :0.000	Min. :61.08	Min. : 777	Min. : 0.00
1st Qu.:116.2	1st Qu.:0.000	1st Qu.:71.17	1st Qu.: 939	1st Qu.: 35.00
Median :231.5	Median :0.000	Median :75.00	Median : 990	Median : 81.50
Mean :231.5	Mean :0.381	Mean :75.47	Mean : 986	Mean : 80.12
3rd Qu.:346.8	3rd Qu.:1.000	3rd Qu.:79.65	3rd Qu.:1031	3rd Qu.:137.00
Max. :462.0	Max. :1.000	Max. :95.00	Max. :1207	Max. :137.00

```
gender      cageentry
Min. :0.00  Min. : -14.3885
1st Qu.:1.00 1st Qu.: -4.3052
Median :1.00 Median : -0.4719
Mean :0.79  Mean :  0.0000
3rd Qu.:1.00 3rd Qu.:  4.1740
Max. :1.00  Max. : 19.5281
>
> # 2b: False. It would be a male of average age.
>
```

```
> # 2b: (i) - (vi)
> mod = coxph( Surv(time,death) ~ cageentry + gender, data=retired); summary(mod)
```

```
Call:
```

```
coxph(formula = Surv(time, death) ~ cageentry + gender, data = retired)
```

```
n= 462, number of events= 176
```

	coef	exp(coef)	se(coef)	z	Pr(> z)
cageentry	0.08549	1.08925	0.01259	6.790	1.12e-11 ***
gender	-0.37591	0.68666	0.17191	-2.187	0.0288 *

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

	exp(coef)	exp(-coef)	lower .95	upper .95
cageentry	1.0892	0.9181	1.0627	1.1165
gender	0.6867	1.4563	0.4902	0.9618

```
Concordance= 0.647 (se = 0.023 )
```

```
Likelihood ratio test= 49.48 on 2 df, p=2e-11
```

```
Wald test = 51.51 on 2 df, p=7e-12
```

```
Score (logrank) test = 52.6 on 2 df, p=4e-12
```

```
> summary(mod)
```

```
Call:
```

```
coxph(formula = Surv(time, death) ~ cageentry + gender, data = retired)
```

```
n= 462, number of events= 176
```

	coef	exp(coef)	se(coef)	z	Pr(> z)
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```
---
```

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Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
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```

```
Score (logrank) test = 52.6 on 2 df, p=4e-12
```

```
>
```

```
> # 2b(vii)
```

```
> summary(coxph( Surv(time,death) ~ ageentry + gender, data=channing) )
```

```
Call:
```

```
coxph(formula = Surv(time, death) ~ ageentry + gender, data = channing)
```

```
n= 462, number of events= 176
```

	coef	exp(coef)	se(coef)	z	Pr(> z)
ageentry	0.007124	1.007149	0.001049	6.790	1.12e-11 ***
gender	-0.375915	0.686661	0.171907	-2.187	0.0288 *

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

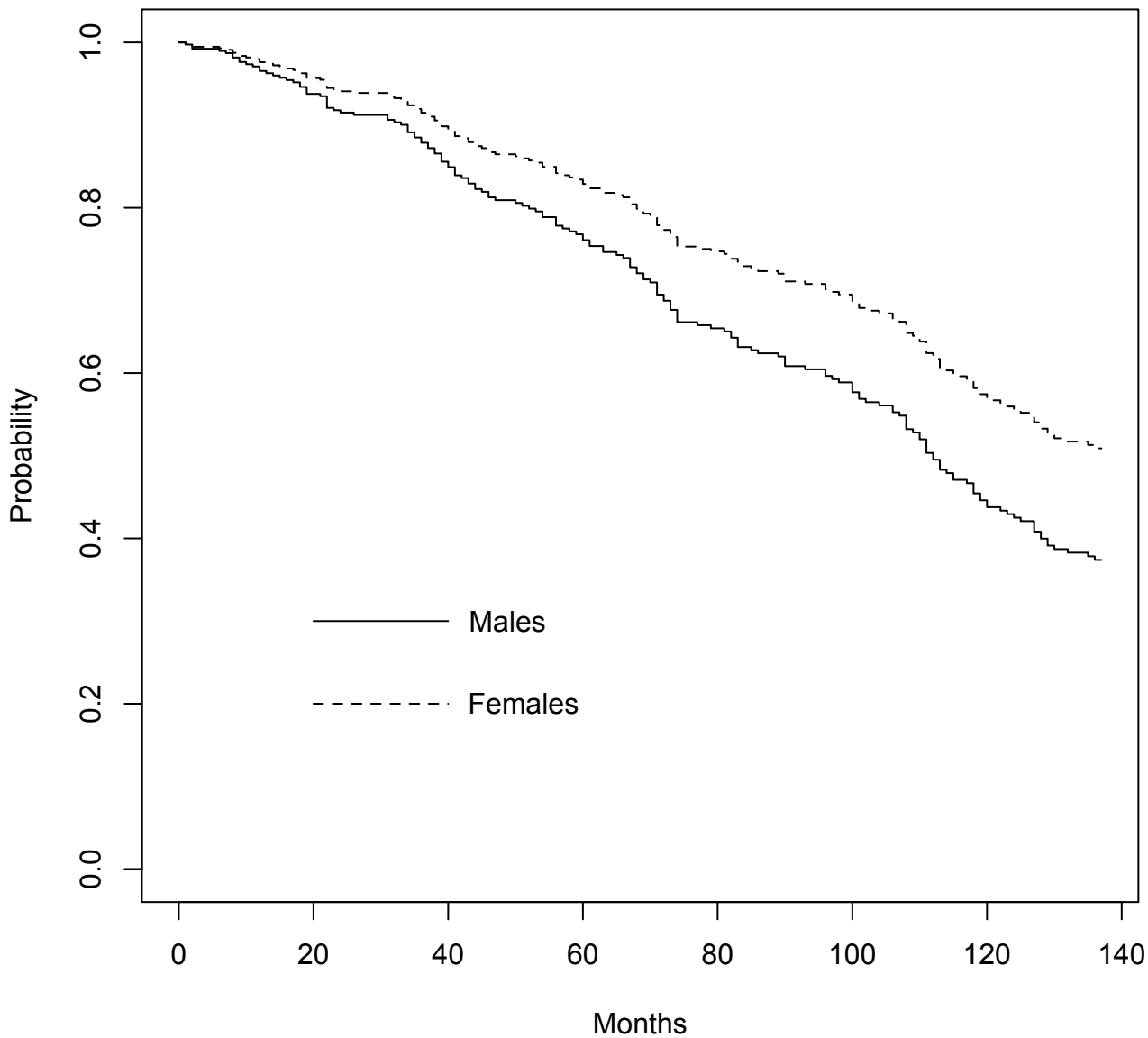
	exp(coef)	exp(-coef)	lower .95	upper .95
ageentry	1.0071	0.9929	1.0051	1.0092
gender	0.6867	1.4563	0.4902	0.9618

Concordance= 0.647 (se = 0.023)
Likelihood ratio test= 49.48 on 2 df, p=2e-11
Wald test = 51.51 on 2 df, p=7e-12
Score (logrank) test = 52.6 on 2 df, p=4e-12

```
>
>
> # 2b(viii)
> # Just like in PH2 with R
> guy = data.frame(cageentry=0, gender=0); gal = data.frame(cageentry=0, gender=1)
> sexcomp = rbind(guy,gal); rownames(sexcomp) = c("M","F"); sexcomp
  cageentry gender
M          0      0
F          0      1
> s = survfit(mod,newdata=sexcomp); s
Call: survfit(formula = mod, newdata = sexcomp)

      n events median 0.95LCL 0.95UCL
M 462    176    112     98     NA
F 462    176     NA    124     NA
>
>
> # Make a nice plot
> plot(s,lty = c(1,2),xlab="Months", ylab="Probability")
> title('Probability of Surviving More Than x Months at the Home')
> xm = c(20,40); ym = c(0.3,0.3); lines(xm,ym,lty=1)
> text(50,0.3,"Males ")
> xf = c(20,40); yf = c(0.2,0.2); lines(xf,yf,lty=2)
> text(50,0.2," Females")
>
>
```

Probability of Surviving More Than x Months at the Home



- vii. Suppose you had left gender as 1-2 and you left age in months, uncentered. Would this affect the test statistics and the conclusions? Run a quick example to find out. I used the original channing data set.
- viii. Estimate the median survival times for men of average age and women of average age. Oops! I can't get an estimate for women.
- ix. Make a well-labelled plot showing the estimated survival functions for a man and a woman of average age. There should be two different line types, like maybe solid and dashed. Print the graph and bring it to the quiz. From the picture, can you see why there was no estimated median for women?
- x. The following is just a comment, not a question. A crucial assumption of the censored data model is that the censoring mechanism is independent of the outcome. However, lots of nursing home residents leave the home to go to hospital, and they die there rather than at the home. In fact, they are censored *because* the staff thought they were probably about to die. To me, this helps explain the high estimated probabilities of lasting a long time that we see in these data.

3. It is very natural and tempting to use age as a time-varying covariate. This question tries to illustrate why it won't work. Consider a model with an indicator for gender, and also age. Age = $x + t$, where x is age at entry to the study. The partial likelihood can be written

$$\begin{aligned}
 \text{PL}(\beta) &= \prod_{i=1}^D \left(\frac{e^{x(t_i)^T \beta}}{\sum_{j \in R(i)} e^{x(t_i, j)^T \beta}} \right) \\
 &= \prod_{i=1}^D \left(\frac{e^{\beta_1 s(i) + \beta_2 (x(i) + t(i))}}{\sum_{j \in R(i)} e^{\beta_1 s(i, j) + \beta_2 (x(i, j) + t(i))}} \right)
 \end{aligned}$$

Note that the event happens to individual (i) at time $t(i)$, and the explanatory variables in the risk set are all assessed at exactly that same instant $t(i)$. What happens?

$$\begin{aligned}
 &= \prod_{i=1}^n \frac{e^{\beta_1 s(i) + \beta_2 x(i) + \beta_2 t(i)}}{\sum_{j \in R(i)} e^{\beta_1 s(i, j) + \beta_2 x(i, j) + \beta_2 t(i)}} \\
 &= \prod_{i=1}^n \frac{e^{\beta_1 s(i) + \beta_2 x(i)}}{e^{\beta_2 t(i)} \sum_{j \in R(i)} e^{\beta_1 s(i, j) + \beta_2 x(i, j)}}
 \end{aligned}$$

Q4

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```
> #Q4: Area 51
>
> library(survival)
> area51 = read.table("https://www.utstat.toronto.edu/brunner/data/legal/area51.data.txt")
> alien = coxph(Surv(time1,time2,taken) ~ age + sex + hat, data=area51)
> summary(alien)
```

Call:

```
coxph(formula = Surv(time1, time2, taken) ~ age + sex + hat,
      data = area51)
```

n= 4244, number of events= 103

	coef	exp(coef)	se(coef)	z	Pr(> z)
age	-0.00068	0.99932	0.01020	-0.067	0.94683
sexM	0.15126	1.16330	0.19811	0.764	0.44515
hat	-0.56542	0.56812	0.20478	-2.761	0.00576 **

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

	exp(coef)	exp(-coef)	lower .95	upper .95
age	0.9993	1.0007	0.9795	1.0195
sexM	1.1633	0.8596	0.7890	1.7152
hat	0.5681	1.7602	0.3803	0.8487

Concordance= 0.591 (se = 0.028)

Likelihood ratio test= 8.48 on 3 df, p=0.04

Wald test = 8.2 on 3 df, p=0.04

Score (logrank) test = 8.4 on 3 df, p=0.04

```
>
```

```
>
```