

Mathematical Statistics II

STA2212H S LEC9101

Week 4

February 3 2021

Start recording!



Calling Bullshit @callin_bull · Jan 30

Guys. Time for some causal graph theory.



Hillary Clinton @HillaryClinton · Jan 28

Data proving @GretaThunberg right—"you are never too small to make a difference."



Geoffrey Supran @GeoffreySupran · Jan 26

"The @GretaThunberg Effect" is now an empirically demonstrated, peer-reviewed phenomenon:

"We find that those who are more familiar with Greta Thunberg have higher intentions of taking collective actions to reduce global warming."



The Greta Thunberg Effect: Familiarity with Greta Thunberg predicts intentions to engage in climate activism in the United States

Anandita Sabherwal¹  | Matthew T. Ballew²  | Sander van der Linden¹  |
Abel Gustafson³  | Matthew H. Goldberg⁴  | Edward W. Maibach⁵  |
John E. Kotcher⁵  | Janet K. Swim⁶  | Seth A. Rosenthal⁴  | Anthony Leiserowitz⁴ 



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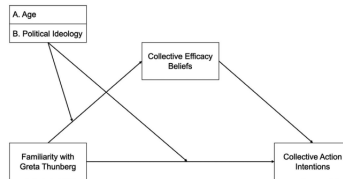


FIGURE 1 Conceptual model for hypotheses. Model tests the effect of familiarity with Greta Thunberg on intentions to take collective action through collective efficacy beliefs, as a simple mediation (Hypothesis 1), moderated by age (A; Hypotheses 2a and 2b), and moderated by political ideology (B; Hypotheses 3a and 3b), respectively



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- Choosing test statistics – optimality, likelihood, convenience

- Simple and composite hypotheses

$$\theta = \theta_0; \quad \theta \in \Theta_0$$

- Generalized likelihood ratio tests

$$W(\theta_0) = 2\{\ell(\hat{\theta}) - \ell(\tilde{\theta}_0)\}$$

- Wald tests

$$(\hat{\theta} - \theta_0)/\hat{se}$$

- significance tests

$$p\text{-value: } \Pr(T > t^{obs}; H_0)$$

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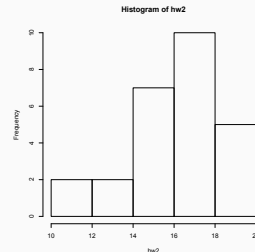
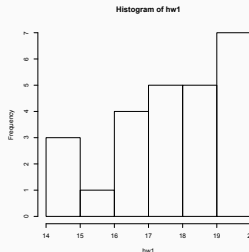
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1. Homework 2
2. confidence intervals
3. goodness-of-fit tests
4. multiple testing



- February 8 3.00 – 4.00 Paul McNicholas
- “Selected Problems in Classification” [Link](#)

Data Science and Applied Research Series



Coefficients:

	Estimate	Std. Error	z value	Pr(> z)	
(Intercept)	-34.103704	6.530014	-5.223	1.76e-07	***
zn	-0.079918	0.033731	-2.369	0.01782	*
indus	-0.059389	0.043722	-1.358	0.17436	
chas	0.785327	0.728930	1.077	0.28132	
nox	48.523782	7.396497	6.560	5.37e-11	***
rm	-0.425596	0.701104	-0.607	0.54383	
age	0.022172	0.012221	1.814	0.06963	.
dis	0.691400	0.218308	3.167	0.00154	**
rad	0.656465	0.152452	4.306	1.66e-05	***
tax	-0.006412	0.002689	-2.385	0.01709	*
ptratio	0.368716	0.122136	3.019	0.00254	**
black	-0.013524	0.006536	-2.069	0.03853	*
lstat	0.043862	0.048981	0.895	0.37052	
medv	0.167130	0.066940	2.497	0.01254	*

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

logistic regression

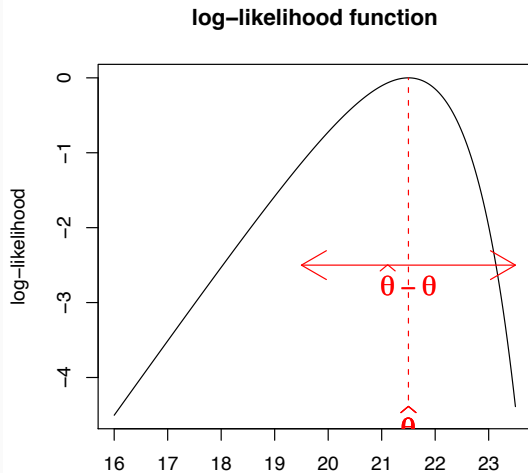
```
> summary(lm1)
Call:
lm(formula = log(cost) ~ date + log(t1) + log(t2) + log(cap) +
    pr + ne + ct + bw + log(cum.n) + pt, data = nuclear)
Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept) -14.24198    4.22880  -3.368  0.00291 **
date          0.20922    0.06526   3.206  0.00425 **
log(t1)       0.09187    0.24396   0.377  0.71025
log(t2)       0.28553    0.27289   1.046  0.30731
log(cap)      0.69373    0.13605   5.099 4.75e-05 ***
pr           -0.09237    0.07730  -1.195  0.24542
ne            0.25807    0.07693   3.355  0.00300 **
ct            0.12040    0.06632   1.815  0.08376 .
bw            0.03303    0.10112   0.327  0.74715
log(cum.n)   -0.08020    0.04596  -1.745  0.09562 .
pt           -0.22429    0.12246  -1.832  0.08125 .
---

```

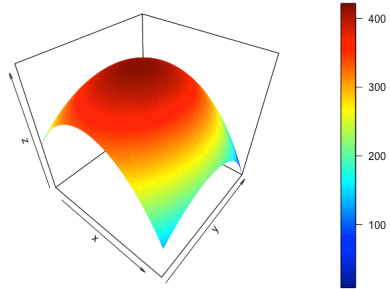
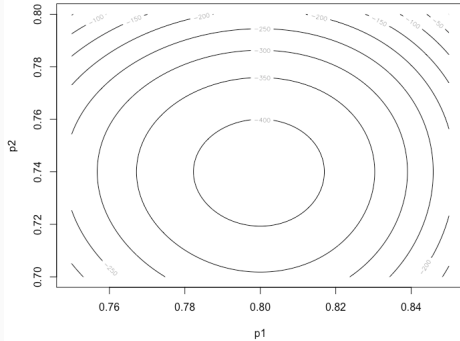
```
log(cap)      0.69373    0.13605    5.099 4.75e-05 ***
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pt           -0.22429    0.12246   -1.832  0.08125 .
---
Residual standard error: 0.1645 on 21 degrees of freedom
Multiple R-squared:  0.8717, Adjusted R-squared:  0.8106
F-statistic: 14.27 on 10 and 21 DF,  p-value: 3.081e-07
```

Example: Likelihood confidence intervals

... likelihood confidence intervals



... likelihood confidence intervals



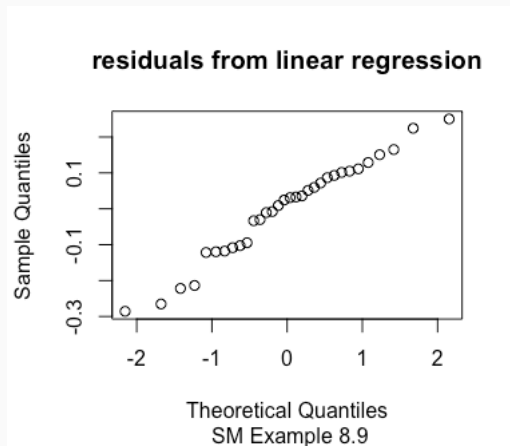
- interpretation of confidence interval or confidence bound random endpoint(s)
- size of test $\leftrightarrow$ confidence level
- power of test $\leftrightarrow$ width of confidence interval

- data $x_1, \dots, x_n$ independent, identically distributed
- test statistic $t = t(\mathbf{x})$, observed value t^{obs}

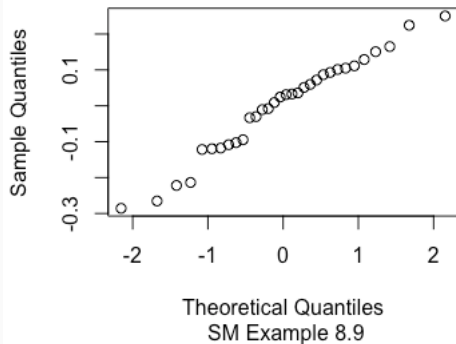
-

$$p^{obs} = \Pr(T \geq t^{obs}; H_0)$$

- Example: sign test



residuals from linear regression



Maize data SM Ex 7.24

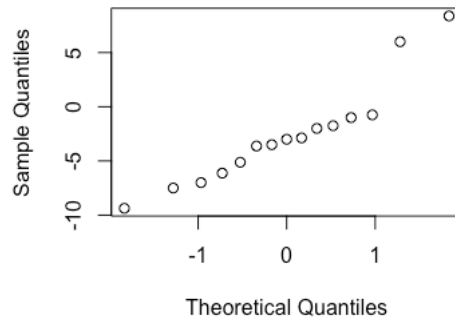
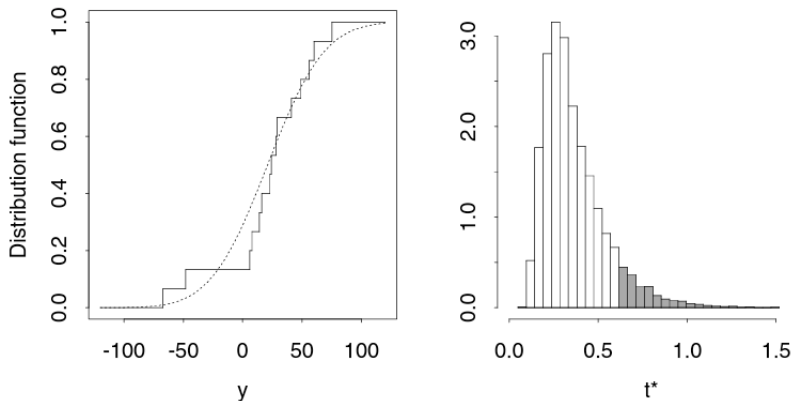


Figure 7.5 Analysis of maize data. Left: empirical distribution function for height differences, with fitted normal distribution (dots). Right: null density of Anderson–Darling statistic T for normal samples of size $n = 15$ with location and scale estimated. The shaded part of the histogram shows values of T^* in excess of the observed value t_{obs} .



SM Example 7.24 testing $N(\mu, \sigma^2)$ distribution

cumulative d.f.

• $X_1, \dots, X_n$ i.i.d. $F(\cdot)$; $H_0 : F = F_0$

• $\hat{F}_n(t) = \frac{1}{n} \sum_{i=1}^n \mathbf{1}\{X_i \leq t\}$

• test statistic:

1. $\sup_t |\hat{F}_n(t) - F_0(t)|$

2. $\int \{\hat{F}_n(t) - F_0(t)\}^2 dt$

3. $\int \frac{\{\hat{F}_n(t) - F_0(t)\}^2}{\hat{F}_n(t)\{1 - \hat{F}_n(t)\}} dt$

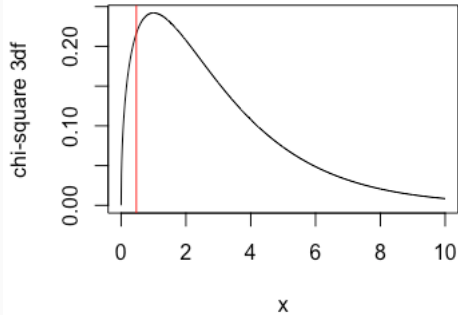
4. χ^2 tests

AoS 10.4

• SM Example 7.24 testing $N(\mu, \sigma^2)$ distribution

• SM Example 7.23; 6.14 testing $U(0, 1)$ distribution

AoS Example 10.18

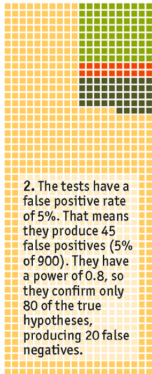
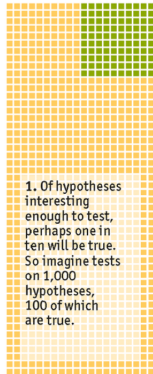


Multiple testing

Unlikely results

How a small proportion of false positives can prove very misleading

False True False negatives False positives



Source: *The Economist*

False Discovery Rates

Benjamini-Hochberg

