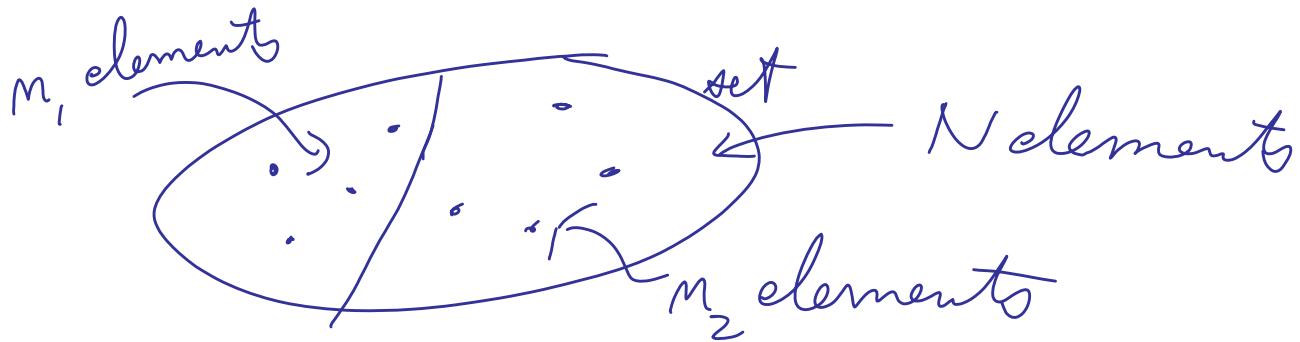


Some Basic Models

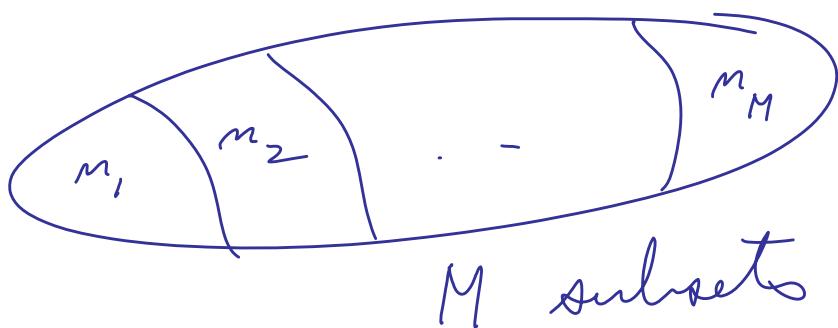


of ways of partitioning $\Rightarrow m_1$ in
1st + m₂ in 2nd

$$= \binom{N}{m_1, m_2} = \frac{N!}{m_1! m_2!} \quad \left(= \binom{N}{m_1} \text{ or } \binom{N}{m_2} \right)$$

total # of partitions

$$= \sum_{0 \leq m_1 + m_2 \leq N} \binom{N}{m_1, m_2} = 2^N$$



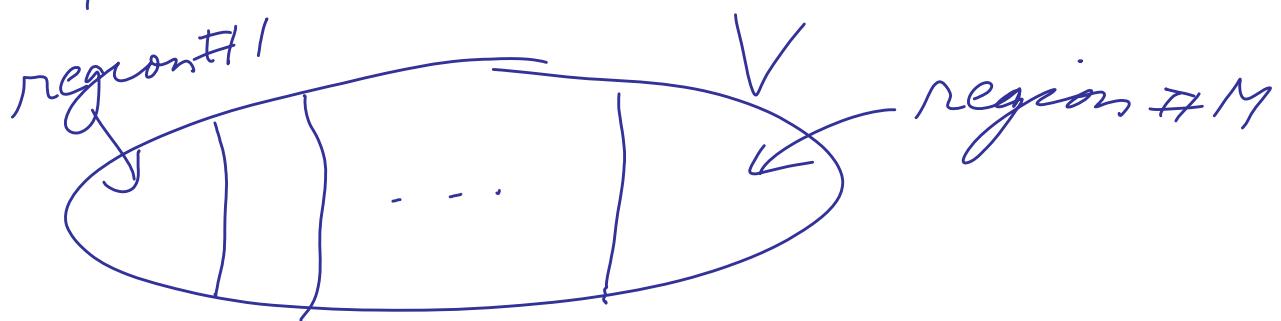
of ways of partitioning with m_1 , in 1st, ...; m_M in the M th is

$$\binom{N}{m_1, \dots, m_M} \quad \text{multinomial coefficient}$$

$$\begin{aligned} \# \text{ of partitions} &= \sum_{0 \leq m_1 + \dots + m_M = N} \binom{N}{m_1, \dots, m_M} = \underbrace{(1 + \dots + 1)^N}_{M \text{ terms}} \\ &= M^N \end{aligned}$$

Application

N - particles into M regions



particles placed in V .

of arrangements of particles = M^N

Let $\underset{N \times 1}{\tilde{X}}$ keep track of which regions the particles are in.

Notice x_i can only take on values $1, \dots, M$. \nearrow i th component of \tilde{x}
 $\underbrace{\quad\quad\quad}_{\text{relates to } i\text{th particle}}$

There are M^N possible "values" of \tilde{x} .

Assume particles are placed in some random way such that

$$(*) E[h(\tilde{x})] = \frac{1}{M^N} \sum_{\substack{\text{all possible} \\ \tilde{x}}} h(\tilde{x}), \forall h$$

\downarrow
real-valued
 $(h: \mathbb{R}^N \rightarrow \mathbb{R})$

(say particles have no preference in regions)

Special h

$$h(\tilde{x}) = h_1(x_1) h_2(x_2) \cdots h_N(x_N)$$

Plug this into (*) to get

$$E[h_1(x_1) \cdots h_N(x_N)] = \left[\frac{1}{M} \sum_{x_1} h_1(x_1) \right] \cdots \left[\frac{1}{M} \sum_{x_N} h_N(x_N) \right]$$

Def'n of \tilde{X}

$$= E[h_1(\tilde{X}_1)] \cdots E[\tilde{h}_N(\tilde{X}_N)]$$

Defn $E[h_1(\tilde{X}_1)] = \frac{1}{M^N} \sum_{\substack{\text{possible} \\ \tilde{X}}} h_1(x_i) = \frac{1}{M} \sum_{x_i=1}^M h_1(x_i)$

Def'n X_1, X_2, \dots are independent if

$$E[h_1(X_1) h_2(X_2) \cdots]$$

$$= E[h_1(X_1)] E[h_2(X_2)] \cdots, \text{"if" } h_i \text{ is}$$

Def'n A_1, A_2, \dots are independent if
 $I(A_1), I(A_2), \dots$ are.

e.g. Suppose A_1, A_2 are ind. Then

$$E[I(A_1) I(A_2)] = E[I(A_1)] E[I(A_2)]$$

$$\Rightarrow P(A_1 A_2) = P(A_1) P(A_2)$$

Notice A_1, A_2, A_3 independent

$$\Rightarrow E[I(A) I(A_2) \underset{\substack{\uparrow \\ \gamma_3}}{I} \bar{I}(A_3)] \\ \equiv 1$$

$$= E(I(A) I(A_2) \times 1) = P(A) P(A_2)$$

In fact if A_1, A_2, \dots are ind then

$$P(A_{i_1} A_{i_2} \dots A_{i_k}) = P(A_{i_1}) P(A_{i_2}) \dots P(A_{i_k})$$

(say $1, 2, 3, \dots$ form a subsequence of provides
 $i_1 < i_2 < i_3 < \dots$
 $\uparrow \uparrow \uparrow$)

Some applications

Call X a counting rv if its range $\subset \{0, 1, 2, \dots\}$.

$$f(x) = P(X=x) \quad \text{--- } \cancel{\text{pdf}}$$

Notice $f(x) \geq 0$
 $\sum_x f(x) = 1$

$$G(z) = E(z^X) \quad \text{--- } \cancel{\text{pgf}}$$

Note — $|G(z)| \leq E(|z|^X) \leq 1 \quad \text{if } |z| \leq 1$

$$- G(z) = \sum_{k=0}^{\infty} z^k P(X=k)$$

— G determines the dist'n of X

e.g. $X \sim \text{Bernoulli}(p)$, $0 < p < 1$.

$$\left. \begin{array}{l} P(X=1) = p \\ P(X=0) = 1-p \quad (=q) \end{array} \right\} \begin{array}{l} \text{pdf is} \\ G(z) = q + pz \end{array}$$

Let X_1, X_2, \dots be iid Bernoulli (p) rv's.

$$Y = X_1 + \dots + X_N \sim \text{binomial}(N, p)$$

$$\begin{aligned} G_Y(z) &= E(z^Y) = E(z^{X_1 + \dots + X_N}) \\ &= E(z^{X_1} \cdots z^{X_N}) \end{aligned}$$

$$= E(z^{X_1}) \cdots E(z^{X_N})$$

$$= (q + pz)^N$$

binomial theorem

$$= \sum_{k=0}^N \binom{N}{k} (pz)^k q^{N-k}$$

$$= \sum_{k=0}^N \binom{N}{k} p^k q^{N-k} z^k$$

$P(Y=k)$

pgf of a binomial(N, p) is

$$(q + pz)^N = E(z^Y) = G_Y(z)$$

$$\frac{d}{dz} G_Y(z) = \underset{\substack{\uparrow \\ \text{in general}}}{E}\left(\frac{d}{dz} z^Y\right) = E(Y z^{Y-1})$$

need to justify

$$\frac{d^2}{dz^2} G_Y(z) = E\left(\frac{d^2}{dz^2} z^Y\right) = E(Y(Y-1)z^{Y-2})$$

put $z=1$ to get $E(Y) + \underbrace{E[Y(Y-1)]}_{E(Y^2) - E(Y)}$

$$\frac{d}{dz} (q+pz)^N = N(q+pz)^{N-1} \stackrel{z=1}{=} Np$$

$$\frac{d^2}{dz^2} (q+pz)^N = N(N-1)(q+pz)^{N-2} \stackrel{z=1}{=} N(N-1)p^2$$

$$\begin{aligned} \therefore \quad & E(Y) = Np \\ & E(Y^2) - E(Y) = N(N-1)p^2 \end{aligned} \quad \left. \begin{array}{l} \Rightarrow E(Y^2) \\ = N(N-1)p^2 + Np \end{array} \right.$$

$$\begin{aligned} \text{Var}(Y) &= \underline{E(Y^2)} - (\underline{E(Y)})^2 = N(N-1)p^2 + Np - N^2 p^2 \\ &= Npq \end{aligned}$$

$$G(z) = P(X=0) + P(X=1)z + \dots$$

$$\Rightarrow G(0) = P(X=0)$$

$$G''(0) = P(X=1)$$

$$G^{(2)}(0) = 2! P(X=2)$$

:

$$G^{(k)}(0) = k! P(X=k)$$

Poisson(λ) probabilities

$$\sum_{k=0}^{\infty} e^{-\lambda} \frac{\lambda^k}{k!} z^k$$

Back to

$$(q + pz)^N = \left[1 + \frac{\lambda}{N} \right]^N \approx e^{\lambda(z-1)}$$

for large N .

If $X \sim \text{Poisson}(\lambda)$

$$E(X) = \text{Var}(X) = \lambda$$

$$f(x) = e^{-\lambda} \frac{\lambda^x}{x!}, x=0, 1, \dots$$

$$= 0, \text{ otherwise}$$

Remark: Finite sum of ind Poisson r.v.'s will be Poisson.

binomial

iid Bernoulli trials yielding

$$\tilde{X}_1, \tilde{X}_2, \dots, \tilde{X}_N \rightarrow \text{iid } \tilde{X}$$

$\tilde{X}_1 \sim \begin{cases} X_1 & z_1 \\ 1 - X_1 & z_2 \end{cases}$

$$P_1 + P_2 = 1$$

Note possible values of \tilde{X} are $\begin{pmatrix} 0 \\ 1 \end{pmatrix}$ or $\begin{pmatrix} 1 \\ 0 \end{pmatrix}$

if pgf of the components of \tilde{X} or just the

the pgf of \tilde{X}

$$G_{\tilde{X}}(\tilde{z}) = E\left(\tilde{z}^{\tilde{X}}\right) = E\left(z_1^{1^{\text{st comp of } \tilde{X}}} z_2^{2^{\text{nd comp of } \tilde{X}}}\right)$$

$$= z_2 P_2 + z_1 P_1$$

$$= P_1 z_1 + P_2 z_2$$

= pgf of a "vector Bernoulli"

Now set

$$\begin{aligned} \tilde{Y} &= \tilde{X}_1 + \dots + \tilde{X}_N \\ \Rightarrow G_{\tilde{Y}}(\tilde{z}) &= (P_1 z_1 + P_2 z_2)^N = \sum_{y_1+y_2=N} \binom{N}{y_1, y_2} P_1^{y_1} P_2^{y_2} z_1^{y_1} z_2^{y_2} \\ \Rightarrow P(\tilde{Y}=y) &= \binom{N}{y_1, y_2} P_1^{y_1} P_2^{y_2} \end{aligned}$$

Note $z_2=1 \Rightarrow (p_1 z_1 + p_2)^N$ is the pgf of the 1st component of \underline{X} (this is the binomial(N, p_1) as you know it).

Extend to the multinomial

$$\underbrace{X_1}_{M \times 1}, \underbrace{X_2}_{M \times 1}, \dots, \underbrace{X_N}_{M \times 1} \text{ iid } \underline{X}$$

\underline{X} has one component = 1 + the rest 0. The probability that the 1 is in the i th place we call p_i ($i=1, \dots, M$). The pgf of \underline{X} is

$$(p_1 z_1 + p_2 z_2 + \dots + p_M z_M)$$

* letting

$$\underline{Y} = \underline{X}_1 + \dots + \underline{X}_N$$

we get

$$G_{\underline{Y}}(\underline{z}) = (p_1 z_1 + \dots + p_M z_M)^N$$

$$\Rightarrow P(\underline{Y} = \underline{y}) = \underbrace{\binom{N}{y_1, \dots, y_M}}_{\left(\begin{array}{c} N \\ y_1, \dots, y_M \end{array} \right)} p_1^{y_1} \cdots p_M^{y_M}$$

$$\text{Note } \frac{\partial^2}{\partial z_1 \partial z_2} G(\underline{z}) = E(Y, z_1^{Y_1-1} z_2^{Y_2-1} z_3^{Y_3-1} \dots) \quad \left. \begin{array}{l} \text{works} \\ \text{for} \\ \text{all counting} \\ \text{rvectors} \end{array} \right\}$$

$\stackrel{\underline{z} = ?}{=} E(Y, Y)$

Can use $\frac{\partial^2}{\partial z_1 \partial z_2} G(\underline{z})$ to get the $\text{cov}(Y_i, Y_j)$ for counting rvec's etc.

Properties of covariance ($\text{cov}(X, Y) = E(XY) - E(X)E(Y)$)

$$\left. \begin{array}{l} \text{cov}(X, X) = \text{Var}(X) \\ \text{cov}(X+c, Y+d) = \text{cov}(X, Y) \\ \text{cov}(aX, bY) = ab \text{cov}(X, Y) \\ \text{cov}\left(\sum_i X_i, \sum_j Y_j\right) = \sum_{i,j} \text{cov}(X_i, Y_j) \end{array} \right\} \text{vert}$$

e.g. $X_1 \sim \text{Poisson}(\lambda_1)$, $X_2 \sim \text{Poisson}(\lambda_2)$,

$X_3 \sim \text{Poisson}(\lambda_3)$

$$\left. \begin{array}{l} U = X_1 + X_2 \\ V = X_2 + X_3 \end{array} \right\} \text{easy to get the pgf}$$

$$= E(z_1^U z_2^V z_3^{X_3})$$

$$\begin{aligned}
 &= E(z_1^{X_1} (z_1 z_2)^{X_2} z_2^{X_3}) \\
 &= E(z_1^{X_1}) E(z_1 z_2)^{X_2} E(z_2^{X_3}) \\
 &= e^{\lambda_1(z_1-1)} e^{\lambda_2(z_1 z_2 - 1)} e^{\lambda_3(z_2 - 1)}
 \end{aligned}$$

- Use this to get $E(UV) - \cancel{\text{imp}}$ } 15 minutes
 * then get $\text{cov}(U, V)$

Another way

$$\begin{aligned}
 \text{cov}(U, V) &= \text{cov}(X_1 + X_2, X_2 + X_3) \\
 &= \text{cov}(X_1, X_2) + \text{cov}(X_1, X_3) \\
 &\quad + \text{cov}(X_2, X_2) + \text{cov}(X_2, X_3) \\
 &= \text{cov}(X_2, X_2) = \text{Var}(X_2) = \lambda_2
 \end{aligned}$$

Note X, Y ind $\Rightarrow E(XY) = E(X)E(Y) \Rightarrow \text{cov}(X, Y) = 0$

e.g. $Z \sim N(0, 1)$. Let $X = Z$, $Y = Z^2$. Then

X & Y are dependent but

$$E(XY) = E(ZZ^2) = E(Z^3) = 0$$

& so X & Y are uncorrelated.

$$\underbrace{E(X)}_0 \underbrace{E(Y)}_1$$