

Should Bayesians and frequentists  
calibrate their parameters?

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# Calibrate?

Calibrate parameters?  $B \mid (\mu, \sigma)$   $\frac{d\mu d\sigma}{\sigma}$   
 $f \mid \sigma \quad \sigma^2 \quad \log \sigma^2$

Conceptualize parameter? What is real interest parameter?

Measure of departure of  
 data from parameter value?

Common issues...  $B$  &  $f$ ?

Claim: If Bayesians & frequentists calibrate  
 then they'll agree... same conclusions

xcxc = western dec 5

Does it matter?

Does it make a difference?

- Examples
  - There are issues
  - Calibration is central
  - There are methods
- 

Thanks

Acknowledgements To

Many colleagues —

# B Example 1

b1

## One Normal

$y_1, \dots, y_n \sim N(\mu, \sigma^2)$   
Interest in  $\dots \sigma$

$$\text{or } \underline{y} = X\beta + \sigma \underline{z}$$

$N(0, 1)$   
Student(7)

Calibrate?  $\sigma$   $\sigma^2$   $\log \sigma^2$  "same" parameter

$$\hat{\sigma}^2 = n^{-1} \sum (y_i - \hat{y}_i)^2$$

$$s_y^2 = (n-p)^{-1} \sum (y_i - \bar{y})^2$$

} N case

| Unbiasedness for  $\sigma^2$ , ... for  $\sigma$  ?

## II Examples

b2

### 1) One Normal

 $y_1, \dots, y_n \quad N \quad \mu \quad \sigma^2$ 

$$\text{or } \underline{y} = X\beta + \sigma \underline{z}$$

 $N \quad 0 \quad 1$   
Student (7)Interest in ...  $\sigma$ 

f Calibrate?     $\sigma$      $\sigma^2$      $\log \sigma^2$     "same" parameter

$$\hat{\sigma}^2 = \bar{n}^{-1} \sum (y_i - \hat{y}_i)^2$$

$$s_y^2 = (n-p)^{-1} \sum (y_i - \bar{y})^2$$

} N case

} Unbiasedness for  $\sigma^2$ , ... for  $\sigma$  ?Bayes ?

$$\sigma^{-2} d\mu d\sigma$$

$$\sigma^{-1} d\mu d\sigma$$

(Left inv., Right inv.)

# II Examples

## 1) One Normal

$y_1, \dots, y_n \sim N(\mu, \sigma^2)$   
Interest in ...  $\sigma$

or  $\underline{y} = X\beta + \sigma \underline{z}$

N O 1  
Student (7)

Calibrate?  $\sigma$   $\sigma^2$   $\log \sigma^2$  "same" parameter

$$\hat{\sigma}^2 = \bar{n}^{-1} \sum (y_i - \hat{y}_i)^2$$

$$s_y^2 = (n-p)^{-1} \sum (y_i - \bar{y})^2$$

} N case  
Unbiasedness for  $\sigma^2$  ... for  $\sigma$  ?

Bayes?  $\sigma^{-2} d\mu d\sigma$

$\sigma^{-1} d\mu d\sigma$

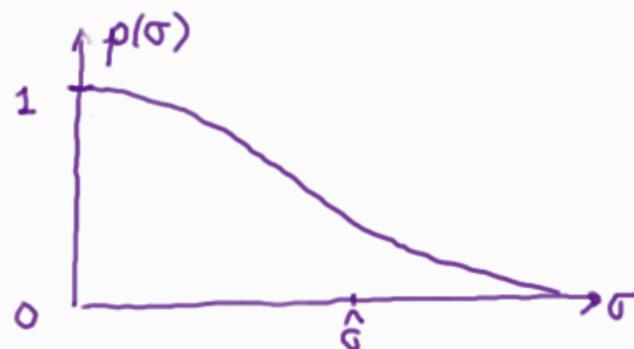
(Left inv., Right inv.)

p-value

$$p(\sigma) = P \left\{ \overset{\text{pivot}}{\text{pivot}} \leq \text{obs. pivot} \right\}$$

$$= P \left\{ \frac{\sum (y_i - \hat{y}_i)^2}{\sigma^2} \leq \frac{\sum (y_i^0 - \hat{y}_i^0)^2}{\sigma^2} \right\}$$

$\uparrow \chi^2_{n-p}$  in N case



Measure of departure

... built into pivot

Intrinsic (continuity)

p-value = confidence distribution function

# C Example 2

C1

## Two normals

$y_{11}, \dots, y_{1n} \quad N \mu_1 \sigma_1^2$       Interest in  $\delta = \mu_1 - \mu_2$   
 $y_{21}, \dots, y_{2m} \quad N \mu_2 \sigma_2^2$

Calibration?      Difference in means ... !      Obvious!

$$f \quad z = \frac{\bar{y}_1 - \bar{y}_2}{\left\{ \frac{\sum (y_{1i} - \hat{y}_{1i})^2}{n(n-1)} + \frac{\sum (y_{2i} - \hat{y}_{2i})^2}{m(m-1)} \right\}^{1/2}} \quad \text{cf " ? " (Long standing issue!)}$$

## II Examples

c2

### 2) Two normals

$$\begin{array}{l} y_{11}, \dots, y_{1n} \quad N, \mu_1, \sigma_1^2 \\ y_{21}, \dots, y_{2m} \quad N, \mu_2, \sigma_2^2 \end{array} \quad \text{Interest in } \delta = \mu_1 - \mu_2$$

Calibration? Difference in means ... ! Obvious!

$$f \quad z = \frac{\bar{y}_1 - \bar{y}_2}{\left\{ \frac{\sum (y_{1i} - \hat{y}_{1i})^2}{n(n-1)} + \frac{\sum (y_{2i} - \hat{y}_{2i})^2}{m(m-1)} \right\}^{1/2}} \quad \text{cf ? Long standing issue}$$

$$B \quad \pi(\theta) = d\mu_1 d\mu_2 \frac{d\sigma_1 d\sigma_2}{\sigma_1^2 \sigma_2^2} \quad \text{or} \quad d\mu_1 d\mu_2 \frac{d\sigma_1 d\sigma_2}{\sigma_1 \sigma_2}$$

Left inv. / Jeffreys

Right inv. / modified Jeffreys / Fisher

Behrens 1929 Fisher 1934 Jeffreys 1939 1946

## II Examples

c3

### 2) Two normals

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Calibration? Difference in means ... ! Obvious!

$$f \quad z = \frac{\bar{y}_1 - \bar{y}_2}{\left\{ \frac{\sum (y_{1i} - \hat{y}_{1i})^2}{n(n-1)} + \frac{\sum (y_{2i} - \hat{y}_{2i})^2}{m(m-1)} \right\}^{1/2}} \quad \text{cf ? Long standing issue}$$

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Left inv. / Jeffreys

Right inv. / modified Jeffreys / Fisher

$$\pi(\theta) = d\mu_1 d\mu_2 \frac{d\sigma_1 d\sigma_2}{\sigma_1^3 \sigma_2^3 \left( \frac{\sigma_1^2}{n} + \frac{\sigma_2^2}{m} \right)^{-1}}$$

Ghosh Kim (2001)

Behrens 1929 Fisher 1934 Jeffreys 1939 1946 Ghosh Kim 2001 (CJS)

Example Behrens-Fisher 1929 1934 Ghosh Kim 2001

$y_{11}, \dots, y_{1n} \sim N(\mu_1, \sigma_1^2)$

Interest  $\psi = \mu_1 - \mu_2$

$y_{21}, \dots, y_{2n} \sim N(\mu_2, \sigma_2^2)$

GK:  $\pi = \sigma_1^{-3} \sigma_2^{-3} \left( \frac{\sigma_1^2}{n} + \frac{\sigma_2^2}{m} \right)$

vs Jeffreys(right):  $\sigma_1^{-1} \sigma_2^{-1}$

Some simulation numbers: Case

| $n$ | $m$ | $\sigma_1^2$ | $\sigma_2^2$ | $\psi$ |
|-----|-----|--------------|--------------|--------|
| 2   | 2   | 2            | 1            | 2      |

Nominal 5% 95% points by various methods

Example Behrens-Fisher 1929 1934 Ghosh Kim 2001

$$y_{11}, \dots, y_{1n} \quad N \quad \mu_1 \quad \sigma_1^2$$

Interest  $\psi = \mu_1 - \mu_2$

$$y_{21}, \dots, y_{2n} \quad N \quad \mu_2 \quad \sigma_2^2$$

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us Jeffreys (right):  $\sigma_1^{-1} \sigma_2^{-1}$

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Nominal 5% 95% points by various methods

|   | <u>Nominal</u> | <u>5%</u> | <u>95%</u> |
|---|----------------|-----------|------------|
| B | Jeffreys       | 0.7%      | 99.1%      |
| B | Kim Ghosh      | 1.7%      | 97.9%      |
| f | Lik ratio      | 13.2%     | 86.9%      |

Example Behrens-Fisher 1929 1934 Ghosh Kim 2001

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Some simulation numbers: Case 

|     |     |              |              |        |
|-----|-----|--------------|--------------|--------|
| $n$ | $m$ | $\sigma_1^2$ | $\sigma_2^2$ | $\psi$ |
| 2   | 2   | 2            | 1            | 2      |

Nominal 5% 95% points by various methods including flat prior

|  | <u>Nominal</u>     | <u>5%</u>    | <u>95%</u>    |                               |
|--|--------------------|--------------|---------------|-------------------------------|
|  | Jeffreys           | 0.7%         | 99.1%         |                               |
|  | Kim Ghosh          | 1.7%         | 97.9%         |                               |
|  | Lik ratio          | 13.2%        | 86.9%         |                               |
| flat prior }<br>data dep. }<br>$O(n^{-3/2})$ }<br>$B \equiv f$ } | $p(\psi), s(\psi)$ | 4.23%        | 95.8%         | <u><math>N=100,000</math></u> |
|  | Sim 95% limits     | (4.86, 5.14) | (94.9, 95.14) |                               |

# D Data dependent priors

dl

- Sounds like a contradiction?

- Box & Cox 1964  $y^{\wedge} = X\beta + \sigma \tilde{z}$   $z_i \sim N(0,1)$

Used data dependent prior

----- Later...

- Cox: Prior depend on data? Of course!

- Wasserman 2000

but Why?

## D Data dependent priors

d2

- Sounds like a contradiction?

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Prior depend on data ?

Prior depend on model ? .... default prior (objective ???)

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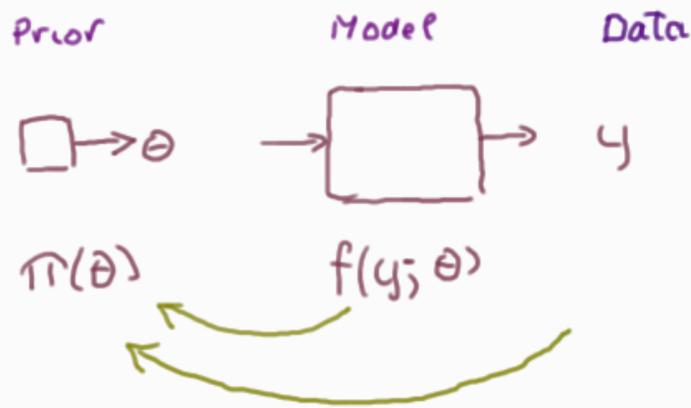
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Prior depend on data ?

Prior depend on model ? .... default prior (objective???)



## D Data dependent priors

Should: Bayesian right tail posterior probabilities  
be equal to p-values i.e. have frequentist properties?

Widely addressed: "matching" in B community!

Judith Rousseau ... "important to be as close as possible  
to the uniform distribution"

## D Matching:

Pure case: Scalar, scalar  $\theta$  $f(y; \theta)$ 

" p-value = right tail posterior probability

" Implications!

$$p^{\circ}(\theta) = F(y^{\circ}; \theta) = \int_{\theta}^{\infty} L^{\circ}(\theta) \pi(\theta) d\theta$$

## D Matching:

Pure case: Scalar, scalar  $\theta$   $f(y; \theta)$

p-value = right tail posterior probability

$$p^{\circ}(\theta) = F(y^{\circ}; \theta) = \int_{\theta}^{\infty} L^{\circ}(\theta) \pi(\theta) d\theta$$

$$F(y^{\circ}; \theta) = \int_{\theta}^{\infty} c f(y^{\circ}; \theta) \pi(\theta) d\theta$$

Differentiate:

$$F_{;\theta}(y^{\circ}; \theta) = -c F_y(y^{\circ}; \theta) \pi(\theta)$$

$$\pi(\theta) = -c \frac{F_{;\theta}(y^{\circ}; \theta)}{F_y(y^{\circ}; \theta)}$$

Needed  
Data-dependent  
model-based  
prior

# D Matching:

d7

Pure case: Scalar, scalar  $\theta$   $f(y; \theta)$

p-value = right tail posterior probability

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Differentiate

$$F_{;\theta}(y^{\circ}; \theta) = -c F_y(y^{\circ}; \theta) \pi(\theta)$$

$$\pi(\theta) = -c \frac{F_{;\theta}(y^{\circ}; \theta)}{F_y(y^{\circ}; \theta)} = \left. \frac{dy}{d\theta} \right|_{\text{data fixed pivot}}$$

Need: Data-dependent prior  $\Rightarrow$  "matching"

Note - prior is flat if model is location

-  $F(y; \theta) \sim U(0, 1)$  .... is pivotal

- Prior =  $\left. \frac{dy}{d\theta} \right|_{y^{\circ}}$  .. (with fixed pivot)

# E Inference elements

e1

$$f(y; \theta) + y^o \Rightarrow ?$$

Case: Scalar  $y$  &  $\theta$  --- Where things should start!

|          |                        |               |                  |
|----------|------------------------|---------------|------------------|
| Obvious: | $f^o = f(y^o; \theta)$ | $L^o(\theta)$ | Prob <u>at</u>   |
|          | $F^o = F(y^o; \theta)$ | $p^o(\theta)$ | Prob <u>left</u> |

# E Inference elements

e2

$$f(y; \theta) + y^o \Rightarrow ?$$

Case: Scalar  $y$  &  $\theta$  --- Where things should start!

Obvious:

|                        |               |                  |
|------------------------|---------------|------------------|
| $f^o = f(y^o; \theta)$ | $L^o(\theta)$ | Prob <u>at</u>   |
| $F^o = F(y^o; \theta)$ | $p^o(\theta)$ | Prob <u>left</u> |

Q

Is there anything more?

# E Inference elements

e3

$$f(y; \theta) + y^o \Rightarrow ?$$

Case: Scalar  $y$  &  $\theta$  ... Where things should start!

Obvious:  $f^o = f(y^o; \theta)$        $L^o(\theta)$       Prob at  
 $F^o = F(y^o; \theta)$        $p^o(\theta)$       Prob left

Q

Is there anything more?

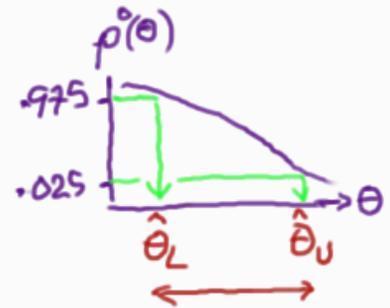
$u = F(y; \theta)$  is a pivot

Invert: get all confidence bounds, conf. int's

$$95\% \text{ CI} = \{p^{-1}(.975), p^{-1}(.025)\}$$

A

It's all there!



# Inference & what more than likelihood

e4

$$f(y; \theta) + y^o \Rightarrow ?$$

Case: Scalar  $y$  &  $\theta$  ... Where things should start!

Obvious:

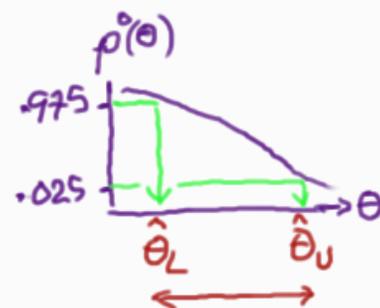
|                        |               |                  |
|------------------------|---------------|------------------|
| $f^o = f(y^o; \theta)$ | $L^o(\theta)$ | Prob <u>at</u>   |
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Is there anything more?

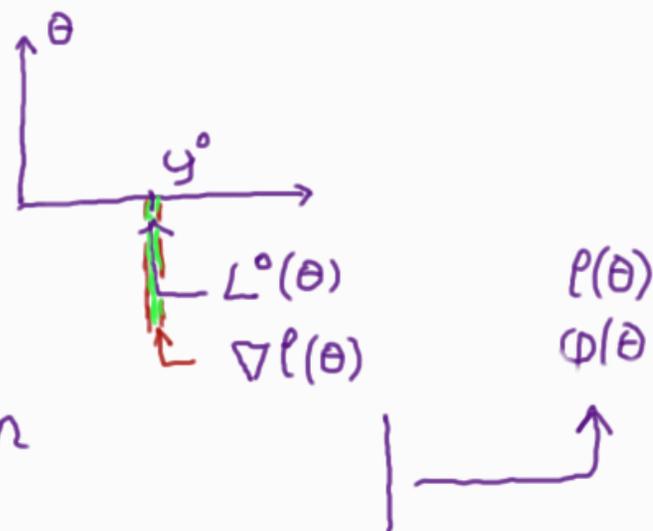
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$$95\% \text{ CI} = \{p^{-1}(.975), p^{-1}(.025)\}$$



Can do in general;  
Min. assumptions!



But: you need more than  
likelihood

Inference: Need more than Likelihood!

e5

$$f(y; \theta) + y^0 \Rightarrow ?$$

$$y \text{ in } \mathbb{R}^n$$

$$\theta \text{ in } \mathbb{R}^p$$

moderate regularity

Inference: Need more than Likelihood!

e6

$$f(y; \theta) + y^0 \Rightarrow ?$$

$$y \text{ in } \mathbb{R}^n \quad \theta \text{ in } \mathbb{R}^p$$

moderate regularity

Need:  $l(\theta; y^0) = a + \log f(y^0; \theta)$

$\varphi(\theta) =$  Gradient likelihood in directions  $V = (\underline{v}_1, \dots, \underline{v}_p)$   
tangent to "conditioning"

Inference: Need more than Likelihood!

e7

$$f(y; \theta) + y^0 \Rightarrow ?$$

$y$  in  $\mathbb{R}^n$      $\theta$  in  $\mathbb{R}^p$     moderate regularity

Need:  $l(\theta; y^0) = a + \log f(y^0; \theta)$

$\varphi(\theta) =$  Gradient likelihood in directions  $V = (\underline{v}_1, \dots, \underline{v}_p)$   
tangent to "conditioning"

Get p-value for any scalar  $\psi(\theta)$  3rd

Likelihood " " " or vector  $\psi(\theta)$  3rd

p-value for any vector  $\psi(\theta)$  2nd

uses just mle's, info's & CPU ...

## F Calibrate: priors

Get: 3rd order inference using

$$V = \frac{dy}{d\theta} \Big|_{(y^0, \hat{\theta})}$$

Differentiate for fixed pivotal

Get: 3rd order matching prior

$$V(\theta) = (\underline{v}_1(\theta), \dots, \underline{v}_p(\theta)) = \frac{dy}{d\theta} \Big|_{(y^0, \theta)}$$

$$\text{Prior} = c | M' V(\theta) |$$

$$\begin{aligned} M &= \frac{\partial^2}{\partial y \partial \theta} \ell(\theta; y) \Big|_{\hat{\theta}^0, y^0} \\ &= \underline{L}_{y, \theta}(\hat{\theta}^0; y^0) \end{aligned}$$

## F Calibrate: priors

Got: 3rd order inference using

$$V = \frac{dy}{d\theta} \Big|_{(y^0, \hat{\theta})}$$

Differentiate for fixed pivotal

Get: 3rd order matching prior

$$V(\theta) = (\underline{v}_1(\theta), \dots, \underline{v}_p(\theta)) = \frac{dy}{d\theta} \Big|_{(y^0, \theta)}$$

$$\text{Prior} = c |M' V(\theta)| d\theta$$

$$M = \frac{\partial^2}{\partial y \partial \theta} \ell(\theta; y) \Big|_{\hat{\theta}^0, y^0}$$

$$= \underline{\ell}_{y, \theta}(\hat{\theta}^0; y^0)$$

Example:  $\underline{y} = X\beta + \sigma \underline{z}$

$\underline{z}_i$ : Student or Extreme value

$$\text{Prior} = c |M' V(\theta)| d\theta$$

$$= d\beta \frac{d\sigma}{\sigma}$$

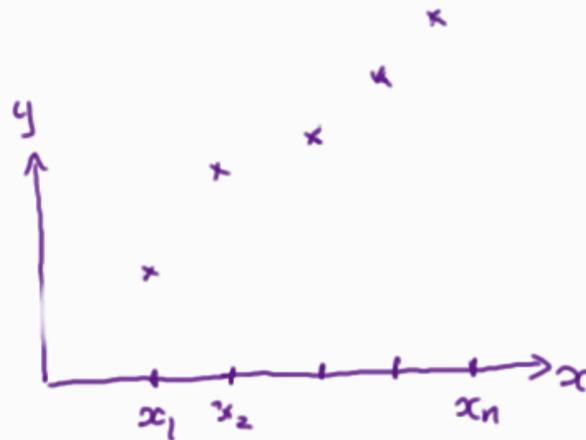
and also gives confidence

Fraser, Reid, Mavras (Rome) Yi (Waterloo) 2005

G Example: Box & Cox 1964

Finding the parameter

Model  $y^{\lambda} = \alpha + \beta x + \sigma z$

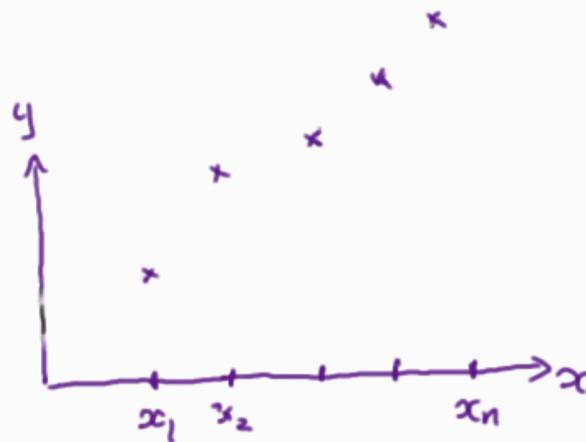


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Finding the parameter

Model  $y^{\lambda} = \alpha + \beta x + \sigma z$

a) If you knew  $\lambda$   
just ordinary regression

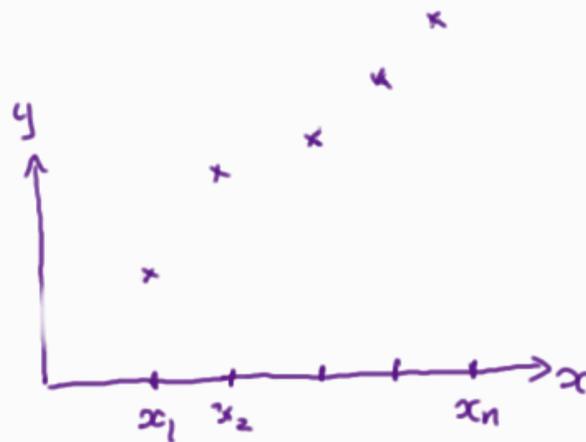


G Example: Box & Cox 1964

Finding the parameter

Model  $y^\lambda = \alpha + \beta x + \sigma z$

a) If you knew  $\lambda$   
just ordinary regression



Psychologist is uncertain as to how to express/calibrate/scale response  $y$   
Possibly:  $y^2$  or  $y^{1/2}$  or  $\log y$  to get: Lin model, Normality, homosced. ...  $y > 0$   
try  $y^\lambda$

Model:  $y = (\alpha + \beta x + \sigma z)^{1/\lambda}$

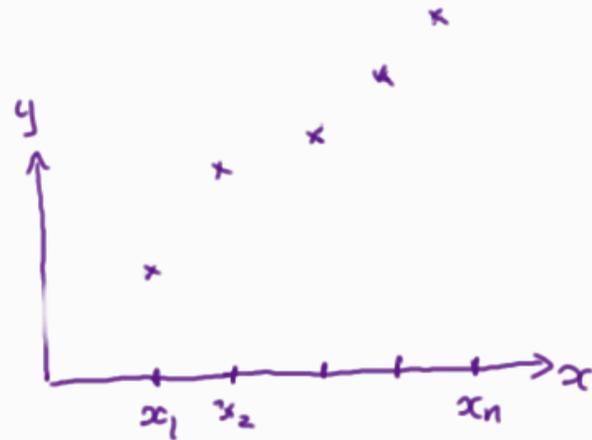
... say  $z \sim N$

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Model  $y^\lambda = \alpha + \beta x + \sigma z$

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" $y > 0$ "

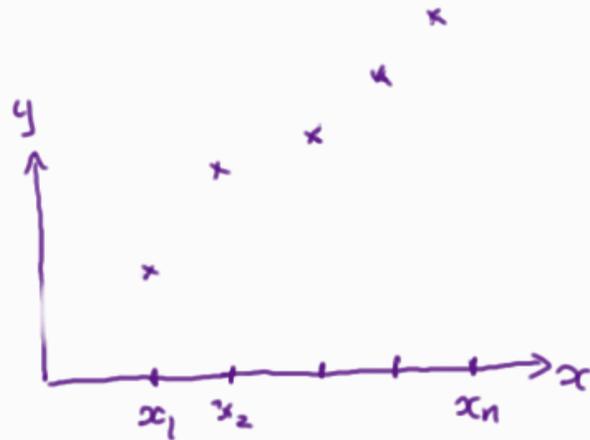
but what does  $\beta$  mean?      Box Cox 1964    Bickel Doksum 1981    Box Cox 1982 !

G Example: Box & Cox 1964

Finding the parameter

Model  $y^\lambda = \alpha + \beta x + \sigma z$

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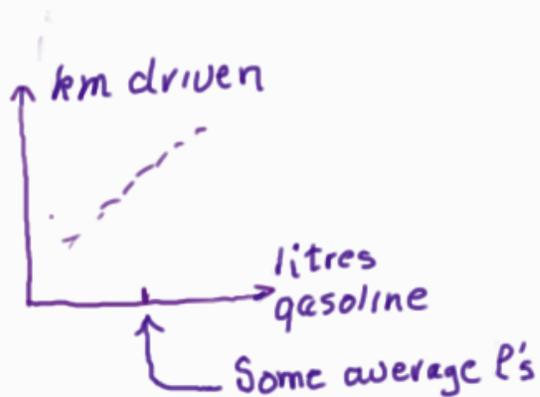
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try  $y^\lambda$

Model:  $y = (\alpha + \beta x + \sigma z)^{1/\lambda}$

--- say  $z \sim N$  "y > 0"

Example



Interest?

How "km" change with  $x$  change

$\psi(\theta) = \frac{d}{dx} (\alpha + \beta x)^{1/\lambda} \Big|_{x_0}$  real parameter

So get p-value.

# A message:

## Calibrate parameter

- 1) Need: likelihood, gradient of likelihood  
 get: 3rd order p-value f
  
- 2) Need: likelihood, pivotal B  
 get: 3rd order posterior tail probs
  
- 3) Bayes & likelihood agree (f B) ♡