

Priors from a differential viewpoint:

How Bayes can deliver 2nd order Accuracy!

D A S Fraser
Statistical Sciences
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2014 April 10

www.utstat.toronto.edu/dfraser/documents/UWO2014.pdf
some references as: ... [~/documents/xxxxx.pdf](#) where xxxx =

Priors from a differential viewpoint:

How Bayes can deliver 2nd order Accuracy!

With a long history

great collaborators:

Nancy

A Wong York

M Bédard U de Montréal

W Lin Toronto

A M Fraser UBC

M J Fraser Toronto

(Preliminary report)

Background:

2nd Order Bayes ?

- u?

Background:

1) Science 2014

Reproducibility & Statistics

2a) Science 2011

"Data" & the 'dust-up'

2b) but...

Retractions and Reproducibility

3a) Science 2013

Efron on..... Reproducibility

3b) Science 2013

Laplace had confidence.... Reproducibility

2nd Order Bayes ?

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- 1) Science 2014 Reproducibility & Statistics
- 2a) Science 2011 "Data" & the 'dust-up'
- 2b) but... Retractions and Reproducibility
- 3a) Science 2013 Efron on..... Reproducibility
- 3b) Science 2013 Laplace had confidence.... Reproducibility

2nd Order Bayes ?

- 1 Scalar parameter: Welch-Peers 1963 B + 200
Example
- 2 Scalar linear interest parameter
Example
- 3 Scalar rotating interest:
Example
- 4 Scalar curved interest:
Example

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Discussion

Science 2014

① View
from Science

Editorial: Marcia McNutt Editor-in-chief
"Science" 17 January 2014

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EDITORIAL

Reproducibility

Marcia McNutt is Editor-in-Chief of *Science*.

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— Marcia McNutt
10.1126/science.1250475



*S. C. Landis et al., *Nature* 490, 187 (2012).

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Low key ...

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Indifference!

Just... Experimental design



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Also: Statistics more generally

Low key ...

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Indifference ...

Just Experimental design

from Science: LHC $p < 3 \cdot 10^{-7}$ 5σ

But there is a lot of background ...

a) ne "Data"

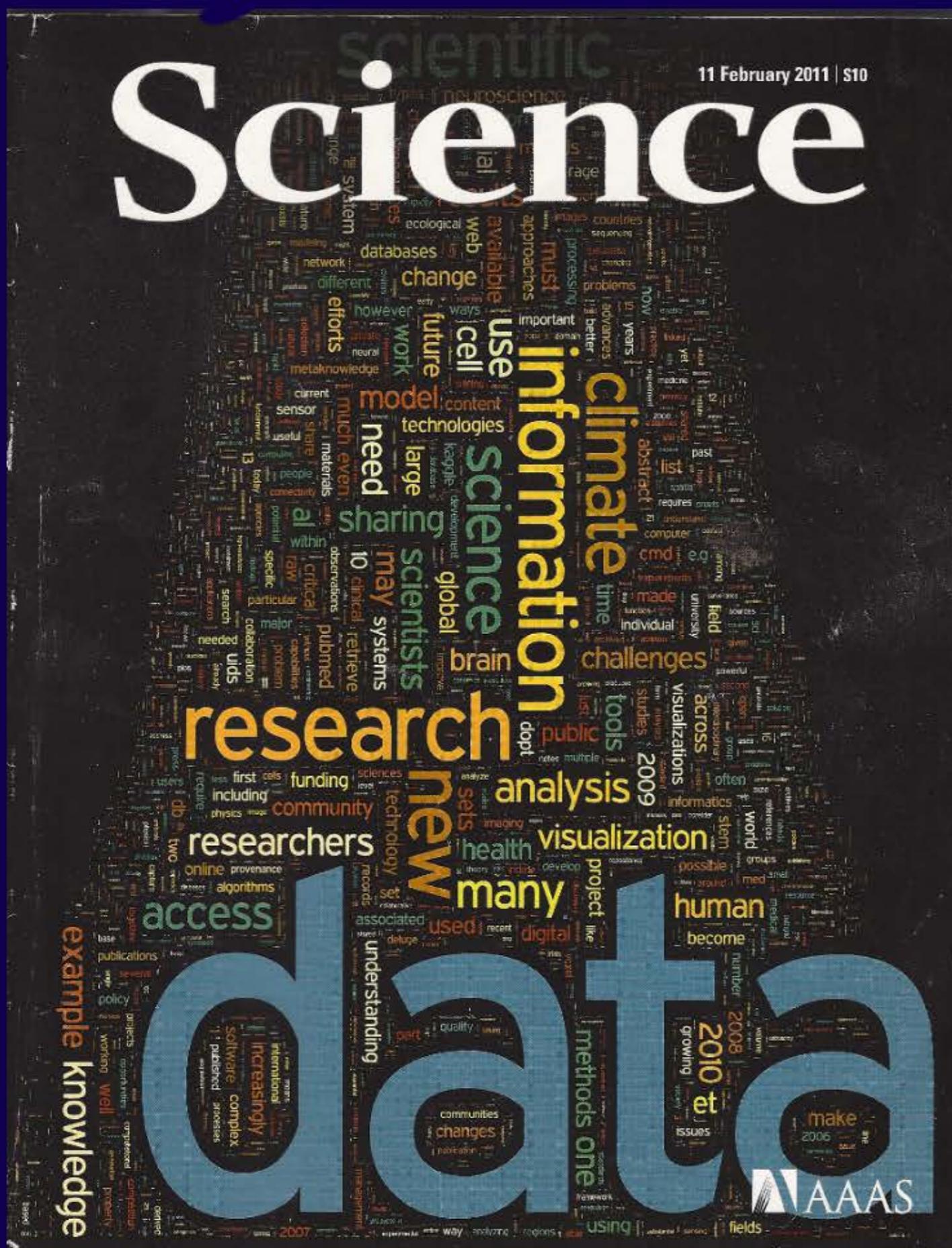
b) ne "Retraction"

So, not as innocent or high principled as it seems ...

2a

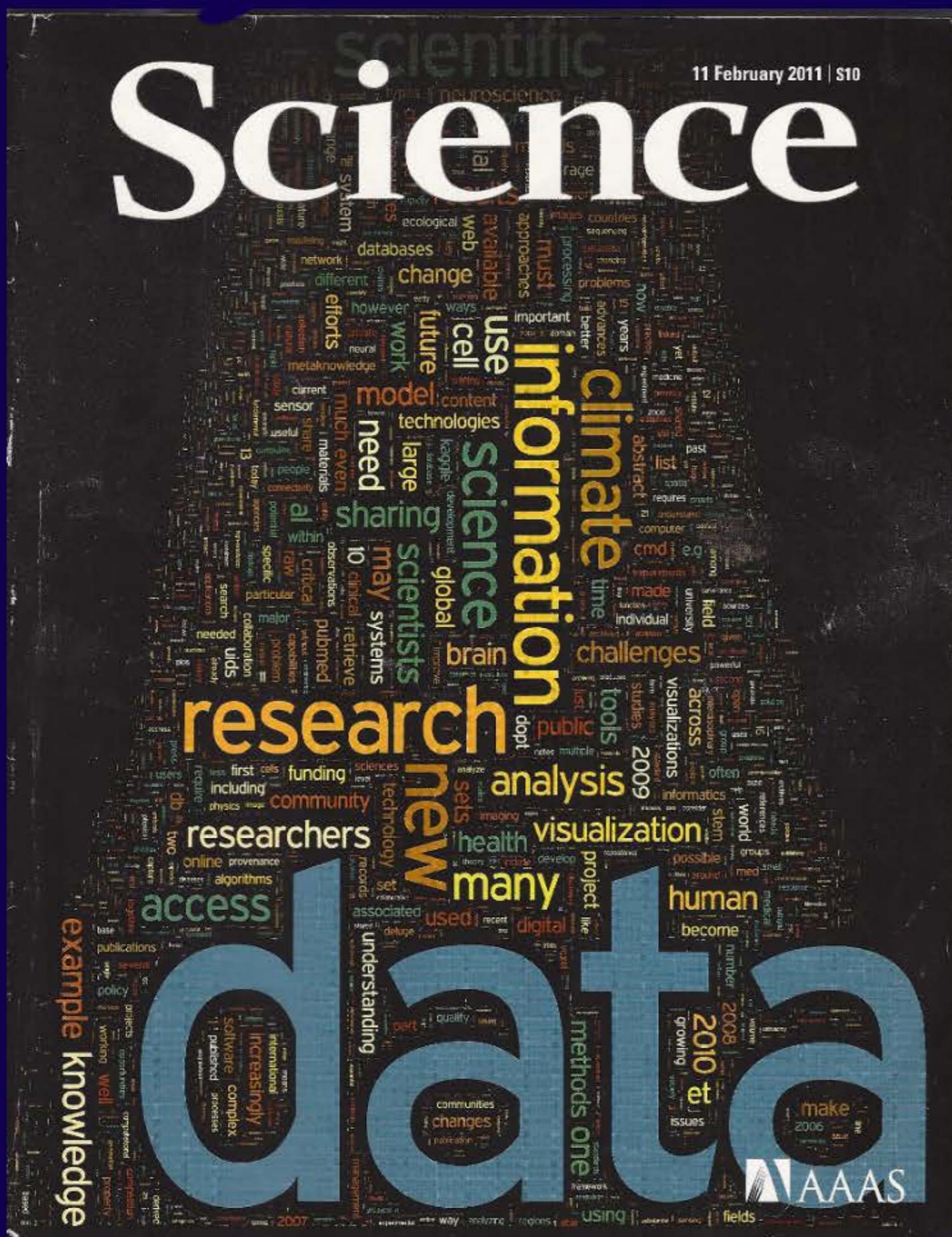
The "Data" Dust-up

Science again:
11 February 201



2a

The "Data" Dust-up

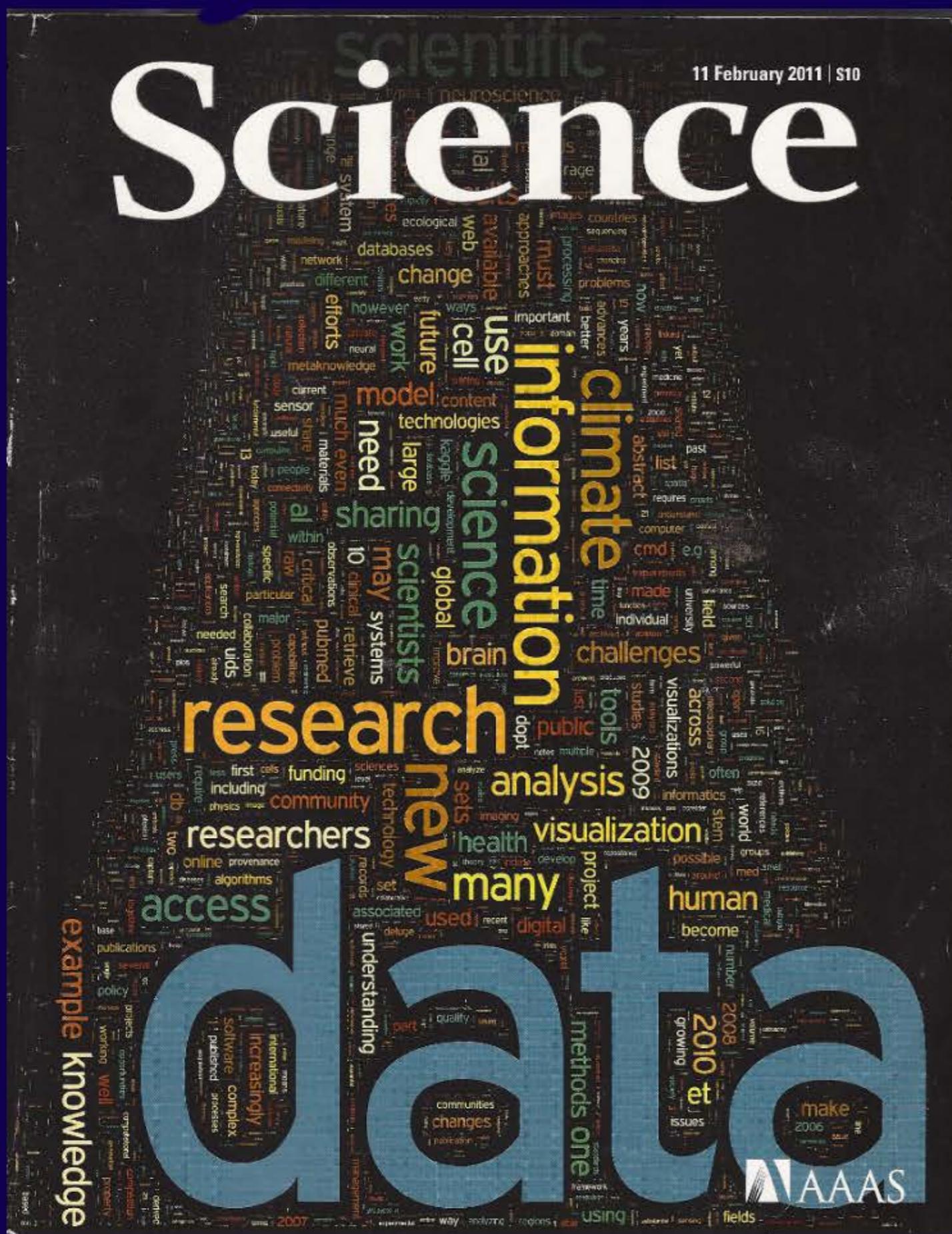


Science again: 11 February 2011

Full issue on Data
Early for 'Big Data'

2a

The "Data" Dust-up



Science again: 11 February 2011

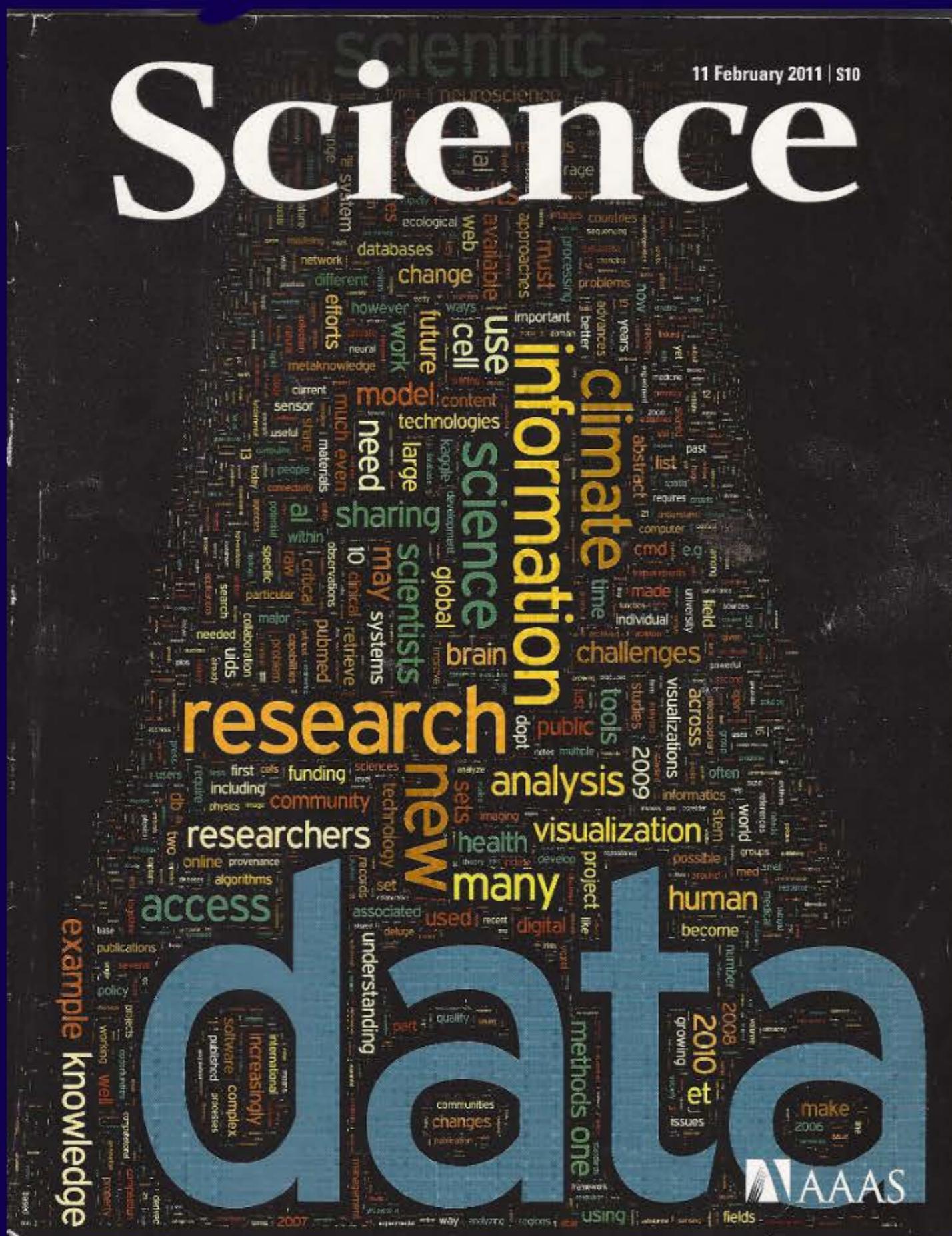
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But look at word cloud

"Statistics" doesn't appear!

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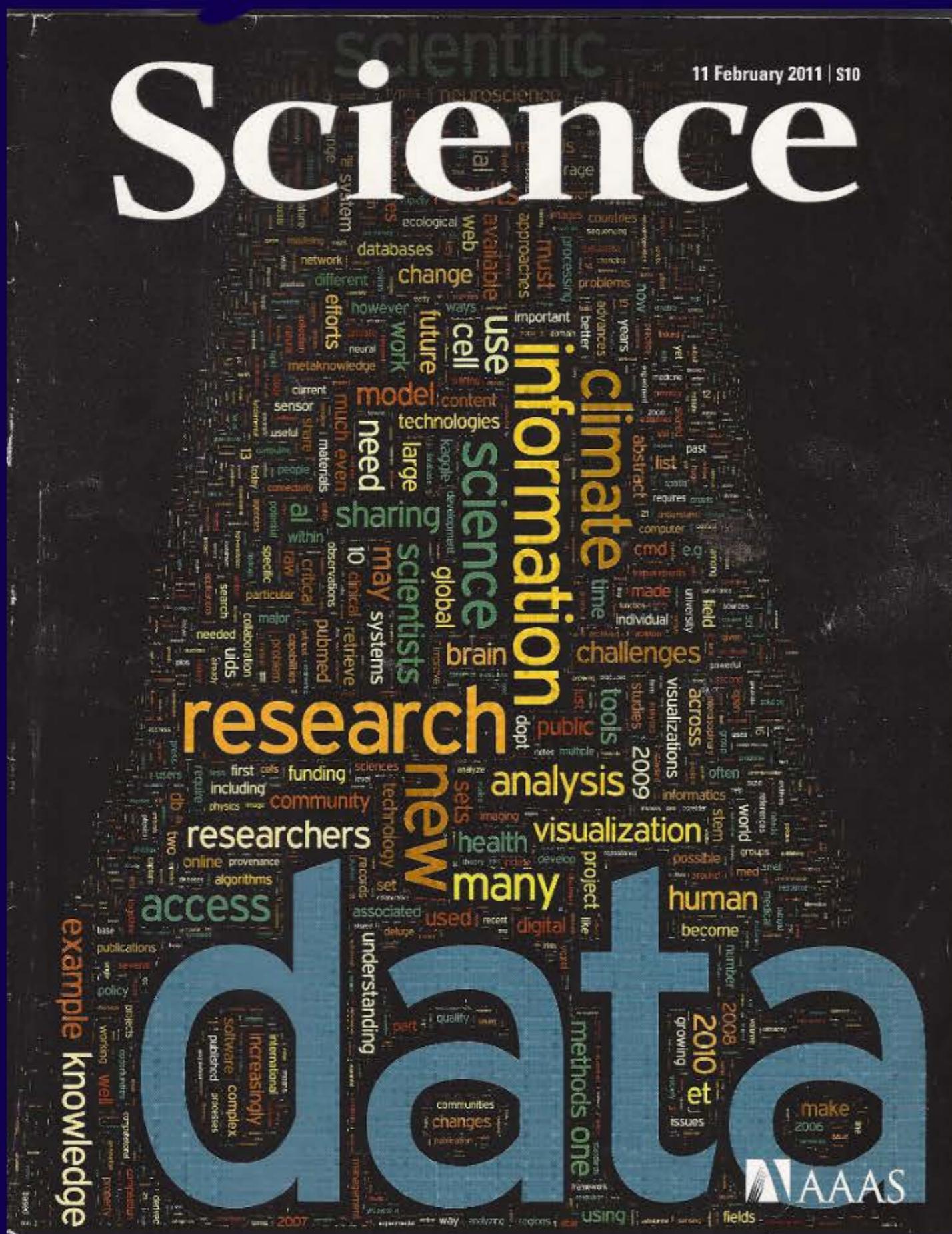
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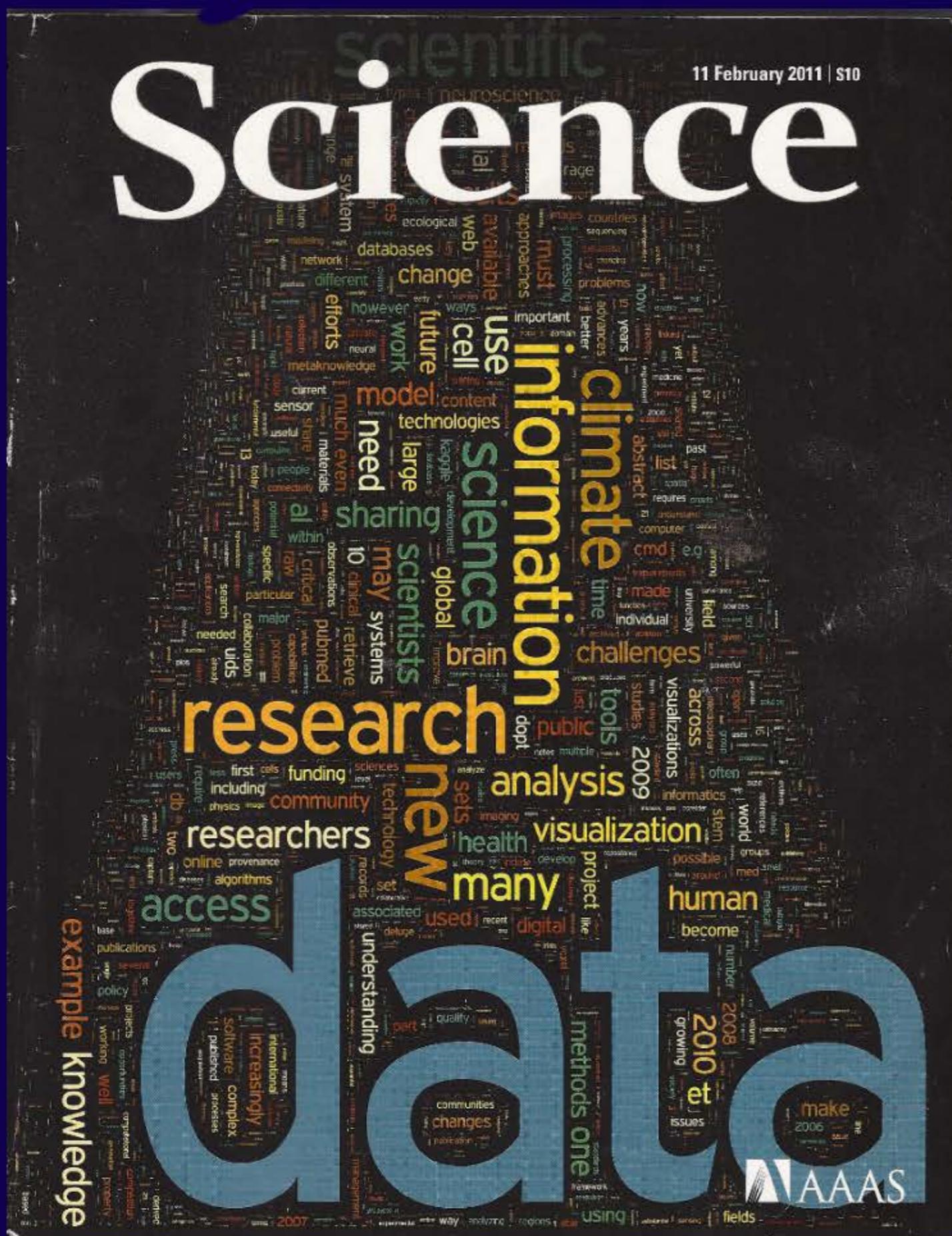
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Two senior statisticians complained to Science/AAAS

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A dismissive response

"You statisticians have an
Image problem!"

②b

"Retraction"

But the article on

Reproducibility

②b

"Retraction"

but the article on Reproducibility

also mentioned Retraction

No Journal likes to retract papers

②b

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So: they were rethinking their "dismissal of statistics"?

... they had their own Image problem!

②b

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So: they were rethinking their "dismissal of statistics"?

... they had their own Image problem!

... they were back-tracking fast !

②b

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... on high principle!

②b

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But...

③a

and in Statistics

... Statistics isn't immune to all of this!

③a

and in Statistics

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Statistics has two theories (^{Bayes}_{freq.}) and they are contradictory | xxcc =
| 266

③a)

and in Statistics

... Statistics isn't immune to all of this!

Statistics has two theories (^{Bayes}_{freq.}) and they are contradicting | xx^{xx}=
and no one cares ... | 266

③a

and in Statistics

... Statistics isn't immune to all this...

Statistics has two theories (^{Bayes}_{freq.}) and they are contradicting

| XXXX =
| 266

and no one cares ...

"we are just exploring..."

John Doyle, G&M, April 1

③a)

and in Statistics

... Statistics isn't immune to all this...

Statistics has two theories (^{Bayes}_{freq.}) and they are contradicting | xx^{xx} = 266

and no one cares ...

"we are just exploring..."

but sometimes it's for "real"

LHC; L'Aquila; VLOxx

③a) and in Statistics

... Statistics isn't immune to all this...

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| 266

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"we are just exploring..."

but sometimes it's for "real"

Should results "mean what they say?" Courts (legal)?

LHC; L'Aquila; VLO | ↑

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MATHEMATICS

Bayes' Theorem in the 21st Century

Bradley Efron

The term "controversial theorem" sounds like an oxymoron, but Bayes' theorem has played this part for two-and-a-half centuries. Twice it has soared to scientific celebrity, twice it has crashed, and it is currently enjoying another boom. The theorem itself is a landmark of logical reasoning and the first serious triumph of statistical inference, yet is still treated with suspicion by most statisticians. There are reasons to believe in the staying power of its current popularity, but also some signs of trouble ahead.

Here is a simple but genuine example of Bayes' rule in action (see sidebar) (1). A physicist couple I know learned, from sonograms, that they were due to be parents of twin boys.

Department of Statistics, Stanford University, Stanford, CA 94305, USA. E-mail: brad@stat.stanford.edu

Bayes' theorem plays an increasingly prominent role in statistical applications but remains controversial among statisticians.

They wondered what the probability was that their twins would be identical rather than fraternal. There are two pieces of relevant evidence. One-third of twins are identical; on the other hand, identical twins are twice as likely to yield twin boy sonograms, because they are always same-sex, whereas the likelihood of fraternal twins being same-sex is 50:50. Putting this together, Bayes' rule correctly concludes that the two pieces balance out, and that the odds of the twins being identical are even. (The twins were fraternal.)

Bayes' theorem is thus an algorithm for combining prior experience (one-third of twins are identical) with current evidence (the sonogram). Followers of Nate Silver's FiveThirtyEight Web blog got to see the rule in spectacular form during the 2012 U.S. presidential campaign: The algorithm updated prior poll results with new data on

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Bayes' 1763 paper was an impeccable exercise in probability theory. The trouble and the subsequent busts came from overenthusiastic application of the theorem in the absence of genuine prior information, with Pierre-Simon Laplace as a prime violator. Suppose that in the twins example we lacked the prior knowledge that one-third of twins are identical. Laplace would have assumed a uniform distribution between zero and one for the unknown prior probability of identical twins, yielding 2/3 rather than 1/2 as the answer to the physicists' question. In modern parlance, Laplace would be trying to assign an "uninformative prior" or "objective prior" (2), one having only neutral effects on the output of Bayes' rule (3). Whether or not this

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empirical

2) mathematical
"Pierre-Simon Laplace"

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Support | genuine priors (1)

Avoid | Laplace & Opinion (2) (3)

(36) "Laplace had confidence":

Science

Sept 2013

Fraser

LETTERS

Low Marks for Education Funding Priorities

ANYONE INVOLVED SUBSTANTIVELY IN SCIENCE education during the past five decades will see the irony in the decision by the Office of Management and Budget (OMB) to trim the federal government's science, technology, engineering, and mathematics (STEM) programs on the grounds that many of them lack evaluation data on efficacy ("An invisible hand behind plan to realign U.S. science education," J. Mervis, *News Focus*, 26 July, p. 338). Although federal funding often supported formative evaluation (assessment in the pilot phase to improve the program itself) during the development of new curricula, it was virtually impossible to secure funding for summative evaluation (assessment of effectiveness after implementation) because of the costs and time frames involved. At the Biological Sciences Curriculum Study (*J*), where the value of summative evaluation always has been self-evident, we often lamented that the federal government funded a series of 90-meter dashes, supporting development of new instructional materials but not their evaluation. Funding from the Institute for Education Sciences for efficacy trials (*2*) that provide one type of summative evaluation constitutes some progress, but it is not enough.

It is perverse for OMB to blame STEM projects for deficiencies that were inherent in the government's funding priorities. Perhaps an evaluation of those priorities is in order.

JOSEPH D. MCINERNEY
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References

1. Biological Sciences Curriculum Study (www.bscs.org).
2. J. K. Sprybrook, S. W. Raudenbush, *Educ. Eval. Pol. Anal.* **31**, 298 (2009).

Bayes' Confidence

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The Wellcome Trust's Basic Scientist Career Tracker (*J*) demonstrates the disproportionate number of women exiting academia early in their careers. Although an academic research career brings rewards, it remains a risky long-term career choice (*2*), and as McNutt describes, childbearing years typically coincide with the time when a faculty member needs to build a strong portfolio and gain tenure, thereby securing a less risky future.

Academia needs to attract and retain high-quality, highly trained researchers; research funders such as the Wellcome Trust can play an important role by following these steps: (i) Funders need to ensure that career awareness and mentorship are integral components of their training provision; (ii) Funders must ensure that their eligibility and/or funding guidelines do not discriminate against certain researchers (for example, a bias in funding decisions toward grant applications that include a move between institutions may inadvertently discriminate against those with established local ties); (iii) Funders need to promote and develop opportunities for researchers to use their funding flexibly, including options for career breaks, reentry fellowships, opportunities to work in posts other than as a principal investigator, and part-time schedules; (iv) We need to expand the opportunities for female role models working across academia to tell their story; this should be a core component of training programs.

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DAVID LYNN
Strategic Planning and Policy Unit, Wellcome Trust, London, NW1 2BE, UK.

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Reference

1. Wellcome Trust, "Wellcome Trust Basic Science Career Tracker: Results of Wave 4 (2012)" (2013); www.wellcome.ac.uk/funding/Biomedical-science/Career-tracker/Basic-tracker/index.htm.
2. Ipsos MORI, "Risks and rewards: How PhD students choose their careers" (Ipsos MORI, London, 2013).

CORRECTIONS AND CLARIFICATIONS

This Week in Science: "Pushy black hole" (6 September, p. 1041). The last line should be "possibly limiting star formation and galaxy growth" instead of "possibly contributing to star formation and galaxy growth." The HTML and PDF versions online have been corrected.

Reports: "Pandoravirus: Amoeba viruses with genomes up to 2.5 Mb reaching that of parasitic eukaryotes" by N. Philippe et al. (19 July, p. 281). In the first sentence of the legend to Fig. 1, the "1D" and "1D" should not have been italicized, as they refer to panels A1/A2 and B1/B2 and not to references 1 and 2. In the legend to Fig. 1E, the "a" and "b" labels should have been transposed. In addition, a reference to panels B1 and B2 is now included. In the acknowledgments, the GenBank accession numbers were incorrectly listed. They should read KC977571 and KC977570 (not KC977471 and KC977470). Also, the financial support of the Provence-Côte-d'Azur Région was missing. The HTML and PDF versions online have been corrected.

Letters to the Editor

Letters (~300 words) discuss material published in *Science* in the past 3 months or matters of general interest. Letters are not acknowledged upon receipt. Whether published in full or in part, Letters are subject to editing for clarity and space. Letters submitted, published, or posted elsewhere, in print or online, will be disqualified. To submit a letter, go to www.submit2science.org.

(36) Laplace had confidence:

Science

Sept 2013

Fraser

Recall:

Efron in Science \downarrow

Priors $\pi(\theta)$

1) frequency
empirical

Efron

Genuine

2) mathematical
"Pierre-Simon Laplace" "trouble"

3) opinion "trouble"

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NEITHER THE PERSPECTIVE "BAYES' THEOREM in the 21st century" (B. Efron, 7 June, p. 1177) nor the responding Letter "A statistically significant future for Bayes' rule" (R. van Hulst, 26 July, p. 343) refer to the mystical flavor often associated with Bayes in their discussions of the theorem's popularity.

Bayes had a propensity to use names that suggest something more than what is directly being described. For example, "Bayes' rule" is just conditional probability applied in a specialized context. The "controversial theorem" is nothing more than a formula for conditional probability.

"Perhaps more disconcerting in Bayes is the term "objective prior" for the uninformative priors used by Pierre-Simon Laplace. Such priors, of course, are just imagined; they are not in fact objective themselves, but rather aim to produce objective conclusions. Indeed, many of Laplace's calculations of posterior probability using uninformative priors are numerically equal to frequentist calculations of confidence."

van Hulst mentions that the life sciences need a "synthesis of multiple categories of evidence." Certainly Bayes provides a simple and accessible means of combining different data results: Just multiply the likelihoods together. But this option is also available to the frequentist: Just combine the likelihoods and ignore what's left. The typical frequentist, however, realizes that this method would lose information and is unwilling to make this tradeoff for simplicity. Thus, he would choose an exact confidence interval when available.

Treating Bayes as a route to approximate confidence could go a long way toward resolving the presence of two theories in statistical inference.

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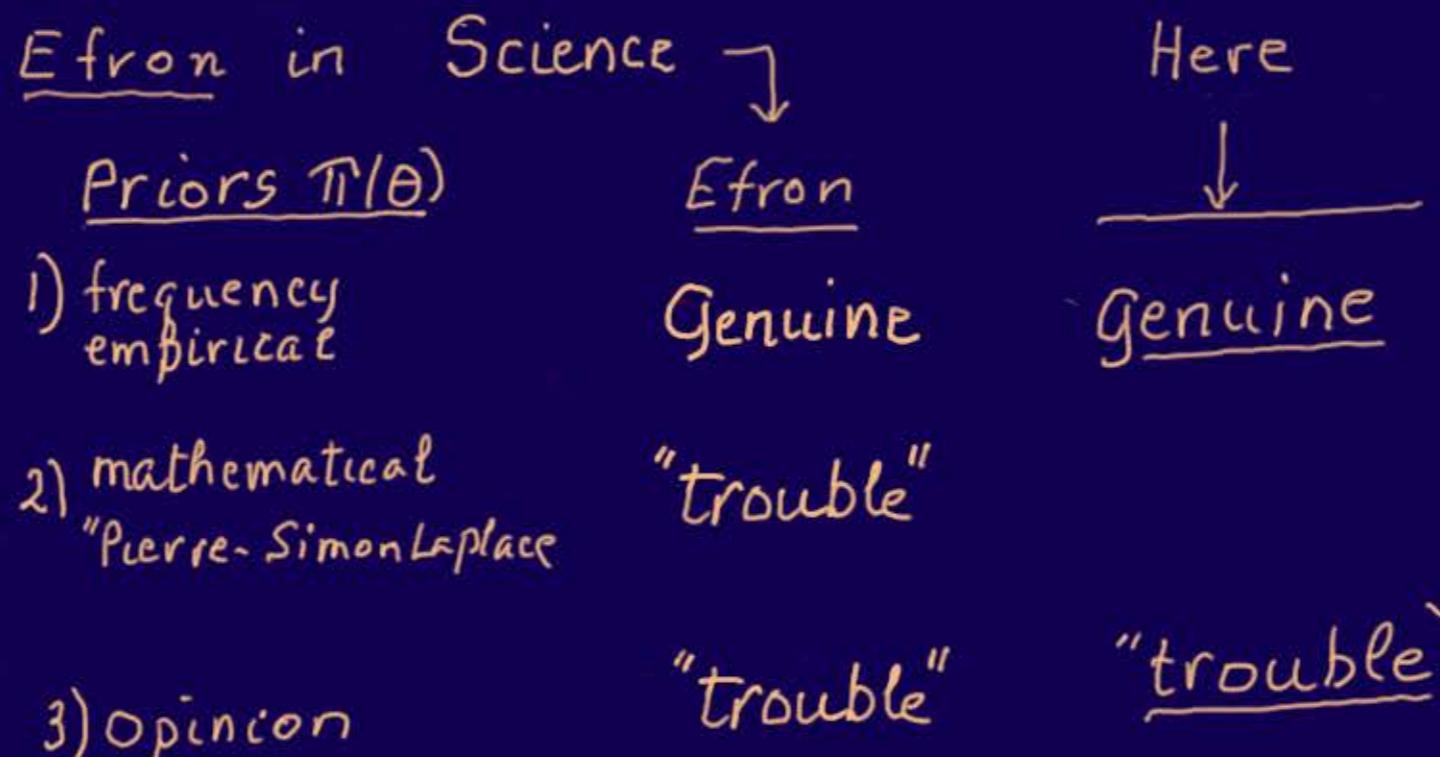
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3) opinion

Efron

Genuine

"trouble"

Here

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"trouble"

Summary:

1) If you have opinion, let's hear it!

but don't use it to analyze data!

otherwise
"misconduct"!

Display it in parallel. Let user see both (opinion separately
analysis

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"Laplace" | o/w "trouble"

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Bayes' Confidence

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Bayes had a propensity to use names that suggest something more than what is directly being described. For example, "Bayes' rule" is just conditional probability applied in a specialized context. The "controversial theorem" is nothing more than a formula for conditional probability.

Perhaps more disconcerting in Bayes is the term "objective prior" for the uninformative priors used by Pierre-Simon Laplace. Such priors, of course, are just imagined; they are not in fact objective themselves, but rather aim to produce objective conclusions. Indeed, many of Laplace's calculations of posterior probability using uninformative priors are numerically equal to frequentist calculations of confidence.

van Hulst mentions that the life sciences need a "synthesis of multiple categories of evidence." Certainly Bayes provides a simple and accessible means of combining different data results: Just multiply the likelihoods together. But this option is also available to the frequentist: Just combine the likelihoods and ignore what's left. The typical frequentist, however, realizes that this method would lose information and is unwilling to make this tradeoff for simplicity. Thus, he would choose an exact confidence interval when available.

Treating Bayes as a route to approximate confidence could go a long way toward resolving the presence of two theories in statistical inference.

D. A. S. FRASER

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Efron in Science \rightarrow

Priors $\pi(\theta)$

1) frequency
empirical

2) mathematical
"Pierre-Simon Laplace"

3) opinion

Efron

Genuine

"trouble"

Here

↓

genuine

"trouble" \Rightarrow - Laplace (when reproducible) ✓
- otherwise "trouble" ✗

"trouble"

Summary:

1) If you have opinion, let's hear it!
but don't use it to analyze data! ↴ "misconduct!"
Display it in parallel. Let user see both (opinion separately
analysis)

2) Bayes mathematical | OK if reproducible (Laplace had confidence first)
"Laplace" | o/w "trouble"

3) Question: How to get reproducible Bayes? (confidence!)
How to get Objective (truly) Bayes

③b Laplace had confidence:

Science

Sept 2013

Fraser

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✗

✓

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How to get Objective (truly) Bayes

(Term is already in use "without reproducibility" in B-community)

④a) Can Bayes/Jeffreys give reproducible inference ?

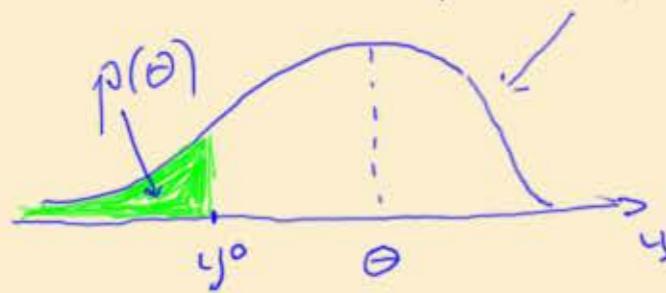
Try: Scalar $f(y-\theta)$... very simple case

④a) Can Bayes/Jeffreys give reproducible inference ?

Try: Location, Scalar $f(y - \theta)$

1 The Case for Bayes

a) Motivating example $f(y - \theta)$: Data y^* ; assess θ



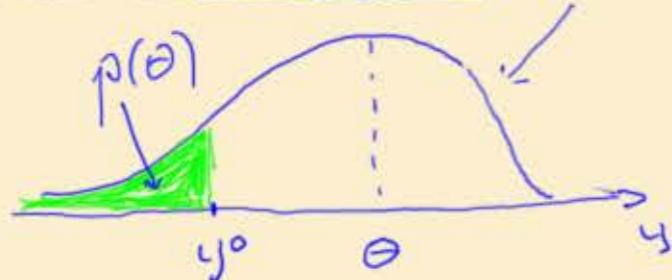
$$\begin{aligned}\text{Observed p-value} &= p^*(\theta) = \int_{-\infty}^{y^*} f(y - \theta) dy \\ &= \% \text{age position of } y^* \text{ re } \theta = F^*(\theta)\end{aligned}$$

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Try: Location, Scalar $f(y - \theta)$

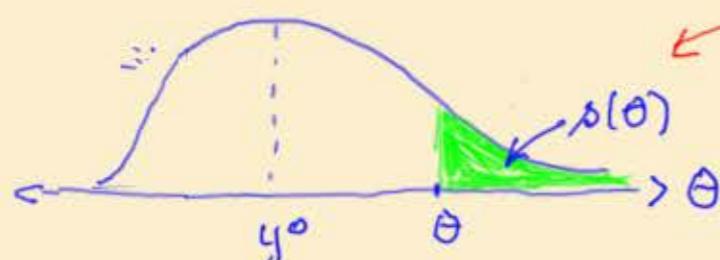
1 The Case for Bayes - Location

a) Motivating example $f(y - \theta)$: Data y^* ; assess θ



$$\text{Observed p-value} = p^*(\theta) = \int_{-\infty}^{y^*} f(y - \theta) dy \\ = \% \text{age position of } y^* \text{ re } \theta$$

b) Bayes (Location / flat prior): pdf at y^* ; flipped



$$\pi(\theta|y^*) = 1 \cdot f(y^* - \theta)$$

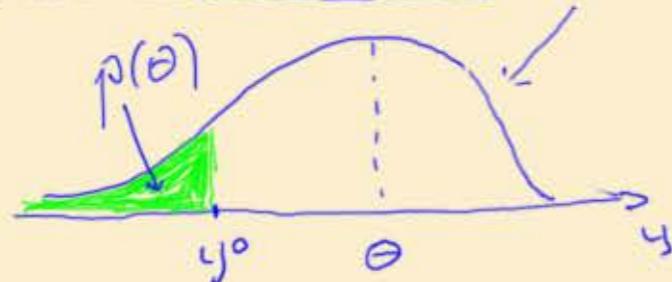
$$\text{Bayes survival} = S(\theta) = \int_{\theta}^{\infty} 1 \cdot f(y^* - \theta) d\theta$$

④a) Can Bayes/Jeffreys give reproducible inference?

Try: Location, Scalar $f(y - \theta)$ Bayes give confidence

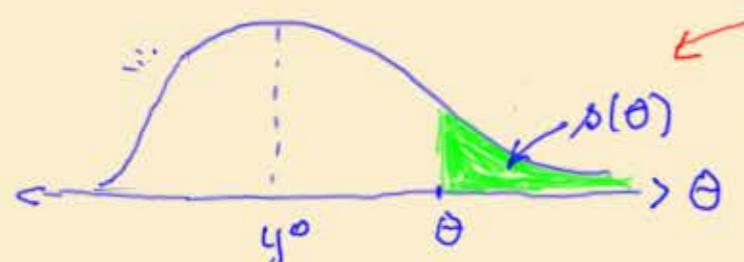
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c) Thus: $p(\theta) = s(\theta)$... "Reflection" Just a "calculus recalculation"

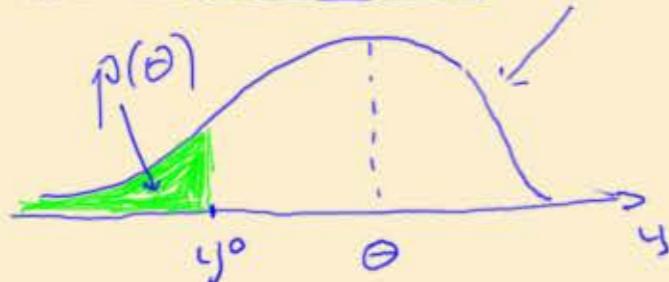
Bayes here gives confidence!

④a) Can Bayes/Jeffreys give reproducible inference?

Try: Location, Scalar $f(y - \theta)$ Bayes give confidence

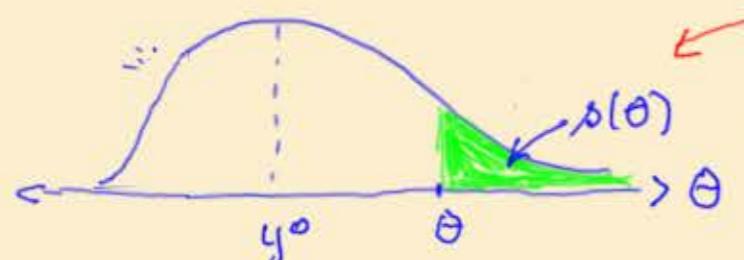
1 The Case for Bayes - Location

a) Motivating example $f(y - \theta)$: Data y^* ; assess θ



$$\text{Observed p-value} = p^*(\theta) = \int_{-\infty}^{y^*} f(y - \theta) dy \\ = \% \text{ age position of } y^* \text{ re } \theta$$

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Bayes here gives confidence

Does this generalize?

④b) Scalar parameter regular ... $f(y; \theta)$

④b) Scalar parameter regular ... $f(y; \theta)$

i) Can always be rewritten as exponential (3rd order)

$$f(y; \theta) = \exp\{\varphi A - k(\varphi)\} h(y) \quad \text{Likelihood, asymptotics} \quad \mathcal{O}(n^{-\frac{3}{2}})$$

④b Scalar parameter regular ... $f(y; \theta)$

i) Can always be rewritten as exponential (3rd order)

$$f(y; \theta) = \exp\{\varphi s - k(\varphi)\} h(y)$$

2) Can always be standardized $\mathcal{O}(n^{-1})$

$$f(s; \varphi) = \exp\left\{\varphi s - \frac{\varphi^2}{2} - \gamma \frac{\varphi^3}{6n^2}\right\} h(s) \quad \text{Taylor 2nd order}$$

(4b)

Scalar parameter regular ... $f(y; \theta)$

1) Can always be rewritten as exponential (3rd order)

$$f(y; \theta) = \exp\{\varphi s - k(\varphi)\} h(y)$$

2) Can always be standardized

$$\begin{aligned} f(s; \varphi) &= \exp\left\{\varphi s - \frac{\varphi^2}{2} - \gamma \frac{\varphi^3}{6n^{1/2}}\right\} h(s) \\ &= \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{(s-\varphi)^2}{2} - \gamma \frac{\varphi^3}{6n^{1/2}} + \gamma \frac{\varphi^3}{6n^{1/2}}\right\} \frac{1}{(1 - \gamma s / 2n^{1/2})} \end{aligned}$$

Expanded
Determined by "pdf"

(4b)

Regular, Scalar parameter $f(y; \theta)$

1) Can always be rewritten as exponential (3rd order)

$$f(y; \theta) = \exp\{\varphi s - k(\varphi)\} h(y)$$

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$$\begin{aligned} f(s; \varphi) &= \exp\left\{ \varphi s - \frac{\varphi^2}{2} - \gamma \frac{\varphi^3}{6n^{1/2}} \right\} h(s) \\ &= \frac{1}{\sqrt{2\pi}} \exp\left\{ -\frac{(s-\varphi)}{2} - \gamma \frac{\varphi^3}{6n^{1/2}} + \gamma \frac{\varphi^2}{2n^{1/2}} \right\} (1 - \gamma s / 2n^{1/2}) \end{aligned}$$

3) Can use a constant-info parameter $\beta = \varphi + \gamma \varphi^2 / 2n^{1/2}$

(4b)

Regular, Scalar parameter $f(y; \theta)$

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3) Can use a constant-info parameter $\beta = \varphi + \gamma \frac{\varphi^2}{2n^{1/2}}$

$$f(\hat{\beta}; \beta) = \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{(\hat{\beta}-\beta)^2}{2} - \gamma \frac{(\hat{\beta}-\beta)^3}{6n^{1/2}}\right\} \cdot d\hat{\beta} \quad \text{Rewrite as } \underline{\text{location}}$$

(4b)

Regular, Scalar parameter $f(y; \theta)$

1) Can always be rewritten as exponential (3rd order)

$$f(y; \theta) = \exp\{\varphi s - k(\varphi)\} h(y)$$

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$$\begin{aligned} f(s; \varphi) &= \exp\{\varphi s - \varphi^2/2 - \gamma \varphi^3/6n^{1/2}\} h(s) \\ &= \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{(s-\varphi)}{2} - \gamma \varphi^3/6n^{1/2} + \gamma \varphi^2/2n^{1/2}\right\} (1 - \gamma s/2n^{1/2}) \end{aligned}$$

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4) Jeffreys "automatic" for location model: posterior is

$$f(\beta; \hat{\beta}) = \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{(\hat{\beta}-\beta)^2}{2} - \gamma (\hat{\beta}-\beta)^3/6n^{1/2}\right\} \cdot d\beta$$

(4b)

Regular, Scalar parameter $f(y; \theta)$

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$$f(y; \theta) = \exp\{\varphi s - k(\varphi)\} h(y)$$

2) Can always be standardized

$$\begin{aligned} f(s; \varphi) &= \exp\{\varphi s - \varphi^2/2 - \gamma \varphi^3/6n^{1/2}\} h(s) \\ &= \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{(s-\varphi)}{2} - \gamma \varphi^3/6n^{1/2} + \gamma \delta^3/6n^{1/2}\right\} (1 - \gamma s^2/2n^{1/2}) \end{aligned}$$

3) Can use a constant-info parameter $\beta = \varphi + \varphi^2/2n^{1/2}$

$$f(\hat{\beta}; \beta) = \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{(\hat{\beta}-\beta)^2}{2} - \gamma (\hat{\beta}-\beta)^3/6n^{1/2}\right\} \cdot d\hat{\beta}$$

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$$f(\beta; \hat{\beta}) = \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{(\hat{\beta}-\beta)^2}{2} - \gamma (\hat{\beta}-\beta)^3/6n^{1/2}\right\} \cdot d\beta$$

5) Reexpress $d\beta = (1 + \gamma \varphi/n^{1/2}) d\varphi$

$$= \text{Likelihood} \cdot \underbrace{(\text{root info})}_{\text{Jeffreys}} \cdot d\varphi$$

- Rewrite posterior differential
- 2nd order
- \Rightarrow pure confidence

(4b)

Regular, Scalar parameter $f(y; \theta)$

1) Can always be rewritten as exponential (3rd order)

$$f(y; \theta) = \exp\{\varphi s - k(\varphi)\} h(y)$$

2) Can always be standardized $O(n^{-1})$

$$\begin{aligned} f(s; \varphi) &= \exp\{\varphi s - \varphi^2/2 - \gamma \varphi^3/6n^{1/2}\} h(s) \\ &= \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{(s-\varphi)}{2} - \gamma \varphi^3/6n^{1/2} + \gamma \varphi^2/6n^{1/2}\right\} (1 - \gamma \varphi^2/2n^{1/2}) \end{aligned}$$

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5) Reexpress $d\beta = (1 + \gamma \varphi/n^{1/2}) d\varphi$

$$= \text{Likelihood} \cdot (\text{root info}) \cdot d\varphi$$

Jeffreys is 2nd order Accurate Welch Peers 1963
 is reproducible Brown Cai DasGupta 2001 from Bayes to Jeffreys

Scalar parameter model:

Example 1

Setup

- Model: $Y \sim \text{Gamma}(\alpha, \beta)$ where α is shape and β is rate

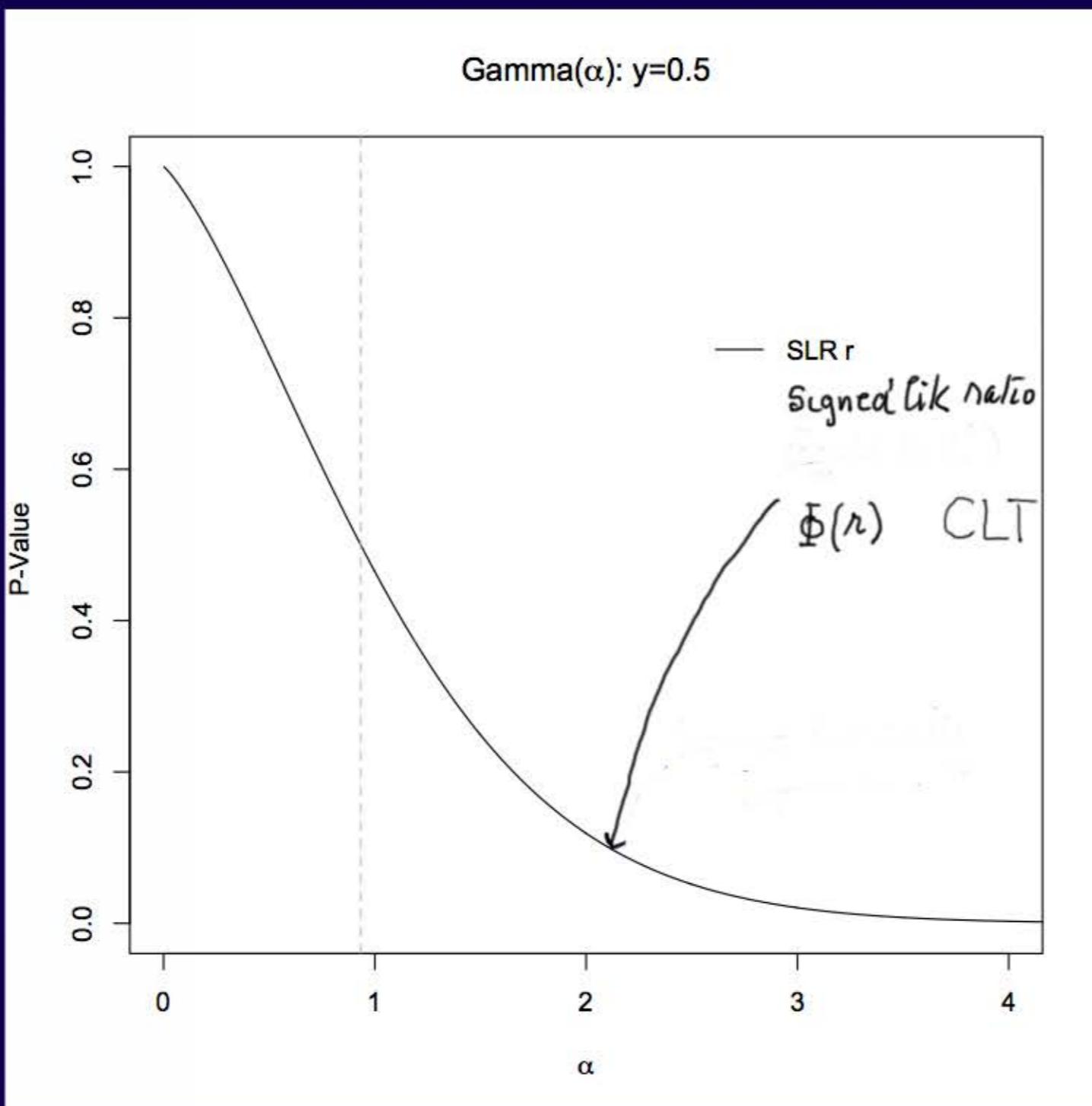
pdf is $\frac{\beta^\alpha}{\Gamma(\alpha)} y^{\alpha-1} e^{-\beta y}$

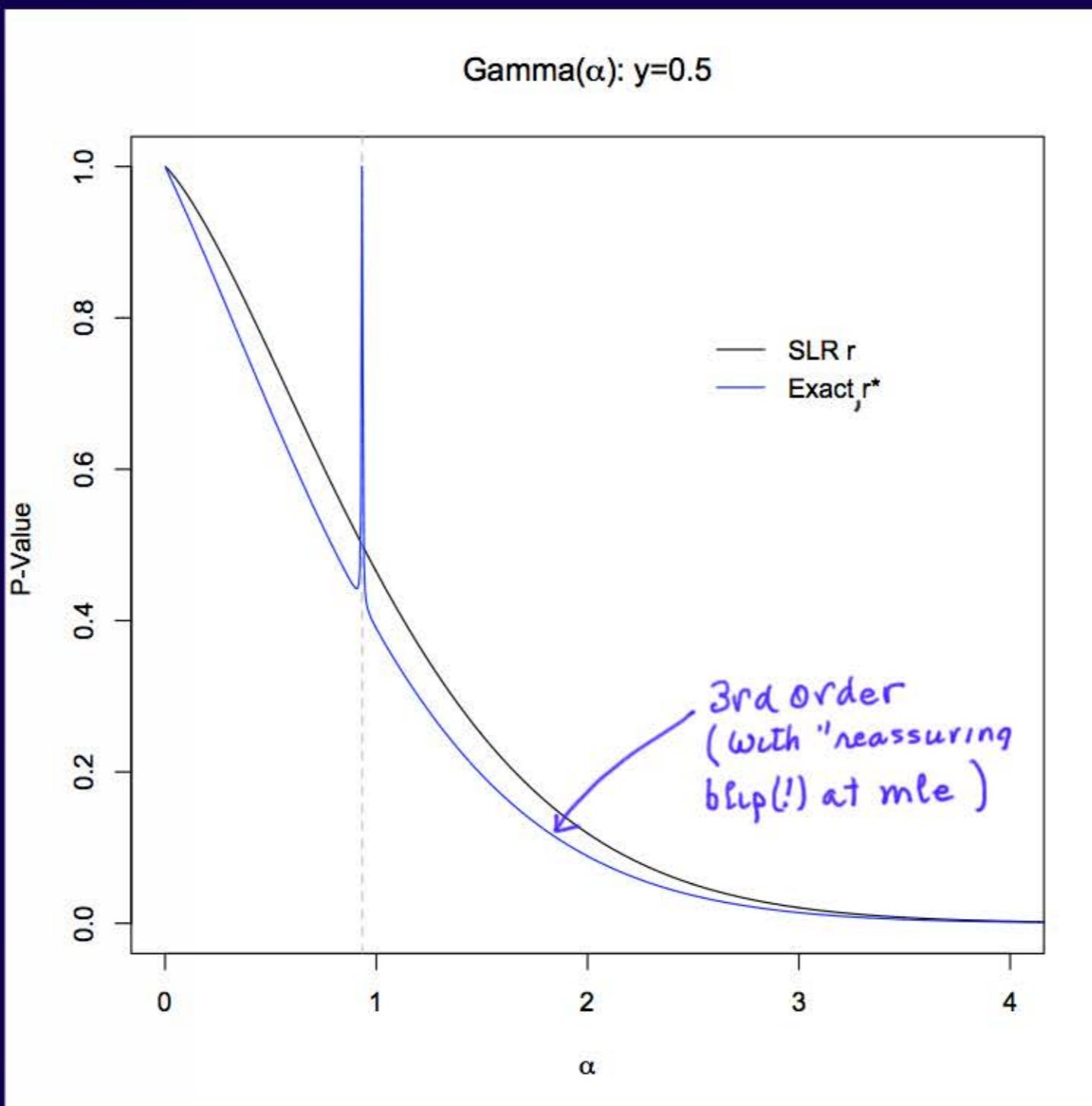
Let $n = 1$ and data is $y^o = .5$

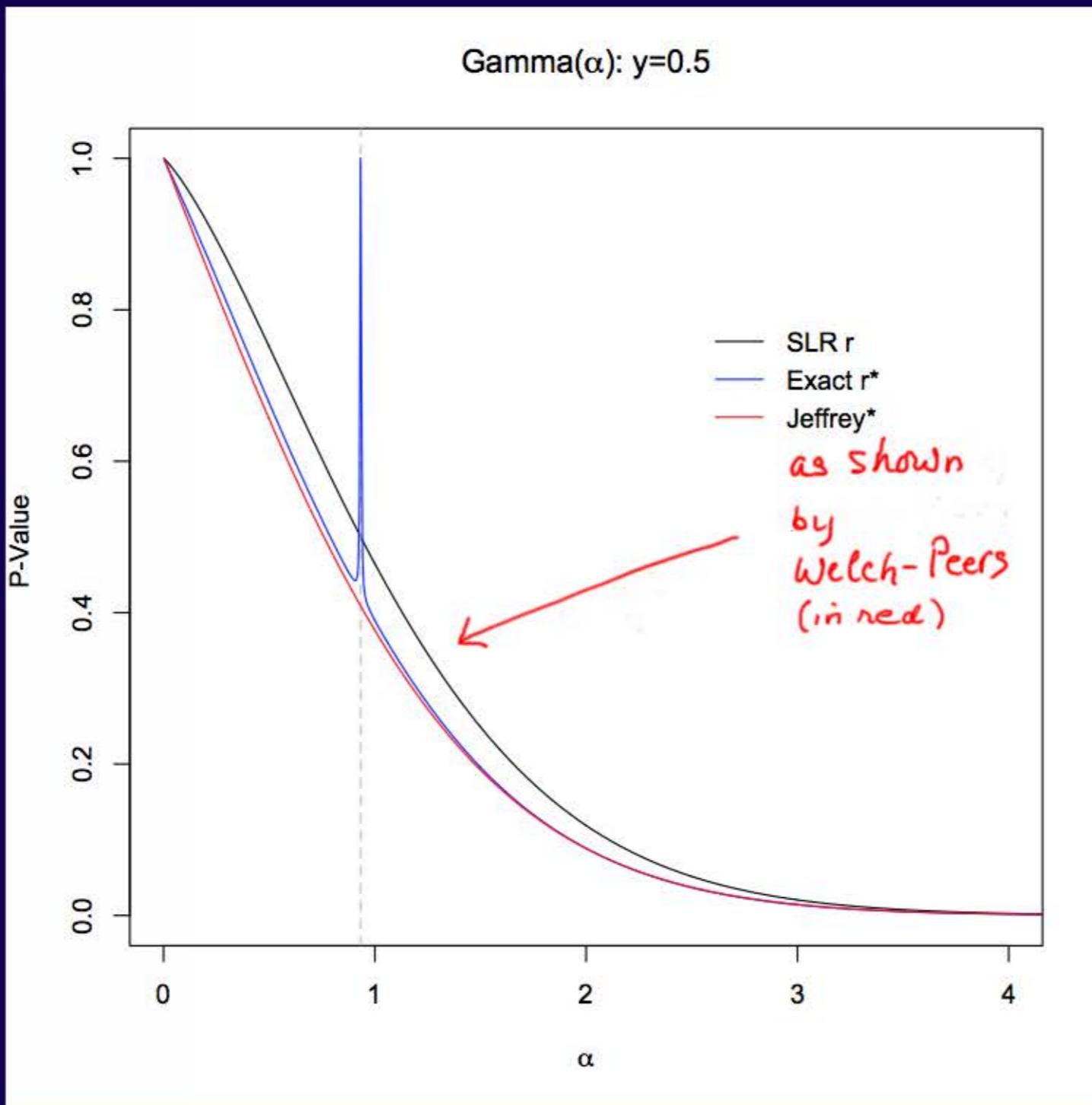
Parameter of interest is α

Fix $\beta = 1$

$$\Gamma'(\alpha) y^{\alpha-1} e^{-y} \cdot dy$$
$$y^o = .5$$







reproducible!

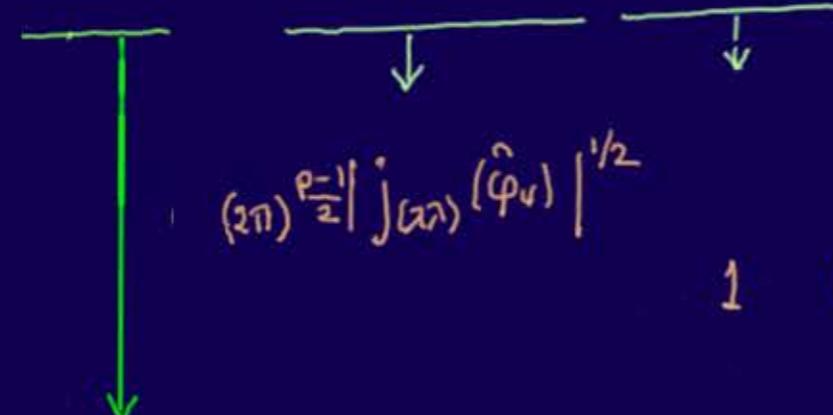
④c) Regular Statistical model: $f(y; \varphi)$ φ canonical; $\psi(\varphi)$ = scalar interest

1) Can always be re-written as exponential model: likelihood asymptotics
re-parameterization

$$f(y; \varphi) = \exp\{\varphi's - k(\varphi)\} h(y)$$

2) Can always be separated: $f(s; \varphi) \cdot g(t | s; \varphi, \lambda)$

$$= f(s; \varphi) \cdot \frac{|J_{\lambda\lambda}(\hat{\varphi}_s)|^{-1/2}}{(2\pi)^{p-1}} \exp\{\ell - \tilde{l}\} \quad \text{re } ds dt \quad \text{Saddle point analysis}$$

3) Prior to eliminate  direct multiplication
2nd factor $(2\pi)^{p-1} |J_{\lambda\lambda}(\hat{\varphi}_s)|^{1/2}$
3rd factor 1 on profile
or SP integration of (λ) given φ
re Exp model form for (λ)
over section fixed φ

4) Prior to calibrate via W-P 1st factor $|J_{\lambda\lambda\lambda}(\varphi)|^{1/2}$ Welch-Peers plus Lik. Asy.

5) Combine (1st 2nd) $\frac{|J_{\varphi\varphi}(\hat{\varphi}_s)|^{1/2} |J_{\lambda\lambda}(\hat{\varphi}_s)|^{1/2}}{|J_{\varphi\varphi}(\hat{\varphi}_s)|^{1/2} |J_{\lambda\lambda}(\hat{\varphi}_s)|^{1/2}}$ into φ at $\hat{\varphi}_s$
" " x "

6) 3rd factor eliminates λ by Laplace
(also by full Bayes)

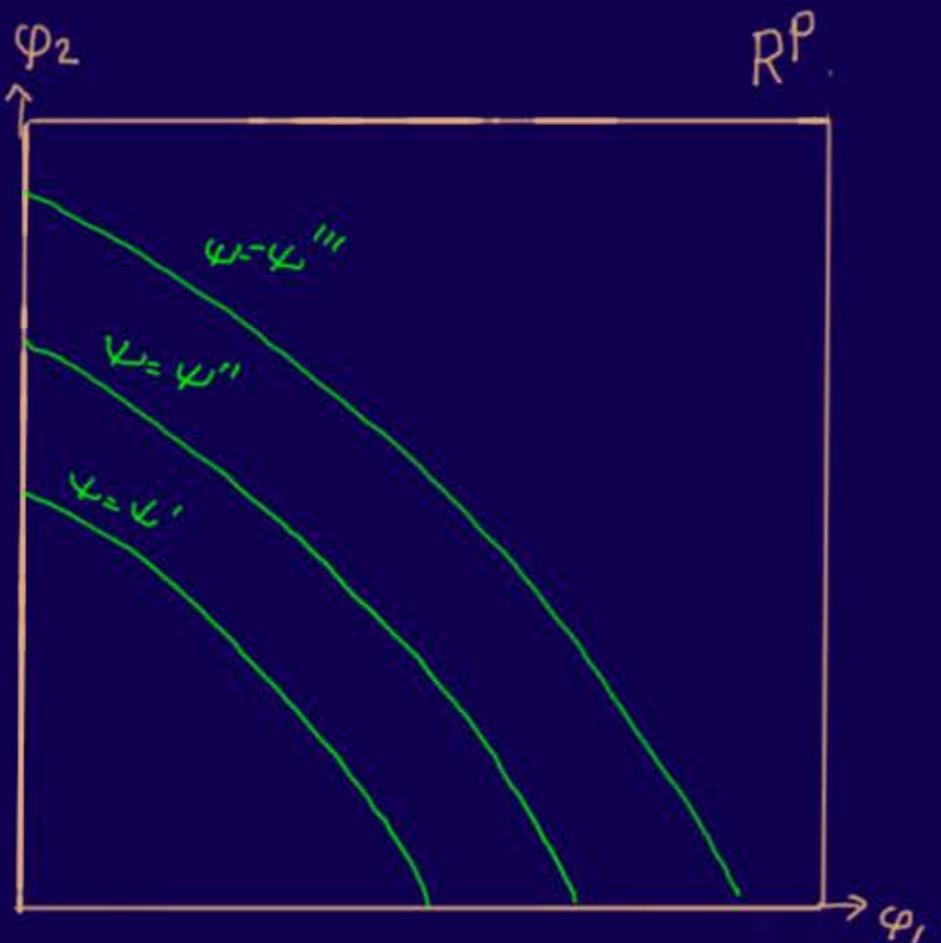
5) Gives full Jeffreys $|J_{\varphi\varphi}(\hat{\varphi}_s)|^{1/2}$ on profile(φ) + "curvature" correction

Cases: a) Linear φ b) Rotating φ c) Curved φ & calculate $d(\varphi)$

The geometry:

1 Parameter space
(canonical) $\bar{\Phi}$

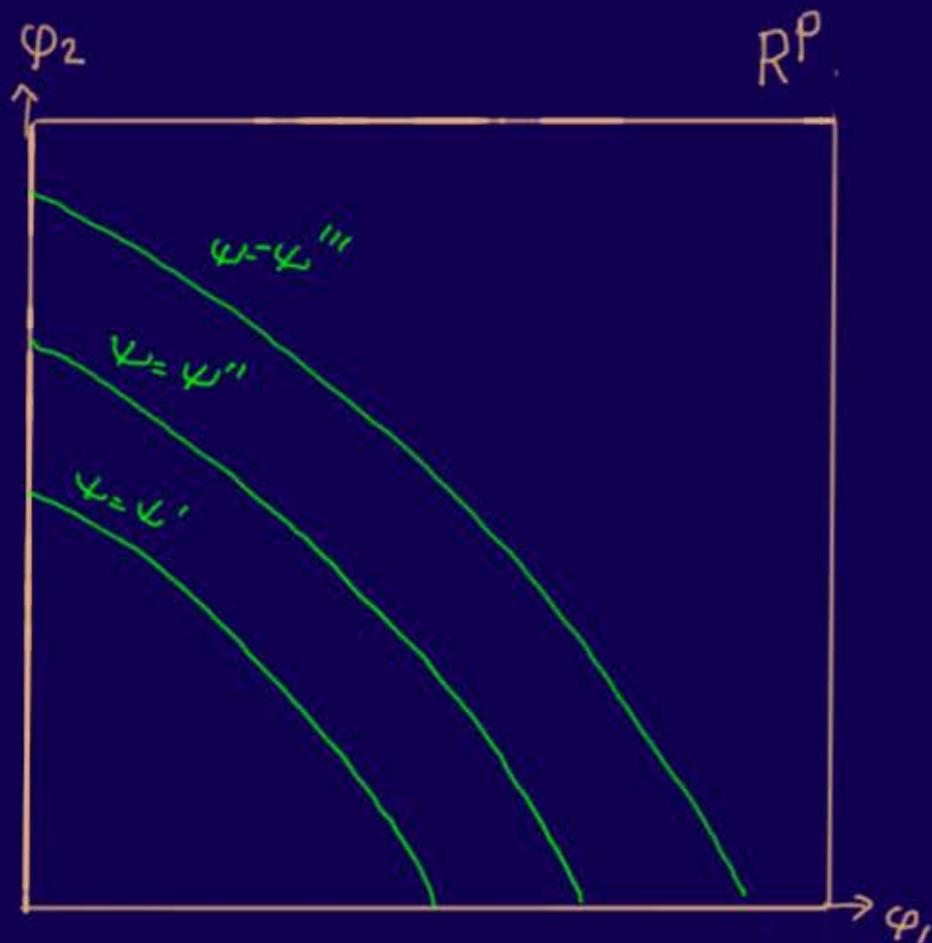
Interest $\psi(\varphi)$



The geometry:

1 Parameter space
(canonical) $\bar{\Phi}$

Interest $\psi(\varphi)$



Jeffreys (usual)

$$\text{Likelihood} = L^\circ(\varphi)$$

$$\text{J. prior} = \pi(\varphi) = |J\varphi(\varphi)|^{1/2}$$

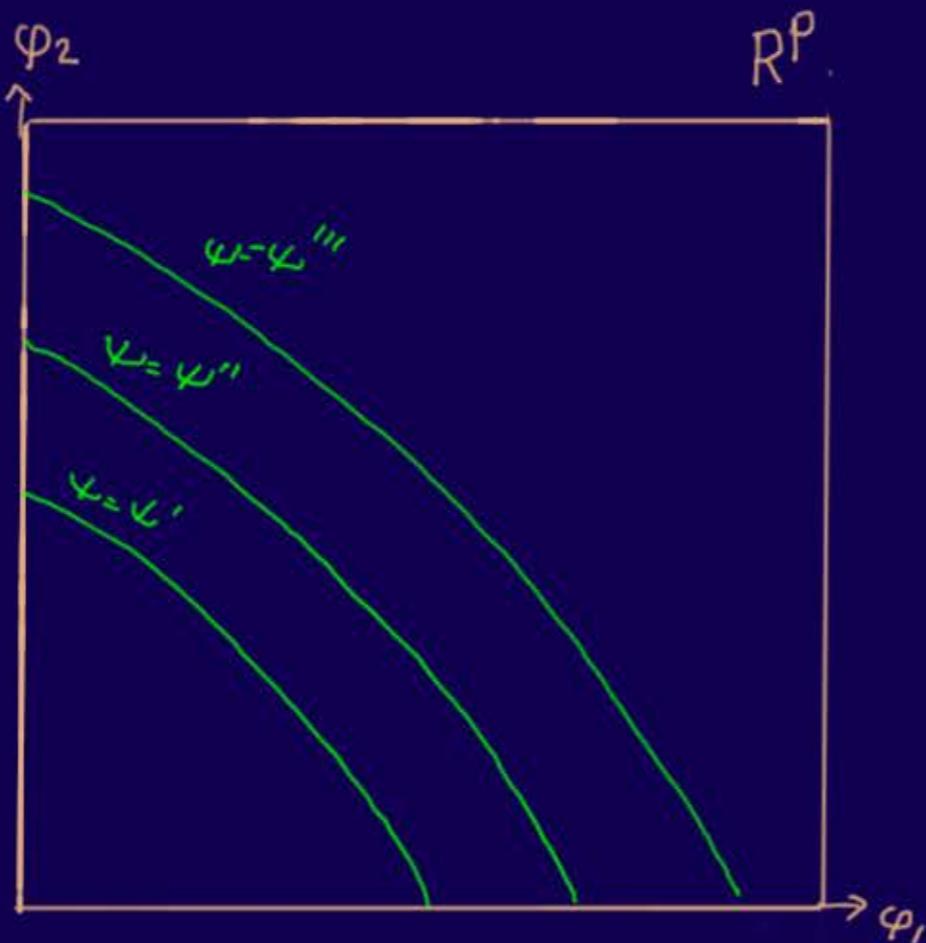
$$\text{Posterior} = L^\circ(\varphi) \pi(\varphi)$$

& integrate up to contour " "
for dist'n of φ But...

The geometry:

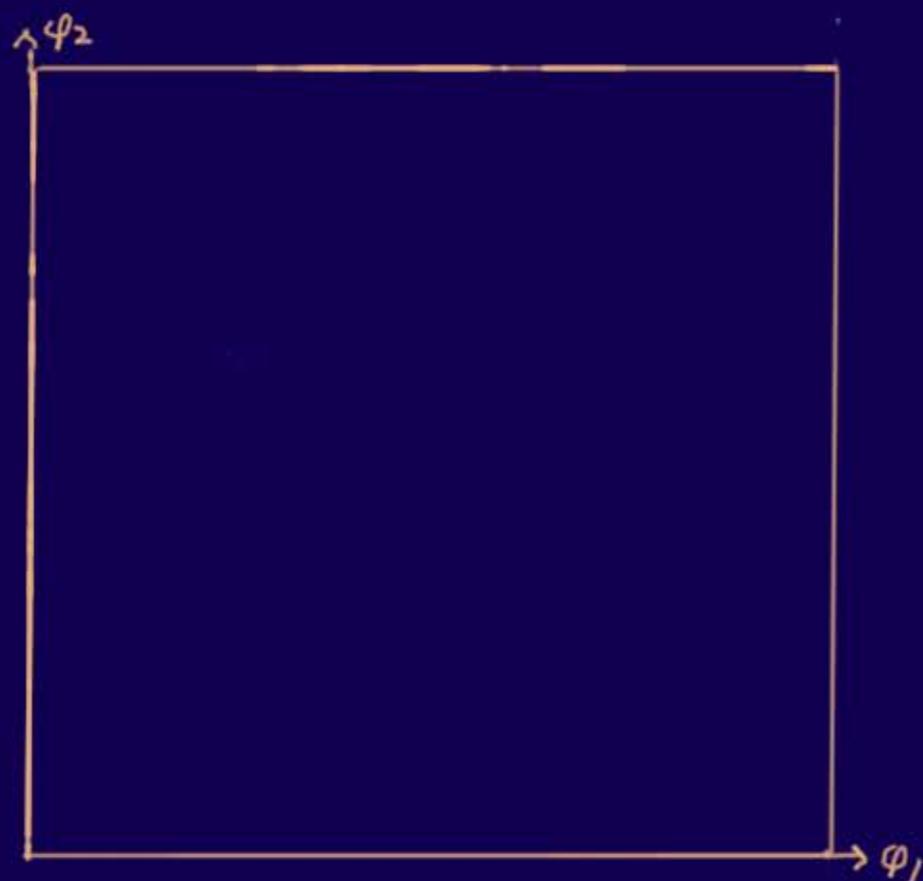
1 Parameter space
(canonical) Φ

Interest $\psi(\varphi)$



2. Parameter space
(Pointly sym. re ψ°) $\bar{\Phi}$

Interest $\psi(\varphi)$



Jeffreys (usual)

$$\text{Likelihood} = L^\circ(\varphi)$$

$$\text{J. prior} = \pi(\varphi) = |g_\varphi(\varphi)|^{1/2}$$

$$\text{Posterior} = L^\circ(\varphi) \pi(\varphi)$$

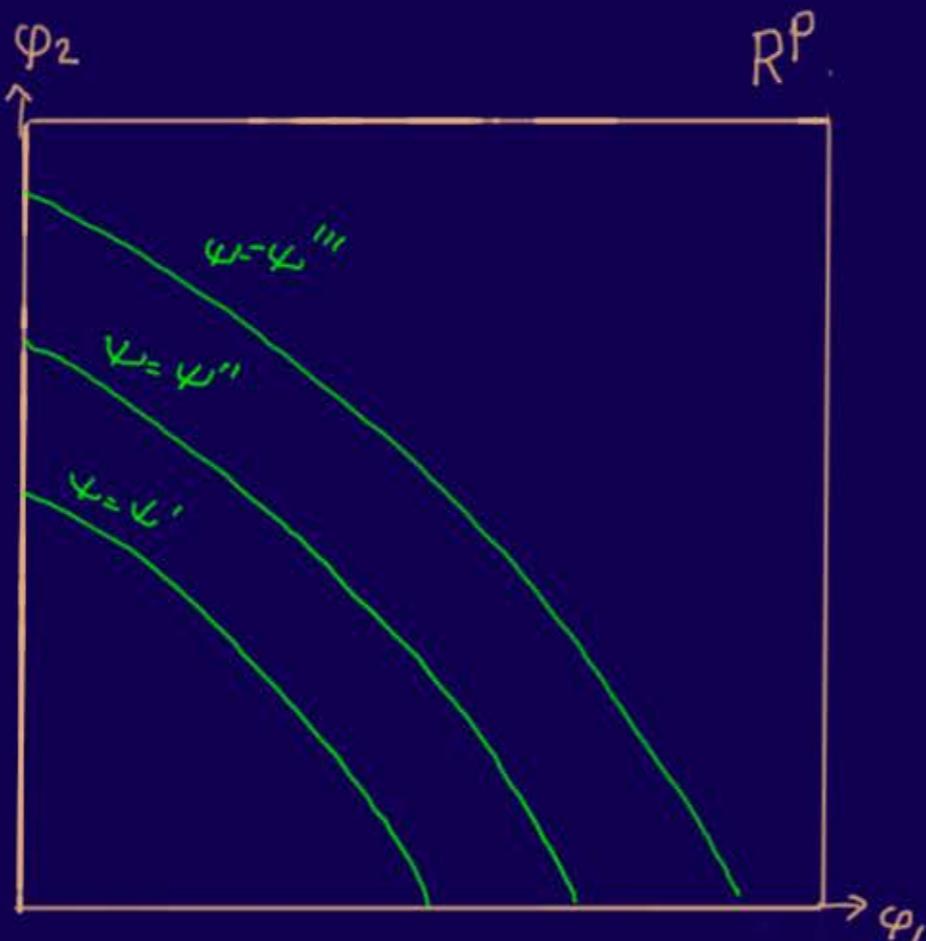
& integrate up to contour " " for dist'n of ψ

Accelerated Jeffreys: J^*

The geometry:

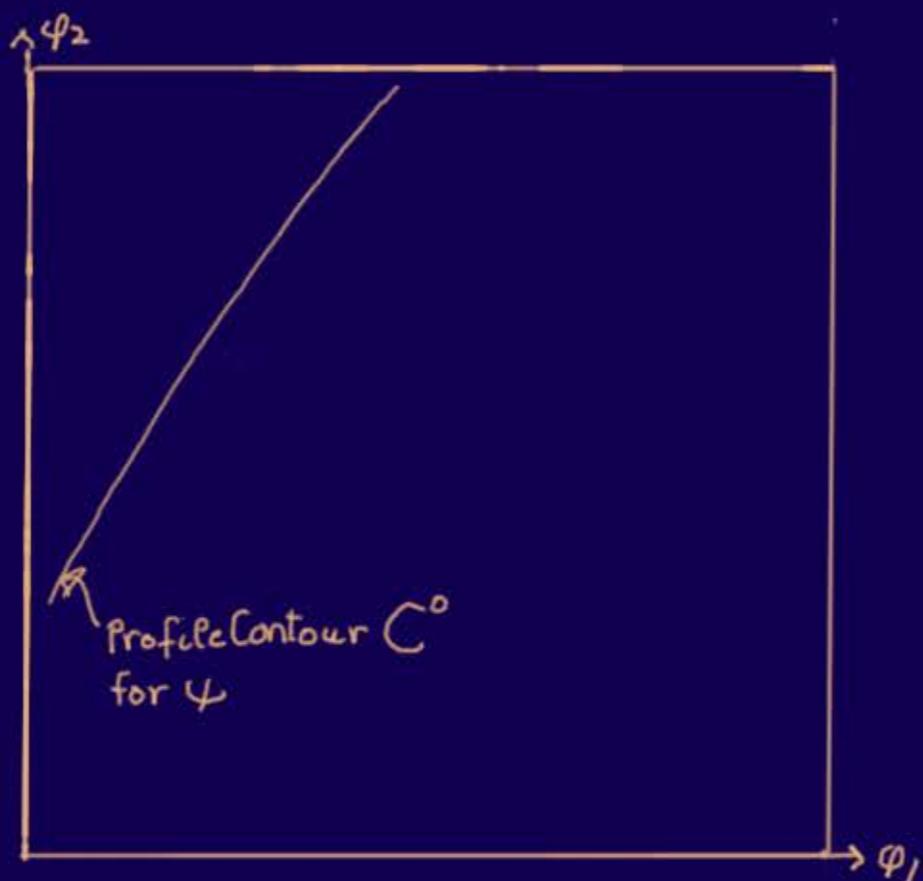
1 Parameter space
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Interest $\psi(\varphi)$



2. Parameter space
(Point symmetric re ψ°) $\bar{\Phi}$

Interest $\psi(\varphi)$



Jeffreys (usual):

Likelihood = $L^\circ(\varphi)$

$$J.\text{ prior} = \pi^\circ(\varphi) = |J_{\varphi\varphi}(\varphi)|^{1/2}$$

$$\text{Posterior} = L^\circ(\varphi) \pi^\circ(\varphi)$$

& integrate up to contour " " for dist'n of ψ

accelerated Jeffreys J^*

Use full Jeffreys

but just on profile curve C° re ψ

$$\text{Posterior} = L^\circ(\varphi) \pi_\psi^\circ(\varphi) \text{ on } C^\circ \text{ (one dimensional)}$$

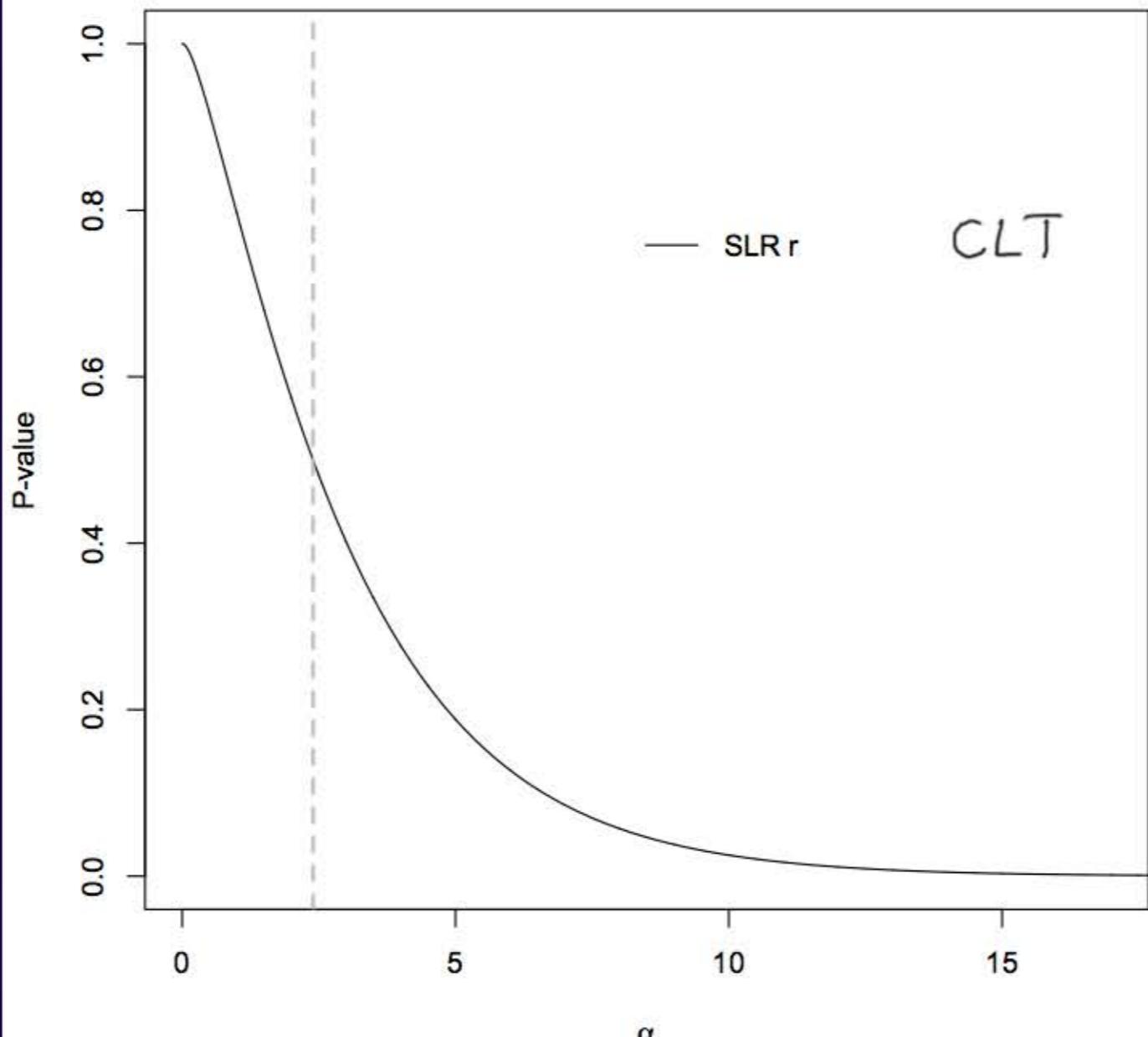
Vector parameter (α, β) ; Scalar of interest α
Example 2

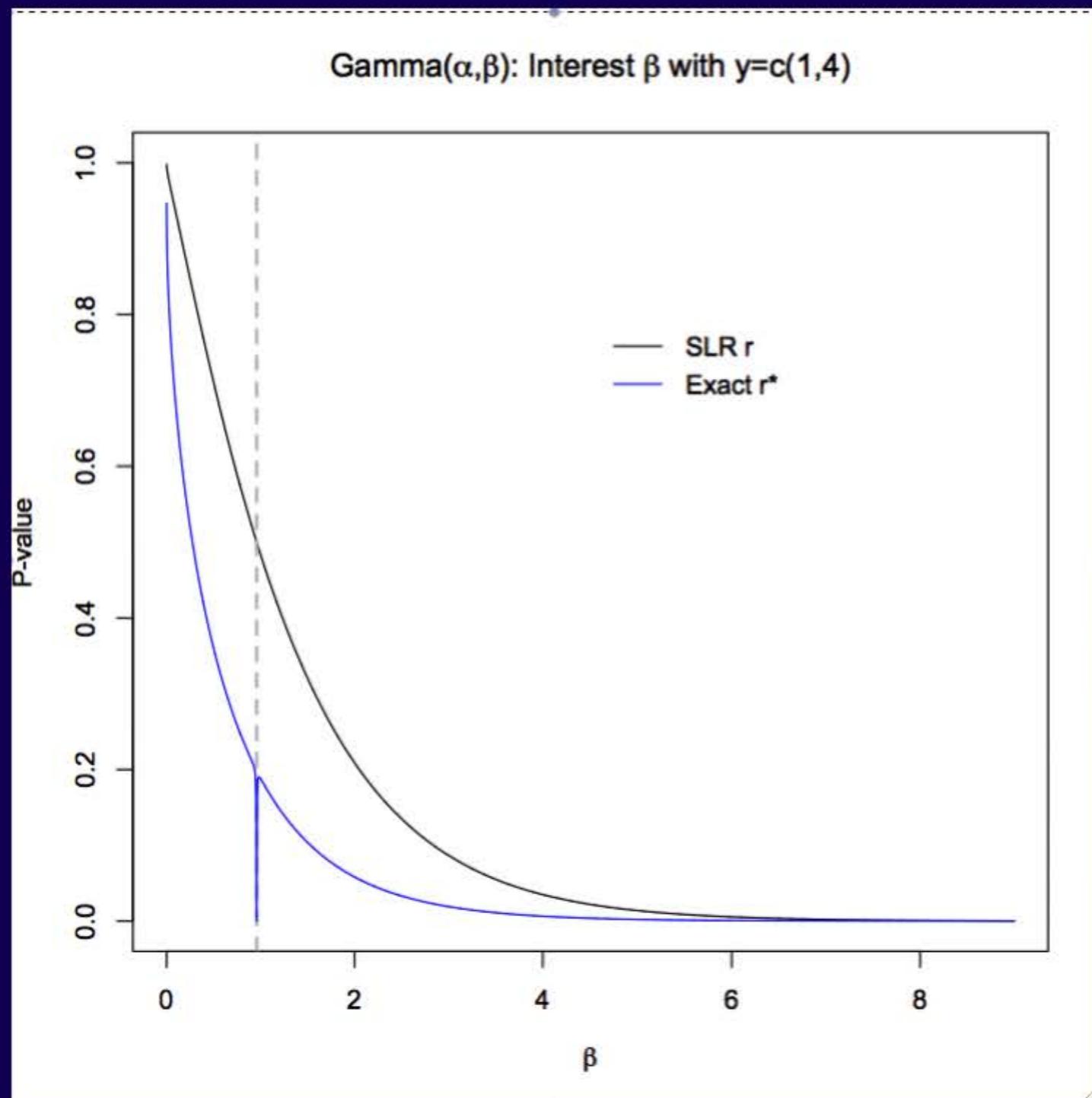
Setup

- Model: $Y \sim \text{Gamma}(\alpha, \beta)$ where α is shape and β is rate
pdf is $\frac{\beta^\alpha}{\Gamma(\alpha)} y^{\alpha-1} e^{-\beta y}$
Let $n = 2$ and data is $(y_1, y_2) = (1, 4)$

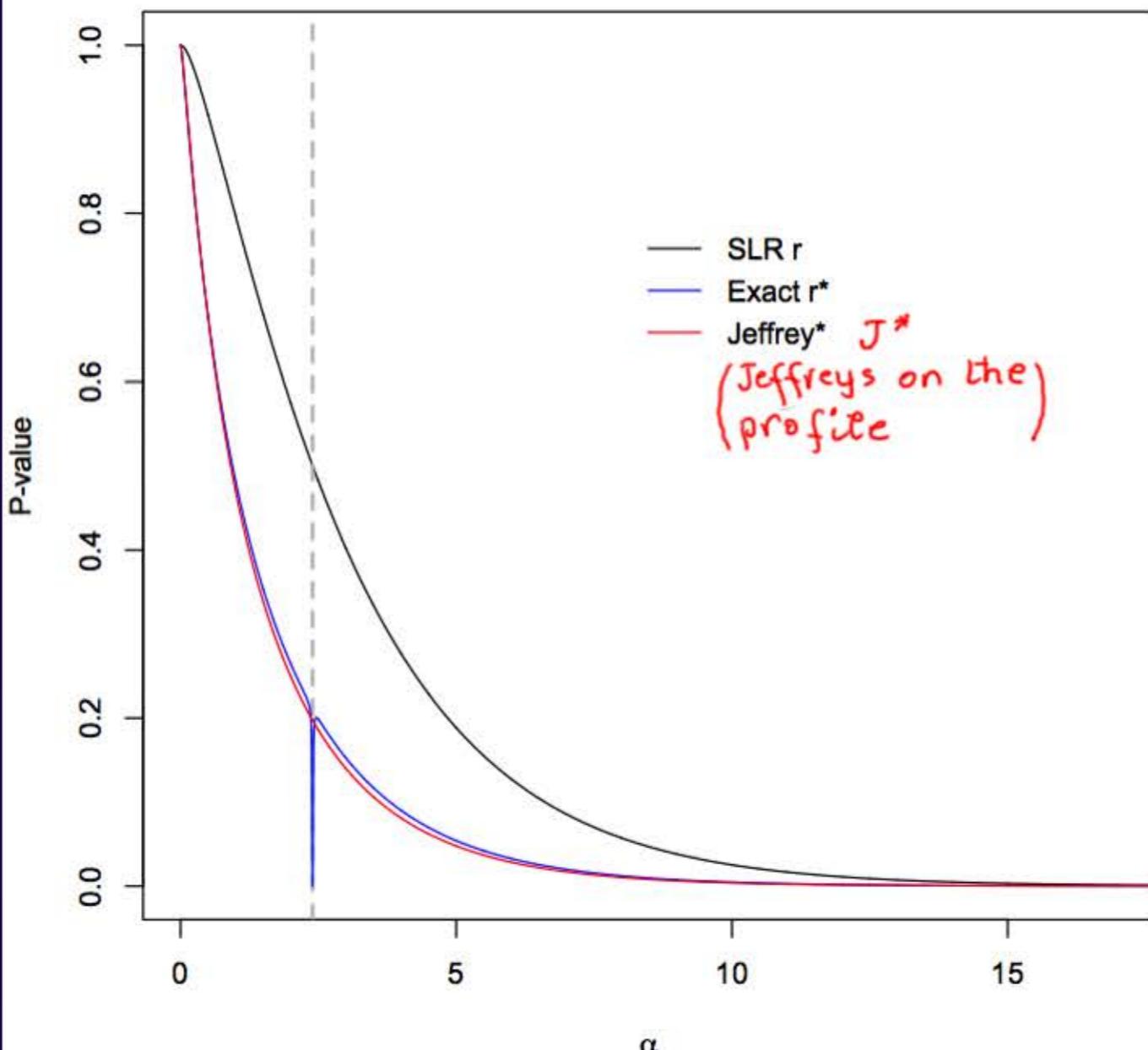
Parameter of interest is α , β free nuisance

Gamma(α, β): Interest α with $y=c(1,4)$





Gamma(α, β): Interest α with $y=c(1,4)$



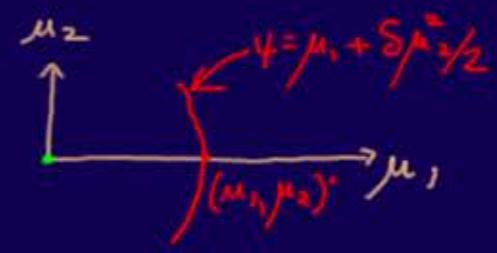
Example 4

$$\mathcal{N}(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}; I)$$

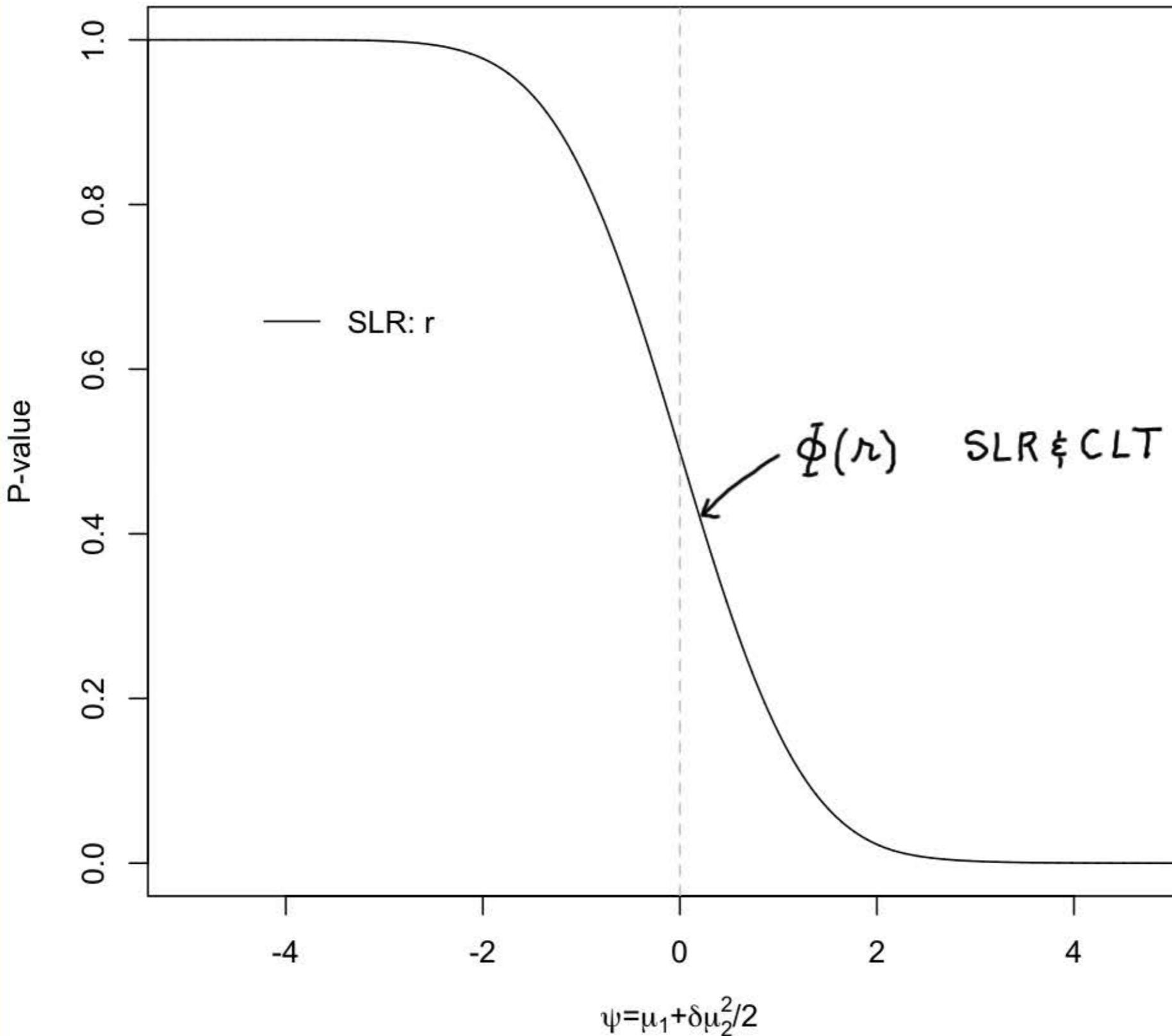
$$y^0 = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$$

$$\psi = \mu_1 + \delta \mu_2^2 / 2$$

Curved Interest (A ^{asymptotic} interest in itself!)

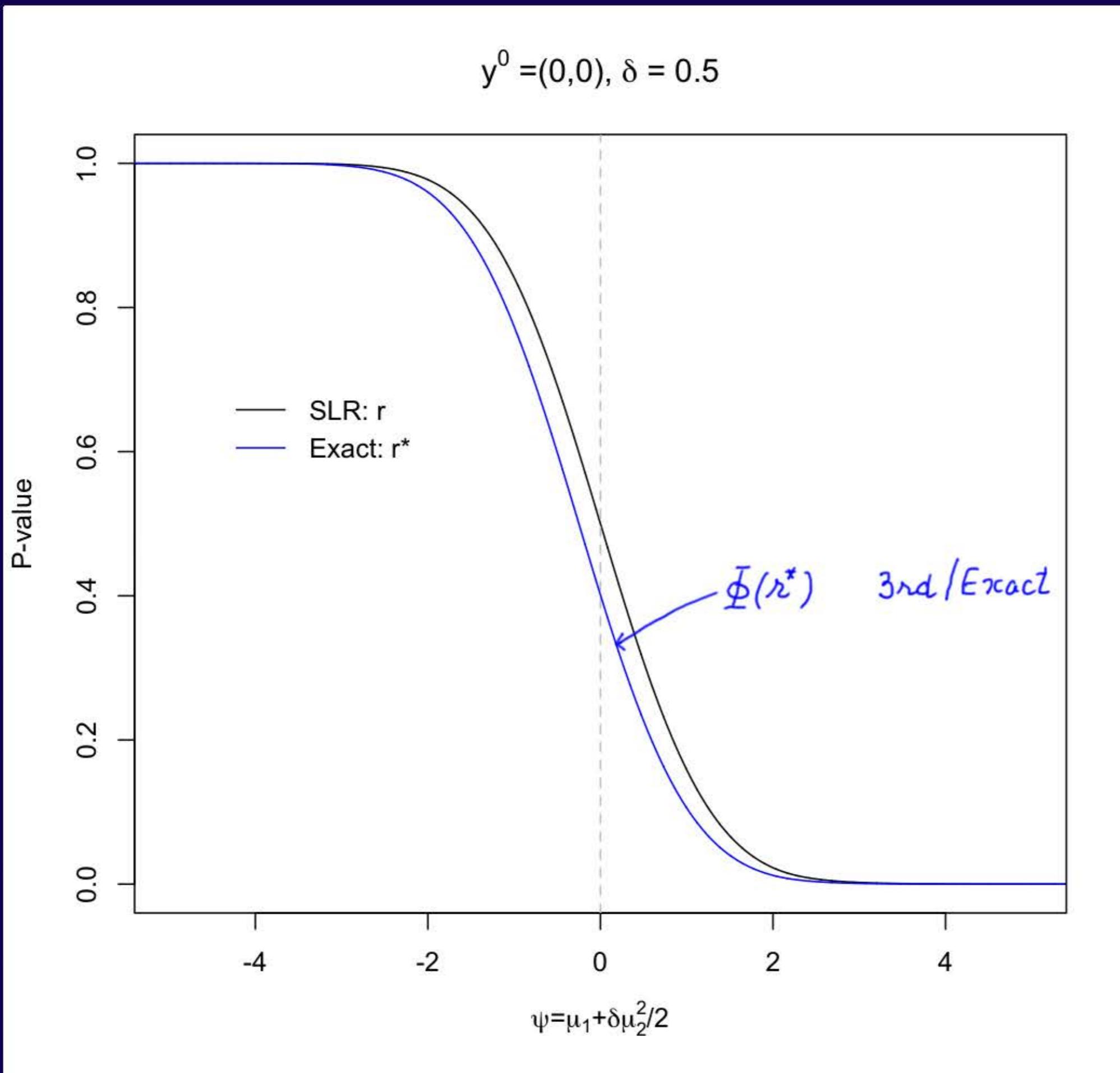


$$y^0 = (0,0), \delta = 0.5$$



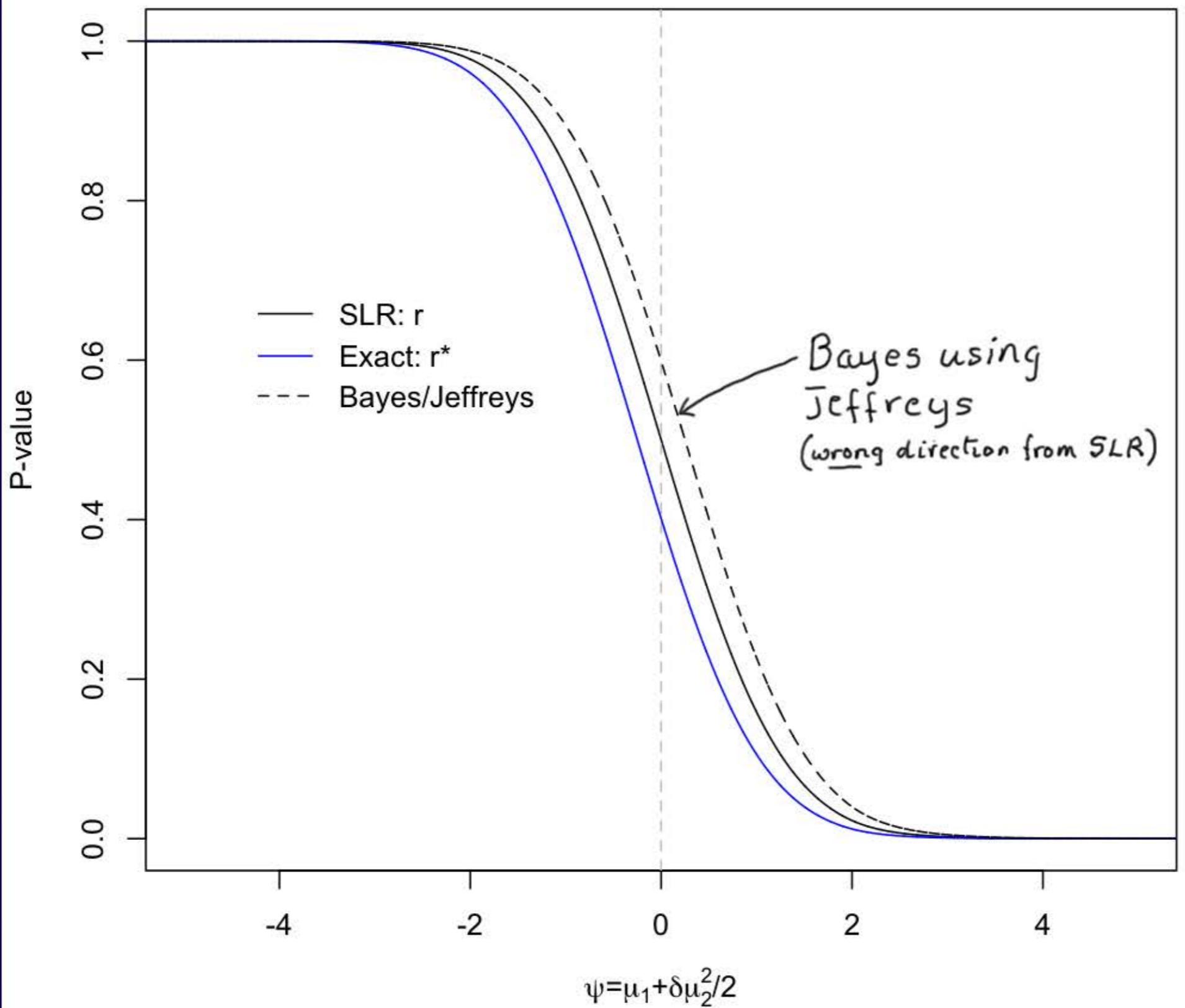
$$\text{Example 4} \quad \mathcal{N}(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}; I) \quad \psi = \mu_1 + \delta \mu_2^2 / 2$$

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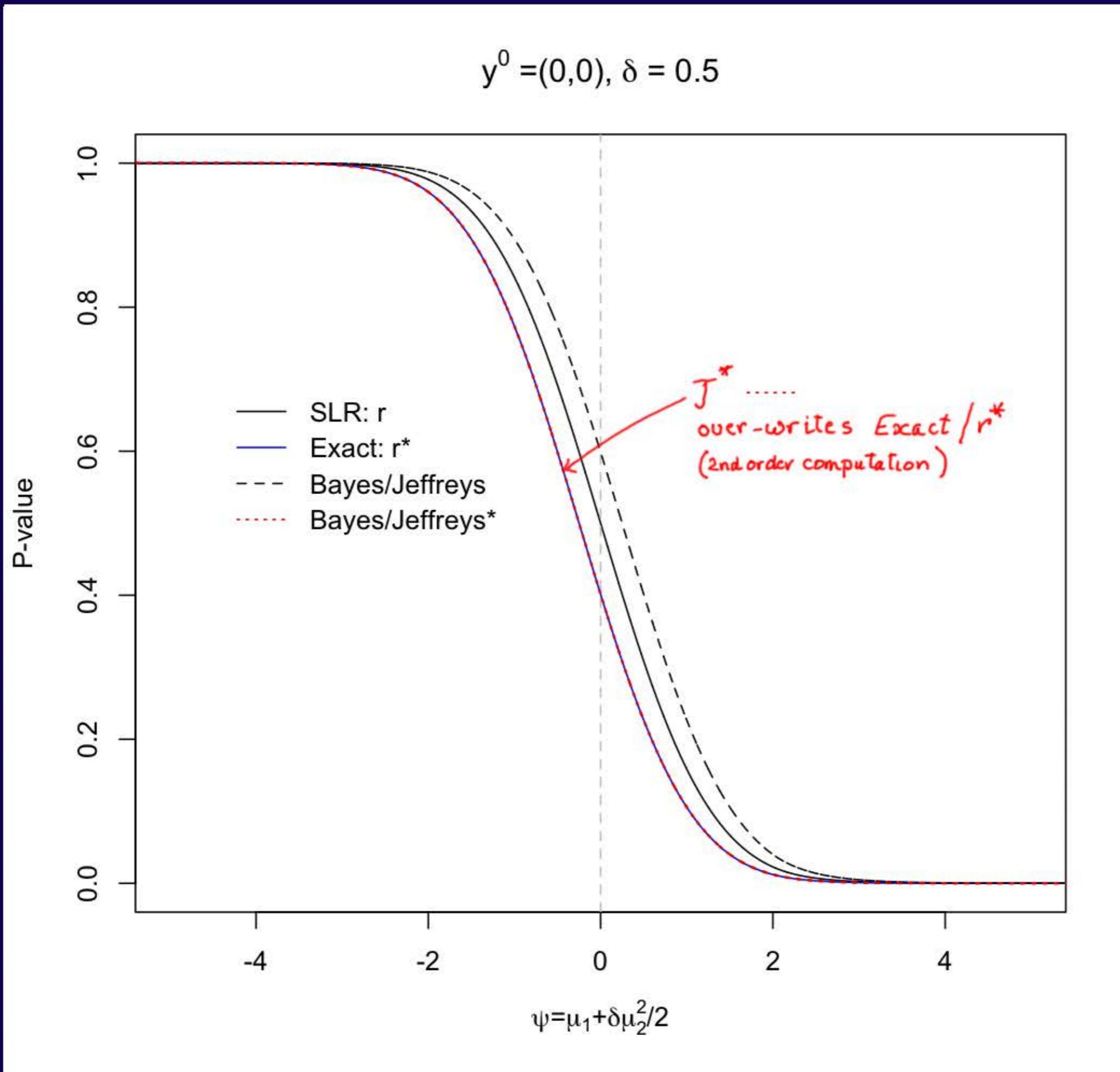
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BIG DATA

The Parable of Google Flu: Traps in Big Data Analysis

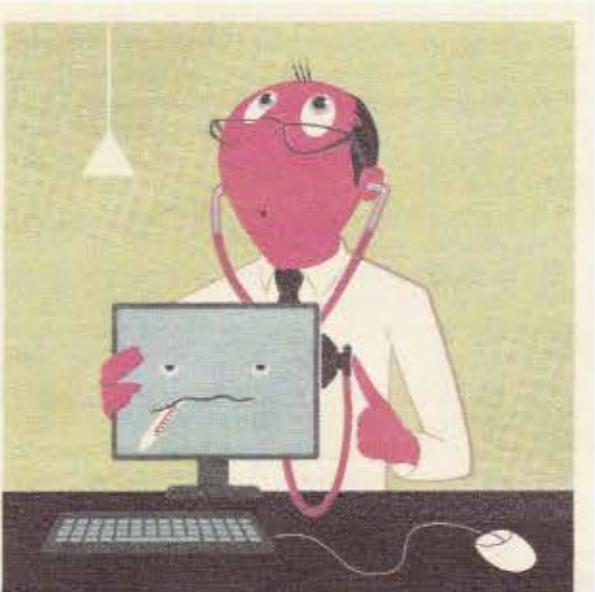
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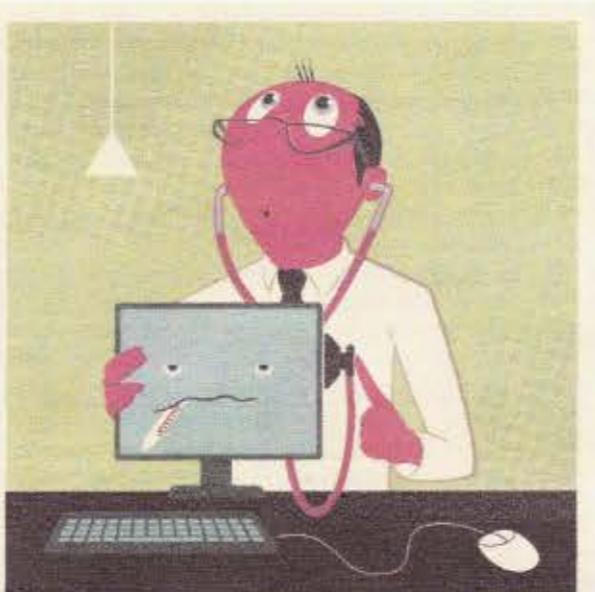
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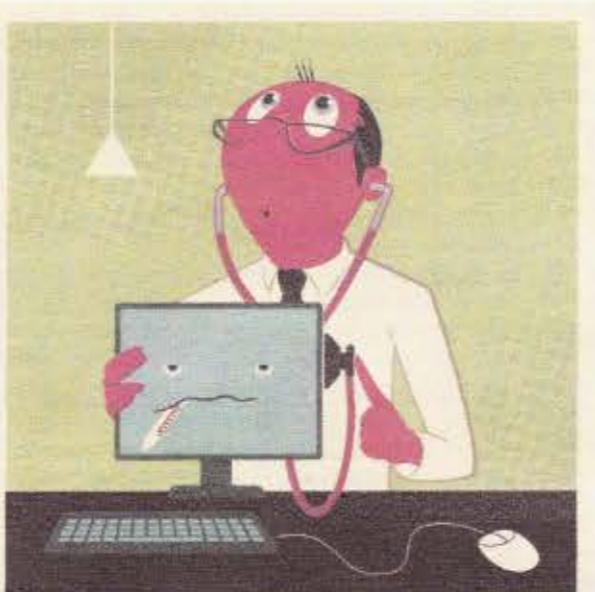
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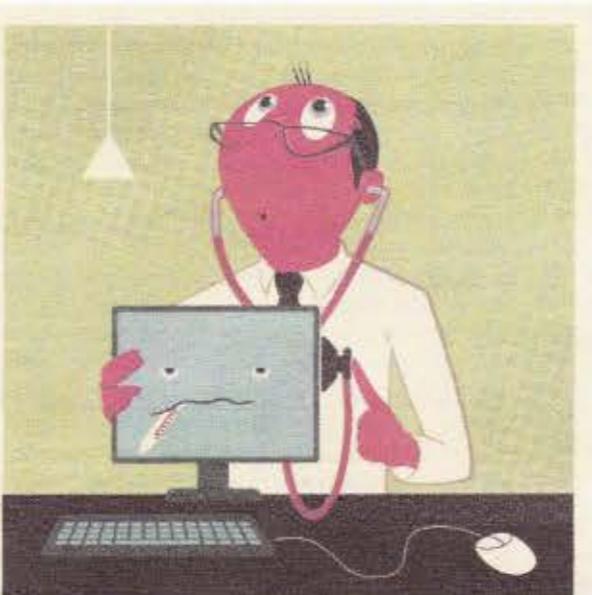
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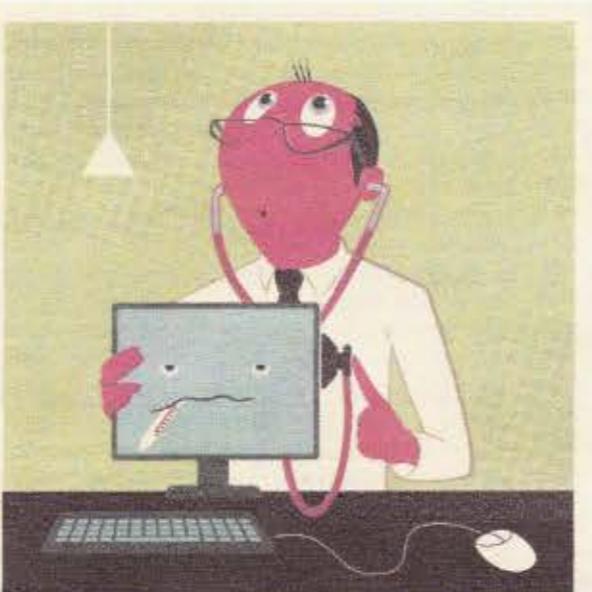
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BIG DATA

⇒ Reproducibility!

A cautionary Tale

Discussion:

1. $f(y - \theta)$ with Jeffreys \Rightarrow Exact inference, Reproducibility
2. $f(y; \theta)$, vector θ , y° ,
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 $y = y(\theta; z)$

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2. $f(y; \theta)$, vector θ , y^* ,
regular, continuity, \Rightarrow unique quantile
indep. components representation \Rightarrow Directions
 $y = y(\theta; z)$ where y "measures" θ at y^*

$$\nabla \cdot (n_1, \dots, n_p) = \frac{dy}{d\theta} \Big|_{y^*} \Rightarrow$$

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Rescale so $\mathcal{J}_{\varphi\varphi}(\hat{\varphi}^\circ) = I$

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scalar interest $\psi(\theta)$ \Rightarrow Jeff = $\pi(\psi) = |\lambda_{\psi\psi}(\psi) = -\ell_{\psi\psi}(\theta)|^{1/2}$ Use $\pi(\psi)$ only on
Rescale so $\lambda_{\psi\psi}(\hat{\psi}^\circ) = I$ curve $C_\psi = \{\psi : \psi(\psi) = \hat{\psi}^\circ\}$

Discussion:

1. $f(y - \theta)$ with Jeffreys \Rightarrow Exact inference, Reproducibility
2. $f(y; \theta)$, vector θ, y° ,
regular, continuity
indep. components \Rightarrow Unique quantile
representation \Rightarrow Directions
 $y = y(\theta; z)$
 $v = (v_1, \dots, v_p) = \frac{dy}{d\theta} \Big|_{y^\circ}$ \Rightarrow Canonical
where y "measures" θ at y° of exponential model
3. $\ell^\circ(\theta), \varphi(\theta), \text{data } y^\circ, \Rightarrow$
scalar interest $\psi(\theta)$ \Rightarrow Unique $O(n^{3/2})$
p-value fr $p(\theta)$ $\left(\begin{array}{l} \text{Sufficiency; Ancillarity} \\ \text{not needed / wanted} \end{array} \right) \Rightarrow$ unique reproducible
inference
4. $\ell^\circ(\theta), \varphi(\theta), \text{data } y^\circ, \Rightarrow$
scalar interest $\psi(\theta)$ \Rightarrow Jeff = $\pi(\psi) = |\int_{\varphi^\circ}^{\psi^\circ} \ell_{\varphi\varphi}(\varphi) - \ell_{\varphi\varphi}(\theta^\circ)|^{1/2}$
Rescale so $\int_{\varphi^\circ}^{\hat{\psi}^\circ} \ell_{\varphi\varphi}(\varphi) = I$ Use $\pi(\psi)$ only on One dim. integration:
curve $C_\psi = \{\varphi : \psi(\varphi) = \hat{\psi}^\circ\}$ $O(n^1)$ if $\psi(\varphi)$ linear

Discussion:

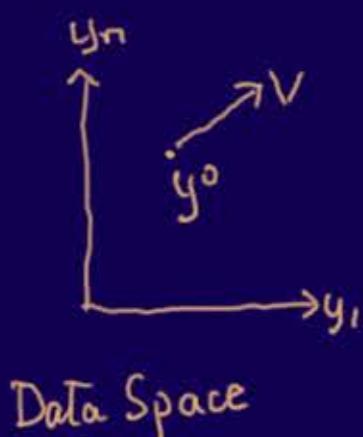
1. $f(y - \theta)$ with Jeffreys \Rightarrow Exact inference, Reproducibility
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3. $\ell^\circ(\theta), \varphi(\theta), \text{data } y^\circ, \Rightarrow$
scalar interest $\varphi(\theta) \Rightarrow$ Unique $O(n^{3/2})$
p-value fr $p(\theta)$ (Sufficiency; Ancillarity)
not needed / wanted \Rightarrow unique reproducible
inference
4. $\ell^\circ(\theta), \varphi(\theta), \text{data } y^\circ, \Rightarrow$
scalar interest $\varphi(\theta) \Rightarrow$ Jeff = $\pi(\varphi) = |\int_{\varphi^\circ}^{\varphi} \ell_{\varphi\varphi}(\varphi) = -\ell_{\varphi\varphi}(\theta)|^{1/2}$
Rescale so $\int_{\varphi^\circ}^{\hat{\varphi}} \ell_{\varphi\varphi}(\varphi) = 1$ Use $\pi(\varphi)$ only on One dim. integration:
curve $C_\varphi = \{\varphi : \varphi(\varphi) = \hat{\varphi}\}$ $O(n^1)$ if $\varphi(\varphi)$ linear
5. If $\varphi(\varphi)$ is curved \Rightarrow
simple adjustment to $\pi(\varphi)$

Discussion:

1. $f(y - \theta)$ with Jeffreys \Rightarrow Exact inference, Reproducibility
2. $f(y; \theta)$, vector θ, y° ,
regular, continuity
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 $y = y(\theta; z)$
 $V = (n_1, \dots, n_p) = \frac{dy}{d\theta} \Big|_{y^\circ}$ \Rightarrow Canonical
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curve $C_\psi = \{\varphi : \psi(\varphi) = \hat{\psi}^\circ\}$ $O(n^1)$ if $\psi(\varphi)$ linear
5. If $\psi(\varphi)$ is curved \Rightarrow
simple adjustment to $\pi(\psi)$
6. Geometry: Expt'l model;
can var u ; can par. φ

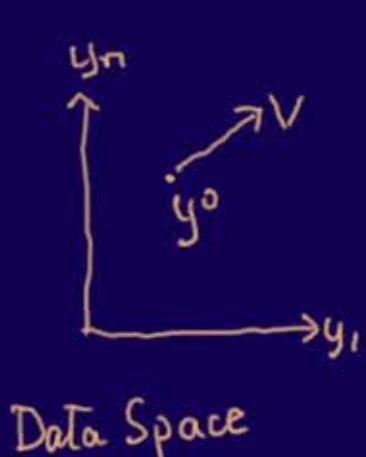
Discussion:

1. $f(y - \theta)$ with Jeffreys \Rightarrow Exact inference, Reproducibility
2. $f(y; \theta)$, vector θ, y^* ,
regular, continuity
indep. components \Rightarrow Unique quantile representation \Rightarrow Directions
 $y = y(\theta; z)$
 $v = (v_1, \dots, v_p) = \frac{dy}{d\theta} \Big|_{y^*}$ \Rightarrow Canonical of exponential model
where y "measures" θ at y^*
3. $\ell^*(\theta), \varphi(\theta), \text{data } y^*,$
scalar interest $\psi(\theta) \Rightarrow$ Unique $O(n^{3/2})$
p-value fr $p(\theta)$ $\left(\begin{array}{l} \text{Sufficiency; Ancillarity} \\ \text{not needed/wanted} \end{array} \right) \Rightarrow$ unique reproducible inference
4. $\ell^*(\theta), \varphi(\theta), \text{data } y^*,$
scalar interest $\psi(\theta) \Rightarrow$ Jeff = $\pi(\psi) = |\int_{\varphi(\theta)} \ell_{\varphi\varphi}(\theta)|^{1/2}$
Rescale so $\int_{\varphi(\theta)} \ell_{\varphi\varphi}(\hat{\psi}^*) = I$ Use $\pi(\psi)$ only on One dim. integration:
curve $C_\psi = \{\psi : \psi(\psi) = \hat{\psi}^*\}$ $O(n^1)$ if $\psi(\psi)$ linear
5. If $\psi(\psi)$ is curved \Rightarrow simple adjustment to $\pi(\psi)$
6. Geometry: Expt'l model;
can var u ; can par. φ



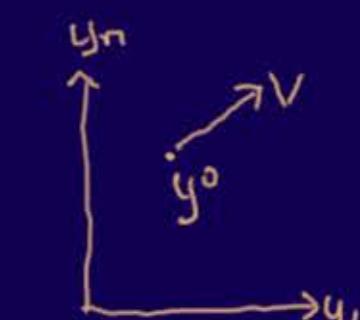
Discussion:

1. $f(y - \theta)$ with Jeffreys \Rightarrow Exact inference, Reproducibility
2. $f(y; \theta)$, vector θ, y^* , regular, continuity, indep. components \Rightarrow Unique quantile representation $y = y(\theta; z)$ Directions $v = (v_1, \dots, v_p) = \frac{dy}{d\theta} \Big|_{y^*}$ Canonical of exponential model where y "measures" θ at y^*
3. $\ell^*(\theta), \varphi(\theta), \text{data } y^*, \text{ scalar interest } \psi(\theta) \Rightarrow$ Unique $O(n^{1/2})$ p-value for $\psi(\theta)$ (Sufficiency; Ancillarity) $\left(\begin{array}{l} \text{not needed / wanted} \end{array} \right) \Rightarrow$ unique reproducible inference
4. $\ell^*(\theta), \varphi(\theta), \text{data } y^*, \text{ scalar interest } \psi(\theta) \Rightarrow$ Jeff = $\pi(\psi) = |\int_{\varphi(\theta)} \ell^*(\theta)|^{1/2}$ Use $\pi(\psi)$ only on curve $C_\psi = \{\psi: \psi(\theta) = \hat{\psi}^*\}$ One dim. integration: Rescale so $\int_{\varphi(\theta)} \ell^*(\theta) = 1$ $O(n^1)$ if $\psi(\theta)$ linear
5. If $\psi(\theta)$ is curved \Rightarrow simple adjustment to $\pi(\psi)$
6. Geometry: Expt'l model; can var u ; can par. φ



Discussion:

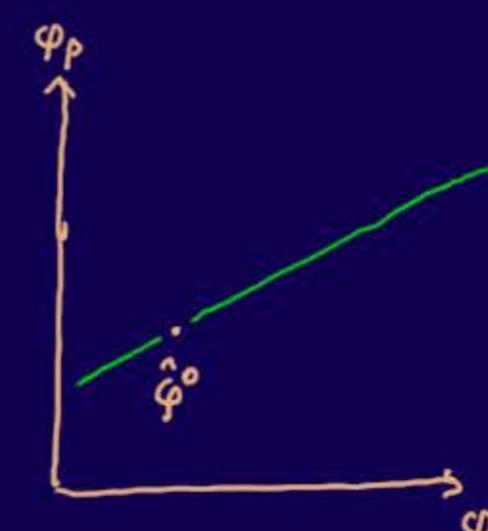
1. $f(y - \theta)$ with Jeffreys \Rightarrow Exact inference, Reproducibility
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5. If $\varphi(\varphi)$ is curved \Rightarrow simple adjustment to $\pi(\varphi)$
6. Geometry: Exponential model; can var u ; can par. φ



Data Space



Can. var.
space

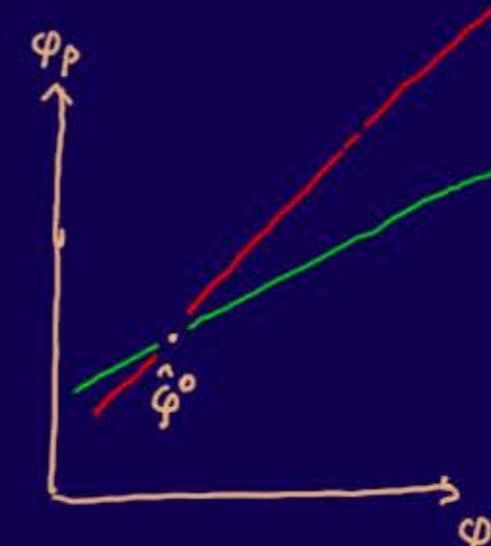
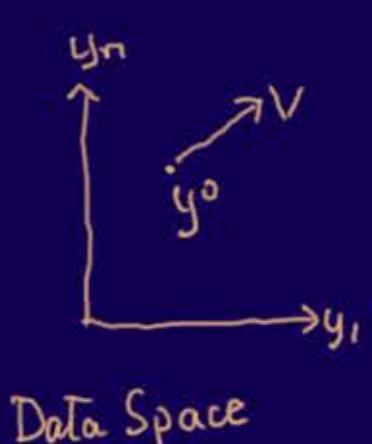


Can. par.
space

Interest in $\varphi^*(\varphi)$:
Profile contour for $\varphi^*(\varphi)$ C_φ
(use full Jeffreys)

Discussion:

1. $f(y - \theta)$ with Jeffreys \Rightarrow Exact inference, Reproducibility
2. $f(y; \theta)$, vector θ, y^* , regular, continuity, indep. components \Rightarrow Unique quantile representation $y = y(\theta; z)$ Directions $v = (v_1, \dots, v_p) = \frac{dy}{d\theta} \Big|_{y^*}$ Canonical of exponential model where y "measures" θ at y^*
3. $\ell^*(\theta), \varphi(\theta)$, data y^* , scalar interest $\varphi(\theta)$ \Rightarrow Unique $O(n^{1/2})$ p-value for $p(\theta)$ (Sufficiency; Ancillarity) $\not\rightarrow$ unique reproducible inference (not needed/wanted)
4. $\ell^*(\theta), \varphi(\theta)$, data y^* , scalar interest $\varphi(\theta)$ \Rightarrow Jeff = $\pi(\varphi) = |\int_{\varphi^*}^{\varphi} \ell_{\varphi\varphi}(\varphi) d\varphi|^{1/2}$ Use $\pi(\varphi)$ only on curve $C_\varphi = \{\varphi : \varphi(\varphi) = \hat{y}^*\}$ One dim. integration: Rescale so $\int_{\varphi^*}^{\varphi} \ell_{\varphi\varphi}(\varphi) d\varphi = I$ $O(n^1)$ if $\varphi(\varphi)$ linear
5. If $\varphi(\varphi)$ is curved \Rightarrow simple adjustment to $\pi(\varphi)$
6. Geometry: Exptl model; can var u ; can par. φ



Interest in $\varphi''(\varphi)$:
Profile contour for $\varphi''(u)$ $C_{\varphi''}$
(use full Jeffreys on line)

Interest in $\varphi'(\varphi)$:
Profile contour for $\varphi'(u)$ $C_{\varphi'}$
(use full Jeffreys)

Summary:

1. All info (2nd/3rd) for scalar $\psi(\theta)$ is on profile curve C_ψ (1 dim)
use full Jeffreys on $C_\psi \dots \underline{\text{not}}$ on full space

Summary:

1. All info (2nd/3rd) for scalar $\psi(\theta)$ is on profile curve C_ψ (1 dim)
use full Jeffreys on C_ψ ... not on full space
2. Different $\psi(\theta)$... different curve C_ψ

Summary:

1. All info (2nd/3rd) for scalar $\psi(\theta)$ is on profile curve C_ψ (1 dim)
use full Jeffreys on C_ψ ... not on full space
2. Different $\psi(\theta)$... different curve C_ψ
3. Gives 2nd order inference

Summary:

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use full Jeffreys on C_ψ ... not on full space
2. Different $\psi(\theta)$... different curve C_ψ
3. Gives 2nd order inference
but if "curved", simple curvature adjustment available for 2nd order

Summary:

1. All info (2nd/3rd) for scalar $\psi(\theta)$ is on profile curve C_ψ (1 dim)
use full Jeffreys on C_ψ ... not on full space
2. Different $\psi(\theta)$... different curve C_ψ
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but if "curved", simple curvature adjustment available for 2nd order

Thank you ...