

Priors and Inference: A differential view

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www.utstat.toronto.edu/dfraser/documents/Oxford2012.pdf

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Index for some references

Outline:

- 1 Priors & Inference: Is there a problem?
- 2 Data dependent prior? Why?
- 3 Welch Peers: Exponential as location model
- 4 Exponential model: as an approximation
- 5 Saddlepoint: for exponential & for general
- 6 Conditional / Marginal for linear parameter component
- 7 Directional Jeffreys
- 8 Curved parameter: How Bayes misses!

Some Q & A at talk

Some Q & A at discussion later

1 Priors & Inference: Is there a problem?

- Bayes & freq. give different results!
- A discipline with 2 contradictory methodologies!

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- Does accuracy matter?

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- Does accuracy matter?
- Does the discipline care?

1 Priors & Inference: Is there a problem?

- Bayes & freq. give different results!
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- Does accuracy matter? | Merck, Vioxx, \$10b profit, Recalled in 2006
- Does the discipline care? | FDA penalty \$5b Est. # deaths 40K / uwol2.paf

Q Need for accuracy?

A Available statistics had shown high risks for Vioxx. Some information had been suppressed, other information had been ignored, as in the "Shuttle's" O-rings; so Accuracy is needed broadly

Both: In choice and use of statistics & In the external messages from statistics
as with { Vioxx mentioned above.
O-rings

1 Priors & Inference: Is there a problem?

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- Does accuracy matter?

- Does the discipline care?

Merck, Vioxx, \$10b profit, Recalled in 2006

FDA penalty \$5b Est. # deaths 40K / [uwol2.paf](#)

Take a simple case: scalar

$$\ast \text{ If } \underline{f(y-\theta)} \quad + \quad p(\theta) = \int_{-\infty}^{y^0} f(y-\theta) dy$$

$$\text{B } s(\theta) = \int_{\theta}^{\infty} \pi(\theta) f(y^0-\theta) d\theta$$

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Take a simple case: scalar

* If $f(y-\theta)$ + $p(\theta) = \int_{-\infty}^y f(y-\theta) dy$

Equal if

Nice pure case, but

$\pi(\theta) = c$

Vector case \Rightarrow Doesn't work!

A: Paradigm
for when Bayes
is objectively
repetition-valid

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B $s(\theta) = \int_{\theta}^{\infty} \pi(\theta) f(y-\theta) d\theta$ $\pi(\theta) = c$ Vector case \Rightarrow Doesn't work!

\leftarrow calculus \searrow (bdry conditions)

If $f(y; \theta)$
stoch. inc.

$$p(\theta) = F(y^0; \theta) \equiv \int_{\theta}^{\infty} f_{y; \theta}(y; \theta) d\theta$$
$$s(\theta) = \int_{\theta}^{\infty} \pi(\theta) F_y(y^0; \theta) d\theta \quad \text{"agree?"}$$

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← calculus (boundary conditions)

If $f(y; \theta)$
stoch. inc.

$$p(\theta) = F(y^0; \theta) \equiv \int_{\theta}^{\infty} f_{y; \theta}(y; \theta) d\theta$$

$$s(\theta) = \int_{\theta}^{\infty} \pi(\theta) F_y(y^0; \theta) d\theta$$

"agree?"

Q: Clarify:

A: To have $p(\theta) = s(\theta)$

and thus have Bayes repetition-valid one needs to have:

Equal if $\pi(\theta) = - \frac{F_{y; \theta}(y^0; \theta)}{F_y(y^0; \theta)}$

prior \rightarrow

$$\pi(\theta) = \left. \frac{dy}{d\theta} \right|_{y^0, \theta^0}$$

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If $f(y; \theta)$
stoch. inc.

$p(\theta) = F(y^0; \theta) \equiv \int_{\theta}^{\infty} f_{; \theta}(y^0; \theta) d\theta$ df | $F(y; \theta) = u$ dist'n

$s(\theta) = \int_{\theta}^{\infty} \pi(\theta) F_y(y^0; \theta) d\theta$ "agree?" gf | $y = y(u; \theta)$ quantile

Equal if $\pi(\theta) = - \frac{F_{; \theta}(y^0; \theta)}{F_y(y^0; \theta)}$

prior \rightarrow

1 Priors & Inference: Is there a problem?

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If $f(y; \theta)$
stoch. inc.

$p(\theta) = F(y^0; \theta) \equiv \int_{\theta}^{\infty} f_{y; \theta}(y^0; \theta) d\theta$ df | $F(y; \theta) = u$ *dist'n*

$s(\theta) = \int_{\theta}^{\infty} \pi(\theta) F_y(y^0; \theta) d\theta$ "agree?" gf | $y = y(u; \theta)$ *quantile*

Equal if $\pi(\theta) = - \frac{F_{y; \theta}(y^0; \theta)}{F_y(y^0; \theta)}$ $\pi(\theta) = \frac{\partial}{\partial \theta} y(u; \theta) \Big|_{y^0, \hat{\theta}^0}$ = "How θ moves y near y^0 "

prior \rightarrow

\Rightarrow Data-dependent prior!

2 Data-dependent prior? Why?

-Contrary to "classical" Bayes view!

2

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- But well recognized ... in places

Box & Cox 1964

Wasserman 2000 JRSSB

Freid MY 2010 JRSSB

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- B. & freq. differ? Data-dependent can make it right? Sometimes!

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- B. & freq. differ? Data-dependent can make it right? Sometimes!

- How to evaluate B vs freq? Use bounds, confidence vs. posterior

Confidence: level always right! ← under model

Posterior: level right if $\pi(\theta)$ describes source!

otherwise wrong!

2 Data-dependent prior? Why?

- Contrary to "classical" Bayes view!

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- B. & freq. differ? Data-dependent can make it right? Sometimes!

- How to evaluate Bus freq? Use bounds, confidence vs. posterior

Confidence: level always right!

Posterior: level right if $\pi(\theta)$ describes source!

otherwise wrong!

- Opinion prior?

⇒ Get confidence result first; put Opinion prior in parallel

e.g. Multiple Conf bounds (or Confidence dist's maybe?)

See later

- Don't use Opinion prior to analyze model!

⇒ an Admission of bankruptcy!

3 Welch-Peers 1963 Scalar Exponential: As a location model

Q Just Exponential?

A also "general asymptotic" in W-P using
Expected info $i(\theta)$ but interest here
in Exponential model (also see next section)

3 Welch-Peers 1963 Scalar Exponential: As a location model

Root info prior \Rightarrow Posterior bound is confidence bound $O(n^{-1})$

$$f(y; \theta) = \exp\{s(y)\varphi(\theta) + K(\theta)\} H(y)$$

\swarrow $\dim = n$

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$$f(y; \theta) = \exp\{s(y)\varphi(\theta) + K(\theta)\} H(y)$$

$$g(s; \varphi) = \exp\{s\varphi - K(\varphi)\} g(s)$$

Sufficiency

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$$g(s; \varphi) = \exp\{s\varphi + \ell(\varphi)\} g(s)$$

Sufficiency

← Center & scale $J_{\varphi\varphi}^{1/2}(\hat{\varphi}^0)(\varphi - \hat{\varphi}^0)$ write as φ
 $J_{\varphi\varphi}^{-1/2}(\hat{\varphi}^0)(s - s^0)$ write as s

$$J_{\varphi\varphi}^0 = \text{obs info}$$

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$$g(s; \varphi) = \exp\{s\varphi - K(\varphi)\} g(s)$$

$$g(s; \varphi) = \exp\{s\varphi + \ell^0(\varphi)\} g(s)$$

$$= \exp\left\{-\frac{\varphi^2}{2} - \gamma \frac{\varphi^3}{6n^{1/2}} + s\varphi\right\} g(s)$$

\leftarrow dim = n

Obs. likelihood but in new coordinates

Sufficiency

\leftarrow Center $J_{\varphi\varphi}^{1/2}(\hat{\varphi}^0)(\varphi - \hat{\varphi}^0)$ write as φ
 & scale $J_{\varphi\varphi}^{-1/2}(\hat{\varphi}^0)(s - s^0)$ write as s

$J_{\varphi\varphi}^0 = \text{obs info}$

\leftarrow Taylor expand log model

\leftarrow missing terms ^{then} come from "pdf" property

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$$= \exp\left\{-\frac{\varphi^2}{2} - \gamma \frac{\varphi^3}{6n^{1/2}} + s\varphi\right\} g(s)$$

$$= \phi(s - \varphi) \exp\left\{-\gamma \frac{\varphi^3}{6n^{1/2}} + \gamma \frac{s^3}{6n^{1/2}}\right\} \left(1 - \gamma \frac{s}{2n^{1/2}}\right)$$

Sufficiency

\leftarrow Center $\int_{\varphi\varphi}^{1/2}(\hat{\varphi}^0)(\varphi - \hat{\varphi}^0)$ write as φ
 & scale $\int_{\varphi\varphi}^{-1/2}(\hat{\varphi}^0)(s - s^0)$ write as s

$\int_{\varphi\varphi}^0 = \text{obs info}$

\leftarrow Taylor expand log model

Get other terms by "pdf property"

Use If s is $N(\varphi, 1)$ then $E s^3 = \varphi^3 + 3\varphi$

3 Welch-Peers 1963 Scalar Exponential: As a location model

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$$f(y; \theta) = \exp\{s(y)\varphi + K(\varphi)\} H(y) \quad \leftarrow \text{dim} = n$$

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 & scale $\int_{\varphi\varphi}^{-1/2}(\hat{\varphi}^0)(s - s^0)$ write as s

$\int_{\varphi\varphi}^0 = \text{obs info}$

\leftarrow Taylor expand log model

Get other terms by "pdf property"

Use: If s is $N(\varphi, 1)$ then $E s^3 = \varphi^3 + 3\varphi$

Get

Root info parameter $\beta = \int_0^\varphi \int^{1/2}(\varphi) d\varphi = \int_0^\varphi (1 + \gamma\varphi/2n^{1/2}) d\varphi = \varphi + \gamma\varphi^2/4n^{1/2}$

mle Use: $\hat{\varphi} = s - \gamma s^2/2n^{1/2} \Rightarrow \hat{\beta} = s - \gamma s^2/4n$

Location parameter $\beta = \varphi + \gamma\varphi^2/4n^{1/2}$
 (a Quadratic adjustment)

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Root info prior \Rightarrow Posterior bound is confidence bound $O(n^{-1})$

$f(y; \theta) = \exp\{s(y)\varphi(\theta) + K(\theta)\} H(y)$ $\leftarrow \text{dim} = n$

$g(s; \varphi) = \exp\{s\varphi - K(\varphi)\} g(s)$

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Sufficiency
 \leftarrow Center $\int_{\varphi\varphi}^{1/2}(\hat{\varphi}^0)(\varphi - \hat{\varphi}^0)$ write as φ
 & scale $\int_{\varphi\varphi}^{-1/2}(\hat{\varphi}^0)(s - s^0)$ write as s $\int_{\varphi\varphi}^0 = \text{obs info}$

$= \exp\{-\frac{\varphi^2}{2} - \gamma \frac{\varphi^3}{6n^{1/2}} + s\varphi\} g(s)$ \leftarrow Taylor expand log model

$= \phi(s - \varphi) \exp\{-\gamma \frac{\varphi^3}{6n^{1/2}} + \gamma \frac{s^3}{6n^{1/2}}\} (1 - \gamma \frac{s\varphi}{2n^{1/2}})$ \leftarrow Get other terms by "pdf property" Use: If s is $N(\varphi, 1)$ then $E s^3 = \varphi^3 + 3\varphi$

Get

Root info parameter $\beta = \int_0^\varphi \int^{1/2}(\varphi) d\varphi = \int_0^\varphi (1 + \gamma\varphi/2n^{1/2}) d\varphi = \varphi + \gamma\varphi^2/4n^{1/2} = \beta$

mle Use: $\hat{\varphi} = s - \gamma s^2/2n^{1/2}$ $\hat{\beta} = s - \gamma s^2/4n$

$\Rightarrow: f(\hat{\beta}; \beta) d\hat{\beta} = (2\pi)^{-1/2} \exp\{-\frac{(\hat{\beta} - \beta)^2}{2} - \gamma(\hat{\beta} - \beta)^3/12n^{1/2}\} = \text{location model } O(n^{-1})$

Q: Should there be cubic power?
 A: Yes Thank you! Added above

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 \leftarrow Center $\int_{\varphi\varphi}^{1/2}(\hat{\varphi}^0)(\varphi - \hat{\varphi}^0)$ write as φ
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 \leftarrow Taylor expand log model

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Root info parameter $\beta = \int_0^\varphi \int^{1/2}(\varphi) d\varphi = \int_0^\varphi (1 + \gamma\varphi/2n^{1/2}) d\varphi = \varphi + \gamma\varphi^2/4n^{1/2}$

mle Use: $\hat{\varphi} = s - \gamma s^2/2n^{1/2} \Rightarrow \hat{\beta} = s - \gamma s^2/4n$

$$\Rightarrow: f(\hat{\beta}; \beta) d\hat{\beta} = (2\pi)^{-1/2} \exp\left\{-\frac{(\hat{\beta} - \beta)^2}{2} - \gamma(\hat{\beta} - \beta)/2n^{1/2}\right\} = \text{location model } O(n^{-1})$$

Bayes = freq
(Location)

with root-info prior, scalar case
(flat in β)

Vector case?
Does not work!

4 Exponential Model:

as a building block? or As an approximation to a model $O(n^{-1})$

Consider $f(y; \theta)$ regular:

4 Exponential Model:

as a building block? or As an approximation to a model $O(n^{-1})$

Consider $f(y; \theta)$ regular:

$f(y; \theta)$ density, or distn fns, or quantile fns (Often easily available; often used simulations)
 _{n p}

4 Exponential Model:

As a building block? or As an approximation to a model $O(n^{-1})$

Consider $f(y; \theta)$ regular:

$f(y; \theta)$ density, or distn fns, or quantile fns (Often easily available; often used simulations)

$n \times p$ $y = y(u, \theta)$ & data $y^o, \hat{\theta}^o$

$$V = (N_1, \dots, N_p) = \frac{\partial}{\partial \theta} y(u; \theta) \Big|_{y^o, \hat{\theta}^o}$$

4 Exponential Model:

As a building block? or As an approximation to a model $O(n^{-1})$

Consider $f(y; \theta)$ regular:

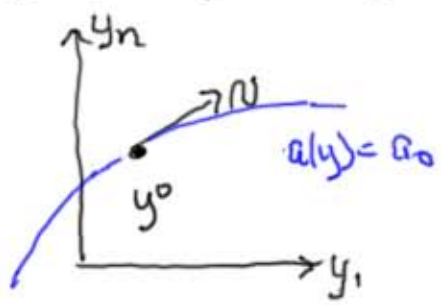
$f(y; \theta)$ density, or distn fns, or quantile fns (Often easily available; often used simulations)
 $n \times p$ \swarrow quantile version

$y = y(u, \theta)$ & data $y^o, \hat{\theta}^o$

but V tangent to ancillary

$V = (N_1, \dots, N_p) = \frac{\partial}{\partial \theta} y(u; \theta) |_{y^o, \hat{\theta}^o}$
 $n \times p$

(continuity: FF Staicu, Bern. 2010)



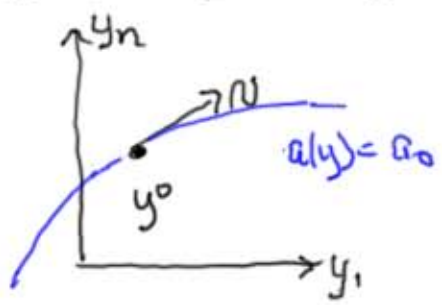
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 $n \times p$ quantile version

$y = y(u, \theta)$ & data $y^o, \hat{\theta}^o$ but V tangent to ancillary



$$V = (N_1, \dots, N_p) = \frac{\partial}{\partial \theta} y(u; \theta) \Big|_{y^o, \hat{\theta}^o} \quad (\text{continuity: FF Staicu, Bern. 2010})$$

$n \times p$

and get $\varphi(\theta) = \frac{\partial}{\partial V} l(\theta; y) \Big|_{y^o} \rightarrow$ Canonical para. of ExpMod, $O(n^{-1})$ approx.

$V = \frac{\partial y}{\partial \theta} \Big|_{\text{fixed "p-value" vector } y^o, \hat{\theta}^o}$
 $\varphi = \frac{\partial}{\partial V} l(\theta; y) \Big|_{y^o}$

} Then use φ -based Exp model for inference!

4 Exponential Model:

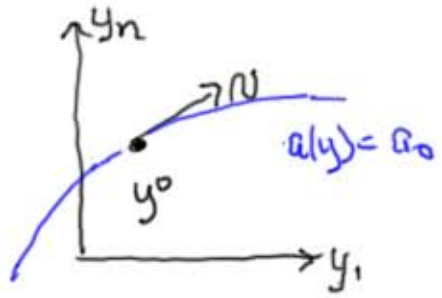
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$y = y(u, \theta)$ & data $y^o, \hat{\theta}^o$ but V tangent to ancillary

$V = (N_1, \dots, N_p) = \frac{\partial}{\partial \theta} y(u; \theta) |_{y^o, \hat{\theta}^o}$ (continuity: FF Staicu, Bern. 2010)



and get $\varphi(\theta) = \frac{\partial}{\partial V} l(\theta; y) |_{y^o} \rightarrow$ Canonical para. of ExpMod, $O(n^{-1})$ approx.

$f_T(\Delta; \varphi) = \exp\{l(\varphi) + \Delta \varphi\} h(\Delta)$ $\Delta^o = 0$ Conditional on ancillary

4 Exponential Model:

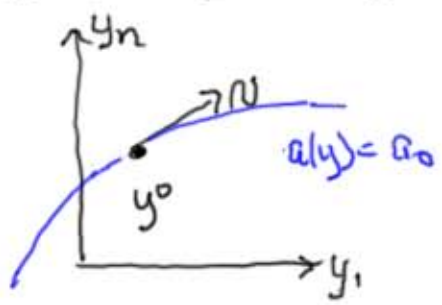
As a building block? or As an approximation to a model $O(\bar{n}')$

Consider $f(y; \theta)$ regular:

$f(y; \theta)$ density, or distn fns, or quantile fns (Often easily available; often used simulations)
 $n \times p$ quantile version

$y = y(u, \theta)$ & data $y^o, \hat{\theta}^o$ but V tangent to ancillary

$V = (N_1, \dots, N_p) = \frac{\partial}{\partial \theta} y(u; \theta) |_{y^o, \hat{\theta}^o}$ (continuity: FF Staicu, Bern. 2010)



and get $\varphi(\theta) = \frac{\partial}{\partial V} l(\theta; y) |_{y^o} \rightarrow$ Canonical para. of ExpMod, $O(\bar{n}')$ approx.

$f_T(\lambda; \varphi) = \exp\{l(\varphi) + \lambda \varphi\} h(\lambda)$ $\lambda^o = 0$ Conditional on ancillary

$l(\varphi; \lambda) = \varphi \lambda + l^o(\varphi)$ 2nd order mod decons $\lambda^o = 0$

Act as if Exptl model $O(\bar{n}')$ and use SP

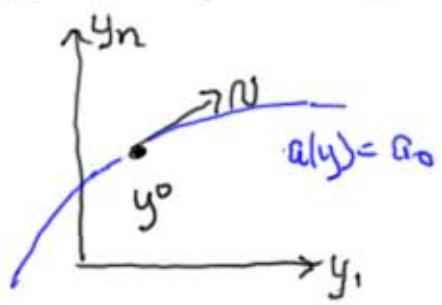
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$V = (N_1, \dots, N_p) = \frac{\partial}{\partial \theta} y(u; \theta) |_{y^o, \hat{\theta}^o}$ (continuity: FF Staicu, Bern. 2010)

and get $\varphi(\theta) = \frac{\partial}{\partial V} l(\theta; y) |_{y^o} \rightarrow$ Canonical para. of ExpMod, $O(\bar{n}')$ approx.

$f_T(\Delta; \varphi) = \exp\{l(\varphi) + \Delta \varphi\} h(\Delta)$ $\Delta^o = 0$ Conditional on ancillary

$l(\varphi; \Delta) = \varphi \Delta + l^o(\varphi)$ 2nd order mod decons Act as if Exptl model $O(\bar{n}')$

Does Welch-Peers extend? Yes! but not as Jeffreys hoped! $\rightarrow |l_{\theta\theta}(\theta)|^{1/2}$

5 Saddlepoint For exponential models (general, by approximation)

Exp. mod.: $f(\eta; \varphi) = \exp\{\eta^T \varphi + l(\varphi)\} h(\eta)$

Centered
coordinates

$$\eta^0 = 0$$

Daniels 1954
B-N Cox 1979

5 Saddlepoint For exponential models (general, by approximation)

Exp. mod.: $f(\delta; \varphi) = \exp\{\delta^T \varphi + l(\varphi)\} h(\delta)$

Centered
coordinates

$$\delta^0 = 0$$

Daniels 1954
B-N Cox 1979

SP: $f(\delta; \varphi) d\delta = \frac{e^{k/n}}{(2\pi)^{p/2}} \exp\left\{-\frac{\tilde{l}^2(\varphi; \delta)}{2}\right\} \cdot \left|\hat{J}_{\varphi\varphi}(\hat{\varphi})\right|^{-1/2} d\delta$
 $\cdot \left|\hat{J}_{\varphi\varphi}(\hat{\varphi})\right|^{1/2} d\hat{\varphi}$

$$O(n^{-3/2})$$

with or without
scaling here

5 Saddlepoint For exponential models (general, by approximation)

Exp. mod.: $f(\delta; \varphi) = \exp\{\delta^T \varphi + \ell(\varphi)\} h(\delta)$

Centered coordinates

$\delta^0 = 0$

Daniels 1954
B-N Cox 1979

SP: $f(\delta; \varphi) d\delta = \frac{e^{k/n}}{(2\pi)^{p/2}} \exp\left\{-\frac{\tilde{\ell}^2(\varphi; \delta)}{2}\right\} \cdot \left| \hat{J}_{\varphi\varphi}(\hat{\varphi}) \right|^{-1/2} d\delta \cdot \left| \hat{J}_{\varphi\varphi}(\hat{\varphi}) \right|^{1/2} d\hat{\varphi}$

$O(n^{-3/2})$

$r^2/2 = \ell(\hat{\varphi}; \delta) - \ell(\varphi; \delta)$
 $= \log \text{lik ratio}$
 $\hat{\varphi} = \text{argsup } \ell(\varphi; \delta)$

Simple pieces \Rightarrow lik ratio \uparrow \curvearrowright Exp'd info / obs info... re φ

5 Saddlepoint For exponential models (general, by approximation)

Exp. mod: $f(s; \varphi) = \exp\{s^T \varphi + l(\varphi)\} h(s)$

Centered coordinates

$$s^0 = 0$$

Daniels 1954
B-N Cox 1979

SP: $f(s; \varphi) ds = \frac{e^{k/n}}{(2\pi)^{p/2}} \exp\left\{-\frac{n^2 l(\varphi; s)}{2}\right\} \cdot \left| \hat{J}_{\varphi\varphi}(\hat{\varphi}) \right|^{-1/2} ds \cdot \left| \hat{J}_{\varphi\varphi}(\hat{\varphi}) \right|^{1/2} d\hat{\varphi}$

$O(n^{-3/2})$ $r^2/2 = l(\hat{\varphi}; s) - l(\varphi; s)$
 $= \log \text{lik ratio}$
 $\hat{\varphi} = \text{argsup } l(\varphi; s)$

Simple pieces \Rightarrow lik ratio \uparrow \uparrow Exp'd info / obs info ... re φ

Accuracy & Separation component parameters: \leftarrow Tilt = $s\varphi + t\lambda$

SP $f(s, t; \varphi, \lambda) = \frac{e^{k/n}}{(2\pi)^{p/2}} \exp\left\{-\frac{n^2 l(\varphi, \lambda; s, t)}{2}\right\} \left| \hat{J}_{\varphi\varphi}(\hat{\varphi}, \hat{\lambda}) \right|^{-1/2}$

$$\begin{aligned} \dim s &= \dim \varphi = d \\ \dim t &= \dim \lambda = p - d \end{aligned}$$

5 Saddlepoint For exponential models (general, by approximation)

Exp. mod: $f(s; \varphi) = \exp\{s^T \varphi + l(\varphi)\} h(s)$

Centered coordinates

$s^0 = 0$

Daniels 1954
B-N Cox 1979

SP: $f(s; \varphi) ds = \frac{e^{k/n}}{(2\pi)^{p/2}} \exp\left\{-\frac{\tilde{l}^2(\varphi; s)}{2}\right\} \cdot \left| \hat{J}_{\varphi\varphi}(\hat{\varphi}) \right|^{-1/2} ds \cdot \left| \hat{J}_{\varphi\varphi}(\hat{\varphi}) \right|^{1/2} d\hat{\varphi}$

$O(n^{-3/2})$

$r^2/2 = l(\hat{\varphi}; s) - l(\varphi; s)$
 $= \log \text{lik ratio}$
 $\hat{\varphi} = \text{argsup } l(\varphi; s)$

Simple pieces \Rightarrow lik ratio \uparrow \uparrow Exp'd info / obs info ... re φ

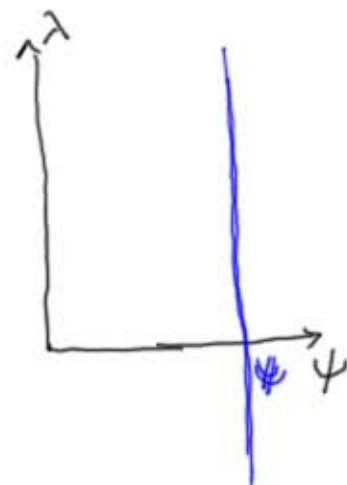
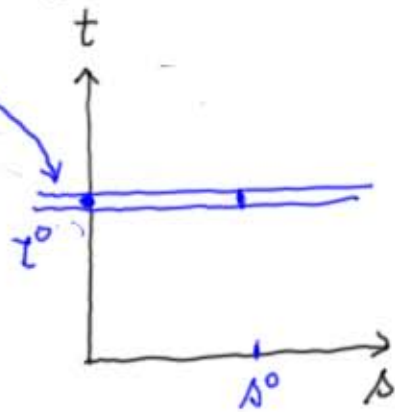
Accuracy & Separation component parameters: \leftarrow Tilt $s\psi + t\lambda$

SP $f(s, t; \psi, \lambda) = \frac{e^{k/n}}{(2\pi)^{p/2}} \exp\left\{-\frac{\tilde{l}^2(\psi, \lambda; s, t)}{2}\right\} \left| \hat{J}_{\varphi\varphi}(\hat{\psi}, \hat{\lambda}) \right|^{-2}$

$\dim s = \dim \psi = d$
 $\dim t = \dim \lambda = p - d$

Interest in ψ $s|t^0$ Exp'l

ψ linear in φ
 e.g. $\varphi = (\psi, \lambda)$



5 Saddlepoint For exponential models (general, by approximation)

Exp. mod: $f(s; \varphi) = \exp\{s^T \varphi + l(\varphi)\} h(s)$

Centered coordinates $s^0 = 0$ Daniels 1954
B-N Cox 1979

SP: $f(s; \varphi) ds = \frac{e^{k/n}}{(2\pi)^{p/2}} \exp\left\{-\frac{\tilde{l}^2(\varphi; s)}{2}\right\} \cdot \left| \hat{J}_{\varphi\varphi}(\hat{\varphi}) \right|^{-1/2} ds \cdot \left| \hat{J}_{\varphi\varphi}(\hat{\varphi}) \right|^{1/2} d\hat{\varphi}$

$O(n^{-3/2})$ $r^2/2 = l(\hat{\varphi}; s) - l(\varphi; s)$
 $= \log \text{lik ratio}$
 $\hat{\varphi} = \text{arg sup } l(\varphi; s)$

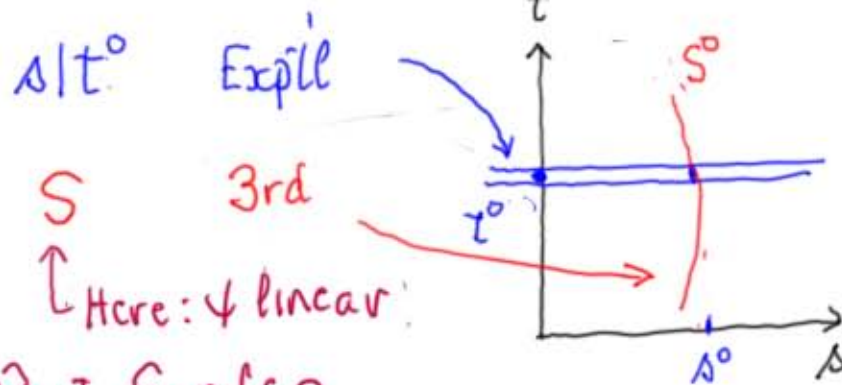
Simple pieces \Rightarrow lik ratio \uparrow \uparrow Exp'd info / obs info ... re φ

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$\dim s = \dim \varphi = d$
 $\dim t = \dim \lambda = p - d$

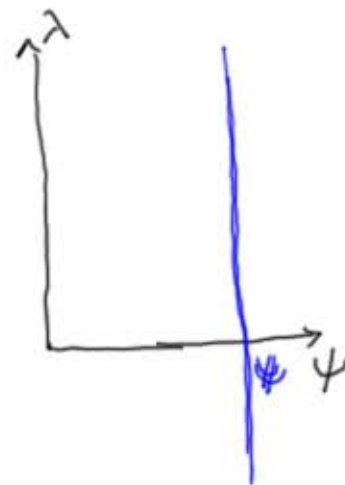
Interest in φ



Interest in λ

\uparrow Here: φ linear

But s also available for curved φ !



... | φ |

5 Saddlepoint For exponential models (general, by approximation)

Exp. mod: $f(s; \varphi) = \exp\{s^T \varphi + l(\varphi)\} h(s)$

Centered coordinates $s^0 = 0$ Daniels 1954
B-N Cox 1979

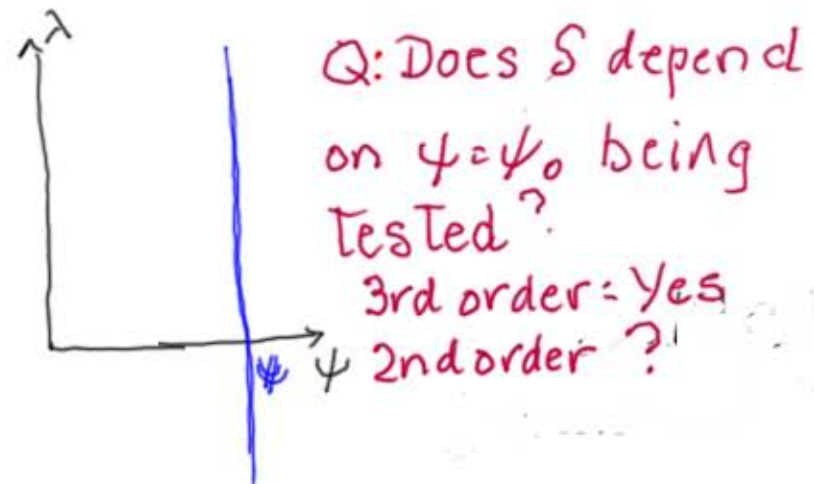
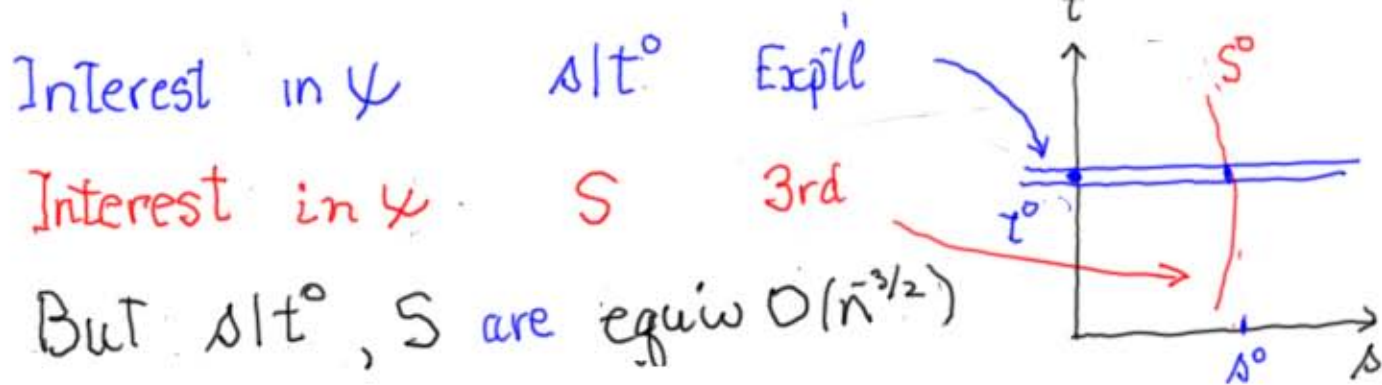
SP: $f(s; \varphi) ds = \frac{e^{k/n}}{(2\pi)^{p/2}} \exp\left\{-\frac{\tilde{l}^2(\varphi; s)}{2}\right\} \cdot \left| \hat{J}_{\varphi\varphi}(\hat{\varphi}) \right|^{-1/2} ds$
 $\cdot \left| \hat{J}_{\varphi\varphi}(\hat{\varphi}) \right|^{1/2} d\hat{\varphi}$ $O(n^{-3/2})$ $r^2/2 = l(\hat{\varphi}; s) - l(\varphi; s)$
 $= \log \text{lik ratio}$
 $\hat{\varphi} = \text{argsup } l(\varphi; s)$

Simple pieces \Rightarrow lik ratio \uparrow \uparrow Exp'd info / obs info ... re φ

Accuracy & Separation component parameters: \leftarrow Tilt $s\varphi + t\lambda$

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$\dim s = \dim \varphi = d$
 $\dim t = \dim \lambda = p - d$



Q: Does S depend on $\varphi = \varphi_0$ being tested?
 3rd order = Yes
 2nd order?

But $s|t^0, S$ are equiv $O(n^{-3/2})$
 SP makes dist'ns available

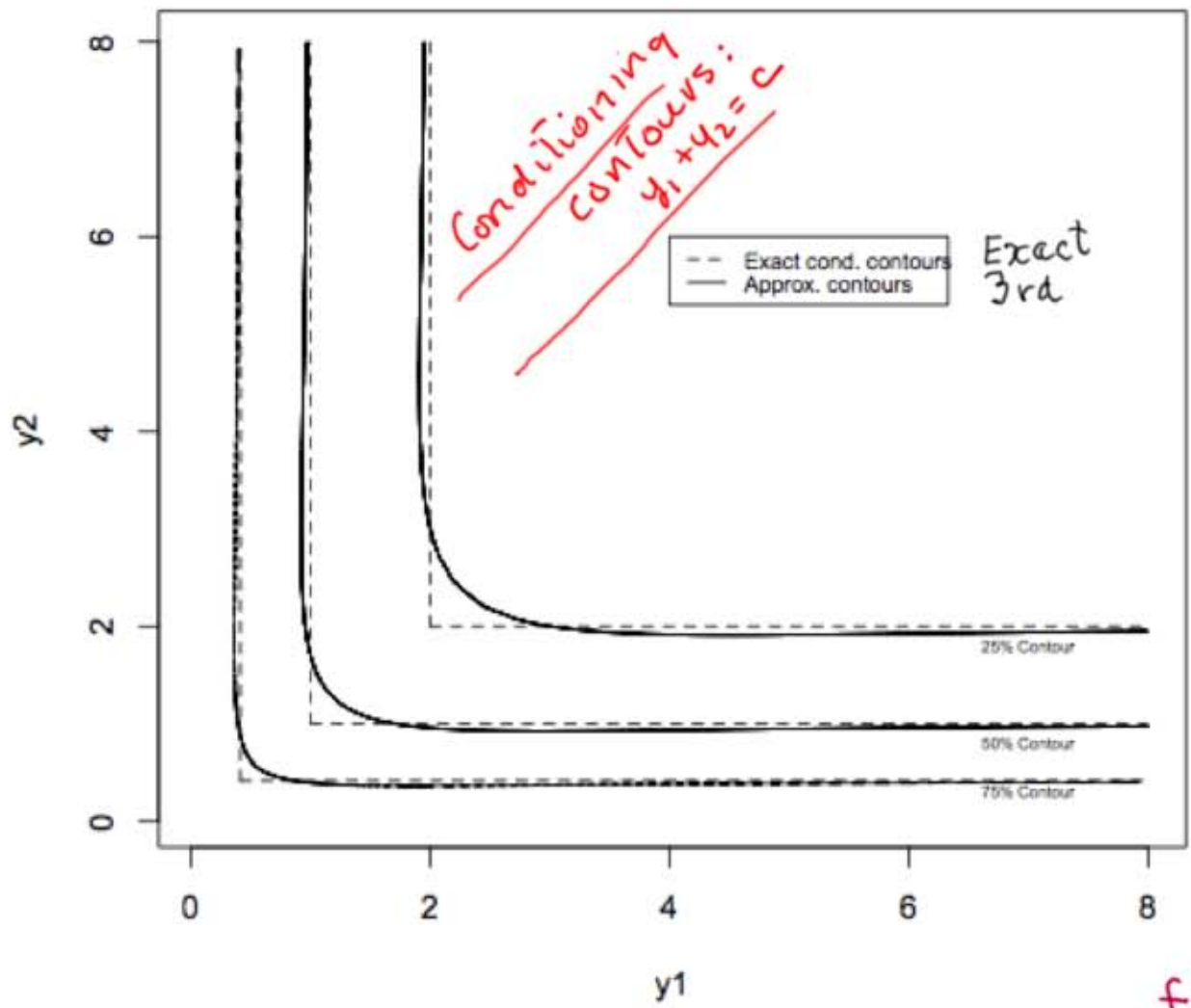


Fig. 1. The exact conditional contours and the third order approximate contours at quantile levels 25%, 50% and 75% for testing $\psi_0 = 0.6931$ in the simple exponential life model

Asymptotics but $n=1$

$$y_1 \sim \varphi_1 e^{-\varphi_1 y_1}$$

$$y_2 \sim \varphi_2 e^{-\varphi_2 y_2}$$

Interest $\psi = \varphi_1 + \varphi_2$

Test $\psi = .4901$

3rd for $s|t$ | Equal
 3rd for S

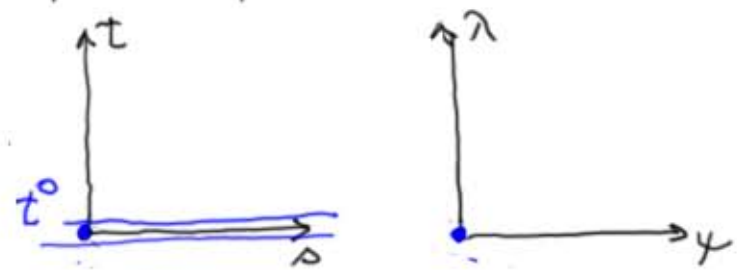
$$f = e^{\log \varphi_1 + \log \varphi_2 - \varphi_1 y_1 - \varphi_2 y_2}$$

S available as power series

6 Conditional & Marginal for linear component parameter

$$f(s, t; \psi, \lambda) = \frac{e^{R/n}}{(2\pi)^{p/2}} \exp\{s\psi + t\lambda + l(\psi, \lambda)\} |\hat{J}(s, t)|^{-1/2}$$

get $s | t^0$



$$|\hat{J}_{(\psi, \lambda)(\psi, \lambda)} \{\hat{\psi}(s, t), \hat{\lambda}(s, t)\}|^{1/2}$$

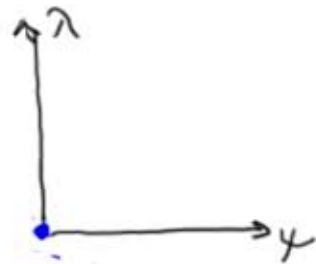
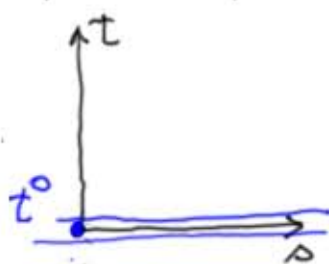
2nd derivative info matrix in canonical parameterization.

6 Conditional & Marginal for linear component parameter

6

$$f(s, t; \psi, \lambda) = \frac{e^{R/n}}{(2\pi)^{p/2}} \exp\{s\psi + t\lambda + l(\psi, \lambda)\} |\hat{J}(s, t)|^{-1/2}$$

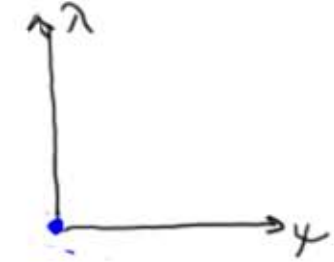
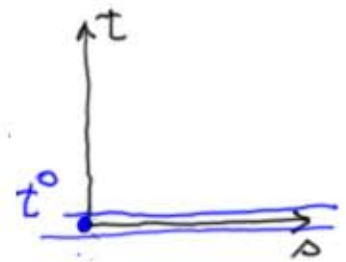
Get: $s | t^0$ Exponential in ψ Use SP \uparrow



6 Conditional & Marginal for linear component parameter

$$f(s, t; \psi, \lambda) = \frac{e^{R/n}}{(2\pi)^{p/2}} \exp\{s\psi + t\lambda + \ell(\psi, \lambda)\} |\hat{j}(s, t)|^{-1/2}$$

$s | t^0$ Exponential in ψ Use SP \uparrow



Case: $\dim \psi = \dim \lambda = 1$ \hat{j} factors Cond'n't info = $\int \lambda \lambda$
 Profile info = $\hat{j}_{\psi\psi}^P$

$\hat{\lambda}_\psi$ free of s
 $\hat{j}_{\lambda\lambda}$ free of s

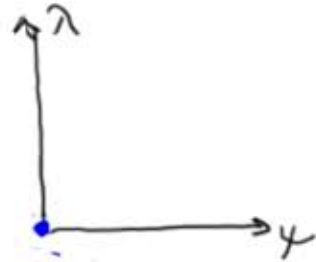
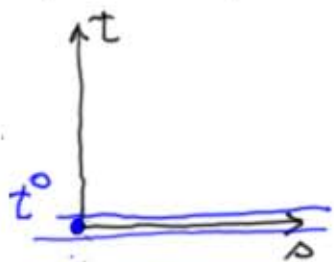
6 Conditional & Marginal for linear component parameter

$$f(s, t; \psi, \lambda) = \frac{e^{R/n}}{(2\pi)^{p/2}} \exp\{s\psi + t\lambda + \ell(\psi, \lambda)\} |\hat{J}(s, t)|^{-1/2}$$

$s | t^0$

Exponential in ψ

Use SP \uparrow



Case: $\dim \psi = \dim \lambda = 1$

\hat{J} factors

Cond'it info = $\int \lambda \lambda$

Profile info = $\int_{\psi} \hat{J}_{\psi}^P$

$\hat{\lambda}_{\psi}$ free of s

$\hat{\lambda}_{\lambda}$ free of s

SP gives

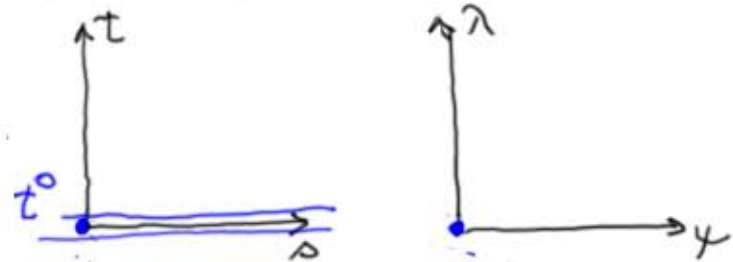
$$f(s | t^0) = \frac{e^{R/n}}{(2\pi)^{1/2}} e^{-\frac{R^2}{2} \frac{\ell_{\psi}(\psi, \lambda)}{\ell_{\psi}(\psi, \lambda)}} \left| \int_{\psi} \hat{J}_{\psi}^P \right|^{-1/2}$$

- 1st: t^0 density
- 2nd: Likelihood adj.

6 Conditional & Marginal for linear component parameter

$$f(s, t; \psi, \lambda) = \frac{e^{R/n}}{(2\pi)^{p/2}} \exp\{s\psi + t\lambda + \ell(\psi, \lambda)\} |\hat{J}(s, t)|^{-1/2}$$

$s | t^0$ Exponential in ψ Use SP \uparrow



Case: $\dim \psi = \dim \lambda = 1$ \hat{J} factors
 Cond'n't info = $\int \lambda \lambda$
 Profile info = $\int \psi \psi$

$\hat{\lambda}_\psi$ free of s
 $\hat{\lambda}_\lambda$ free of s

SP gives

$$f(s | t^0) = \frac{e^{R/n}}{(2\pi)^{1/2}} e^{-\frac{R^2}{2} \frac{\ell(\psi, \Delta)}{\psi}} \left| \int \psi \right|^{-1/2}$$

1st: t^0 density
 2nd: Likelihood adj.

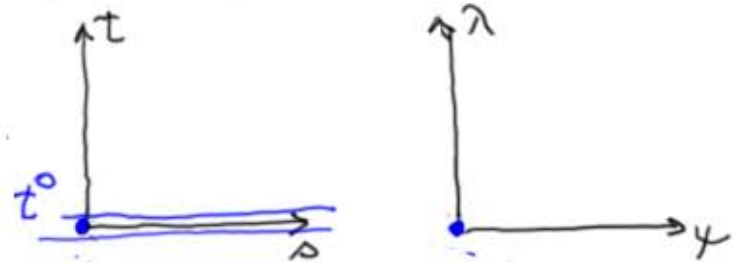
Q: and = $f(s)$ for $S = ?$ A: For linear ψ , available as

1) Inference for ψ Cond'n't and marginal same (3rd) for linear parameters
 composite of conditional $s | t$

6 Conditional & Marginal for linear component parameter

$$f(s, t; \psi, \lambda) = \frac{e^{R/n}}{(2\pi)^{p/2}} \exp\{s\psi + t\lambda + \ell(\psi, \lambda)\} |\hat{J}(s, t)|^{-1/2}$$

$s | t^0$ Exponential in ψ Use SP \uparrow



Case: $\dim \psi = \dim \lambda = 1$ \hat{J} factors
 Cond'n't info = $\int_{\lambda} \hat{J}_{\lambda\lambda}$
 Profile info = $\int_{\psi} \hat{J}_{\psi\psi}^P$

$\hat{\lambda}_\psi$ free of s
 $\hat{\lambda}_\lambda$ free of s

SP gives

$$f(s | t^0) = \frac{e^{R/n}}{(2\pi)^{1/2}} e^{-\frac{r_p^2(\psi, \Delta)}{2}} \left| \int_{\psi} \hat{J}_{\psi\psi}^P \right|^{-1/2}$$

1st: t^0 density
 2nd: Likelihood adj.

1) Inference for ψ Cond'n't and marginal same (3rd)

2) Bayes for ψ $\left| \int_{\psi} \hat{J}_{\psi\psi}(\psi, \lambda^0) \right|^{1/2} d\psi$ Conditional same (2nd) for lin. par.

Same by Welch-Peers

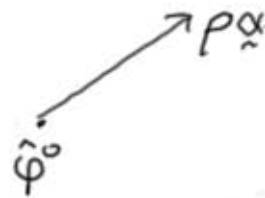
But also in mod. deons $O(n^{-1})$ available from $\ell(\psi)$; see §7 ahead

7 Directional Jeffreys: $O(n^{-1})$ for linear parameters

7

$$\varphi = (\varphi_1, \dots, \varphi_p)$$

$$= \hat{\varphi}_{\text{data}}^0 + \rho \alpha$$



7 Directional Jeffreys: $O(n')$ for linear parameters

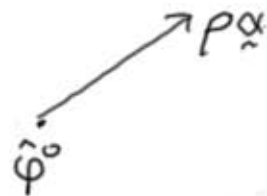
$$\varphi = (\varphi_1, \dots, \varphi_p)$$

$$= \hat{\varphi}^0 + \rho \alpha$$

data

Polar coords; direction α

"Rotated ψ " as ρ ; data $\hat{\rho}^0 = 0$



7 Directional Jeffreys: $O(n')$ for linear parameters

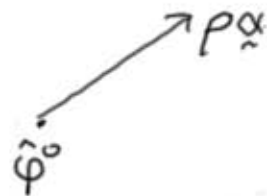
$$\varphi = (\varphi_1, \dots, \varphi_p)$$

Polar coord's; direction α

$$= \hat{\varphi}^0 + \rho \alpha$$

data

"Rotated ψ " as ρ ; data $\hat{\rho}^0 = 0$



From preceding: Marginal dist'n re ρ = Condnl re ρ
& both equal the Bayes by Welch-Peers

$$- \frac{\partial^2 \ell(\varphi)}{\partial \rho^2}$$

- Canonical coordinates

- Diff twice re ρ

with ρ defined re $\hat{\varphi}^0$ (original φ coord's)

7 Directional Jeffreys: $O(n^{-1})$ for linear parameters

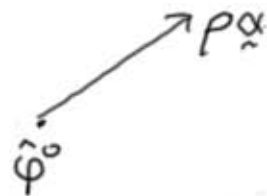
$$\varphi = (\varphi_1, \dots, \varphi_p)$$

Polar coord's; direction α

$$= \hat{\varphi}^0 + \rho \alpha$$

data

"Rotated ψ " as ρ ; data $\hat{\rho}^0 = 0$



From preceding: Marginal dist'n re ρ = Condnl re ρ
& both equal the Bayes by Welch-Peers

Bayes is $\pi(\theta) = \int_{\mathbb{R}^2} p(\alpha) \cdot d\rho$ on line; repetition validity $O(n^{-1})$

Polar coord's re obs mle in φ
Interpretation for "derivation": mod dev'n's about $(s^0, \hat{\varphi}^0)$
thus include just $n^{1/2}$ terms in expansions
Derivatives effectively of the cum. gen. fn.

Q: Would this be used only on line for particular ψ of interest?
A: That would be easiest integration!

Q: Or used on full φ -space?
A: Indications are Yes! Details elsewhere.

7 Directional Jeffreys: $O(n^{-1})$ for linear parameters

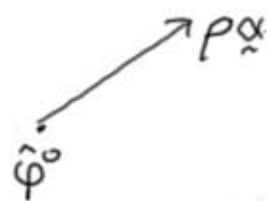
$$\varphi = (\varphi_1, \dots, \varphi_p)$$

Polar coord's; direction α

$$= \hat{\varphi}^0 + \rho \alpha$$

data

"Rotated ψ " as ρ ; data $\hat{\rho}^0 = 0$



From preceding: Marginal dist'n re ρ = Condnl re ρ
& both equal the Bayes by Welch-Peers

Bayes is $\pi(\theta) = \int \ell(\rho \alpha)^{1/2} \cdot d\rho$ on line repetition validity $O(n^{-1})$

Does right things: $N \mu \sigma^2$ $d\mu d\sigma / \sigma$

$y = X\beta + \sigma z$ $d\beta d\sigma / \sigma$

7 Directional Jeffreys: $O(n^{-1})$ for linear parameters

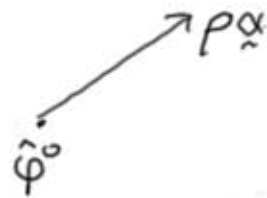
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Does right things: $N \mu \sigma^2 \quad d\mu d\sigma / \sigma$

$$y = X\beta + \sigma z \quad d\beta d\sigma / \sigma$$

In general: Data dependent ... necessarily so !

Priors need to be data-dependent to extract all available info from likelihood function.

Q Fienler-Creasy problem: any light on this?

A Creasy believed she was following Fisher

but the angle (here α) has no $-\infty, +\infty$

Don't think Creasy solution would

fit present analysis

Why should a Bayesian "sell himself short" by using only a prior with definition free of the data point?

7 Directional Jeffreys: $O(n^{-1})$ for linear parameters

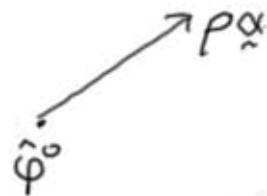
$$\varphi = (\varphi_1, \dots, \varphi_p)$$

Polar coord's; direction α

$$= \hat{\varphi}^0 + \rho \alpha$$

data

"Rotated ψ " as ρ ; data $\hat{\rho}^0 = 0$



From preceding: Marginal dist'n re ρ = Condnl re ρ
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Bayes is $\pi(\theta) = \int_{\mathbb{R}^p} \pi(\rho \alpha) \cdot d\rho$ on line repetition validity $O(n^{-1})$

Does right things: $N \mu \sigma^2 \quad d\mu d\sigma / \sigma$

$$y = X\beta + \sigma z \quad d\beta d\sigma / \sigma$$

In general: Data dependent ... necessarily so!

But just for linear parameters

Curved parameters \Rightarrow Bayes fails

AQ: Even with data-
dependent priors?
Yes! To work with curved
parameters, the prior
needs to be specially
designed for the curved
parameter: draft
paper available.

See Section 8

8 Parameter curvature How Bayes misses!

$y_1 \sim N(\psi, 1)$

Interest in ψ

Data y^o

Suppose $y_i^o - \psi_0 = 1$

$y_2 \sim N(\lambda, 1)$

Nuisance λ

Test ψ_0

8 Parameter curvature How Bayes misses!

$y_1 \sim N(\psi, 1)$

Interest in ψ

Data y_1^o

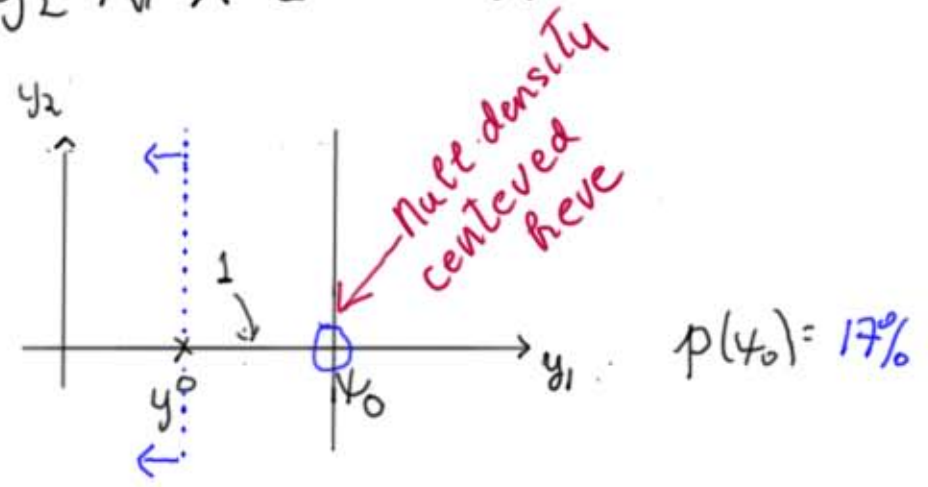
$y_2 \sim N(\lambda, 1)$

Nuisance λ

Test ψ_0

Suppose $y_1^o - \psi_0 = 1$

f



8 Parameter curvature How Bayes misses!

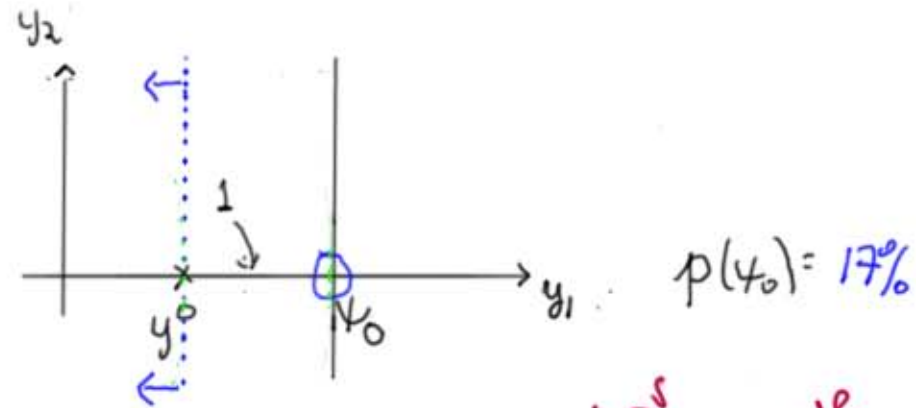
$y_1 \sim N(\psi, 1)$
 $y_2 \sim N(\lambda, 1)$

Interest in ψ
 Nuisance λ

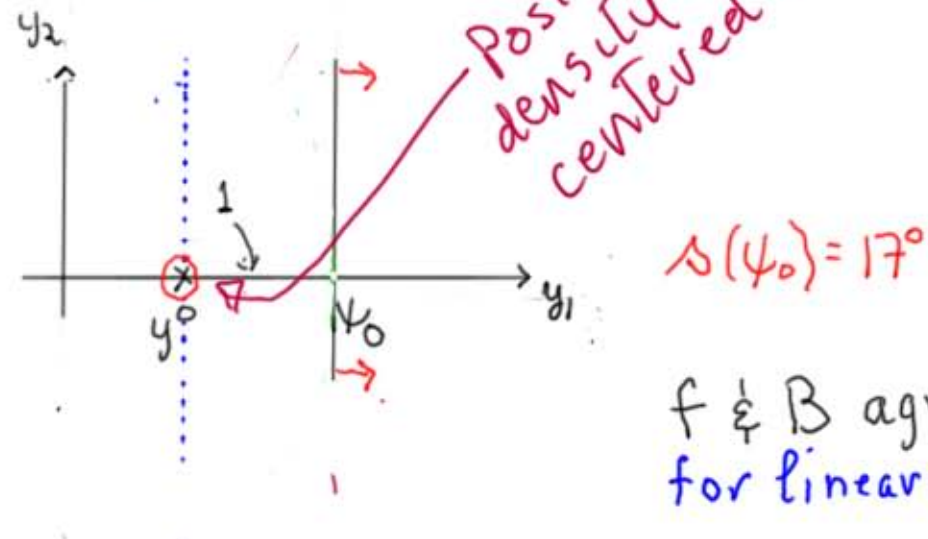
Data y_1^o
 Test ψ_0

Suppose $y_1^o - \psi_0 = 1$

f



B



f & B agree for linear par.

8 Parameter curvature How Bayes misses!

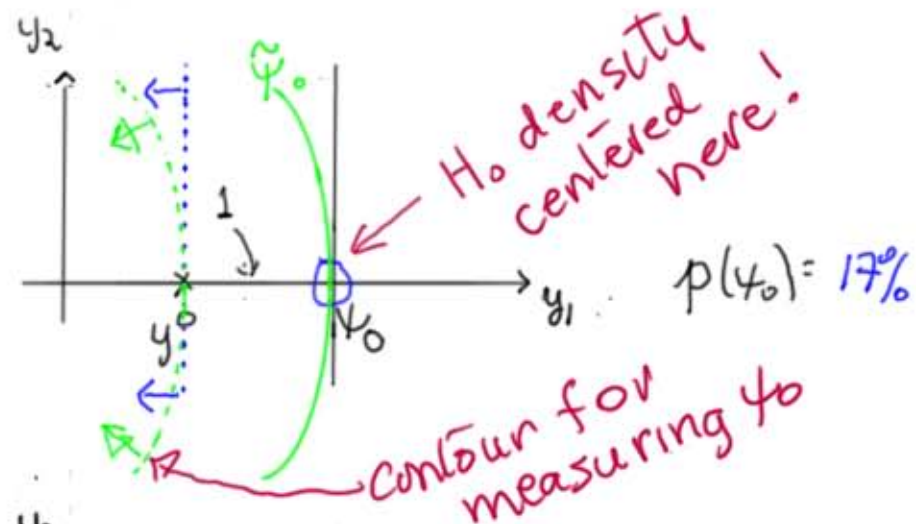
$y_1 \sim N(\psi, 1)$
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Interest in ψ
 Nuisance λ

Data y_1^o
 Test ψ_0

Suppose $y_1^o - \psi_0 = 1$

f

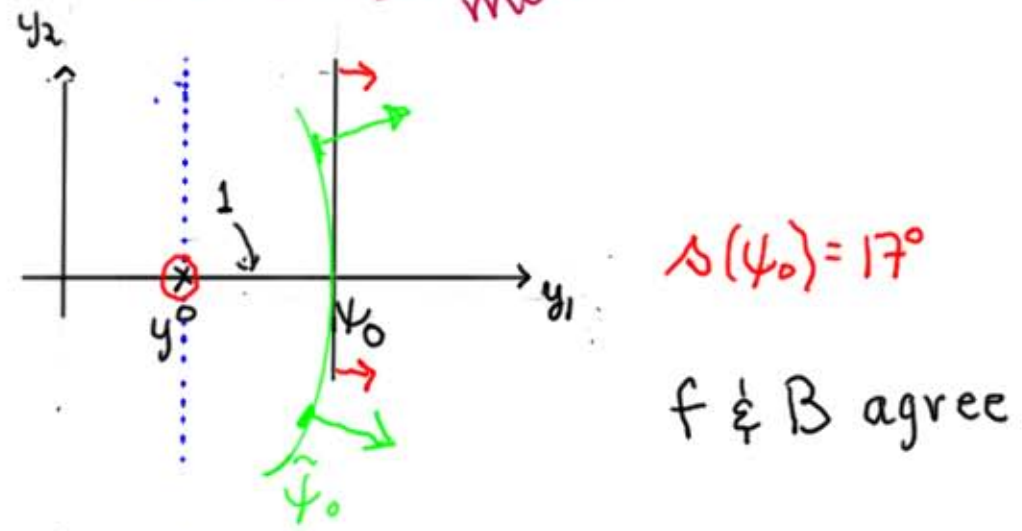


Now consider curved $\tilde{\psi} = \tilde{\psi}_0$

p-value decreases

$p(\tilde{\psi}_0) < 17\%$

B



$\psi_0 = 17\%$
 f & B agree

8 Parameter curvature How Bayes misses!

$y_1 \sim N(\psi, 1)$
 $y_2 \sim N(\lambda, 1)$

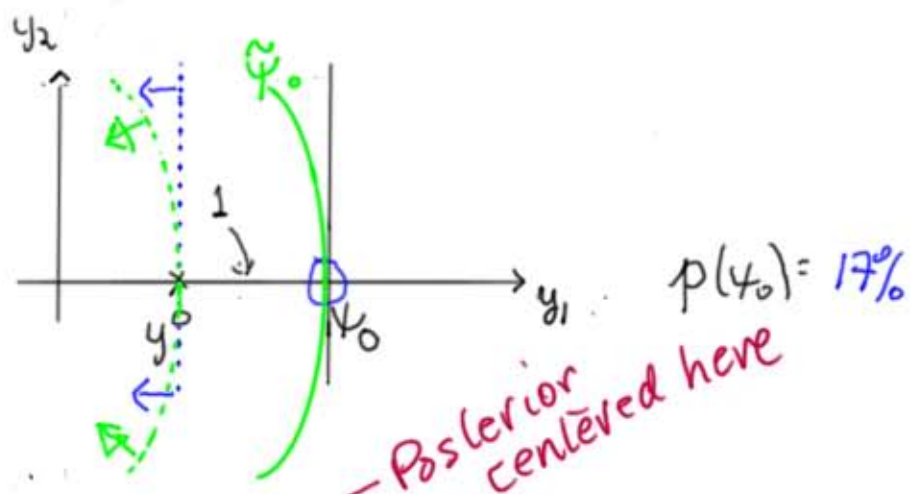
Interest in ψ
 Nuisance λ

Data y_1^o
 Test ψ_0

Suppose $y_1^o - \psi_0 = 1$

Now consider curved $\tilde{\psi} = \tilde{\psi}_0$

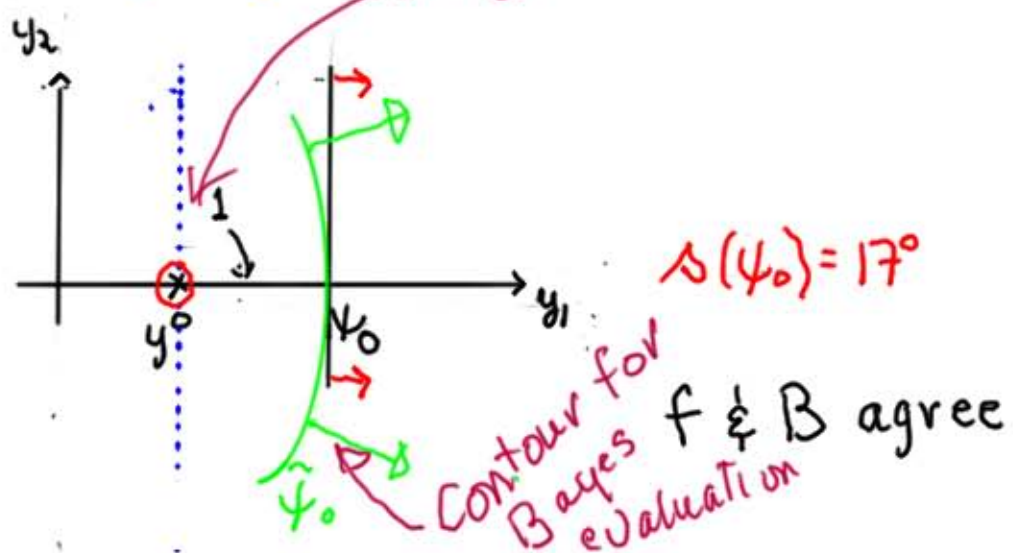
f



p-value decreases

$p(\tilde{\psi}_0) < 17\%$

B



s-value increases

$s(\tilde{\psi}_0) > 17\%$

$f \notin B$ disagree

Go in different directions
 Bayes is wrong
 for curved parameters

- 1 Priors & Inference: Is there a problem?
- 2 Data-dependent prior? Why?
- 3 Welch Peers: Exponential as location model
- 4 Exponential model: as an approximation
- 5 Saddlepoint: for exponential & for general
- 6 Conditional / Marginal for linear parameter component
- 7 Directional Jeffreys
- 8 Curved parameter: How Bayes misses!

Thank you