

Distributional Inference:

Reproducibility and limitations

"Can Bayes give 2nd order inference"

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Statistical Sciences

Univ Toronto

York Univ

2nd ISNPS Conference

Cadiz Spain

2014 June 12-16

www.utstat.toronto.edu/dfraser/documents/Cadiz2014.pdf
: some references ... / xxx.pdf

What is Distributional Inference?

Some history

What are the tools?

Scalar interest parameter

What can Saddlepoint do for Bayes?

Example 1

Scalar interest ψ , nuisance λ (2nd order)

Example 2 α linear

Example 3 μ rotating

Example 4 ψ curved

Concluding...

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Does it do what it says?

Does it behave as claimed?

Is it reproducible ?  Science 2014 Jan 17 Lead Editorial

about-face on "Data" (Science 2011 Feb 11: "Statistics" not present)
Now: all "statistics" & "design"! (Higgs effect)

Bluntly: Is $\Pr\{(\hat{\psi}_n, \infty) \text{ includes } \psi\} = \beta$? Yes or No!

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If "Maybe": "if θ values follow some pattern,..." Genuine pattern?
Efron: Science 2013 Jun

Efron: Science 2013 Jun 30

2 Some history

Two routes:

^I
Pivot Inversion

^{II}
Likelihood inversion

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Theme here: Laplace
& Differential geometry

$\Rightarrow$
?

Reproducible
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/
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1) log-likelihood; 2) a little differential geometry!

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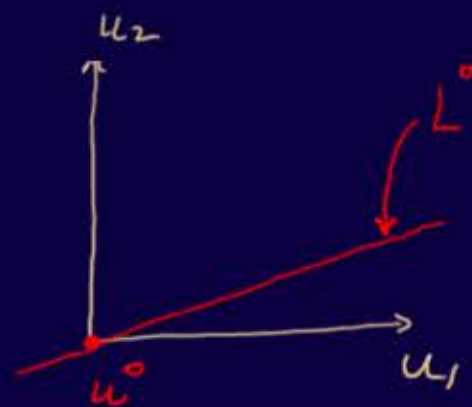
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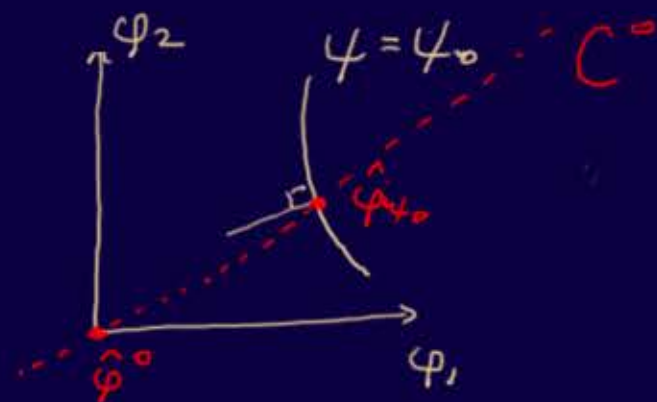
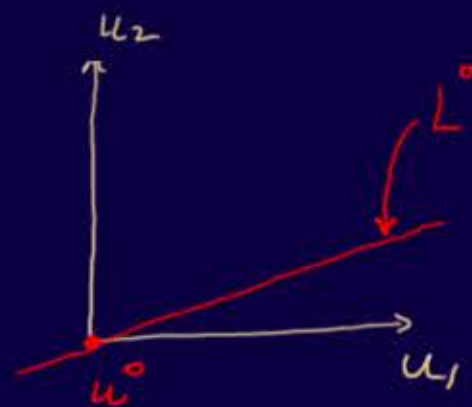
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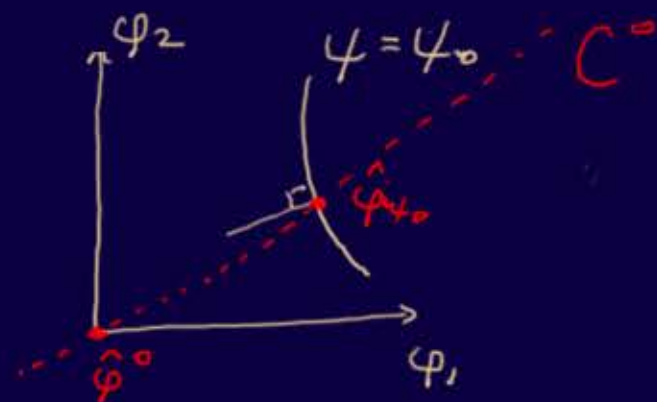
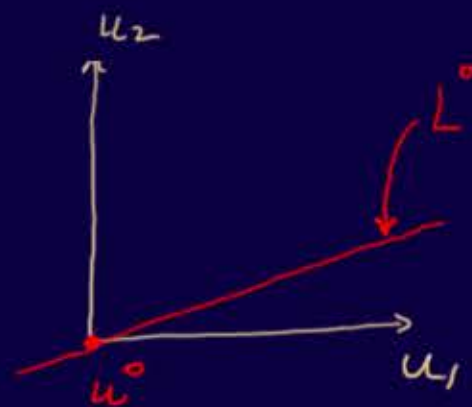
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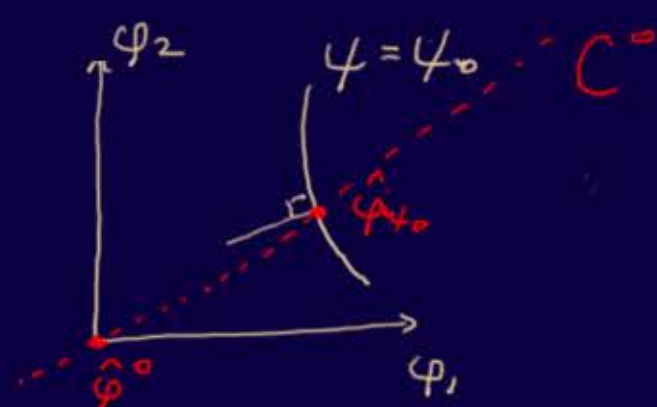
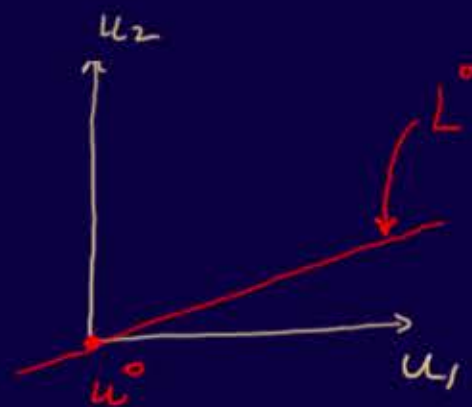
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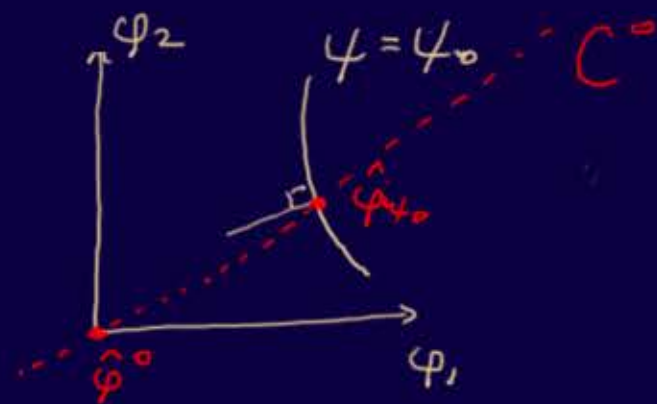
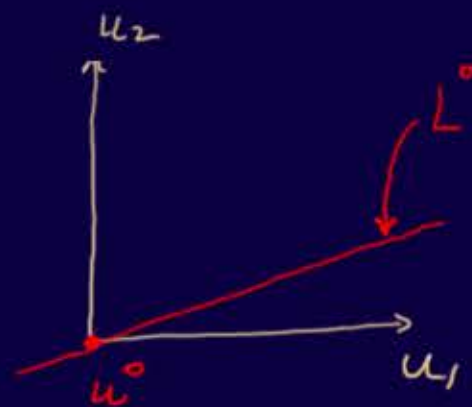
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4) Regular model; scalar parameter; go profile!

Go where the info is! (3rd)

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SP $\frac{e^{n/n}}{(2\pi)^{1/2}} e^{l - \hat{l}}^{-1/2} \int_{\varphi\varphi}(\hat{\varphi}) \cdot du$

Horizontal departure $g = (\hat{\varphi} - \varphi) \hat{l}_{\varphi\varphi}^{1/2} = \text{mle departure}$

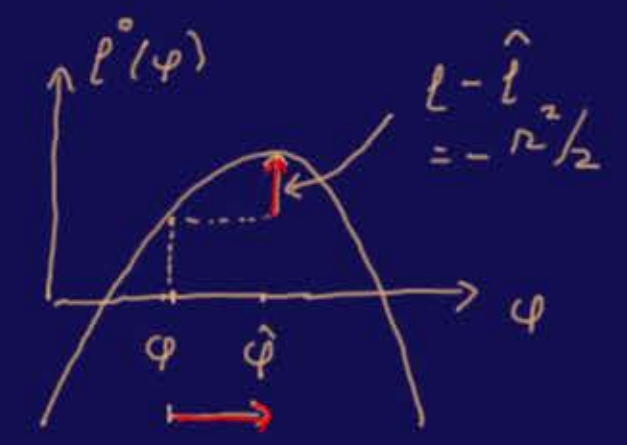
Vertical " $\hat{l} - l = n^2/2 = \text{loglik ratio}$

$n = \pm \sqrt{2(\hat{l} - l)} = \text{SLR} = \text{signed lik. root}$

Density n

$f(n) \cdot dn = e^{n/n} \underbrace{\phi(n) \frac{n}{g}}_{\text{density}} dn$

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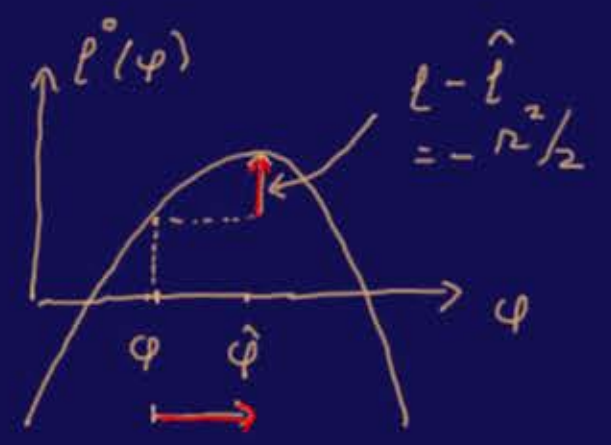
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B-N (1990)

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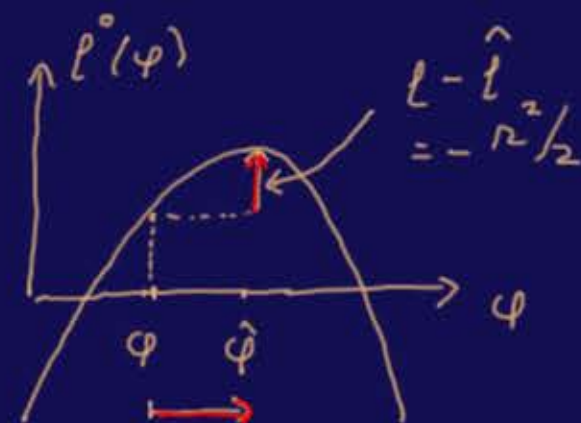
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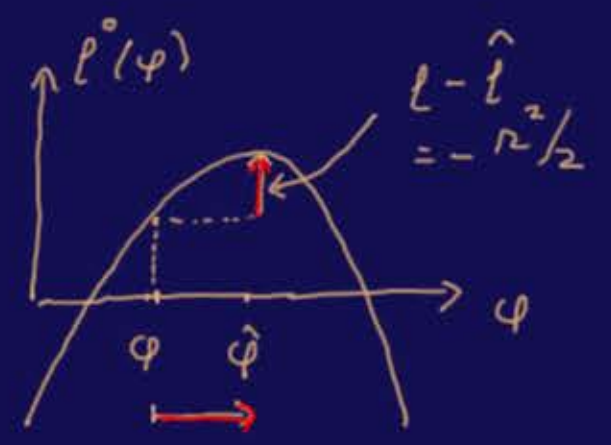
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With nuisance parameter

$g \rightarrow Q = \text{adjusted } g$



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Horizonal departure $g = (\hat{\varphi} - \varphi) \hat{J}_{\varphi\varphi}^{1/2} = \text{mle departure}$

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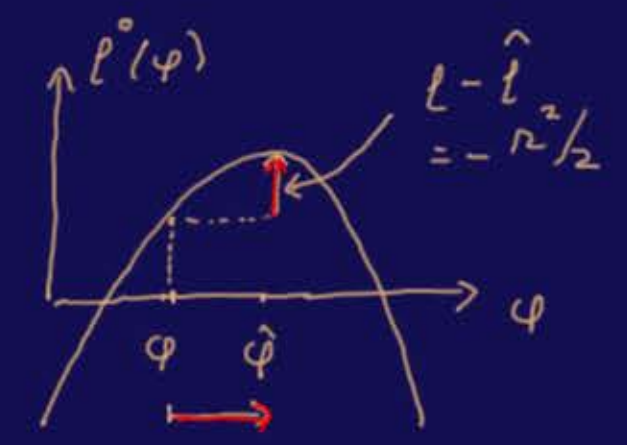
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Above by: Simple change of variable
" integration

Just calculus ! & Answers major questions !



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Expand exponent (Model) $f(\Delta; \varphi) = \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{(\Delta - \varphi)^2}{2} - \gamma \varphi^3 / 6n^{1/2} + \gamma \Delta^3 / 6n^{1/2}\right\} (1 - \gamma \Delta / 2n^{1/2})$

⑤ What can SP do for Bayes? a) Scalar parameter: Welch Peers

Have likelihood $L^o(\theta)$

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Use $\beta = \varphi + \varphi^2 / 2n^{1/2}$ $f(\hat{\beta}; \beta) = \frac{1}{\sqrt{2\pi}} \exp\left\{\frac{(\hat{\beta} - \beta)^2}{2} - \gamma \frac{(\hat{\beta} - \beta)^3}{6n^{1/2}}\right\} \cdot d\hat{\beta}$ Location model
 $\downarrow$ switch

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Location model

Posterior (loc'n flip)

$$f(\beta | \hat{\beta}) = \frac{1}{\sqrt{2\pi}} \exp\left\{\frac{(\hat{\beta} - \beta)^2}{2} - \gamma (\hat{\beta} - \beta)^3 / 6n^{1/2}\right\} \cdot d\beta$$

$$d\beta = \frac{(1 + \gamma \varphi / n^{1/2})}{1} d\varphi$$

root information prior

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root information prior

(Jeffreys prior)

Welch Peers 1963

Brown Lai Das Gupta 2001
(binomial)

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If Bayes was always so easy?

$$\pi(\varphi) = \int \varphi^{\frac{1}{2}}(\varphi) \quad \text{Jeffreys } p=1$$

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$$f(\hat{\beta}; \beta) = \frac{1}{\sqrt{2\pi}} \exp\left\{\frac{(\hat{\beta} - \beta)^2}{2} - \gamma (\hat{\beta} - \beta)^3 / 6n^{1/2}\right\} \cdot d\hat{\beta}$$

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$$\pi(\varphi) = \int_{\varphi}^{1/2}(\varphi)$$

Jeffreys $p=1$

but Jeffreys rejected his prior when $p > 1$

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$$f(\beta | \hat{\beta}) = \frac{1}{\sqrt{2\pi}} \exp\left\{\frac{(\hat{\beta} - \beta)^2}{2} - \gamma (\hat{\beta} - \beta)^3 / 6n^{1/2}\right\} \cdot d\beta$$

$d\beta = \frac{(1 + \gamma \varphi / n^{1/2})}{\text{switch}} d\varphi$

Welch Peers 1963

root information prior

Brown Cai Das Gupta 2001
(binomial)

(Jeffreys prior)

If Bayes was always so easy?

$$\pi(\varphi) = \int_{\varphi}^{1/2}(\varphi)$$

Jeffreys $p=1$

but Jeffreys rejected his prior when $p > 1$

but it maybe better than he expected!

5-1

Scalar parameter

Example 1

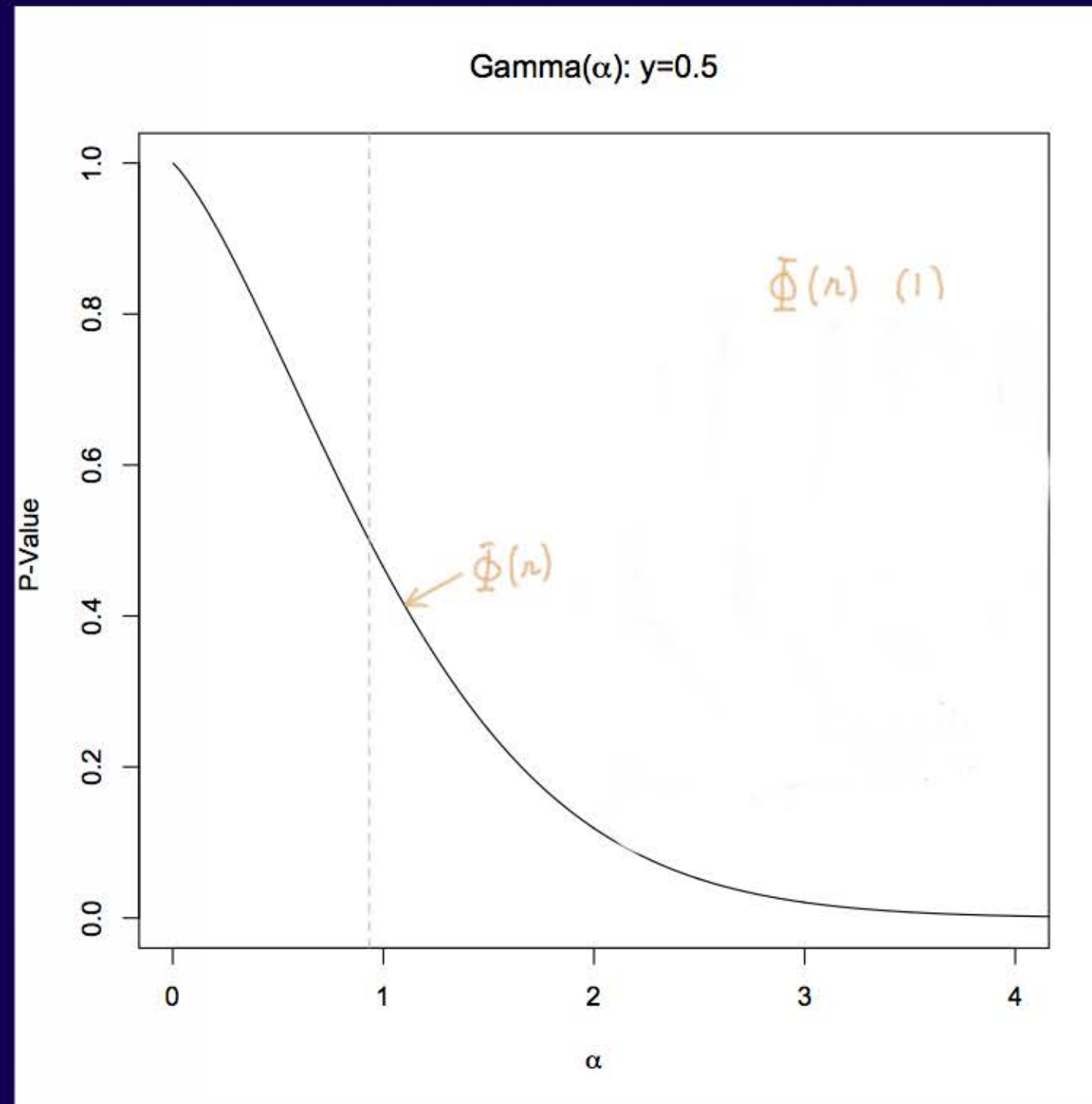
Setup

- Model: $Y \sim \text{Gamma}(\alpha, \beta)$ where α is shape and β is rate
pdf is $\frac{\beta^\alpha}{\Gamma(\alpha)} y^{\alpha-1} e^{-\beta y} \cdot dy \dots$ with $\beta = 1$

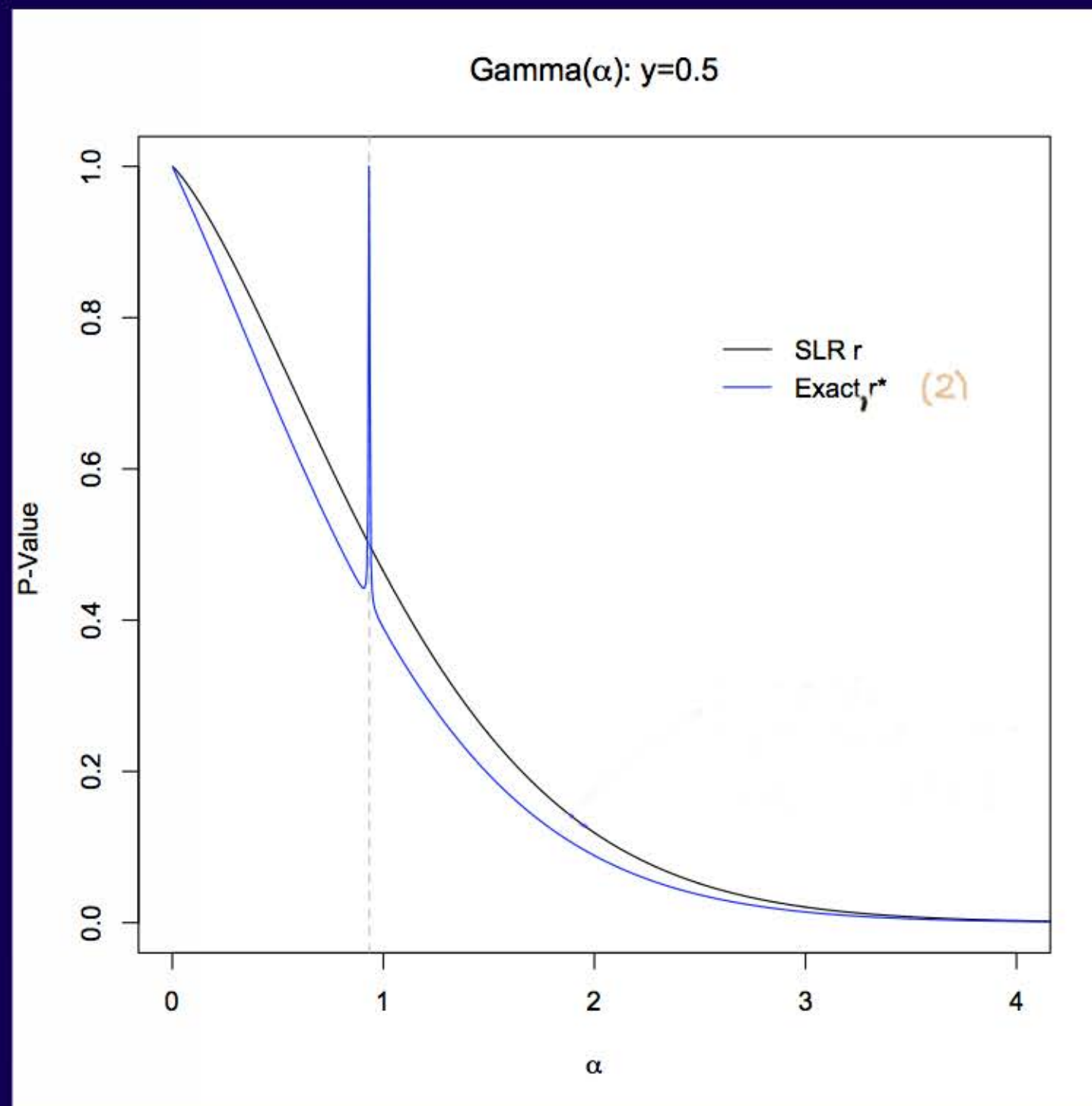
Let $n = 1$

$y^o = .5$

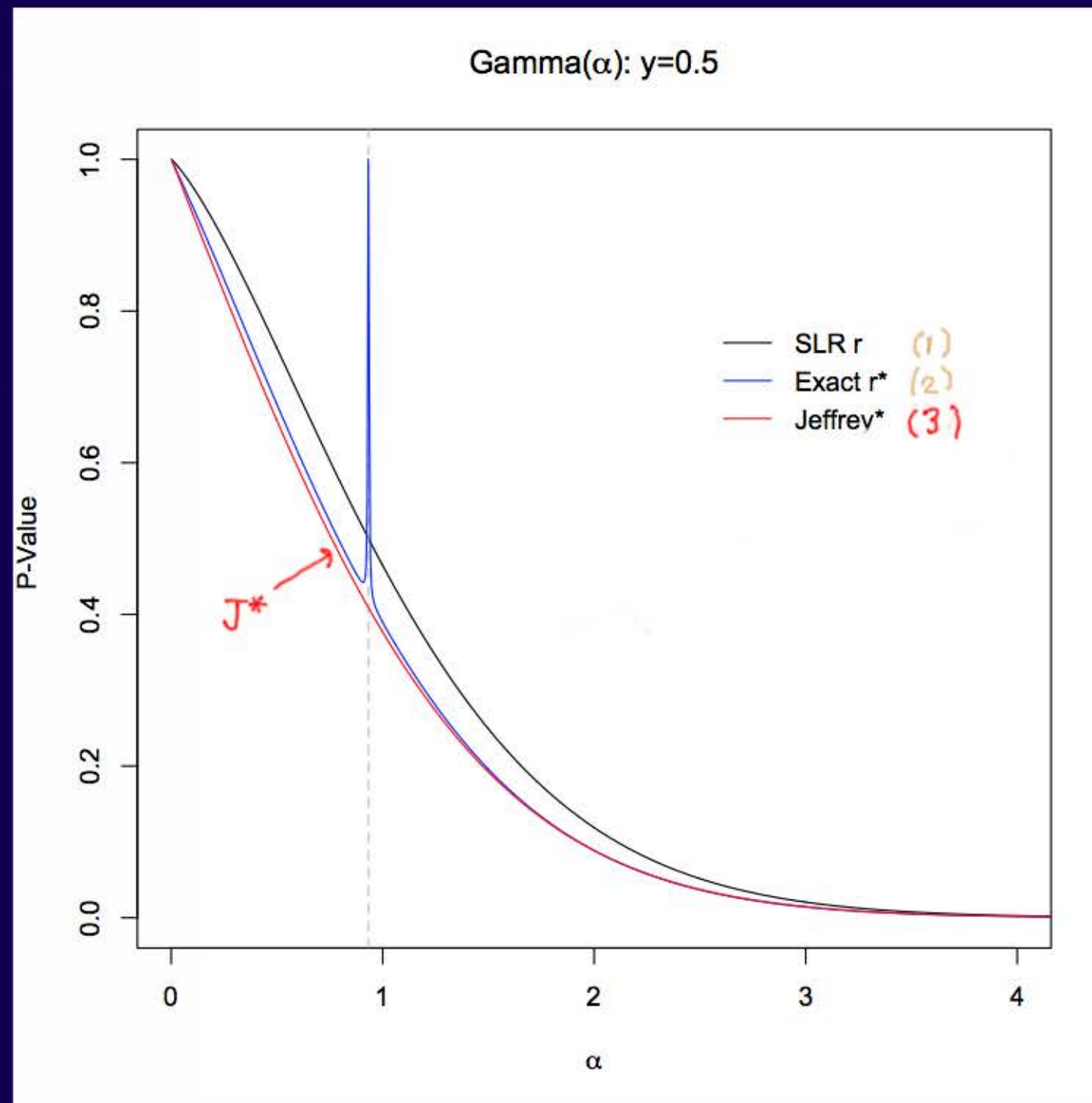
Parameter of interest is α



(1)
 What likelihood
 says: $\Phi(r)$
 $r = SLR$ Φ df N_{01}



- (1)
What likelihood
says: $\Phi(r)$
 $r = \text{SLR} \quad \Phi \text{ df } N_{01}$
- (2)
Third order r^*
Exact



(1)

What likelihood

says: $\Phi(r)$

$r = \text{SLR}$ Φ df $N=1$

(2)

Third order r^*

Exact

(3)

Jeffreys J^* "p=1"

Closely hugs exact

J^* is "reproducible"

6 Bayes: Vector parameter

Interest $\psi = \psi(\varphi)$ Scalar

Nuisance $\lambda = \psi(\varphi)$

SP separates:

$$f(\lambda; \psi) d\lambda \propto \frac{|J(\lambda)(\hat{\psi}_\psi)|^{-1/2}}{(2\pi)^{p_{\lambda'}}} \exp\{\ell - \tilde{\ell}\} dt$$

(continuity, ^{3rd} interest, other)

6 Bayes: Vector parameter

Interest $\psi = \psi(\varphi)$ Scalar

Nuisance $\lambda = \psi(\varphi)$

SP separates:

$$f(\lambda; \psi) d\lambda \cdot \frac{|J(\lambda, \lambda)(\hat{\varphi}_\psi)|^{-1/2}}{(2\pi)^{p_{\lambda'}}} \exp\{\ell - \tilde{\ell}\} dt$$

3rd
(continuity, interest, other)

Bayes: Get rid of nuisance

(1) Norming constant

$$\downarrow$$
$$\frac{|J(\lambda, \lambda)(\hat{\varphi}_\psi)|^{-1/2}}{(2\pi)^{p_{\lambda'}/2}}$$

6 Bayes: Vector parameter

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$$f(\lambda; \psi) d\lambda \propto \frac{|J(\lambda)(\hat{\varphi}_\psi)|^{-1/2}}{(2\pi)^{p_{\lambda'}}} \cdot \exp\{\ell - \tilde{\ell}\} dt$$

3rd
(continuity, interest, other)

Bayes: Get rid of nuisance

(1) Normalizing constant (2) Put on profile

$$\frac{|J(\lambda)(\hat{\varphi}_\psi)|^{-1/2}}{(2\pi)^{p_{\lambda'}}} \cdot 1$$

6 Bayes: Vector parameter

Interest $\psi = \psi(\varphi)$ Scalar

Nuisance $\lambda = \psi(\varphi)$

SP separates:

$$f(s; \psi) ds \cdot \frac{|J(\lambda)(\hat{\varphi}_\psi)|^{-1/2}}{(2\pi)^{p_{\frac{\lambda}{2}}}} \cdot \exp\{\ell - \tilde{\ell}\} dt$$

Bayes: Get rid of nuisance

(1) Normalizing constant

(2) Put on profile

$$\frac{|J(\lambda)(\hat{\varphi}_\psi)|^{-1/2}}{(2\pi)^{p_{\frac{\lambda}{2}}}} \cdot 1$$

Welch-Peers for ψ

$$|j_{(\psi)}^p(\psi)|^{1/2}$$

3rd
(continuity, interest, other)

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Welch-Peers for ψ

$$|j_{\psi\psi}^p(\psi)|^{1/2}$$

Combine (full prior)

- On profile

- One dimensional

$$|j_{\psi\psi}(\hat{\varphi}_\psi)|^{1/2} \cdot \frac{|j(\lambda, \lambda)(\hat{\varphi}_\psi)|^{1/2}}{|j(\lambda, \lambda)(\hat{\varphi}_\psi)|^{1/2}}$$

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Nuisance $\lambda = \psi(\varphi)$

SP separates:

$$f(\lambda; \psi) d\lambda \cdot \frac{|J(\lambda)(\hat{\varphi}_\psi)|^{-1/2}}{(2\pi)^{p_{\lambda}}}} \cdot \exp\{\ell - \tilde{\ell}\} dt$$

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Regular
Jeffreys

Non linearity
adjustment

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Nuisance $\lambda = \psi(\varphi)$

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3rd
(continuity, interest, other)

Bayes: Get rid of nuisance

(1) Normalizing constant (2) Put on profile

$$\frac{|J(\lambda; \hat{\varphi}_\psi)|^{1/2}}{(2\pi)^{p_{\lambda}}/2} \cdot 1$$

Welch-Peers for ψ

$$|j_{\psi\psi}^p(\psi)|^{1/2}$$

Combine (full prior)

- On profile
- One dimensional

$$|j_{\varphi\varphi}(\hat{\varphi}_\psi)|^{1/2} \cdot \frac{|j(\lambda; \hat{\varphi}_\psi)|^{1/2}}{|j(\lambda; \hat{\varphi}_\psi)|^{1/2}}$$

Regular
Jeffreys

Non linearity
adjustment

Root nuisance info re ψ

Root nuisance info re linear χ

6 Bayes: Vector parameter

Interest $\psi = \psi(\varphi)$ Scalar

Nuisance $\lambda = \psi(\varphi)$

SP separates:

$$f(s; \psi) ds \propto \frac{|J(\lambda)(\hat{\varphi}_\psi)|^{-1/2}}{(2\pi)^{p_{\lambda}}/2} \cdot \exp\{\ell - \chi\} dt$$

3rd
(continuity, interest, other)

Bayes: Get rid of nuisance

- (1) Normalizing constant (2) Put on profile

$$\frac{|J(\lambda)(\hat{\varphi}_\psi)|^{1/2}}{(2\pi)^{p_{\lambda}}/2} \cdot 1$$

Welch-Peers for ψ

$$|j_{\psi\psi}^p(\psi)|^{1/2}$$

Combine (full prior)

- On profile
- One dimensional

$$|j_{\psi\psi}(\hat{\varphi}_\psi)|^{1/2} \frac{|j(\lambda)(\hat{\varphi}_\psi)|^{1/2}}{|j(\lambda)(\hat{\varphi}_\psi)|^{1/2}}$$

Root nuisance info re ψ

Root nuisance info re linear χ

Regular
Jeffreys

Non linearity
adjustment

Second order

Jeffreys works but use only on one-dim profile

i.e.
a delta
function

----- on

6-2

Vector parameter (α, β) ; Scalar^{linear} interest α

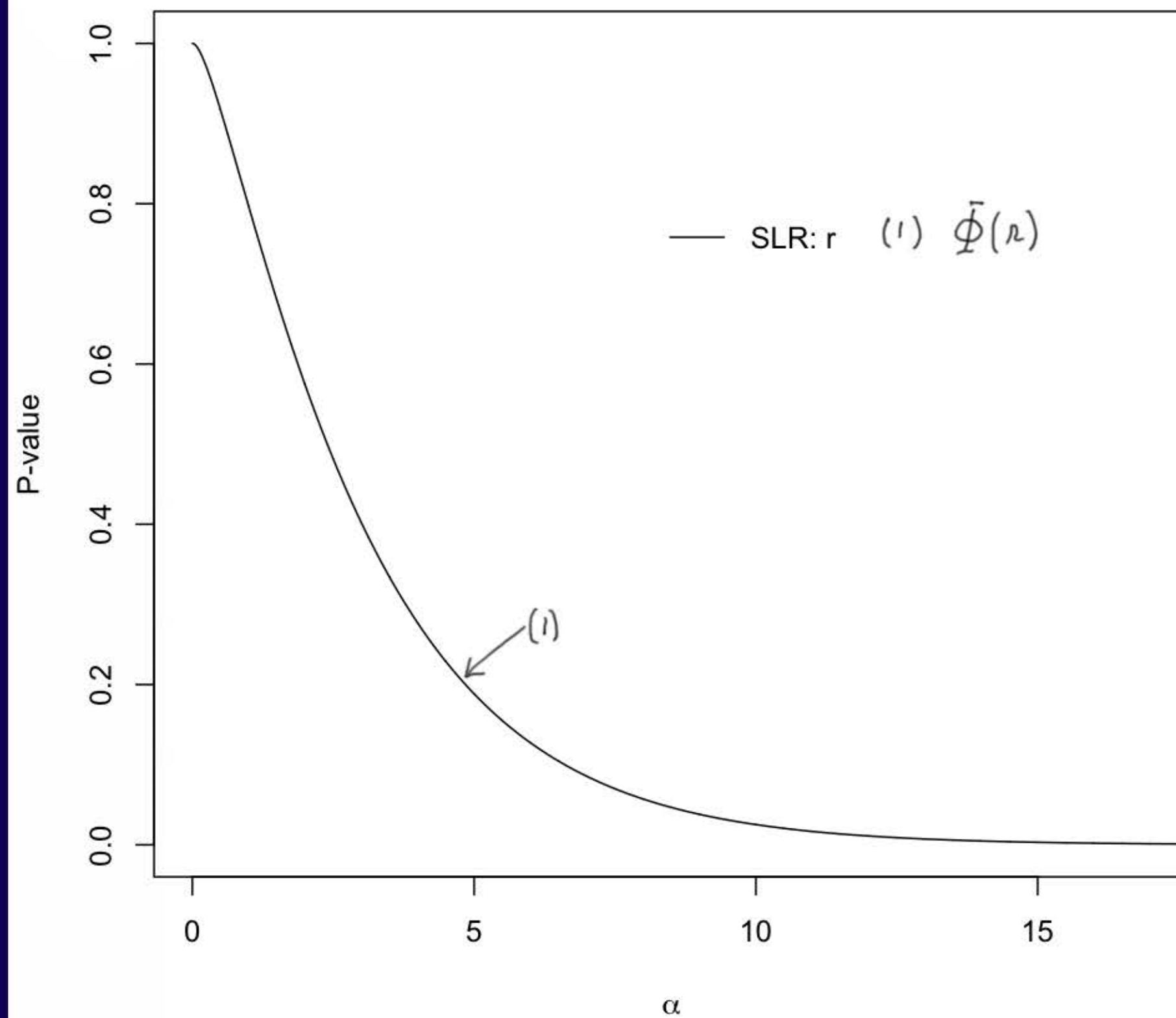
Example 2

Setup

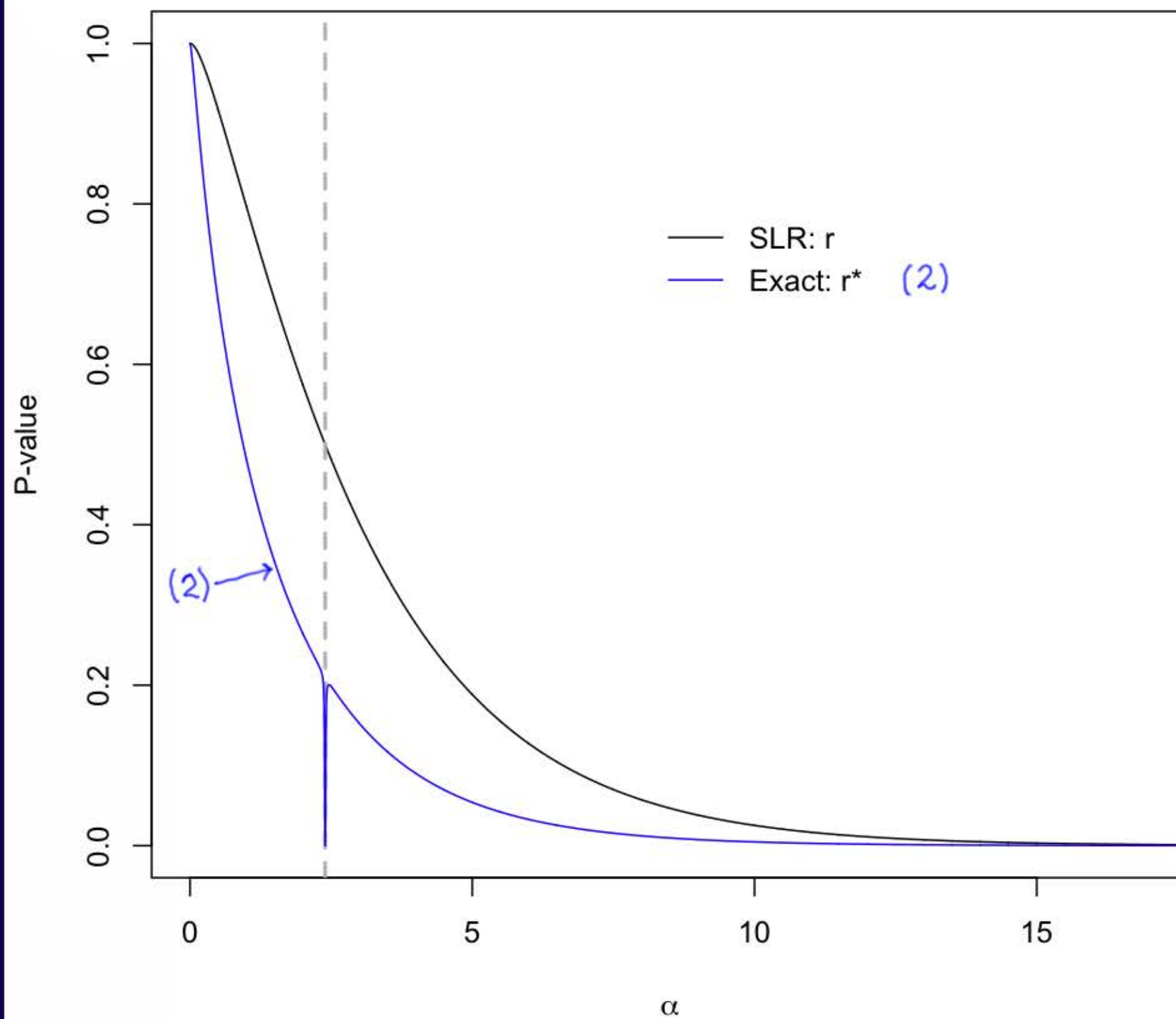
- Model: $Y \sim \text{Gamma}(\alpha, \beta)$ where α is shape and β is rate
pdf is $\frac{\beta^\alpha}{\Gamma(\alpha)} y^{\alpha-1} e^{-\beta y}$
Let $n = 2$ and data is $(y_1, y_2) = (1, 4)$

Parameter of interest is α , β : free nuisance

Gamma(α, β): Interest α with $y=c(1,4)$

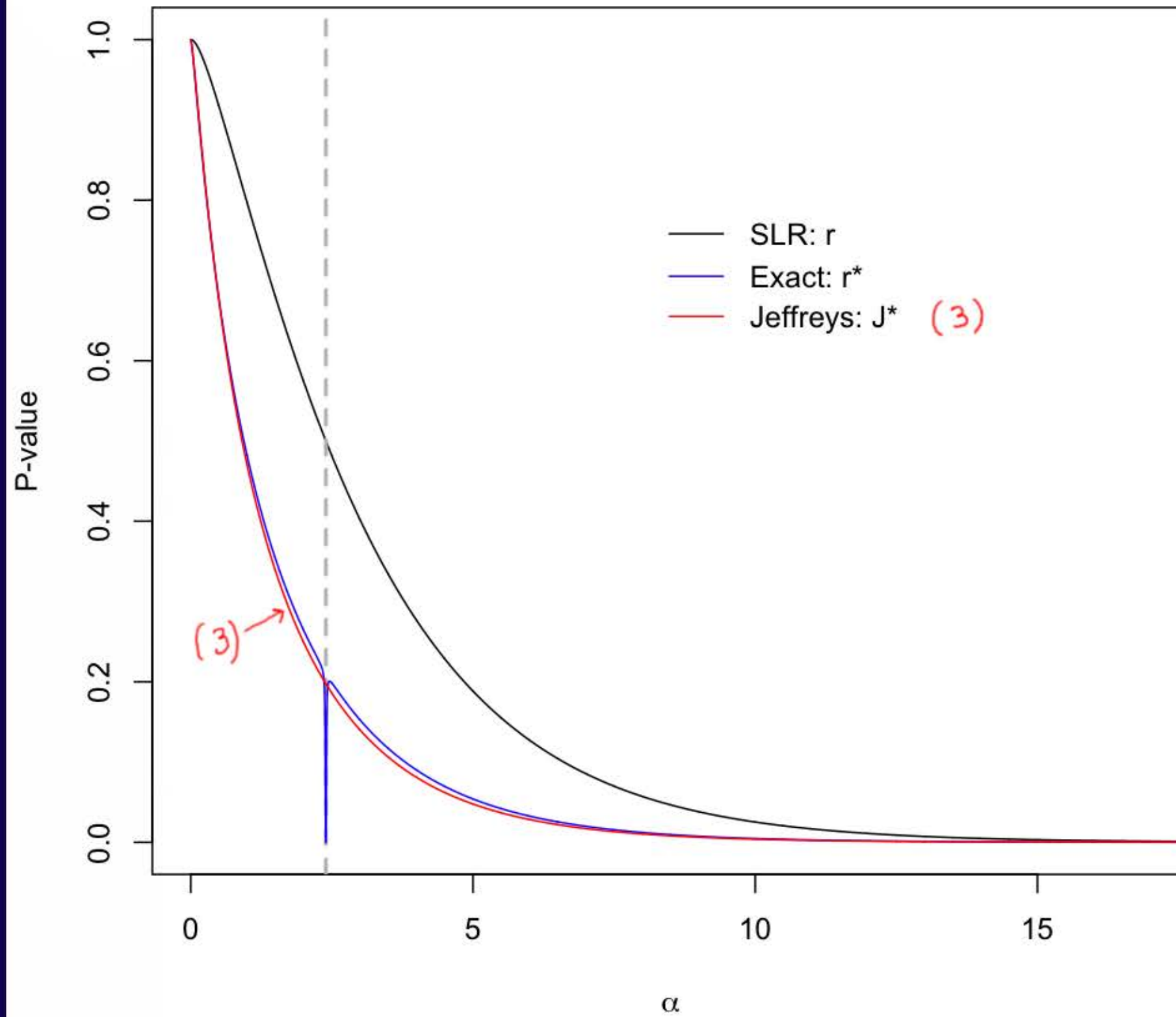


Gamma(α, β): Interest α with $y=c(1,4)$



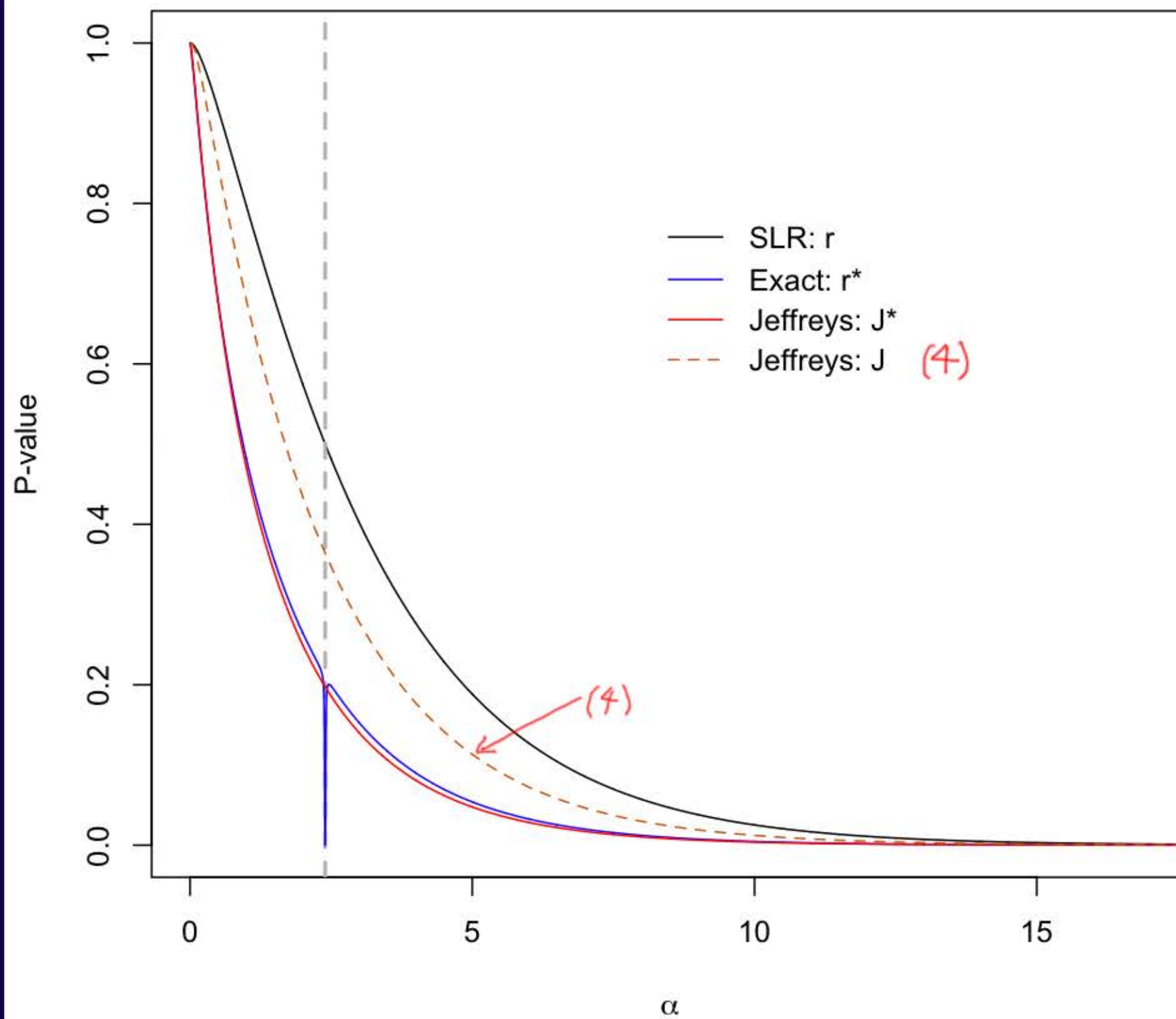
(2) $\Phi(r^*)$

Gamma(α, β): Interest α with $y=c(1,4)$



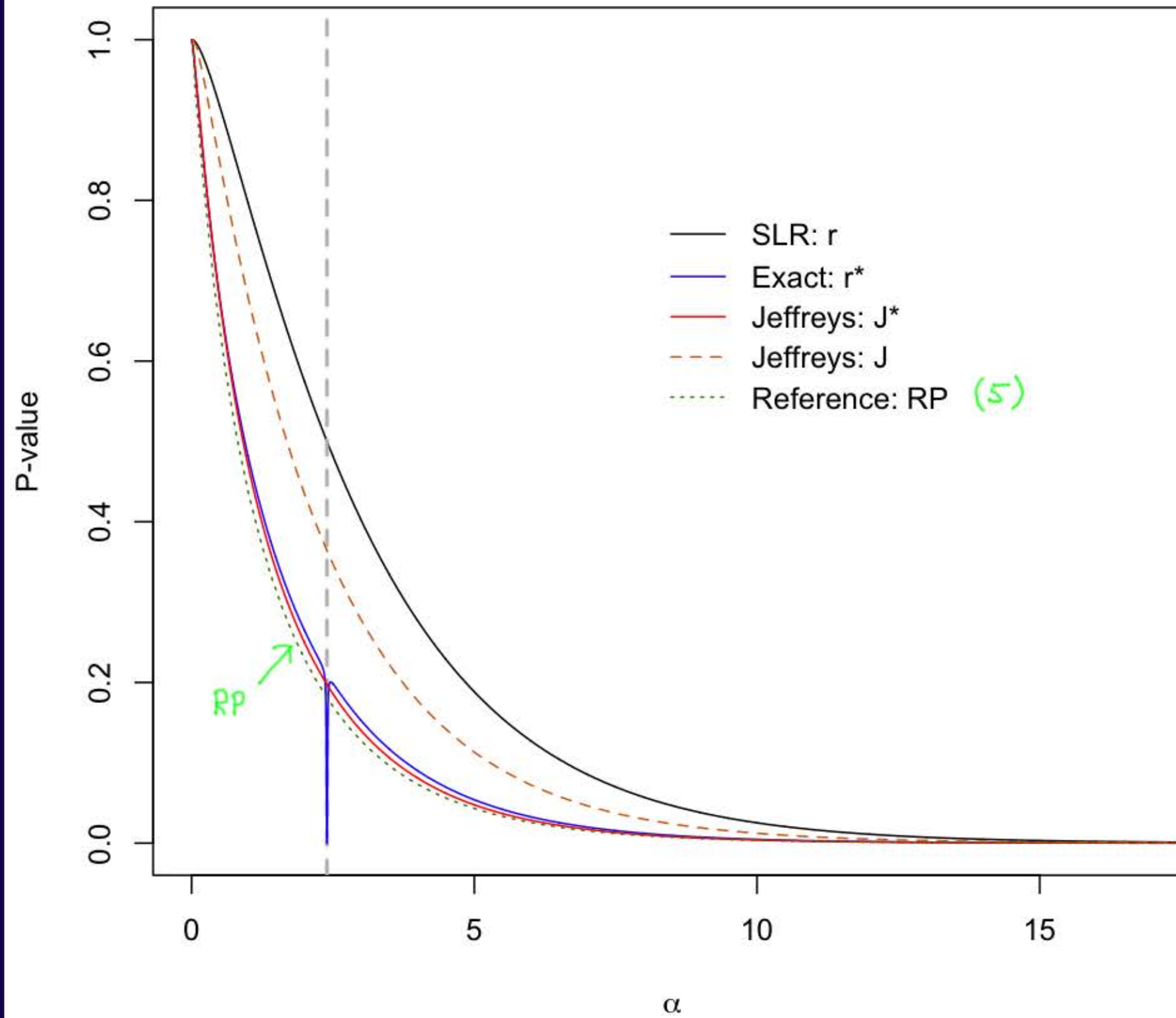
(3) Profile Jeffreys J^*

Gamma(α, β): Interest α with $y=c(1,4)$



(4) Full Jeffreys

Gamma(α, β): Interest α with $y=c(1,4)$



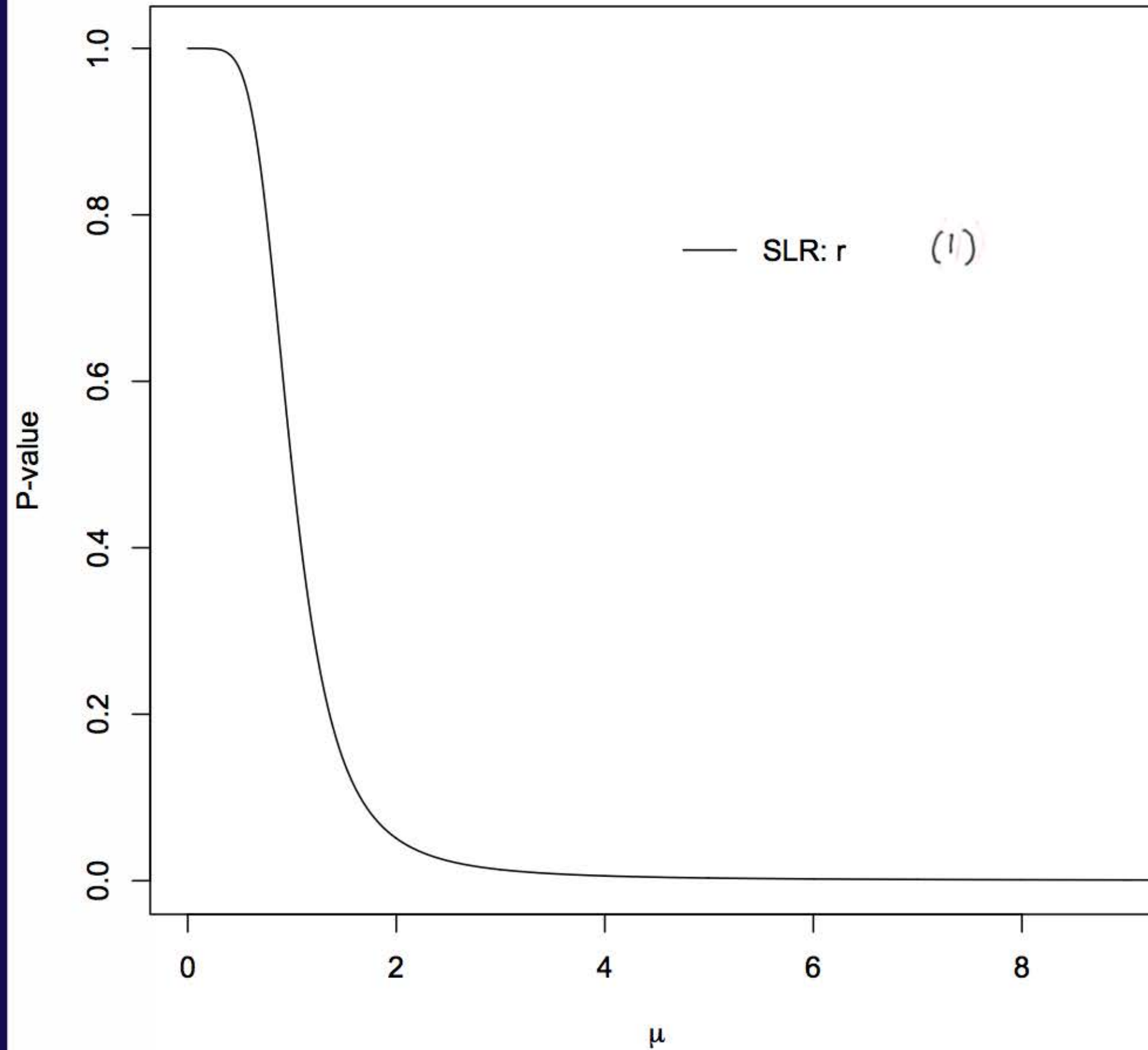
(5) Targeted Reference

6-3

Gamma(α, β) : Interest μ ... rotating

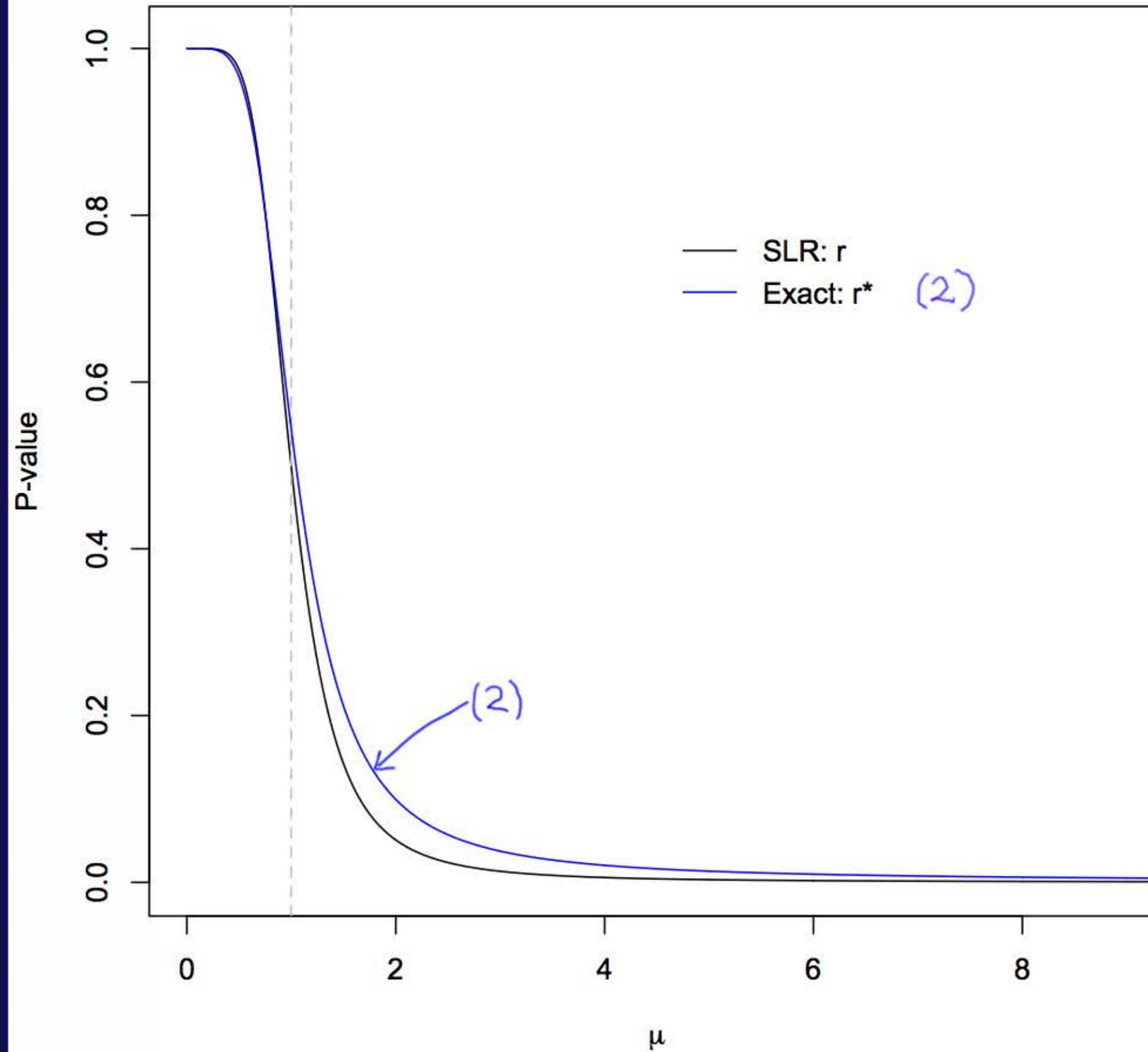
Gamma(α, β): $\mu = \alpha / \beta$ with $y = c(0.2, 0.45, 0.78, 1.28, 2.28)$

Gamma(α, β): $\mu = \alpha / \beta$ with $y = c(0.2, 0.45, 0.78, 1.28, 2.28)$



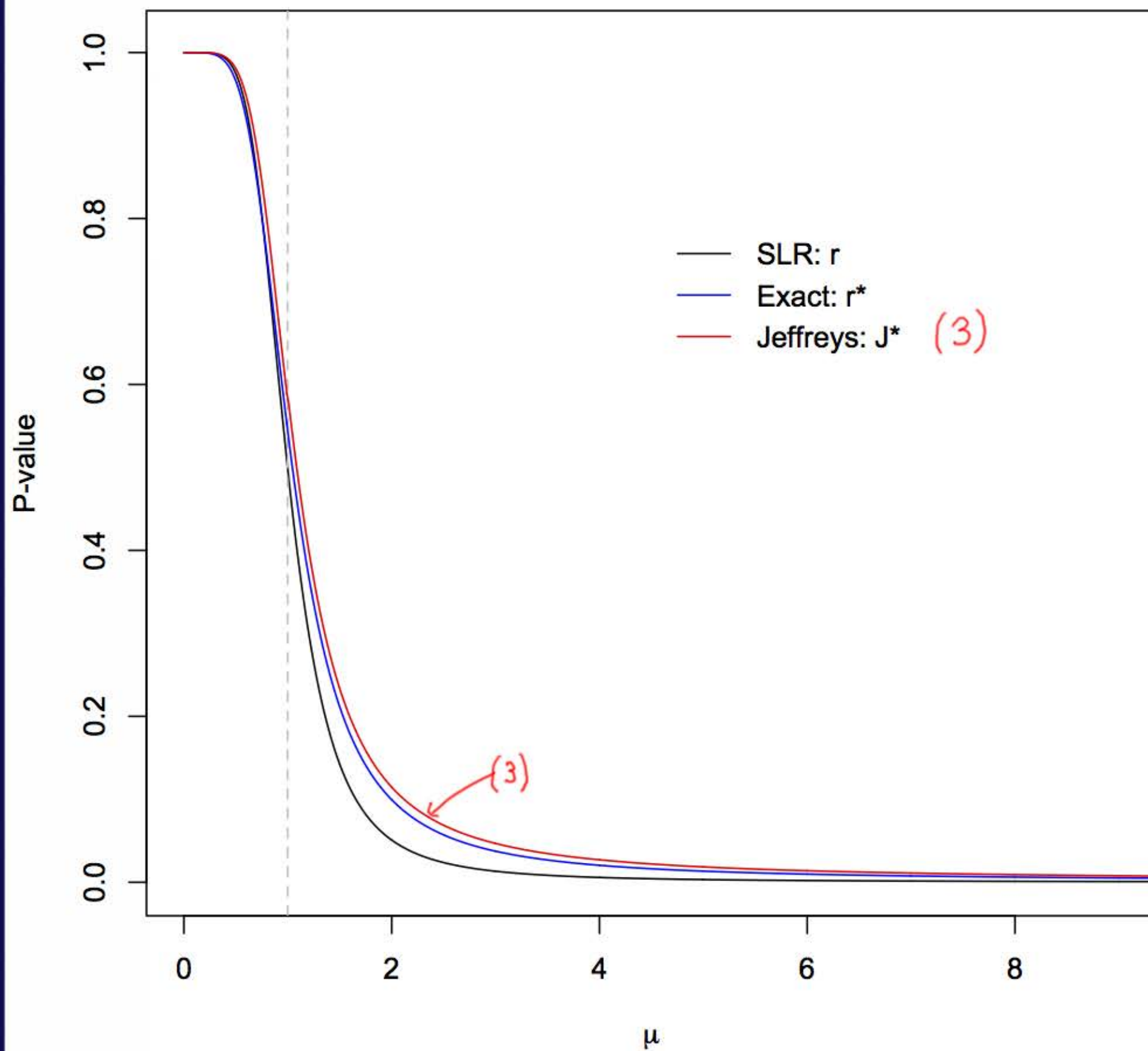
$$(1) \rho(\mu) = \Phi(r)$$

Gamma(α, β): $\mu = \alpha / \beta$ with $y = c(0.2, 0.45, 0.78, 1.28, 2.28)$



$$p(\mu) = \tilde{\Phi}(r^*)$$

Gamma(α, β): $\mu = \alpha / \beta$ with $y = c(0.2, 0.45, 0.78, 1.28, 2.28)$



$$p(\mu) = \text{Jeff}^*$$

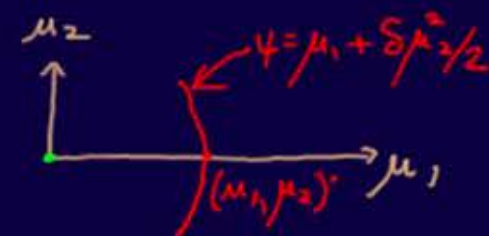
6-4

$$N(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}; I)$$

$$\eta^0 = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$$

$$\psi = \mu_1 + \delta \mu_2^2 / 2$$

Curved Interest (As for ψ in itself!)



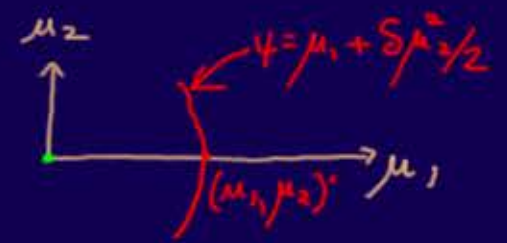
6-4

$$N\left(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}; I\right)$$

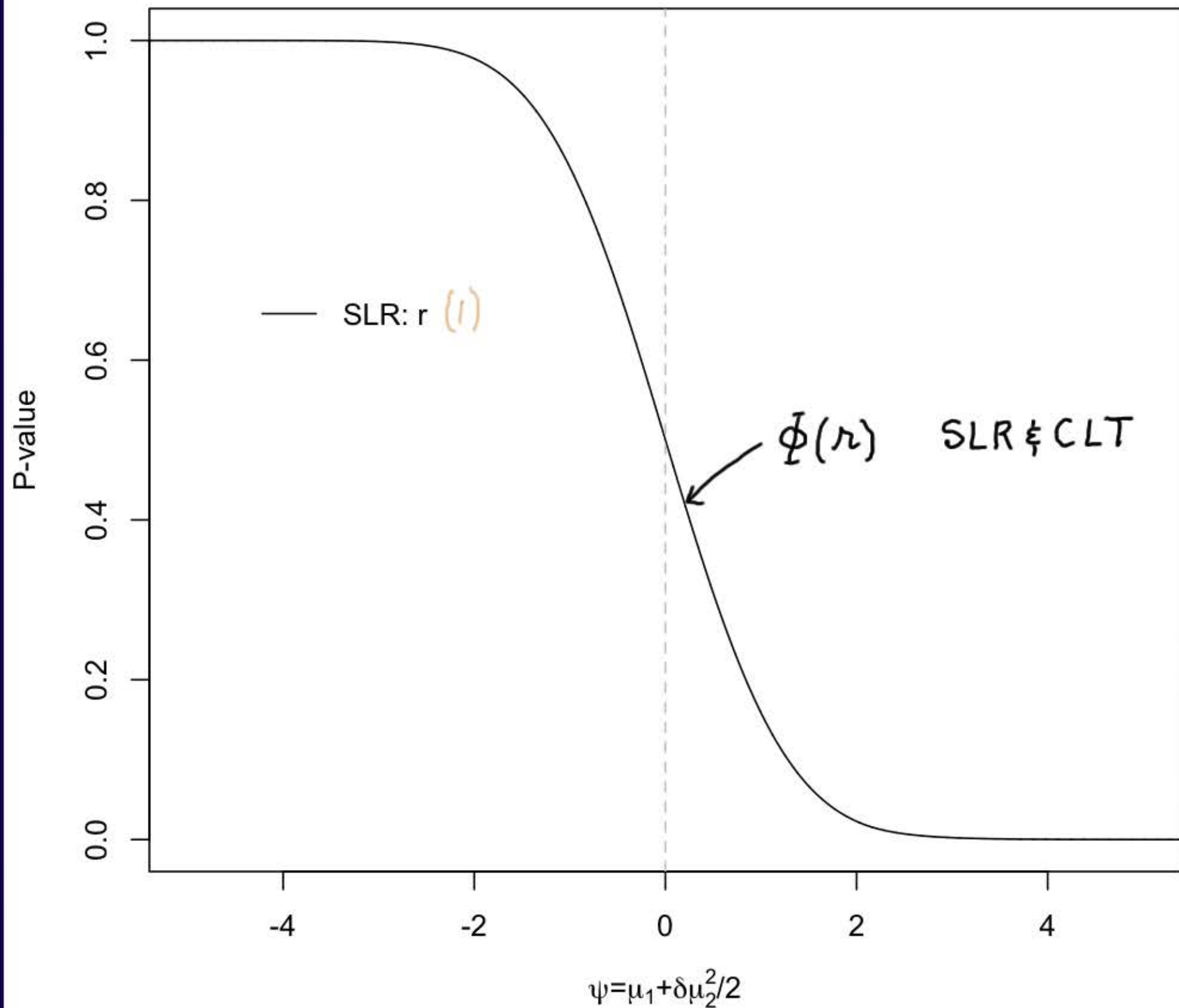
$$\psi^0 = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$$

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Curved Interest (A story in itself!)



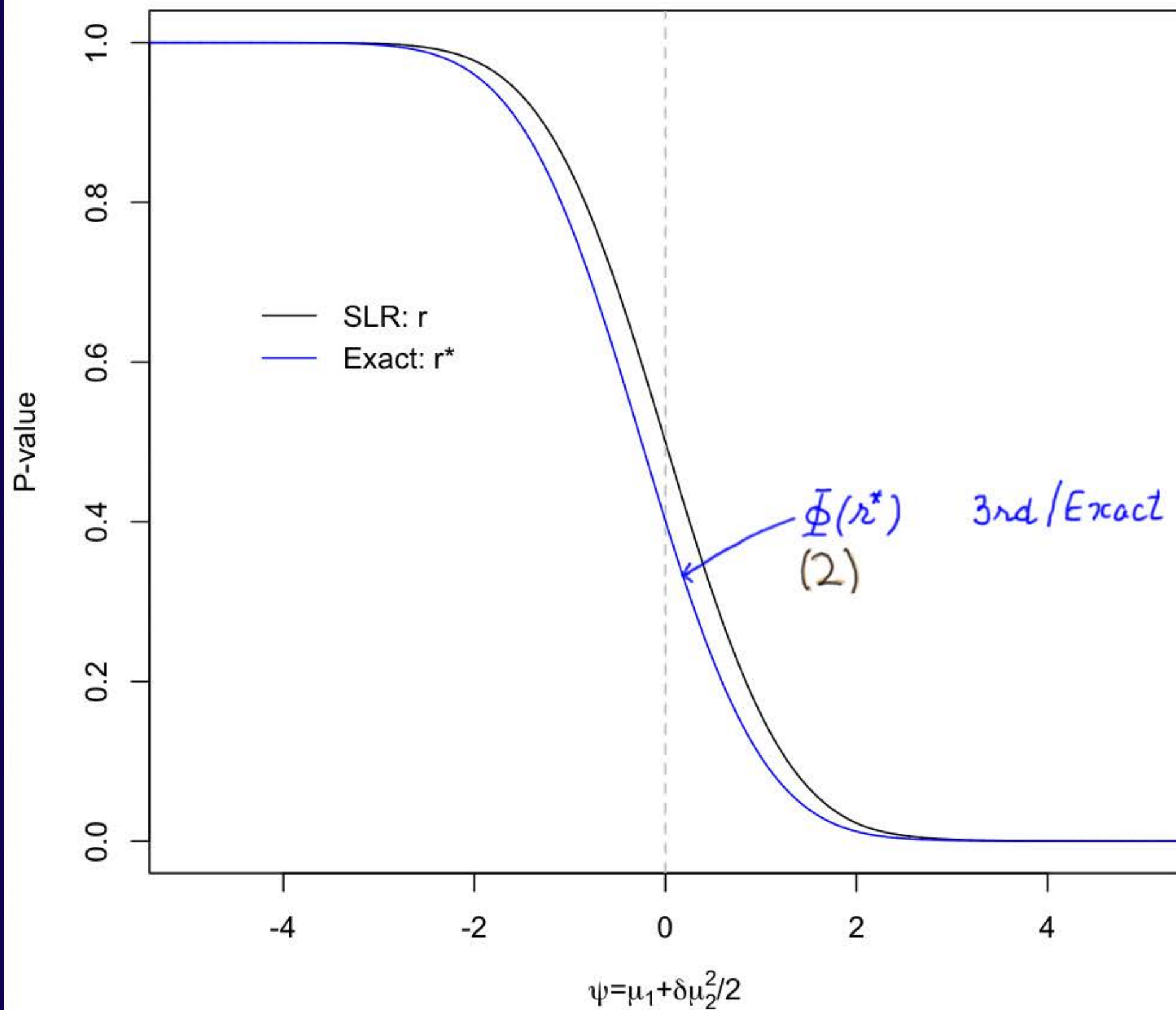
$$y^0 = (0, 0), \delta = 0.5$$

(1) SLR
 $\Phi(r)$

$$N\left(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}; I\right) \quad \psi = \mu_1 + \delta \mu_2^2 / 2$$

$$y^0 = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$$

$$y^0 = (0, 0), \delta = 0.5$$



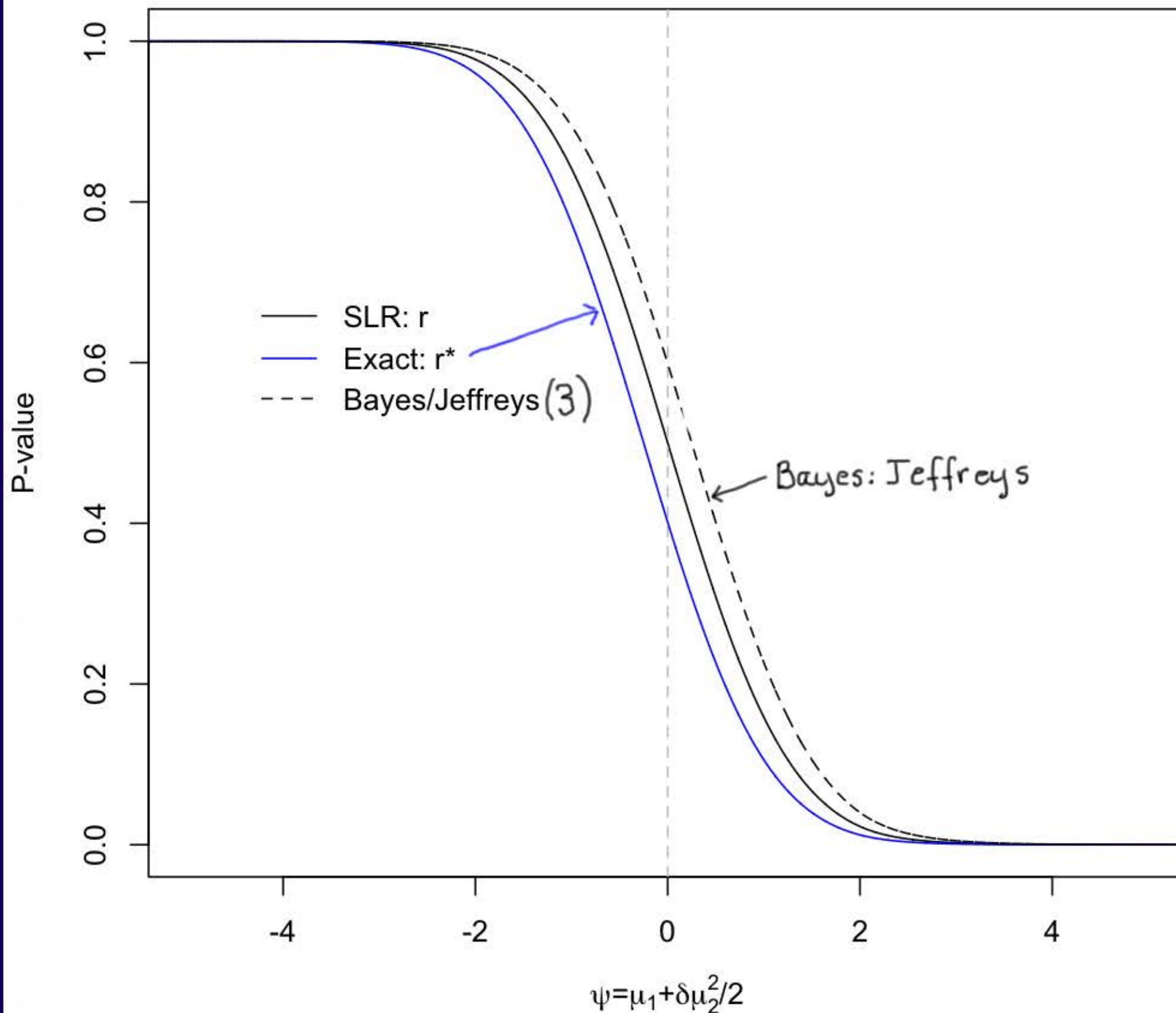
(1) SLR
 $\Phi(r)$

(2) Exact
 $\Phi(r^*)$

$$N\left(\begin{matrix} \mu_1 \\ \mu_2 \end{matrix}; I\right) \quad \psi = \mu_1 + \delta \mu_2^2/2$$

$$y^0 = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$$

$$y^0 = (0,0), \delta = 0.5$$



(1) SLR
 $\Phi(r)$

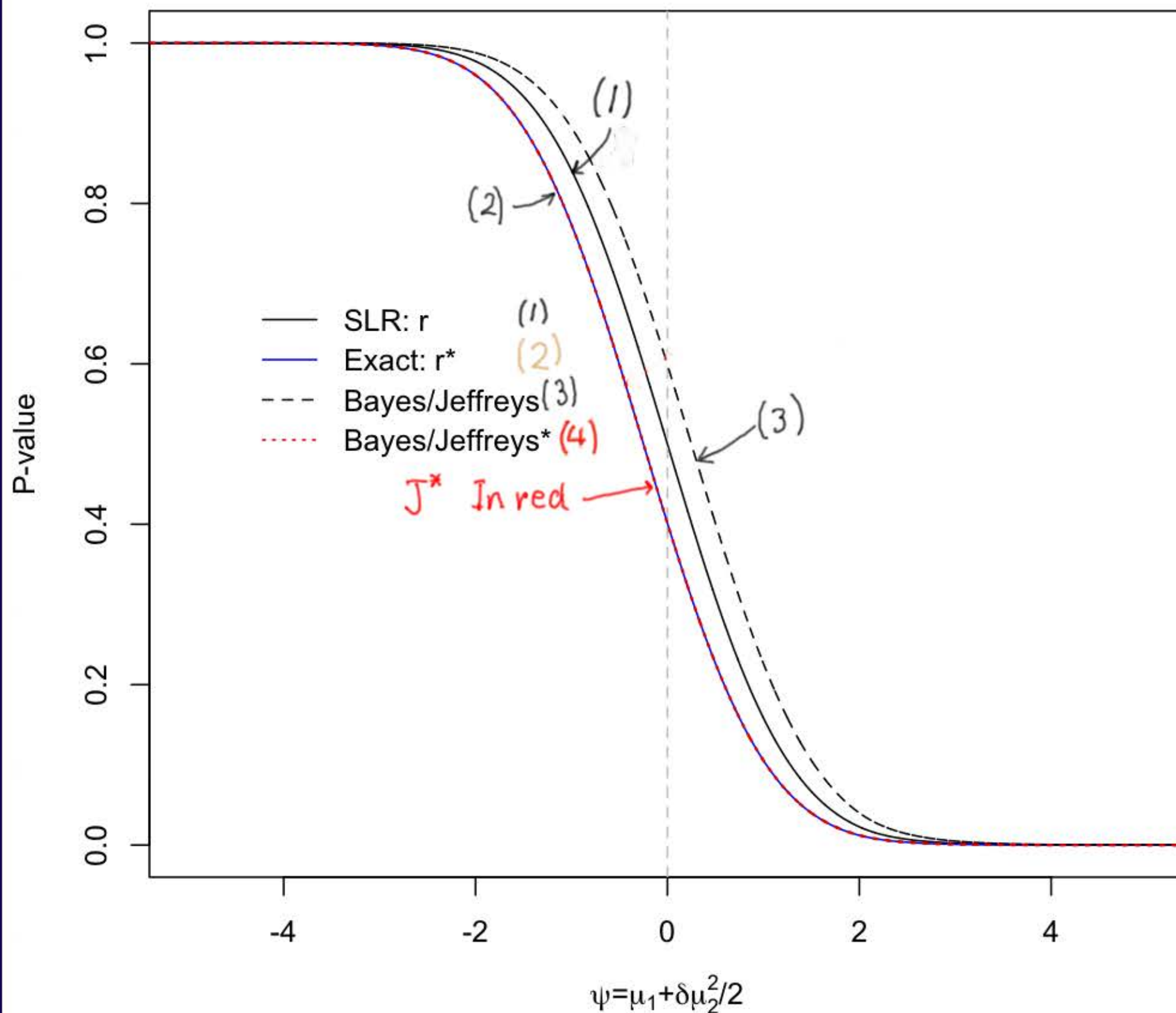
(2) Exact
 $\bar{\Phi}(r^*)$

(3) Bayes
Plain Jeffreys
(goes farther away
from the Exact)
Worse than just L

$$N\left(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}; I\right) \quad \psi = \mu_1 + \delta \mu_2^2/2$$

$$y^0 = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$$

$$y^0 = (0,0), \delta = 0.5$$



(1) SLR
 $\Phi(r)$

(2) Exact
 $\bar{\Phi}(r^*)$

(3) Bayes
Plain Jeffreys
(goes farther away
from the Exact)

(4) J^*
Jeffreys but just
on profile curve
for interest α

J^* almost duplicates
the exact

7 Concluding

1. Bayes is just first order $O(n^{-1/2})$ Safer to use mle, score, SLR

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2. For scalar full parameter? Jeffreys is 2nd order Welch-Peers 1963
3. For scalar interest with nuisance, put Jeffreys(1946) on profile (2nd order)
 - for linear parameter
 - for rotating parameter (Need a Jacobian)
 - for curved parameter (Need curvature adjustment)

| and
combine

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 - Crucial but - include all information!
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 - Crucial but - include all information!
 - respect continuity!
5. Bayes using W-P for interest parameter
 - gives approximate pivotal, thus legitimacy
6. Fisher got it right!

How Bayes can deliver 2nd order Accuracy!

A long history
great collaborators:

Nancy

A Wong York

M Bédard U de Montréal

W Lin Toronto

A M Fraser UBC

M J Fraser Toronto

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