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## Bayesian Copula-based Latent Variable Models

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# Copulas: The Joys

- Copulas are mathematical devices used to model dependence between random variables regardless of their marginals.
- Copulas are useful for data fusion/integration because they lead to coherent joint models, even when the marginals are in different families (e.g., Gaussian, Poisson, Student, etc) or of different types (e.g, discrete, continuous).
- Copulas unlock information contained in the dependence part of the distribution (second-order) that complements the information in the marginals.
- Simply put, copulas allow us to extend statistical methods beyond the use of a multivariate Gaussian or Student.

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#### At the root of it all, a theorem

- If Y<sub>1</sub>, Y<sub>2</sub>,..., Y<sub>K</sub> are continuous r.v.'s with cdfs F<sub>1</sub>, F<sub>2</sub>,..., F<sub>k</sub>, there is an unique copula C : [0, 1]<sup>K</sup> → [0, 1] that links the joint cdf with the marginal ones (Sklar's Theorem).
- ▶ The copula (when K = 2)  $C : [0,1] \times [0,1] \rightarrow [0,1]$  satisfies

$$F_{12}(t,s) = \Pr(Y_1 \le t, Y_2 \le s) = C(F_1(t), F_2(s)).$$

The conditional copula satisfies

$$F_{12|X}(t,s) = \Pr(Y_1 \le t, Y_2 \le s|X) = C(F_{1|X}(t), F_{2|X}(s)|X)$$

► Usually we use parametric families so  $C(u, v) = C_{\theta}(u, v)$  such as Clayton's family:  $C_{\theta}(u, v) = \left[\max\left(u^{-\theta} + v^{-\theta} - 1, 0\right)\right]^{-1/\theta}$ . Frank's family:  $K_{\theta}(u, v) = -\frac{1}{\theta} \ln \left[1 + \frac{(e^{-\theta u} - 1)(e^{-\theta v} - 1)}{e^{-\theta} - 1}\right]$ .

ln a conditional copula,  $\theta$  may depend on X.

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# Latent Variables (LV)

The variable of interest W is sometimes impossible to measure directly

- State of the economy
- Traffic in a city
- State of your health
- State of a complex disease

#### Instead, one measures

- Y = (Y<sub>1</sub>,...,Y<sub>k</sub>)<sup>T</sup> whose components are surrogates of W and each provide partial information about W
- Covariate  $\mathbf{X} \in \mathbb{R}^{p}$

#### ▶ We are often interested in the explanatory power of **X** for *W*.

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### An example

- Cardiotocography (CTG) is a medical procedure that monitors the fetal heart rate.
- The LV is the fetus' underlying state of health during birth, W.
- Our surrogate response is the bivariate vector (Q, Y) where
  - Q is the number of peaks (acceleration followed by a deceleration of heart beats) for the signal recorded by the CTG
  - Y is the log of mean short-term "beat-to-beat" variability (MSTV) where the short-term variability (STV) is obtained by measuring the time between successive R waves (cardiac systoles) of the fetus' electrocardiogram.
- The covariates are FM (fetal movement) and UC (uterine contraction), two continuous variables monitored during birth.

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#### Conditional independence LV model

• A canonical LV model, given  $W_i = X_i\beta + \epsilon$ , is

 $\begin{aligned} Y_i &\perp Q_i | W_i \\ Y_i &\sim \mathcal{N}(\mu_c + \lambda_c W_i, \sigma^2) \\ Q_i &\sim \textit{Poisson}(\exp{(\mu_d + \lambda_d W_i)}) \end{aligned}$ 

- This implies that the two marginal regressions share a common random effect so they are marginally dependent (and conditionally independent)
- ► The induced dependence is not analytically available.

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#### Conditional independence is a Copula LV

• The copula alternative is, conditional on  $W_i$ ,

$$\begin{aligned} H(Y_i, Q_i | W_i) &= C_{\theta_i}(F_Y(Y_i | W_i), F_Q(Q_i | W_i)), \quad \theta_i = \kappa^{-1}(\xi_0 + \xi_1 W_i) \\ Y_i &\sim \mathcal{N}(\mu_c + \lambda_c W_i, \sigma^2); \quad Q_i \sim \textit{Poisson}(\exp(\mu_d + \lambda_d W_i)) \end{aligned}$$

- ► The whole joint distribution of (*Y*, *Q*) is varying with *W* not just the marginals.
- ► The copula captures the residual dependence on *W* after the marginal effects have been accounted for.
- The previous model is obtained when the copula is the independence copula.

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## Why the Conditional Copula?

- $\blacktriangleright Y_i | x \sim N(f_i(x), \sigma_i) x \in \mathbb{R}^2$
- True marginal means:
  - $f_1(x) = 0.6 \sin(5x_1) 0.9 \sin(2x_2)$
  - $f_2(x) = 0.6 \sin(3x_1 + 5x_2)$
  - $\sigma_1 = \sigma_2 = 0.2, X_1 \perp X_2.$
- ▶ Copula: θ(x) = 0.71

Suppose x<sub>2</sub> is not observed so inference is based only on x<sub>1</sub>

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### CTG: The LV Copula Model

•  $(Q_i, Y_i)|W_i$  has joint density

$$f_{(Q,Y)}(q,y) = f_c(y) \cdot \left[ C_{d|c} \left( F_d(q), F_c(y) \right) - C_{d|c} \left( F_d(q-), F_c(y) \right) \right],$$

where

$$C_{d|c}(u_d, u_c) = \frac{\partial}{\partial u_c} C(u_d, u_c).$$

Data Augmentation: Introduce latent variable Z such that

$$Q\stackrel{d}{=} F_d^-(F_Z(Z)),$$

- The copula between (Y, Z) is the same as the copula between (Y, Q)
- We can choose the distribution of *Z* to help the computation.
- For instance if we use a Gaussian copula, it helps to have  $Z \sim N(0, 1)$
- Smith and Khaled (JASA, 2012), C. and Sabeti (JMVA, 2012).

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#### The Augmented LV Copula Model for the CTG Example

The augmented model for CTG data is

$$Z_i \sim \mathcal{N}(0, 1)$$

$$Q_i \mid W_i \sim \text{Poisson} \left(e^{\mu_d + \lambda_d W_i}\right)$$

$$Y_i \mid W_i \sim \mathcal{N}(\mu_c + \lambda_c W_i, \sigma^2)$$

$$(Z_i, Y_i) \mid W_i \sim C^{\text{Gauss}} \left(\Phi(\cdot), \Phi\left(\frac{\cdot - \mu_c - \lambda_c W_i}{\sigma}\right) \mid \theta(W_i, \boldsymbol{\xi})\right)$$

$$W_i \sim \mathcal{N}(x_i \beta, 1),$$

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## CTG: The Augmented LV Copula Model

► The dependence between *Y*, *Z* and *Q* is defined by their joint conditional distribution

$$f_{(Q,Z,Y)|W}(q,z,y \mid w) = h(z,y \mid w,\mu_c,\lambda_c,\psi_c,\boldsymbol{\xi})$$
  
 
$$\cdot \mathbbm{1}_{F_Z^{-1}(F_d(q-|\varphi_d(\mu_d,\lambda_d,w))) \leq z < F_Z^{-1}(F_d(q|\varphi_d(\mu_d,\lambda_d,w)))}$$

▶ Let  $\boldsymbol{\xi} = (\xi_0, \xi_1) \in \mathbb{R}^2$  and  $A(w) = \xi_0 + \xi_1 \cdot w$ . Then we set

$$\theta(w, \xi) = rac{e^{A(w)} - e^{-A(w)}}{e^{A(w)} - e^{-A(w)}}$$

as the correlation parameter of the bivariate Gaussian conditional copula of (Y, Z)|W = w.

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## Some MCMC details

- If the copula and marginals are Gaussian the joint is a multivariate normal so some of the conditional densities are available in closed form.
- ► For other copula families we rely on MwG moves.
- We sample {Z<sub>i</sub> : 1 ≤ i ≤ n} from its conditional distribution and use the samples only to update the copula parameters *ξ*.

#### Good initialization helps:

$$\boldsymbol{\flat} \ \boldsymbol{\beta}^{(0)} = (\mathbf{X}^{\top} \mathbf{X})^{-1} \mathbf{X}^{\top} \mathbf{y}$$

$$\boldsymbol{\mathsf{W}}_{i}^{(0)} = (\boldsymbol{\beta}^{(0)})^{\top} \mathbf{x}_{i}$$

•  $(\mu_d^{(0)}, \lambda_d^{(0)})$  is the MLE based on the marginal likelihood, etc

► Adaptive strategy for all MwG: target an acceptance rate of 44%.

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## Model Selection: WAIC

The WAIC is defined as

$$WAIC(\mathcal{M}) = -2fit(\mathcal{M}) + 2p(\mathcal{M}), \qquad (1)$$

where the model fitness is

$$fit(\mathcal{M}) = \sum_{i=1}^{n} \log \left( \mathbb{E} \left[ \Pr(y_i, q_i | \omega, \mathcal{M}) \right] \right)$$
(2)

and the penalty

$$p(\mathcal{M}) = \sum_{i=1}^{n} Var(\log(Pr(y_i, q_i | \omega, \mathcal{M}))), \qquad (3)$$

where  $\omega$  contains all the parameters and latent variables in the model.

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### Spotlight on dependence: A conditional WAIC

► We use the following two conditional WAICs:

$$\begin{split} \mathsf{CWAIC}_{Y|Q}(\mathcal{M}) &= -2\sum_{i=1}^{n} \log \left( \mathbb{E} \left[ \mathsf{Pr}(y_i | q_i, \omega, \mathcal{M}) \right] \right) + \\ &+ 2\sum_{i=1}^{n} \mathsf{Var} \left( \log \left( \mathsf{Pr}(y_i | q_i, \omega, \mathcal{M}) \right) \right), \\ \mathsf{CWAIC}_{Q|Y}(\mathcal{M}) &= -2\sum_{i=1}^{n} \log \left( \mathbb{E} \left[ \mathsf{Pr}(q_i | y_i, \omega, \mathcal{M}) \right] \right) + \\ &+ 2\sum_{i=1}^{n} \mathsf{Var} \left( \log \left( \mathsf{Pr}(q_i | y_i, \omega, \mathcal{M}) \right) \right), \end{split}$$

► <sup>1</sup>/<sub>2</sub>(CWAIC<sub>1|2</sub> + CWAIC<sub>2|1</sub>) is asymptotically equivalent to CCV for the marginal likelihood

$$\mathsf{CCV}(\mathcal{M}) = \frac{1}{2} \left\{ \sum_{i=1}^{n} \log \left( \mathsf{Pr}(y_i | q_i, \mathcal{D}_{-i}, \mathcal{M}) \right) + \sum_{i=1}^{n} \log \left( \mathsf{Pr}(q_i | y_i, \mathcal{D}_{-i}, \mathcal{M}) \right) \right\}$$

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#### Simulation Experiment

#### Generate data using a Gaussian copula



Gaussian copula

Figure: Bivariate scatterplot of the generated data with Gaussian copula, and Poisson and normal marginals

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## Simulation Experiment

#### • $CWAIC_{Y|Q}$ and $CWAIC_{Q|Y}$ selection criteria

Criteria\Copula	Gaussian	Frank	Gumbel	Clayton	Indep
CWAIC <sub>Y Q</sub>	1627.36	1642.36	2395.17	1637.17	1606.31
CWAIC	950.71	982.42	1673.57	976.05	997.43
Average	1289.04	1312.39	2034.37	1306.61	1301.87

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### Simulation Experiment



Figure: Traceplots for  $\eta$ 's components.

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# Simulation Experiment



	$\beta_1$	$\beta_2$	$\lambda_d$	$\lambda_c$	$\xi_1$
Mean	1.18	0.48	0.90	0.84	3.10
True	1	0.5	1	1	3

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## CTG: The data



What	are	Copulas	
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## CTG: Estimates

- ► WAIC, WAIC<sub>Y|Q</sub> and WAIC<sub>Q|Y</sub> all point to the Gaussian copula (over Gumbel, Frank, Clayton, Independence).
- The posterior means



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## CTG: What does it mean?

- ► A peak in the histogram (counted with Nmax) would typically be produced by an FHR acceleration followed by a deceleration.
- Certain decelerations can be attributed to compression of the baby's head during uterine contractions, so they're not unusual.
- Late decelerations (starting after a uterine contraction begins) and especially variable decelerations often suggest a compromise in the supply of blood and oxygen to the fetus.
- A reduced STV can signify a quiet or sleep phase of the fetus, but also the effects of analgesic drugs given to the mother, fetal hypoxia, prematurity, neurological damage and tachycardia from any cause.
- Interpretation: Extremes values of W are identified with "unhealthy" regimes while small values of |W| correspond to healthy ones.
- It is physiologically plausible that MSTV should be negatively correlated with Nmax.

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## The Past & Future

- Copulas offer a way to bypass the paucity of available joint distributions.
- Copulas allow the integration of multiple (dependent) sources of information/data via joint modeling
- Joint models can be used for prediction/imputation of an expensive variable given values for cheaper ones.
- So far have been used to further empower multivariate regressions, time series, HMMs, LV models, etc
- Computational challenges, especially in higher dimensions
- Papers available here: https://www.utstat.toronto.edu/craiu/Papers/index.html