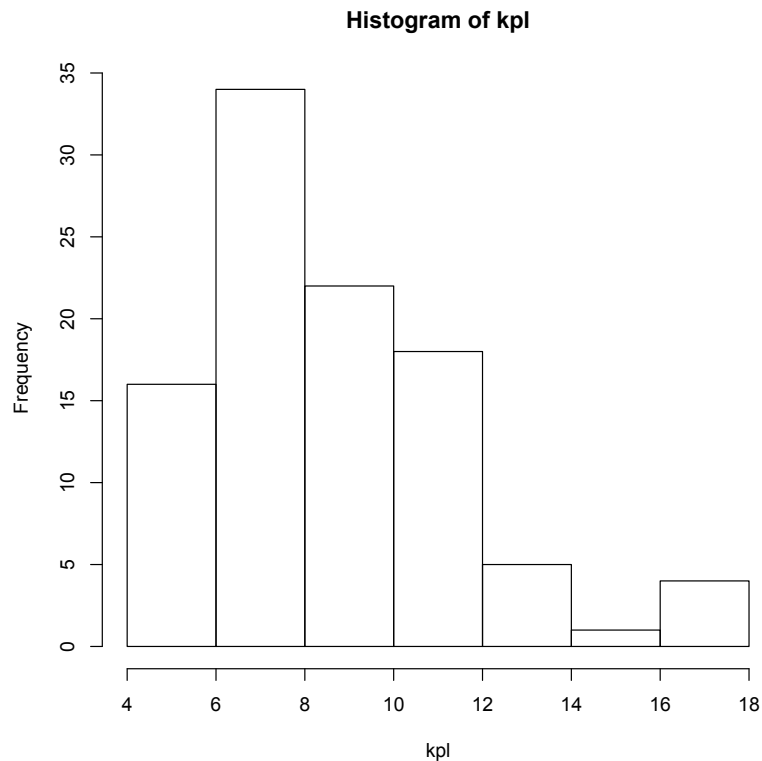


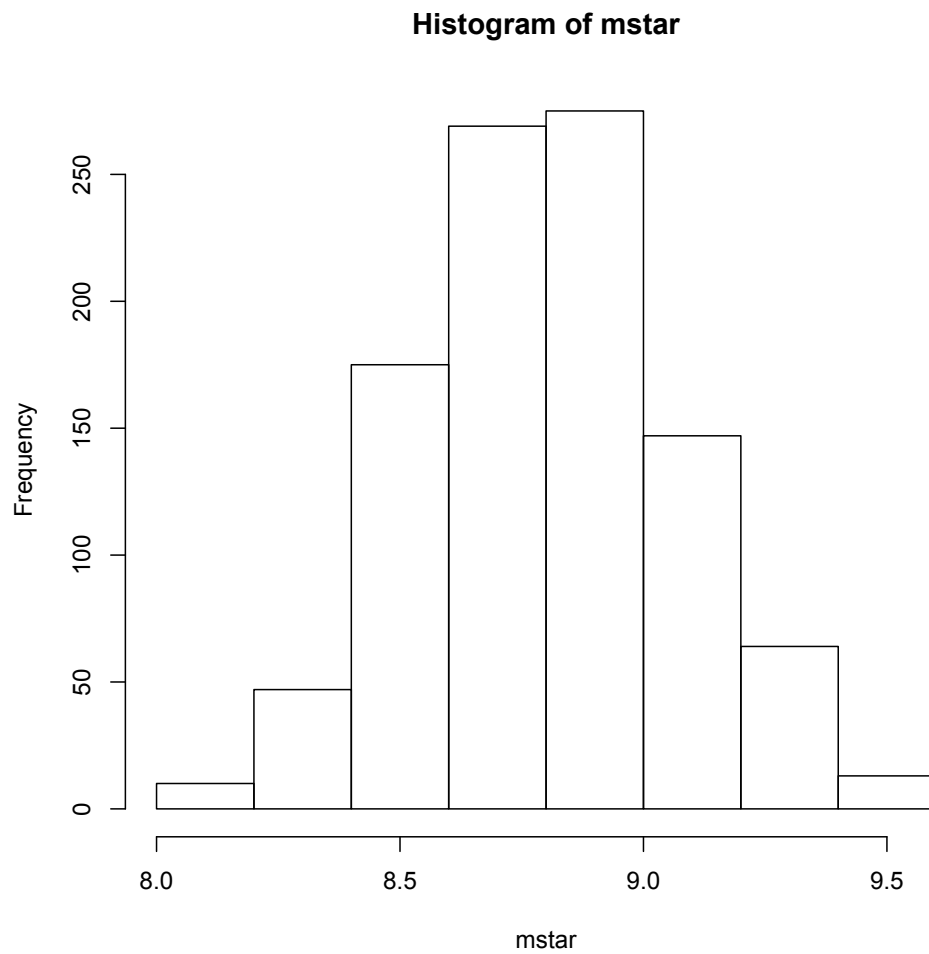
# Bootstrap Regression with R

```
> # Bootstrap regression example
>
> kars =
read.table("http://www.utstat.toronto.edu/~brunner/appliedf12/data/mcars3.data",
+ header=T)
> kars[1:4,]
  Cntry  kpl weight length
1    US  5.04 2.1780 5.9182
2 Japan 10.08 1.0260 4.3180
3    US  9.24 1.1880 4.2672
4    US  7.98 1.4445 5.1054
> attach(kars) # Variables are now available by name
>
> # Before regression, a garden variety univariate bootstrap
> hist(kpl) # Right skewed
```



```
> # Small example for demonstration of R syntax
> set.seed(3244)
> x = kpl[1:10]; x
[1] 5.04 10.08 9.24 7.98 7.98 7.98 9.66 7.56 5.88 10.92
> n = length(x)
> # Sample of size n from the numbers 1:n, with replacement.
> choices = sample(1:n,size=n,replace=T); choices
[1] 2 7 5 1 4 6 9 8 4 10
> x[choices]
[1] 10.08 9.66 7.98 5.04 7.98 7.98 5.88 7.56 7.98 10.92
```

```
> # Now bootstrap the mean of kpl
> n = length(kpl); B = 1000
> mstar = NULL # mstar will contain bootstrap mean values
>
> for(draw in 1:B) mstar = c(mstar,mean(kpl[sample(1:n,size=n,replace=T)]))
> hist(mstar)
```

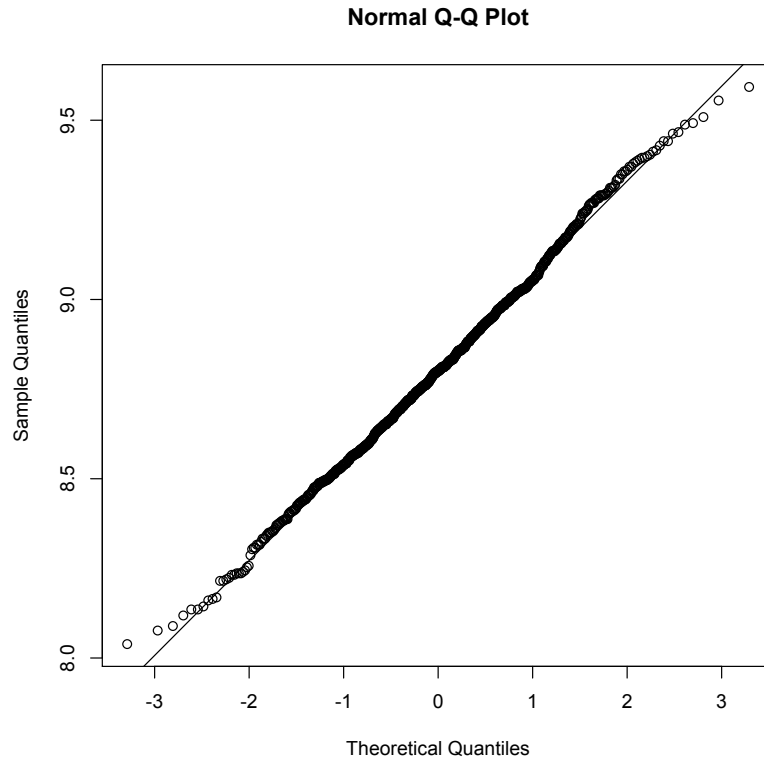


```

> # Look at a normal qq plot. That's a plot of the order statistics against
> # the corresponding quantiles of the (standard) normal. Should be roughly linear
> # if the data are from a normal distribution.

> qqnorm(mstar); qqline(mstar)

```



```

> # Quantile bootstrap CI for mu. Use ONLY if the bootstrap distribution is symmetric.
> sort(mstar)[25]; sort(mstar)[975]
[1] 8.3034
[1] 9.3492
> # Compare the usual CI
> t.test(kpl)

```

#### One Sample t-test

```

data: kpl
t = 32.9363, df = 99, p-value < 2.2e-16
alternative hypothesis: true mean is not equal to 0
95 percent confidence interval:
 8.264966 9.324634
sample estimates:
mean of x
 8.7948

```

```

> # Now regression
> # Compute some polynomial terms
> wsq = weight^2; lsq = length^2; wl = weight*length
> # Bind it into a nice data frame
> datta = cbind(kpl,weight,length,wsq,lsq,wl)
> datta = as.data.frame(datta)
>
> modell = lm(kpl ~ weight + length + wsq + lsq + wl, data=datta)
> summary(modell)

Call:
lm(formula = kpl ~ weight + length + wsq + lsq + wl, data = datta)

Residuals:
    Min       1Q   Median       3Q      Max
-4.0861 -0.8702  0.0490  0.6898  4.4006

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)   79.124     29.121   2.717  0.00784 **
weight        24.336     26.570   0.916  0.36204
length       -33.764     19.350  -1.745  0.08427 .
wsq           11.377      8.531   1.334  0.18556
lsq            5.140      3.410   1.507  0.13508
wl           -12.442     10.174  -1.223  0.22442
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.577 on 94 degrees of freedom
Multiple R-squared:  0.6689,    Adjusted R-squared:  0.6513
F-statistic: 37.98 on 5 and 94 DF,  p-value: < 2.2e-16

> betahat = coef(modell); betahat
(Intercept)      weight      length      wsq      lsq      wl
 79.124214    24.336110   -33.763782    11.376646    5.139649   -12.442449
>
> set.seed(3244)
> bstar = NULL # Rows of bstar will be bootstrap vectors of regression coefficients.
> n = length(kpl); B = 1000
> for(draw in 1:B)
+   {
+     # Randomly sample from the rows of kars, with replacement
+     Dstar = datta[sample(1:n,size=n,replace=T),]
+     model = lm(kpl ~ weight + length + wsq + lsq + wl, data=Dstar)
+     bstar = rbind( bstar,coef(model) )
+   } # Next draw
>
> bstar[1:5,]
(Intercept)      weight      length      wsq      lsq      wl
[1,]    64.73852   15.322187  -25.549389   14.378022   4.437351  -12.779138
[2,]   270.35328  158.868074 -149.690584   26.298031  21.834487  -47.728186
[3,]    30.97446  -0.156246  -7.492504   10.815623   1.853871  -7.490135
[4,]   102.07061  48.638339  -50.062648   15.305469   7.844481 -19.669466
[5,]    74.31620  -4.632140  -23.726771  -1.497523   2.162598   1.070870
>

```

```

> Vb = var(bstar) # Approximate asymptotic covariance matrix of betahat
> Vb
      (Intercept)      weight      length      wsq      lsq      wl
(Intercept)  4009.5755  2805.7433 -2432.9272  403.57428  359.95762 -795.78318
weight       2805.7433  2337.0043 -1816.6783  434.68245  288.94772 -724.67663
length      -2432.9272 -1816.6783  1511.4557 -292.33275 -229.91599  534.81663
wsq         403.5743   434.6825  -292.3327  117.92217  52.99175 -158.04873
lsq         359.9576   288.9477  -229.9160  52.99175  36.19566 -89.15841
wl          -795.7832  -724.6766   534.8166 -158.04873 -89.15841  239.35661
>
> # Test individual coefficients. H0: betaj=0
> se = sqrt(diag(Vb)); Z = betahat/se
> rbind(betahat,se,Z)
      (Intercept)      weight      length      wsq      lsq      wl
betahat  79.124214  24.3361095 -33.7637822  11.376646  5.1396488 -12.4424491
se       63.321209  48.3425725  38.8774443  10.859198  6.0162826  15.4711541
Z        1.249569   0.5034095  -0.8684671   1.047651  0.8542898  -0.8042354
>
> # Now test the product terms all at once
>
> WaldTest = function(L,thetahat,Vn,h=0) # H0: L theta = h
+ # Note Vn is the asymptotic covariance matrix, so it's the
+ # Consistent estimator divided by n. For true Wald tests
+ # based on numerical MLEs, just use the inverse of the Hessian.
+ {
+   WaldTest = numeric(3)
+   names(WaldTest) = c("W","df","p-value")
+   r = dim(L)[1]
+   W = t(L%*%thetahat-h) %*% solve(L%*%Vn%*%t(L)) %*%
+     (L%*%thetahat-h)
+   W = as.numeric(W)
+   pval = 1-pchisq(W,r)
+   WaldTest[1] = W; WaldTest[2] = r; WaldTest[3] = pval
+   WaldTest
+ } # End function WaldTest
>
> Lprod = rbind( c(0,0,0,1,0,0),
+               c(0,0,0,0,1,0),
+               c(0,0,0,0,0,1) )
> WaldTest(Lprod,betahat,Vb)
      W      df  p-value
9.463393 3.000000 0.023724
>
>
> # Normal test for comparison
> model0 = lm(kpl ~ weight + length) # No product terms
> anova(model0,model1) # p = 0.0133
Analysis of Variance Table

Model 1: kpl ~ weight + length
Model 2: kpl ~ weight + length + wsq + lsq + wl
  Res.Df  RSS Df Sum of Sq    F Pr(>F)
1     97 261.81
2     94 233.72  3    28.095 3.7666 0.0133 *
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

Final comment: This is not a typical bootstrap regression. It's more common to bootstrap the residuals. But that applies to a conditional model in which the values of the explanatory variables are fixed constants.