

Logistic Regression

For a binary response variable:
1=Yes, 0=No

This slide show is a free open source document.
See the last slide for copyright information.

Binary outcomes are common and important

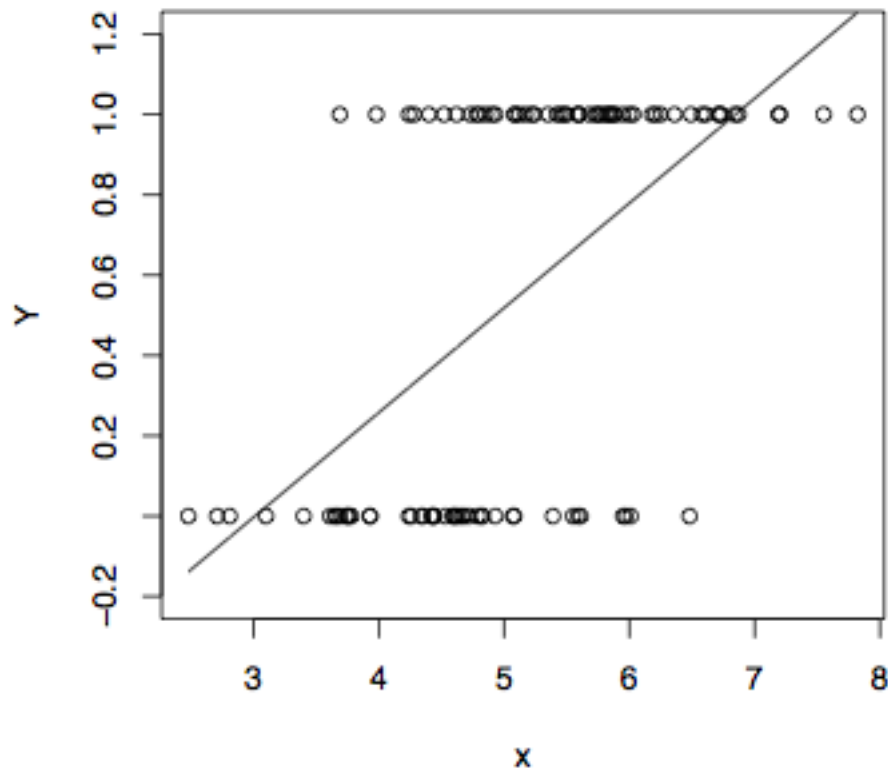
- The patient survives the operation, or does not.
- The accused is convicted, or is not.
- The customer makes a purchase, or does not.
- The marriage lasts at least five years, or does not.
- The student graduates, or does not.

For a binary variable

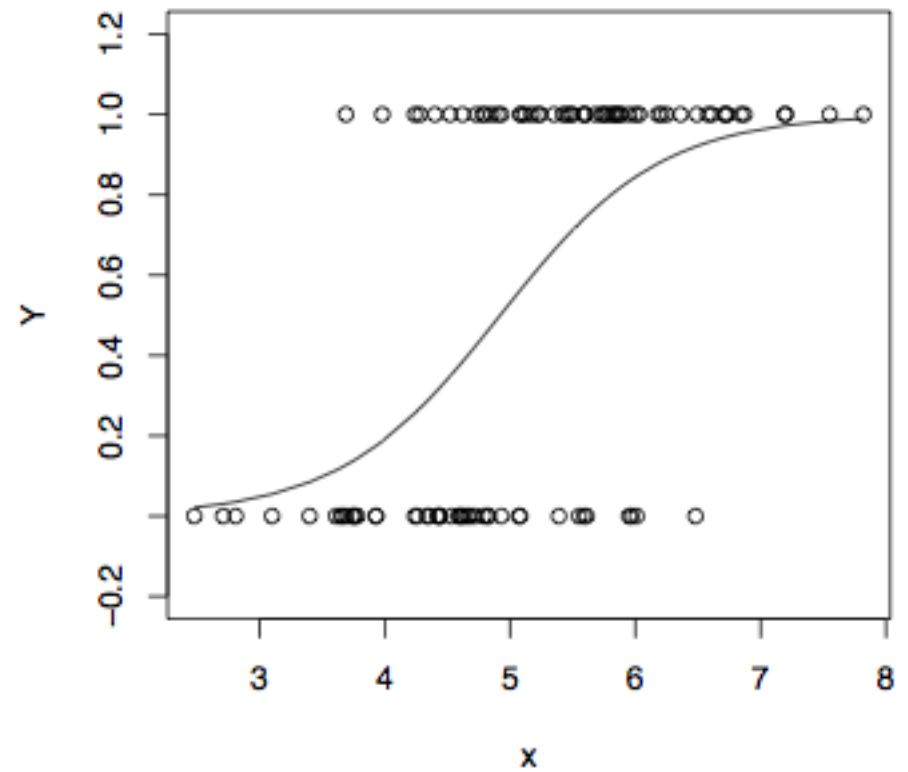
- The population mean $E[Y]$ is the probability that $Y=1$
- Make the mean depend on a set of explanatory variables
- Consider one explanatory variable. Think of a scatterplot

Least Squares vs. Logistic Regression

Least Squares Line



Logistic Regression Curve



The logistic regression curve arises from an indirect representation of the probability of $Y=1$ for a given set of x values.

Representing the probability of an event by π

$$\text{Odds} = \frac{\pi}{1 - \pi}$$

$$\text{Odds} = \frac{\pi}{1 - \pi}$$

- If $P(Y=1)=1/2$, odds = $.5/(1-.5) = 1$ (to 1)
- If $P(Y=1)=2/3$, odds = 2 (to 1)
- If $P(Y=1)=3/5$, odds = $(3/5)/(2/5) = 1.5$ (to 1)
- If $P(Y=1)=1/5$, odds = .25 (to 1)

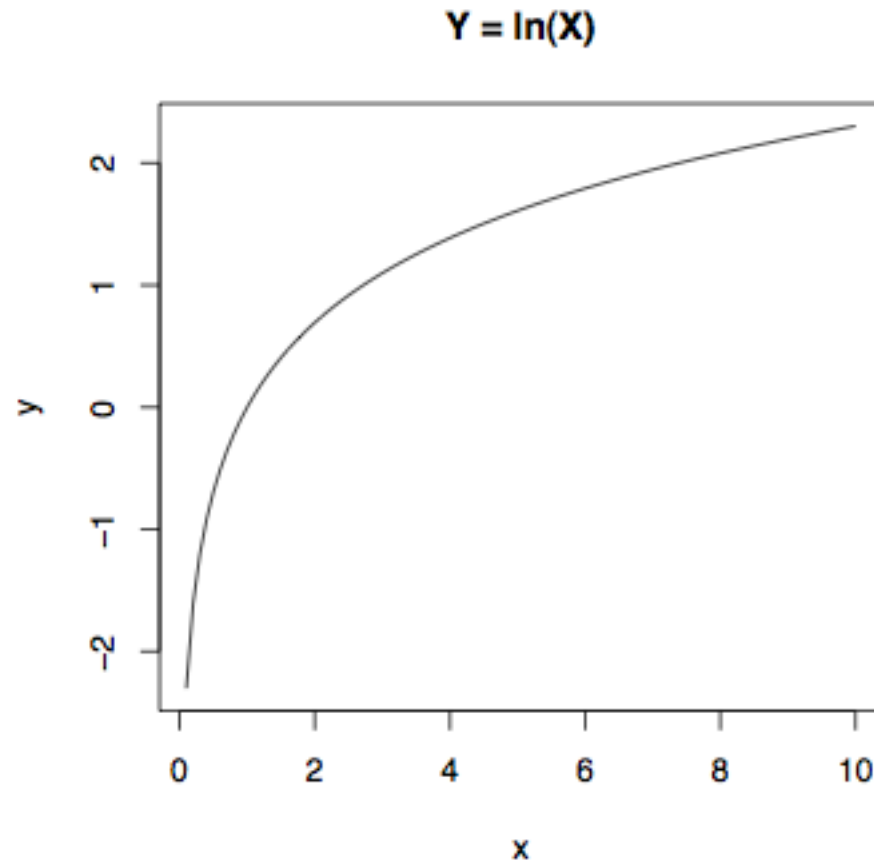
The higher the probability, the greater the odds

$$\text{Odds} = \frac{\pi}{1 - \pi}$$

$$0 \leq \text{Odds} < \infty$$

Linear model for the **log** odds

- Natural log, not base 10
- Symbolized \ln



Some facts about \ln

- The higher the probability, the higher the log odds.
- $\ln(e)=1$, $e = 2.1728\dots$
- Only defined for positive numbers.
- So logistic regression will not work for events of probability exactly zero or exactly one (why not one?)

The log of a product is the
sum of logs

$$\ln(ab) = \ln(a) + \ln(b)$$

$$\ln\left(\frac{a}{b}\right) = \ln(a) - \ln(b)$$

This means the log of an odds *ratio* is the
difference between the two log odds quantities.

Linear regression model for
the log odds of the event $Y=1$

$$\ln \left(\frac{P(Y = 1 | \mathbf{X} = \mathbf{x})}{P(Y = 0 | \mathbf{X} = \mathbf{x})} \right) = \beta_0 + \beta_1 x_1 + \dots + \beta_{p-1} x_{p-1}$$

Equivalent Statements

$$\ln \left(\frac{P(Y = 1 | \mathbf{X} = \mathbf{x})}{P(Y = 0 | \mathbf{X} = \mathbf{x})} \right) = \beta_0 + \beta_1 x_1 + \dots + \beta_{p-1} x_{p-1}$$

$$\begin{aligned} \frac{P(Y = 1 | \mathbf{X} = \mathbf{x})}{P(Y = 0 | \mathbf{X} = \mathbf{x})} &= e^{\beta_0 + \beta_1 x_1 + \dots + \beta_{p-1} x_{p-1}} \\ &= e^{\beta_0} e^{\beta_1 x_1} \dots e^{\beta_{p-1} x_{p-1}} \end{aligned}$$

$$P(Y = 1 | x_1, \dots, x_{p-1}) = \frac{e^{\beta_0 + \beta_1 x_1 + \dots + \beta_{p-1} x_{p-1}}}{1 + e^{\beta_0 + \beta_1 x_1 + \dots + \beta_{p-1} x_{p-1}}}$$

In terms of log odds, logistic regression is like regular regression

$$\ln \left(\frac{P(Y = 1 | \mathbf{X} = \mathbf{x})}{P(Y = 0 | \mathbf{X} = \mathbf{x})} \right) = \beta_0 + \beta_1 x_1 + \dots + \beta_{p-1} x_{p-1}$$

In terms of plain odds,

- Logistic regression coefficients represent *odds ratios*
- For example, “Among 50 year old men, the odds of being dead before age 60 are three times as great for smokers.”

$$\frac{\text{Odds of death given smoker}}{\text{Odds of death given nonsmoker}} = 3$$

Logistic regression

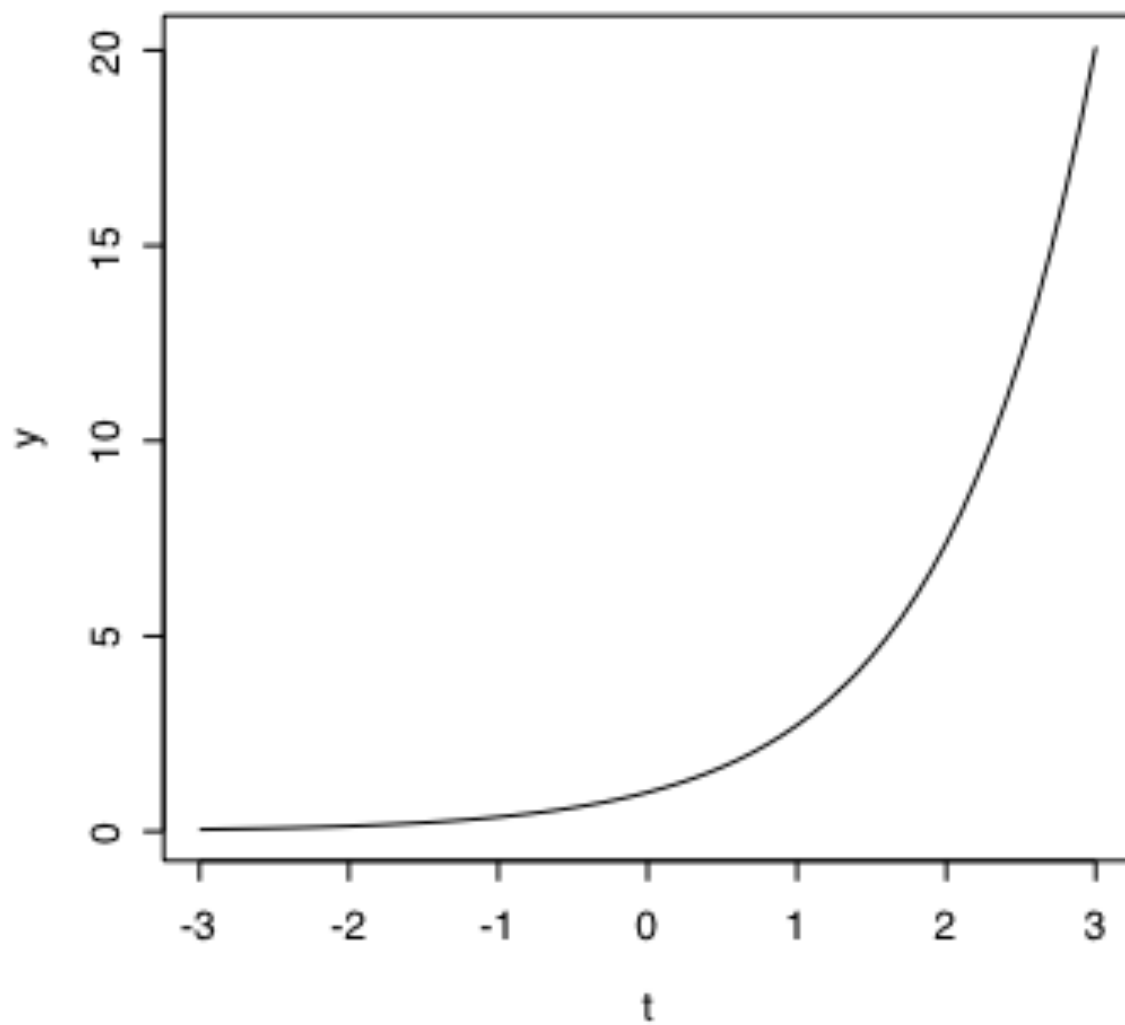
- $X=1$ means smoker, $X=0$ means non-smoker
- $Y=1$ means dead, $Y=0$ means alive
- Log odds of death = $\beta_0 + \beta_1 x$
- Odds of death = $e^{\beta_0} e^{\beta_1 x}$

$$\text{Odds of Death} = e^{\beta_0} e^{\beta_1 x}$$

Group	x	Odds of Death
Smokers	1	$e^{\beta_0} e^{\beta_1}$
Non-smokers	0	e^{β_0}

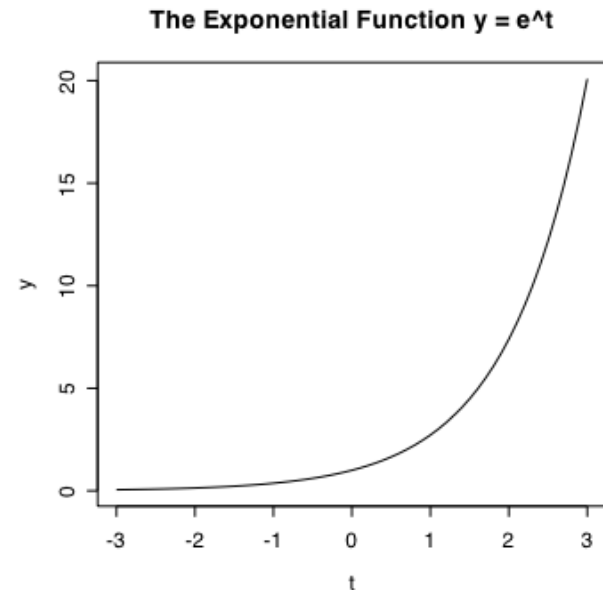
$$\frac{\text{Odds of death given smoker}}{\text{Odds of death given nonsmoker}} = \frac{e^{\beta_0} e^{\beta_1}}{e^{\beta_0}} = e^{\beta_1}$$

The Exponential Function $y = e^t$



Exponential function $f(t) = e^t$

- Always positive
- $e^0=1$, so when $\beta_1 = 0$, the odds ratio equals one (50-50).
- $f(t) = e^t$ is increasing



One more example

$$\text{Log Survival Odds} = \beta_0 + \beta_1 d_1 + \beta_2 d_2 + \beta_3 x$$

Treatment	d_1	d_2	Odds of Survival = $e^{\beta_0} e^{\beta_1 d_1} e^{\beta_2 d_2} e^{\beta_3 x}$
Chemotherapy	1	0	$e^{\beta_0} e^{\beta_1} e^{\beta_3 x}$
Radiation	0	1	$e^{\beta_0} e^{\beta_2} e^{\beta_3 x}$
Both	0	0	$e^{\beta_0} e^{\beta_3 x}$

For any given disease severity
 X ,

$$\frac{\text{Survival odds with Chemo}}{\text{Survival odds with Both}} = \frac{e^{\beta_0} e^{\beta_1} e^{\beta_3 x}}{e^{\beta_0} e^{\beta_3 x}} = e^{\beta_1}$$

In general,

- When x_k is increased by one unit and all other explanatory variables are held constant, the odds of $Y=1$ are multiplied by e^{β_k}
- That is, e^{β_k} is an **odds ratio** --- the ratio of the odds of $Y=1$ when x_k is increased by one unit, to the odds of $Y=1$ when everything is left alone.
- As in ordinary regression, we speak of “controlling” for the other variables.

The conditional probability of $Y=1$

$$P(Y = 1|x_1, \dots, x_{p-1}) = \frac{e^{\beta_0 + \beta_1 x_1 + \dots + \beta_{p-1} x_{p-1}}}{1 + e^{\beta_0 + \beta_1 x_1 + \dots + \beta_{p-1} x_{p-1}}}$$

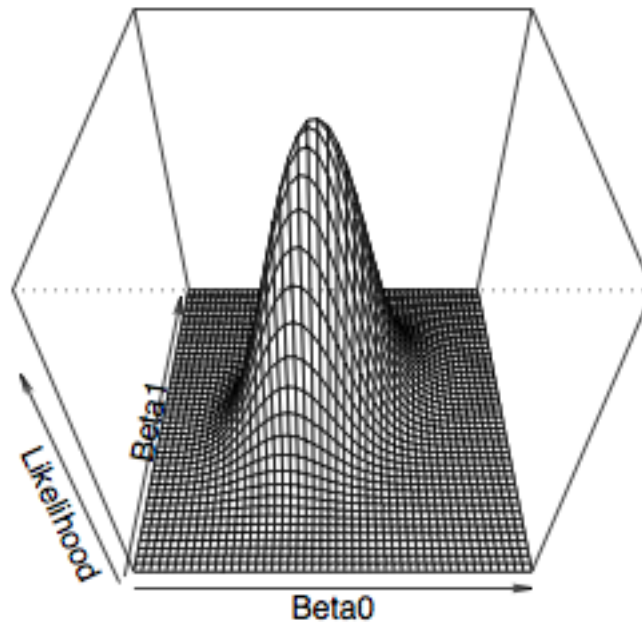
This formula can be used to calculate an estimated $P(Y=1)$
Just replace betas by their estimates (b)

It can also be used to calculate the probability of getting
The sample data values we actually did observe.

Maximum likelihood estimation

- Likelihood = Probability of getting the data values we did observe
- Viewed as a function of the parameters (betas), it's called the "likelihood function."
- Those parameter values for which the likelihood function is greatest are called the *maximum likelihood estimates*.
- Thank you again, Mr. Fisher.

Likelihood Function for Simple Logistic Regression



Maximum likelihood estimates

- Must be found numerically.
- Lead to nice large-sample chi-square tests.
- We will mostly use Wald tests.

Copyright Information

This slide show was prepared by Jerry Brunner, Department of Statistical Sciences, University of Toronto. It is licensed under a Creative Commons Attribution - ShareAlike 3.0 Unported License. Use any part of it as you like and share the result freely. These Powerpoint slides are available from the course website:

<http://www.utstat.toronto.edu/~brunner/oldclass/441s16>