

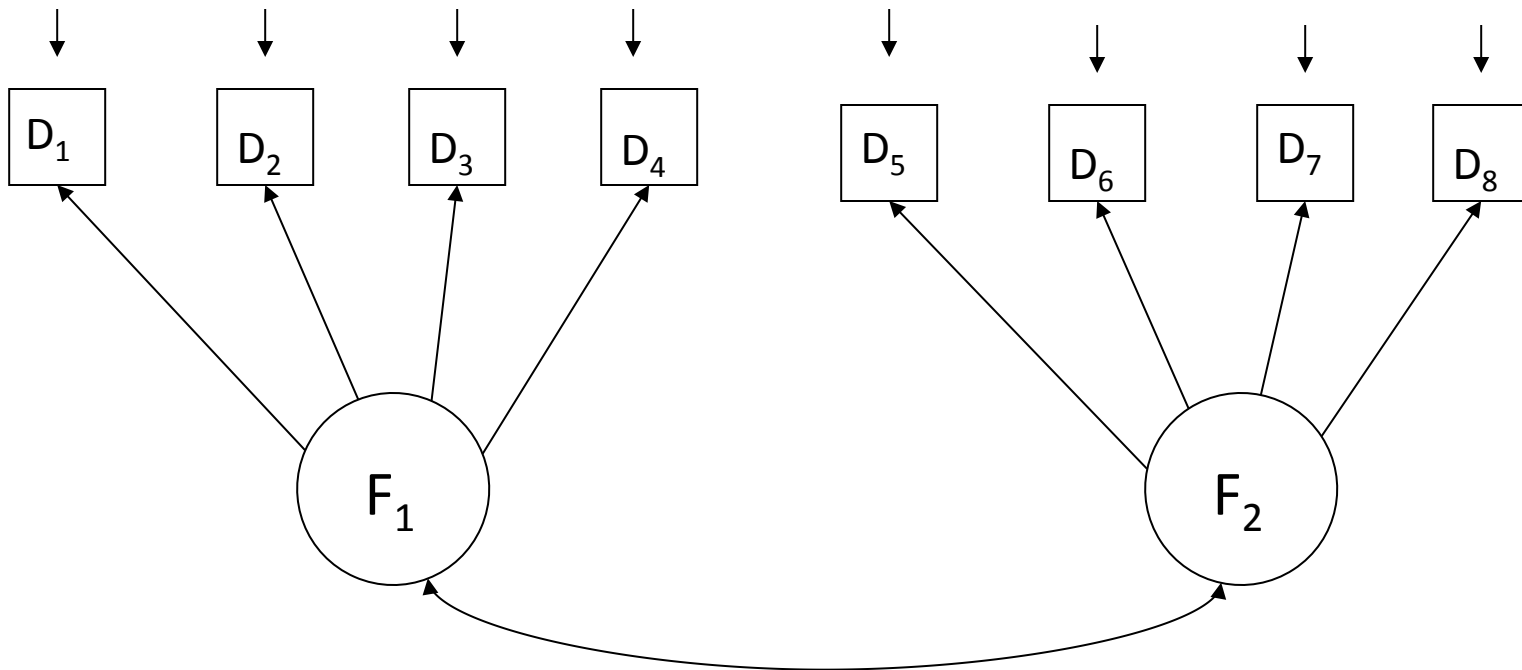
# Exploratory Factor Analysis

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# Factor Analysis: The Measurement Model

$$\mathbf{D}_i = \mathbf{\Lambda} \mathbf{F}_i + \mathbf{e}_i$$



# Example with 2 factors and 8 observed variables

$$\mathbf{D}_i = \mathbf{\Lambda} \mathbf{F}_i + \mathbf{e}_i$$

$$\begin{pmatrix} D_{i,1} \\ D_{i,2} \\ D_{i,3} \\ D_{i,4} \\ D_{i,5} \\ D_{i,6} \\ D_{i,7} \\ D_{i,8} \end{pmatrix} = \begin{pmatrix} \lambda_{11} & \lambda_{12} \\ \lambda_{21} & \lambda_{22} \\ \lambda_{31} & \lambda_{32} \\ \lambda_{41} & \lambda_{42} \\ \lambda_{51} & \lambda_{52} \\ \lambda_{61} & \lambda_{62} \\ \lambda_{71} & \lambda_{72} \\ \lambda_{81} & \lambda_{82} \end{pmatrix} \begin{pmatrix} F_{i,1} \\ F_{i,2} \end{pmatrix} + \begin{pmatrix} e_{i,1} \\ e_{i,2} \\ e_{i,3} \\ e_{i,4} \\ e_{i,5} \\ e_{i,6} \\ e_{i,7} \\ e_{i,8} \end{pmatrix}$$

$$D_{i,1} = \lambda_{11}F_{i,1} + \lambda_{12}F_{i,2} + e_{i,1}$$

$$D_{i,2} = \lambda_{21}F_{i,1} + \lambda_{22}F_{i,2} + e_{i,2} \text{ etc.}$$

The lambda values are called **factor loadings**.

# Terminology

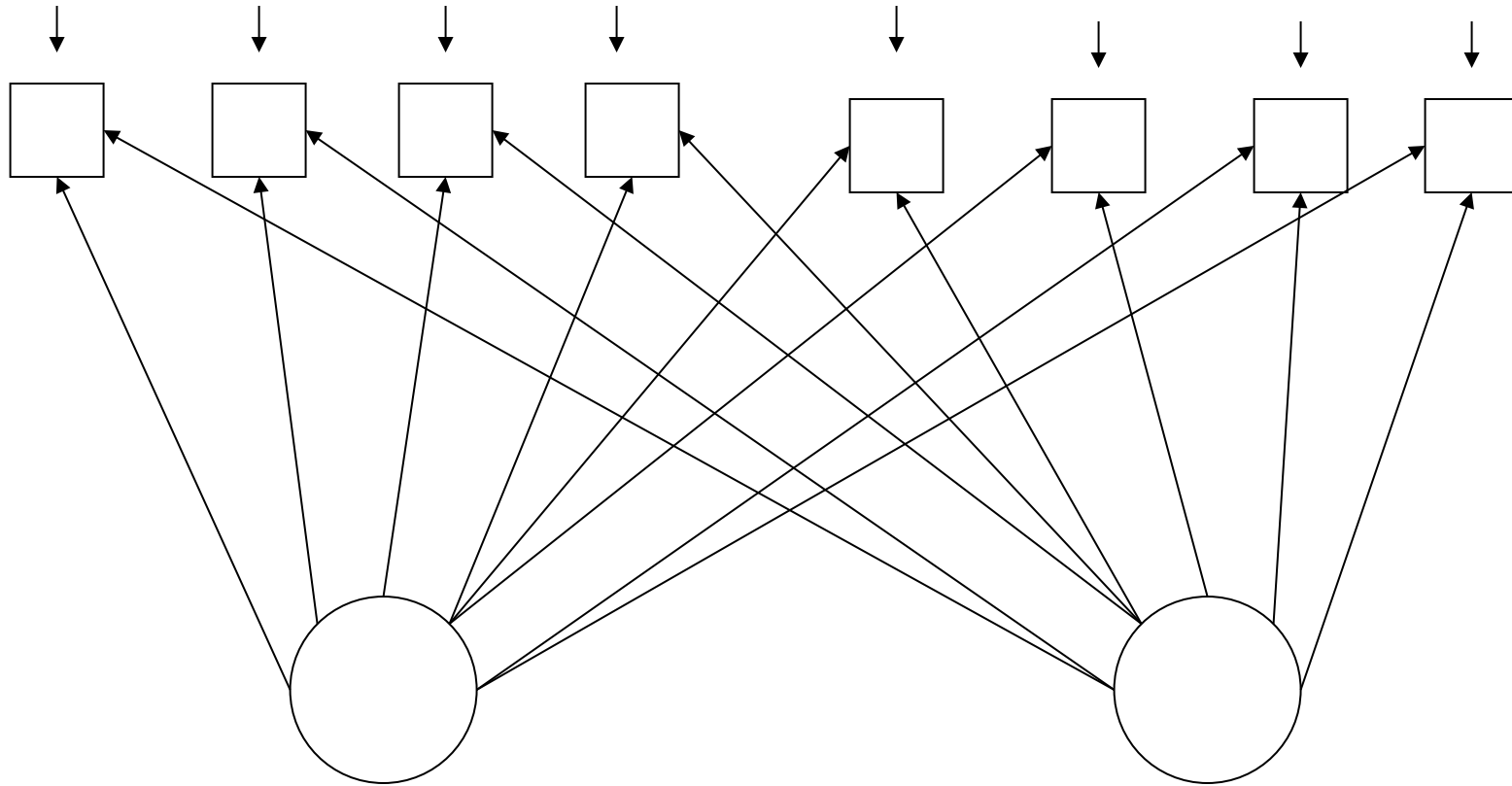
$$D_{i,1} = \lambda_{11}F_{i,1} + \lambda_{12}F_{i,2} + e_{i,1}$$
$$D_{i,2} = \lambda_{21}F_{i,1} + \lambda_{22}F_{i,2} + e_{i,2} \text{ etc.}$$

- The lambda values are called **factor loadings**.
- $F_1$  and  $F_2$  are sometimes called **common factors**, because they influence all the observed variables.
- Error terms  $e_1, \dots, e_8$  are sometimes called **unique factors**, because each one influences only a single observed variable.

# Factor Analysis can be

- **Exploratory:** The goal is to describe and summarize the data by explaining a large number of observed variables in terms of a smaller number of latent variables (factors). The factors are the reason the observable variables have the correlations they do.
- **Confirmatory:** Statistical estimation and testing as usual.

# Unconstrained (Exploratory) Factor Analysis



- Arrows from all factors to all observed variables.
- Massively non-identifiable.
- Reasonable, been going on for around 70-100 years, and completely DOOMED TO FAILURE as a method of statistical estimation.

# Calculate the covariance matrix

$$\mathbf{D}_i = \mathbf{\Lambda} \mathbf{F}_i + \mathbf{e}_i$$

$$\text{cov}(\mathbf{F}_i) = \mathbf{\Phi}$$

$$\text{cov}(\mathbf{e}_i) = \mathbf{\Omega}$$

$\mathbf{F}_i$  and  $\mathbf{e}_i$  independent (multivariate normal)

$$\text{cov}(\mathbf{D}_i) = \mathbf{\Sigma} = \mathbf{\Lambda} \mathbf{\Phi} \mathbf{\Lambda}^\top + \mathbf{\Omega}$$

# A Re-parameterization

$$\Sigma = \Lambda \Phi \Lambda^\top + \Omega$$

Square root matrix:  $\Phi = \mathbf{S}\mathbf{S} = \mathbf{S}\mathbf{S}^\top$ , so

$$\begin{aligned}\Lambda \Phi \Lambda^\top &= \Lambda \mathbf{S}\mathbf{S}^\top \Lambda^\top \\ &= (\Lambda \mathbf{S}) \mathbf{I} (\mathbf{S}^\top \Lambda^\top) \\ &= (\Lambda \mathbf{S}) \mathbf{I} (\Lambda \mathbf{S})^\top \\ &= \Lambda_2 \mathbf{I} \Lambda_2^\top\end{aligned}$$



# Parameters are not identifiable

$$\Sigma = \Lambda \Phi \Lambda^\top + \Omega = \Lambda_2 \mathbf{I} \Lambda_2^\top + \Omega$$

- Two distinct (Lambda, Phi, Omega) sets give the same Sigma, and hence the same distribution of the data (under normality).
- Actually, there are infinitely many. Let  $\mathbf{Q}$  be an arbitrary covariance matrix for  $\mathbf{F}$ .

$$\begin{aligned}\Lambda_2 \mathbf{I} \Lambda_2^\top &= \Lambda_2 \mathbf{Q}^{-\frac{1}{2}} \mathbf{Q} \mathbf{Q}^{-\frac{1}{2}} \Lambda_2^\top \\ &= (\Lambda_2 \mathbf{Q}^{-\frac{1}{2}}) \mathbf{Q} (\mathbf{Q}^{-\frac{1}{2}}^\top \Lambda_2^\top) \\ &= (\Lambda_2 \mathbf{Q}^{-\frac{1}{2}}) \mathbf{Q} (\Lambda_2 \mathbf{Q}^{-\frac{1}{2}})^\top \\ &= \Lambda_3 \mathbf{Q} \Lambda_3^\top\end{aligned}$$

# Restrict the model

$$\mathbf{\Lambda}\mathbf{\Phi}\mathbf{\Lambda}^{\top} = \mathbf{\Lambda}_2\mathbf{I}\mathbf{\Lambda}_2^{\top}$$

- Set  $\mathbf{\Phi}$  = the identity, so  $\text{cov}(\mathbf{F}) = \mathbf{I}$
- All the factors are standardized, as well as independent.
- Justify this on the grounds of simplicity.
- Say the factors are “orthogonal” (at right angles, uncorrelated).

# Another Source of non-identifiability

R is an orthogonal (rotation) matrix

$$\begin{aligned}\Sigma &= \Lambda \Lambda^\top + \Omega \\ &= \Lambda \mathbf{R} \mathbf{R}^\top \Lambda^\top + \Omega \\ &= (\Lambda \mathbf{R})(\mathbf{R}^\top \Lambda^\top) + \Omega \\ &= (\Lambda \mathbf{R})(\Lambda \mathbf{R})^\top + \Omega \\ &= \Lambda_2 \Lambda_2^\top + \Omega\end{aligned}$$

Infinitely many rotation matrices produce the same Sigma.

# A Solution

- Place some restrictions on the factor loadings, so that the only rotation matrix that preserves the restrictions is the identity matrix. For example,  $\lambda_{ij} = 0$  for  $j > i$
- There are other sets of restrictions that work.
- Generally, they result in a set of factor loadings that are impossible to interpret. Don't worry about it.
- Estimate the loadings by maximum likelihood. Other methods are possible but used much less than in the past.
- All (orthogonal) rotations result in the same value of the likelihood function (the maximum is not unique).
- Rotate the factors (that is, post-multiply the estimated loadings by a rotation matrix) so as to achieve a simple pattern that is easy to interpret.
- The result is often satisfying, but has no necessary connection to reality.

# Consulting advice

- When a non-statistician claims to have done a “factor analysis,” ask what kind.
- Usually it was a principal components analysis.
- Principal components are linear combinations of the observed variables. They come from the observed variables by direct calculation.
- In true factor analysis, it’s the observed variables that arise from the factors.
- So principal components analysis is kind of like backwards factor analysis, though the spirit is similar.
- Most factor analysis software (SAS, SPSS, etc.) does principal components analysis by default.

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