

## STA 413F2008 Assignment 7

Please read Sections 5.1, 5.5 and 5.6. The following questions are practice for Test 3 and the Final Examination. They are not to be handed in.

1.  $X_1, \dots, X_n$  be a random sample. For each of the distributions below, give the parameter space  $\Omega$ .
  - (a) Bernoulli
  - (b) Binomial( $m, \theta$ ) with  $m$  known
  - (c) Binomial( $m, \theta$ ) with  $m$  unknown
  - (d) Poisson
  - (e) Geometric
  - (f) Uniform( $\alpha, \beta$ )
  - (g) Exponential
  - (h) Gamma
  - (i) Normal

2. Let  $X_1, \dots, X_n$  be a random sample from a distribution (not necessarily normal) with expected value  $\mu$  and variance  $\sigma^2$ . We will test  $H_0 : \mu = \mu_0$  against  $H_1 : \mu \neq \mu_0$  using the critical region

$$C = \{(x_1, \dots, x_n) : \left| \frac{\bar{x} - \mu_0}{s/\sqrt{n}} \right| > z_{\alpha/2}\}$$

- (a) What is the approximate size of the test?
  - (b) Show that  $H_0$  is rejected if and only if the  $(1 - \alpha)100\%$  confidence interval for  $\mu$  does *not* include  $\mu_0$ . It is easiest to start by writing the set of  $(x_1, \dots, x_n)$  such that  $\mu_0$  is in the confidence interval, and then work on it until it becomes  $C^c$ .
3. Let  $X_1, \dots, X_n$  be a random sample from a normal distribution with expected value  $\mu$  and variance  $\sigma^2$ . We wish to test  $H_0 : \sigma^2 = \sigma_0^2$  against  $H_1 : \sigma^2 > \sigma_0^2$  using the critical region

$$C = \{(x_1, \dots, x_n) : \frac{(n-1)S^2}{\sigma_0^2} > k\}$$

- (a) What is  $\omega_0$ ? Is it simple or composite?
  - (b) What is  $\omega_1$ ? Is it simple or composite?
  - (c) Find  $k$  so that the size of the test is  $\alpha$ . Recall that for random sampling from a normal distribution,  $\frac{(n-1)S^2}{\sigma^2} \sim \chi^2(n-1)$ .
  - (d) Show that this same test is also size  $\alpha$  for testing  $H_0 : \sigma^2 \leq \sigma_0^2$  against  $H_1 : \sigma^2 > \sigma_0^2$ .
  - (e) Write the power function of this test (a function of  $\sigma^2$ ) in terms for the cumulative distribution function of a chi-square random variable.
4. Let  $X_1, \dots, X_{n_1}$  be a random sample from a distribution (not necessarily normal) with expected value  $\mu_1$  and variance  $\sigma_1^2$ , and let  $Y_1, \dots, Y_{n_2}$  be a random sample from a distribution (not necessarily normal) with expected value  $\mu_2$  and variance  $\sigma_2^2$ . The random samples are independent of each other. The Central Limit Theorem tells us that for large

$n_1$ , the distribution of  $\bar{X}_{n_1}$  is approximately  $N(\mu_1, \frac{\sigma_1^2}{n_1})$ . Similarly, the distribution of  $\bar{Y}_{n_2}$  is approximately  $N(\mu_2, \frac{\sigma_2^2}{n_2})$ . So, what should the approximate distribution of  $\bar{X}_{n_1} - \bar{Y}_{n_2}$  be?

5. Let  $X_1, \dots, X_{n_1}$  be a random sample from a (possibly) non-normal distribution with mean  $\mu_1$  and variance  $\sigma_1^2$ , and let  $Y_1, \dots, Y_{n_2}$  be a random sample from a (possibly) non-normal distribution with mean  $\mu_2$  and variance  $\sigma_2^2$ . These are *independent* random samples, meaning that the data are independent between samples as well as within samples. We are interested in testing  $H_0 : \mu_1 = \mu_2$  versus  $H_1 : \mu_1 \neq \mu_2$ . Find the constant  $k$  so that the following critical region will have, for large  $n_1$  and large  $n_2$ , a size of approximately  $\alpha$ .

$$C = \left\{ (\mathbf{x}, \mathbf{y}) : \left| \frac{\bar{X}_{n_1} - \bar{Y}_{n_2}}{\sqrt{\frac{\hat{\sigma}_1^2}{n_1} + \frac{\hat{\sigma}_2^2}{n_2}}} \right| > k \right\},$$

where  $\hat{\sigma}_1^2$  and  $\hat{\sigma}_2^2$  are consistent estimators of  $\sigma_1^2$  and  $\sigma_2^2$ , respectively. There is almost no work to show.

6. Of a random sample of 150 Special Needs students in the Toronto District School Board, 19 were in regular classes, and the rest were in Special Education classes. Of a random sample of 200 Special Needs students in the Toronto Separate School Board, 48 were in regular classes, and the rest were in Special Education classes. Test for difference between the proportions of Special Needs students in regular classes in the two school boards. Use  $\alpha = 0.01$ . What do you conclude? Is the proportion of Special Needs students greater in one of the school boards? Of course you should use the test from the last question.
7. Let  $X_1, \dots, X_{n_1}$  be a random sample from an exponential distribution with parameter  $\theta$ . The null hypothesis is  $H_0 : \theta = \theta_0$  versus  $H_1 : \theta > \theta_0$ .
- Give the critical region  $C_1$  for an approximate size  $\alpha$  test based on the modified Central Limit Theorem. Your answer will use critical values from the standard normal distribution.
  - Write the power function of this test explicitly as a function of the true parameter  $\theta$ . Your answer will involve  $\Phi$ , the cumulative distribution function of a standard normal.
  - Suppose  $\theta_0 = 2$ ,  $\alpha = 0.05$  and  $n = 30$ . What is the approximate power for  $\theta = 2.5$ ? The answer is a single number. I get a power of 0.4129.
  - What is the smallest sample size that will guarantee an approximate power of 0.80 for  $\theta = 2.5$ ? The answer is a single number (an integer, of course). I get  $n = 117$ .
  - Again for  $\theta_0 = 2$ ,  $\alpha = 0.05$  and  $n = 30$ , suppose we observe  $\bar{X}_n = 3$ .
    - What is the value of the test statistic? The answer is a number.
    - Is  $\mathbf{X} \in C$ ? Answer Yes or No.
    - Do you reject  $H_0$ ? Answer Yes or No.
    - Calculate the  $p$ -value. The answer is a number. I get 0.0031.

8. Here is another test for the exponential model of Question 7. Let

$$C = \{\mathbf{x} : \sqrt{n}(\ln(\bar{X}_n) - \ln(\theta_0)) > k\}$$

- (a) Find the value of  $k$  so that this test has approximately size  $\alpha$ . Your answer will involve a critical value of the standard normal distribution. Start by finding the limiting distribution of the test statistic under the null hypothesis.
- (b) Express the power function  $P_\theta\{\mathbf{X} \in C\}$  in terms of the cumulative distribution function of a standard normal random variable.
- (c) Is the power function strictly increasing in  $\theta$ ? Answer Yes or No and prove your answer.
- (d) Suppose the null hypothesis were  $H_0 : \theta \leq \theta_0$ . Is this test still size  $\alpha$ ? Answer Yes or No and justify your answer.
- (e) Suppose  $\theta_0 = 2$ ,  $\alpha = 0.05$  and  $n = 30$ . What is the approximate power for  $\theta = 2.5$ ? The answer is a single number. I get 0.336.

9. Here is yet another test for the exponential model of Question 7. Let

$$C = \{\mathbf{x} : \frac{2\bar{X}_n}{\theta_0} > k\}$$

- (a) Find the value of  $k$  so that this test has *exactly* size  $\alpha$ . Your answer will involve the chi-square distribution. Start by using moment-generating functions to find the distribution of the test statistic under the null hypothesis.
- (b) Express the power function  $P_\theta\{\mathbf{X} \in C\}$  in terms of the cumulative distribution function of a chi-square random variable.
- (c) Is the power function strictly increasing in  $\theta$ ? Answer Yes or No and prove your answer.
- (d) Suppose the null hypothesis were  $H_0 : \theta \leq \theta_0$ . Is this test still size  $\alpha$ ? Answer Yes or No and justify your answer.
- (e) This last item is just a comment. For  $\theta_0 = 2$ ,  $\alpha = 0.05$  and  $n = 30$ , the generic large-sample test had a power of 0.4129 at  $\theta = 2.5$ , while the variance-stabilized test had a power of 0.336. The exact test has a power of 0.362.