

## Weibull Regression with R, Part Two\*

```

> rm(list=ls()); options(scipen=999)
> # install.packages("survival",dependencies=TRUE) # Only need to do this once
> library(survival) # Do this every time
> # install.packages("asaur",dependencies=TRUE) # Only need to do this once
> library(asaur)
> # help(pharmacoSmoking)
> head(pharmacoSmoking)
  id ttr relapse    grp age gender    race employment yearsSmoking
1  21 182      0  patchOnly 36  Male   white          ft           26
2 113  14      1  patchOnly 41  Male   white          other          27
3  39   5      1 combination 25 Female white          other          12
4  80  16      1 combination 54  Male   white          ft           39
5  87   0      1 combination 45  Male   white          other          30
6  29 182      0 combination 43  Male  hispanic        ft           30
  levelSmoking ageGroup2 ageGroup4 priorAttempts longestNoSmoke
1      heavy      21-49      35-49           0           0
2      heavy      21-49      35-49           3          90
3      heavy      21-49      21-34           3          21
4      heavy       50+      50-64           0           0
5      heavy      21-49      35-49           0           0
6      heavy      21-49      35-49           2         1825
> summary(pharmacoSmoking)
  id          ttr          relapse          grp
Min.   : 1.00   Min.   : 0.00   Min.   :0.000   combination:61
1st Qu.: 33.00  1st Qu.:  8.00   1st Qu.:0.000   patchOnly  :64
Median : 67.00  Median : 49.00   Median :1.000
Mean   : 66.15  Mean   : 77.44   Mean   :0.712
3rd Qu.: 99.00  3rd Qu.:182.00  3rd Qu.:1.000
Max.   :130.00  Max.   :182.00  Max.   :1.000

  age          gender          race          employment          yearsSmoking
Min.   :22.00   Female:81   black   :38   ft   :72   Min.   : 9.00
1st Qu.:41.00   Male  :44   hispanic: 8   other:39  1st Qu.:22.00
Median :49.00          other  : 2   pt   :14   Median :30.00
Mean   :48.84          white  :77          Mean   :30.88
3rd Qu.:56.00          Max.   :56.00
Max.   :86.00

  levelSmoking ageGroup2 ageGroup4 priorAttempts longestNoSmoke
heavy:89      21-49:66   21-34:16   Min.   : 0.00   Min.   : 0.0
light:36      50+ :59    35-49:50   1st Qu.: 1.00   1st Qu.:  7.0
          50-64:48   Median : 2.00   Median : 90.0
          65+ :11   Mean   : 12.68  Mean   : 539.7
          3rd Qu.: 5.00  3rd Qu.: 365.0
          Max.   :1000.00  Max.   :6205.0
> attach(pharmacoSmoking)
>
> # Make patch only the reference category
> contrasts(grp) = contr.treatment(2,base=2)
> colnames(contrasts(grp)) = c('Combo') # Names of dummy vars -- just one
> DayOfRelapse = Surv(ttr+1,relapse) # Day of relapse starts with one.
> # Collapse race categories
> Race = as.character(race) # Small r race is a factor. This is easier to modify.
> Race[Race!='white'] = 'blackOther'; Race=as.factor(Race)
> table(race,Race)
      Race
race   blackOther white
black          38     0
hispanic         8     0
other            2     0
white            0    77

```

\* Copyright information is on the last page.

```
> full = survreg(DayOfRelapse ~ grp + age + gender + Race + employment
+
+ yearsSmoking + levelSmoking + priorAttempts,
+
+ dist='weibull')
> summary(full)
```

Call:

```
survreg(formula = DayOfRelapse ~ grp + age + gender + Race +
employment + yearsSmoking + levelSmoking + priorAttempts,
dist = "weibull")
```

	Value	Std. Error	z	p
(Intercept)	1.12177	0.9773	1.1479	0.25102045958
grpCombo	1.09225	0.3819	2.8603	0.00423234508
age	0.08432	0.0341	2.4722	0.01342778276
genderMale	0.03631	0.4142	0.0877	0.93014517788
Racewhite	0.25145	0.3914	0.6424	0.52061740468
employmentother	-1.28799	0.4672	-2.7569	0.00583496922
employmentpt	-1.28482	0.5863	-2.1914	0.02842409501
yearsSmoking	-0.02351	0.0325	-0.7232	0.46955818306
levelSmokinglight	-0.07347	0.4315	-0.1703	0.86480382316
priorAttempts	-0.00105	0.0020	-0.5244	0.59996899163
Log(scale)	0.54194	0.0892	6.0774	0.00000000122

Scale= 1.72

Weibull distribution

Loglik(model)= -463.8    Loglik(intercept only)= -476.5

Chisq= 25.41 on 9 degrees of freedom, p= 0.0025

Number of Newton-Raphson Iterations: 5

n= 125

>

I am thinking about dropping Race, yearsSmoking, levelSmoking and priorAttempts. The last 3 variables all represent smoking history and could be correlated highly enough to wash out each other's effects. Test them simultaneously.

>

```
> # Fit the restricted model: Restricted by H0
```

```
> rest1 = survreg(DayOfRelapse ~ grp + age + gender + Race + employment ,
+
+ dist='weibull')
```

```
> anova(rest1,full) # LR test
```

Terms

1 grp + age + gender + Race + employment

2 grp + age + gender + Race + employment + yearsSmoking + levelSmoking + priorAttempts

	Resid. Df	-2*LL	Test Df	Deviance	Pr(>Chi)
1	117	928.3771	NA	NA	NA
2	114	927.5513	= 3	0.8258271	0.8432801

1    117    928.3771    NA    NA    NA

2    114    927.5513    = 3    0.8258271    0.8432801

>

```
> # Is Race significant with those variables dropped?
```

```
> # Is Race significant with those variables dropped?
> summary(rest1)
```

```
Call:
survreg(formula = DayOfRelapse ~ grp + age + gender + Race +
  employment, dist = "weibull")
      Value Std. Error      z      p
(Intercept)  1.3905    0.8684  1.601 0.10932864665
grpCombo     1.1021    0.3794  2.905 0.00367117788
age          0.0637    0.0190  3.354 0.00079527376
genderMale   0.0561    0.4140  0.136 0.89213377864
Racewhite    0.1880    0.3788  0.496 0.61958001498
employmentother -1.2821  0.4635 -2.766 0.00567301250
employmentpt -1.2251  0.5837 -2.099 0.03582864119
Log(scale)   0.5444    0.0894  6.090 0.00000000113
```

```
Scale= 1.72
```

```
Weibull distribution
Loglik(model)= -464.2  Loglik(intercept only)= -476.5
  Chisq= 24.58 on 6 degrees of freedom, p= 0.00041
Number of Newton-Raphson Iterations: 5
n= 125
```

Decision: Drop race and gender.

```
> full2 = survreg(DayOfRelapse ~ grp + age + employment , dist='weibull')
> summary(full2)
```

```
Call:
survreg(formula = DayOfRelapse ~ grp + age + employment, dist = "weibull")
      Value Std. Error      z      p
(Intercept)  1.4957    0.8414  1.78 0.07545324261
grpCombo     1.1023    0.3793  2.91 0.00365915983
age          0.0643    0.0186  3.45 0.00055474131
employmentother -1.2880  0.4617 -2.79 0.00527676297
employmentpt -1.2123  0.5616 -2.16 0.03088499029
Log(scale)   0.5454    0.0894  6.10 0.00000000105
```

```
Scale= 1.73
```

```
Weibull distribution
Loglik(model)= -464.3  Loglik(intercept only)= -476.5
  Chisq= 24.31 on 4 degrees of freedom, p= 0.000069
Number of Newton-Raphson Iterations: 5
n= 125
```

```
>
> # Test employment status controlling for age and experimental treatment.
> rest2 = survreg(DayOfRelapse ~ grp + age , dist='weibull')
> anova(rest2,full2) # LR test
      Terms Resid. Df    -2*LL Test Df Deviance    Pr(>Chi)
1      grp + age      121 937.9007      NA      NA      NA
2 grp + age + employment      119 928.6554    = 2 9.245333 0.009826558
>
> # Test employment status with a Wald test.
> source("http://www.utstat.toronto.edu/~brunner/Rfunctions/Wtest.txt")
> # function(L,Tn,Vn,h=0) # H0: L theta = h
> # Tn is estimated theta, usually a vector.
> # Vn is the estimated asymptotic covariance matrix of Tn.
> # For Wald tests based on numerical MLEs, Tn = theta-hat,
> # and Vn is the inverse of the Hessian of the minus log
> # likelihood.
```

```

>
> Vhat = vcov(full2); Vhat
      (Intercept)      grpCombo      age employmentother
(Intercept)  0.7079360800 -0.0320256900 -0.0147694486  0.111673731
grpCombo     -0.0320256900  0.1438698739 -0.0004703383  -0.013521493
age          -0.0147694486 -0.0004703383  0.0003472409  -0.003893727
employmentother 0.1116737308 -0.0135214927 -0.0038937268  0.213191081
employmentpt -0.0003554818 -0.0078279548 -0.0013434899  0.077138486
Log(scale)   -0.0098224903  0.0050290739  0.0002048412  -0.003182291
      employmentpt      Log(scale)
(Intercept) -0.0003554818 -0.0098224903
grpCombo     -0.0078279548  0.0050290739
age          -0.0013434899  0.0002048412
employmentother 0.0771384860 -0.0031822913
employmentpt  0.3153999894 -0.0035442716
Log(scale)   -0.0035442716  0.0079888732

> thetahat = full2$coefficients; thetahat
      (Intercept)      grpCombo      age employmentother
      1.4957374      1.1023048      0.0643414      -1.2880472
      employmentpt
      -1.2122529
>

```

Note that the asymptotic covariance matrix includes  $\log(\sigma)$ , but the "coefficients" vector does not.

```

> sigmahat = full2$scale; sigmahat
[1] 1.725305
> thetahat = c(thetahat, log(sigmahat))
>
> # H0: beta3=beta4=0. Express as H0: L theta = h
> eMat = rbind( c(0,0,0,1,0,0),
+              c(0,0,0,0,1,0) )
> Wtest(L=eMat, Tn=thetahat, Vn=Vhat)
      W      df      p-value
9.718885315 2.000000000 0.007754805
> anova(rest2,full2) # Repeating LR test for comparison
      Terms Resid. Df      -2*LL Test Df Deviance      Pr(>Chi)
1 grp + age      121 937.9007      NA      NA      NA
2 grp + age + employment 119 928.6554      = 2 9.245333 0.009826558

>
> # Test part time versus other
> pto = cbind(0,0,0,1,-1,0); pto
      [,1] [,2] [,3] [,4] [,5] [,6]
[1,]    0    0    0    1   -1    0
> Wtest(L=pto, Tn=thetahat, Vn=Vhat)
      W      df      p-value
0.01534747 1.00000000 0.90140640
>

```

Predict the day of relapse for a 50 year old patient who is employed full time and gets the patch-only treatment.

Weibull Regression:  $t_i = \exp\{\beta_0 + \beta_1 x_{i,1} + \dots + \beta_{p-1} x_{i,p-1}\} \cdot \epsilon_i^\sigma = e^{\mathbf{x}_i^\top \boldsymbol{\beta}} \epsilon_i^\sigma$ , where  $\epsilon_i \sim \exp(1)$ .

- $t_i \sim \text{Weibull}$ , with  $\alpha = 1/\sigma$  and  $\lambda = e^{-\mathbf{x}_i^\top \boldsymbol{\beta}}$ .
- $E(t_i) = e^{\mathbf{x}_i^\top \boldsymbol{\beta}} \Gamma(\sigma+1)$ ,  $\text{Median}(t_i) = e^{\mathbf{x}_i^\top \boldsymbol{\beta}} \log(2)^\sigma$ ,  $h_i(t) = \frac{1}{\sigma} \exp\{-\frac{1}{\sigma} \mathbf{x}_i^\top \boldsymbol{\beta}\} t^{\frac{1}{\sigma}-1}$ .
- $S(t) = \exp\left\{-e^{-\frac{1}{\sigma} \mathbf{x}_i^\top \boldsymbol{\beta}} t^{\frac{1}{\sigma}}\right\}$

```
> thetahat
  (Intercept)      grpCombo      age employmentother
    1.4957374      1.1023048      0.0643414      -1.2880472
  employmentpt
    -1.2122529      0.5454037
> x = c(1,0,50,0,0,0)
> xb = sum(x*tethahat)
>
> # a) Use the estimated mean
> exp(xb) * gamma(sigmahat+1)
[1] 175.5516
>
> # b) Use the estimated median
> exp(xb) * log(2)^sigmahat
[1] 59.17273
>
> # I think the median is preferable to mean because the Weibull distribution
> # is skewed. Also, the predict function for Weibull regression works as expected
> # for medians (but not means).
>
> oldguy = data.frame(grp='patchOnly',age=50,employment='ft')
> predict(full2,newdata=oldguy,type='quantile',p=0.5,se=TRUE)
$fit
      1
59.17273

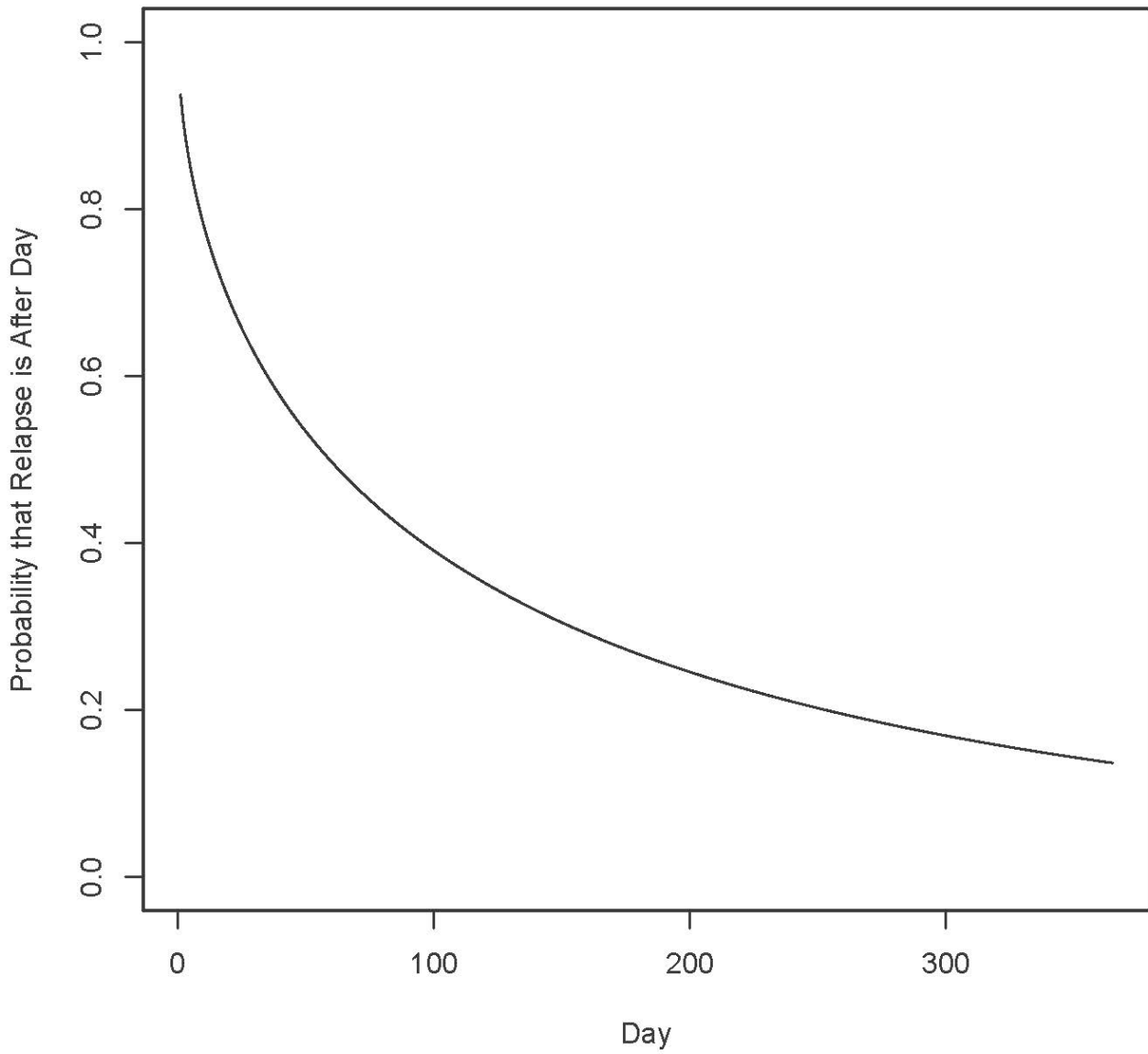
$se.fit
      1
18.87577

> # The 0.5 quantile is the median. se is from the delta method.
>
> # Estimate and plot S(t) for the old guy.
```

The se is straightforward in theory, but messy in practice. G-dot from Mathematica is very ugly. For example, try  $D[\exp(-t^{1/s}) \exp(-(b_0 + 50 b_2)/s), b_0]$

```
> t = 1:365
> Shat = exp(-(exp(-xb/sigmahat)*t^(1/sigmahat)))
>
> plot(t,Shat,type='l',ylim=c(0,1),xlab='Day',
+ ylab='Probability that Relapse is After Day')
> title('Probability of Not Relapsing')
```

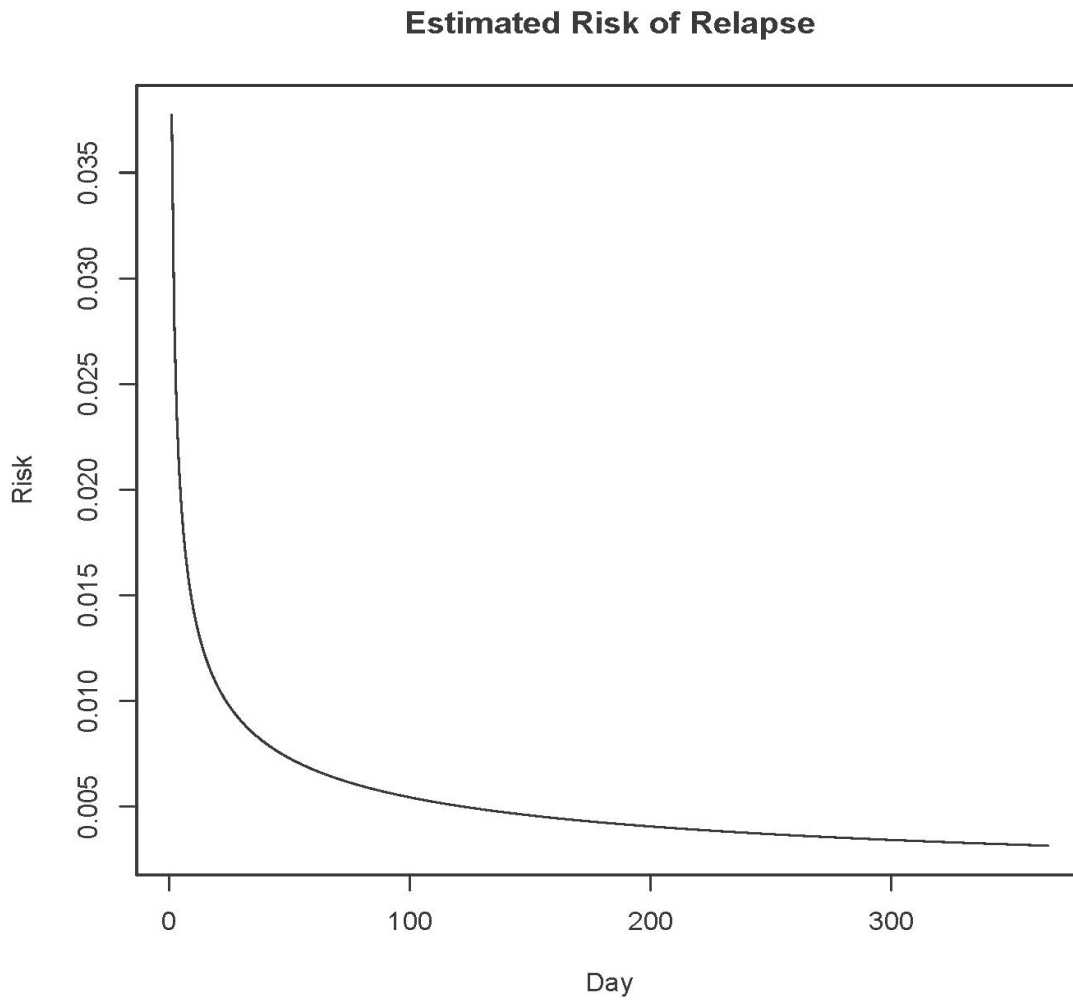
## Probability of Not Relapsing



Plot estimated hazard function for that 50 year old patient who is employed full time and gets the patch-only treatment.

$$\begin{aligned}h(t) &= \frac{f(t)}{S(t)} \\&= \frac{\alpha\lambda(\lambda t)^{\alpha-1} e^{-(\lambda t)^\alpha}}{e^{-(\lambda t)^\alpha}} \\&= \alpha\lambda^\alpha t^{\alpha-1} \\&= \frac{1}{\sigma} e^{-\frac{1}{\sigma} \mathbf{x}^\top \boldsymbol{\beta}} t^{\frac{1}{\sigma}-1}\end{aligned}$$

```
> h = 1/sigmahat * exp(-xb/sigmahat) * t^(1/sigmahat - 1)
> plot(t,h,type='l',xlab='Day',ylab='Risk',main='Estimated Risk of Relapse')
```



# LaTeX code for the record

```

\noindent
Weibull Regression:  $t_i = \exp\{\beta_0 + \beta_1 x_{i,1} + \dots + \beta_{p-1} x_{i,p-1}\} \cdot \epsilon_i^\sigma = e^{\mathbf{x}_i \boldsymbol{\beta}}$ 
 $\boldsymbol{\beta} \epsilon_i^\sigma$ ,
where  $\epsilon_1 \sim \exp(1)$ .
\begin{itemize}
\item  $t_i \sim$  Weibull, with  $\alpha = 1/\sigma$  and  $\lambda = e^{-\mathbf{x}_i \boldsymbol{\beta}}$ .
\item  $E(t_i) = e^{\mathbf{x}_i \boldsymbol{\beta}} \Gamma(\sigma + 1)$ ,
Median( $t_i$ ) =  $e^{\mathbf{x}_i \boldsymbol{\beta}} \log(2)^\sigma$ ,
 $h_i(t) = \frac{1}{\sigma} \exp\{-\frac{1}{\sigma} \mathbf{x}_i \boldsymbol{\beta} t^{\frac{1}{\sigma}-1}\}$ .
\item  $S(t) = e^{-\left( e^{-\frac{1}{\sigma} \mathbf{x}_i \boldsymbol{\beta} t^{\frac{1}{\sigma}}} \right)}$ 
\item  $S(t) = \exp\left\{ - e^{-\frac{1}{\sigma} \mathbf{x}_i \boldsymbol{\beta} t^{\frac{1}{\sigma}}} \right\}$ 
\end{itemize}

% Hazard calculation
\begin{eqnarray*}
h(t) & = & \frac{f(t)}{S(t)} \\
& = & \frac{\alpha \lambda (\lambda t)^{\alpha-1} e^{-(\lambda t)^\alpha}}{e^{-(\lambda t)^\alpha}} \\
& = & \alpha \lambda^\alpha t^{\alpha-1} \\
& = & \frac{1}{\sigma} e^{-\frac{1}{\sigma} \mathbf{x}_i \boldsymbol{\beta} t^{\frac{1}{\sigma}}}
\end{eqnarray*}

```

This document was prepared by [Jerry Brunner](#), University of Toronto. It is licensed under a Creative Commons Attribution - ShareAlike 3.0 Unported License:

[http://creativecommons.org/licenses/by-sa/3.0/deed.en\\_US](http://creativecommons.org/licenses/by-sa/3.0/deed.en_US). Use any part of it as you like and share the result freely. It is available in OpenOffice.org format from the course website:

<http://www.utstat.toronto.edu/~brunner/oldclass/312s19>