

# Assignment 1

1

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> x = -4:4; px = abs(x)/20; y = x^2-1
> cbind(x, x^2, px, y)
      x      px      y
[1,] -4 16 0.20 15
[2,] -3  9 0.15  8
[3,] -2  4 0.10  3
[4,] -1  1 0.05  0
[5,]  0  0 0.00 -1
[6,]  1  1 0.05  0
[7,]  2  4 0.10  3
[8,]  3  9 0.15  8
[9,]  4 16 0.20 15
> Ex = sum(x*px); Ex # You know it's zero by symmetry
[1] 0
> Varx = sum(x^2*px); Varx
[1] 10
> # Use formula for E(g(x)) to get E(y) and Var(y)
> Ey = sum(y*px); Ey
[1] 9
> Vary = sum( (y-Ey)^2 * px ); Vary
[1] 30
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$$(f) P(y=0) = P(x=1) + P(x=-1) = 1/20 + 1/20 = 1/10$$

$$P(y=3) = P(x=2) + P(x=-2) = 2/20 + 2/20 = 2/10$$

$$P(y=8) = P(x=3) + P(x=-3) = 3/20 + 3/20 = 3/10$$

$$P(y=15) = P(x=4) + P(x=-4) = 4/20 + 4/20 = 4/10$$

$$(g) E(y) = \sum_y y P(y) = 0 \cdot \frac{1}{10} + 3 \cdot \frac{2}{10} + 8 \cdot \frac{3}{10} + 15 \cdot \frac{4}{10} \\ = (6 + 24 + 60)/10 = 90/10 = 9$$

$$(h) \text{Var}(y) = \sum_y (y - \mu_y)^2 P(y) \\ = (0-9)^2 \cdot \frac{1}{10} + (3-9)^2 \cdot \frac{2}{10} + (8-9)^2 \cdot \frac{3}{10} + (15-9)^2 \cdot \frac{4}{10} \\ = 81 \cdot \frac{1}{10} + 36 \cdot \frac{2}{10} + 1 \cdot \frac{3}{10} + 36 \cdot \frac{4}{10} \\ = (81 + 72 + 3 + 144)/10 = 300/10 = 30$$

Note (g) & (h) agree with R calculations.

$$\textcircled{2} \textcircled{a) } E(x) = \sum_x x p(x) = a \cdot 1 = a$$

$$\text{Var}(x) = \sum_x (x - \mu_x)^2 p(x) = (a - a)^2 \cdot 1 = 0$$

$$\textcircled{b) } E(y) = \int_{-\infty}^{\infty} a f(x) dx = a \int_{-\infty}^{\infty} f(x) dx = a \cdot 1 = a$$

$$\text{Var}(y) = E(y - E(y))^2 = E\{(g(x) - a)^2\} = E\{h(x)\}$$

$$= \int_{-\infty}^{\infty} (a - a)^2 f(x) dx = \int_{-\infty}^{\infty} 0 \cdot f(x) dx$$

$$= \int_{-\infty}^{\infty} 0 dx = 0$$

(3) (a and b): Write the marginal distributions on the margins

	$x=1$	$x=2$	$x=3$	
$x_1=1$	$3/12$	$1/12$	$3/12$	$7/12$
$x_1=2$	$1/12$	$3/12$	$1/12$	$5/12$
	$4/12$	$4/12$	$4/12$	

(c)  $E(x) = (1+2+3) \cdot \frac{4}{12} = \frac{24}{12} = 2$

(d)

$x$	$x_c = x-2$	$x_c^2 = (x-2)^2$	$P(x)$
1	-1	1	$1/3$
2	0	0	$1/3$
3	1	1	$1/3$

(i)

$x_c$	-1	0	1
$P(x_c)$	$1/3$	$1/3$	$1/3$

(ii)  $E(x_c) = (-1) \cdot \frac{1}{3} + (0) \cdot \frac{1}{3} + 1 \cdot \frac{1}{3} = 0$

(iii)

$x_c^2$	$P(x_c^2)$
0	$1/3$
1	$2/3$

(iv)  $E(x_c^2) = 0 \cdot \frac{1}{3} + 1 \cdot \frac{2}{3} = \frac{2}{3}$

$$(3e) \text{Var}(x) = E(x_c^2) = \frac{2}{3}$$

$$(f) E(y) = 1 \cdot \frac{7}{12} + 2 \cdot \frac{5}{12} = \frac{17}{12}$$

$$(g) \text{Var}(y) = E(y^2) - (E(y))^2$$

$$E(y^2) = 1^2 \cdot \frac{7}{12} + 2^2 \cdot \frac{5}{12} = \frac{7}{12} + \frac{20}{12} = \frac{27}{12}$$

$$= \frac{9}{4}, \text{ so}$$

$$\text{Var}(y) = \frac{27}{12} - \left(\frac{17}{12}\right)^2 = \frac{324}{144} - \frac{289}{144} = \frac{35}{144}$$

(h)

x	y	x+y	P(x,y)
1	1	2	3/12
1	2	3	1/12
2	1	3	1/12
2	2	4	3/12
3	1	4	3/12
3	2	5	1/12

$z_1$	2	3	4	5
P( $z_1$ )	3/12	2/12	6/12	1/12

$$(i) E(z_1) = 2 \cdot \frac{3}{12} + 3 \cdot \frac{2}{12} + 4 \cdot \frac{6}{12} + 5 \cdot \frac{1}{12}$$

$$= (6 + 6 + 24 + 5) / 12 = 41 / 12$$

$$(j) E(x) + E(y) = 2 + \frac{17}{12} = \frac{24 + 17}{12} = \frac{41}{12} \quad \text{yes}$$

3k

$x$	$y$	$xy$	$P(x, y)$
1	1	1	3/12
1	2	2	1/12
2	1	2	1/12
2	2	4	3/12
3	1	3	3/12
3	2	6	1/12

$z_2$	1	2	3	4	6
$P(z_2)$	3/12	2/12	3/12	3/12	1/12

$$\begin{aligned} \textcircled{2} E(z_2) &= 1 \cdot \frac{3}{12} + 2 \cdot \frac{2}{12} + 3 \cdot \frac{3}{12} + 4 \cdot \frac{3}{12} + 6 \cdot \frac{1}{12} \\ &= (3 + 4 + 9 + 12 + 6) / 12 = 34 / 12 = 17 / 6 \end{aligned}$$

$$\textcircled{m} E(x)E(y) = 2 \cdot \frac{17}{12} = 17/6 \quad \text{yes}$$

$$\textcircled{n} \text{Cov}(x, y) = E(xy) - E(x)E(y) = \frac{17}{6} - \frac{17}{6} = 0$$

$\textcircled{o}$   $x$  and  $y$  independent means  $P(x, y) = P(x)P(y)$  for all  $x \neq y$ .  $P(x=1, y=1) = 3/12 = 36/144$ , but  $P(x=1)P(y=1) = \frac{4}{12} \cdot \frac{7}{12} = 28/144$ . So the answer is NO, not independent.

There can be zero covariance without independence.

$$\textcircled{4} \quad E(x_1, x_2) = \int \int x_1, x_2 f_{x_1, x_2}(x_1, x_2) dx_1 dx_2$$

$$= \int \int x_1, x_2 f_{x_1}(x_1) f_{x_2}(x_2) dx_1 dx_2$$

This is where I use independence

$$= \int x_2 \left( \int x_1 f_{x_1}(x_1) dx_1 \right) f_{x_2}(x_2) dx_2$$

$$= \int x_2 E(x_1) f_{x_2}(x_2) dx_2$$

$$= E(x_1) \int x_2 f_{x_2}(x_2) dx_2 = E(x_1) E(x_2)$$

If  $x_1 \neq x_2$  are discrete,

$$E(x_1, x_2) = \sum_{x_1} \sum_{x_2} x_1, x_2 P(x_1, x_2)$$

$$= \sum_{x_1} \sum_{x_2} x_1, x_2 P_1(x_1) P_2(x_2)$$

↑  
indep.

etc.

$$\begin{aligned}
 (5) \quad (a) \quad \text{Var}(Y) &= E\{(Y - \mu_Y)^2\} \\
 &= E\{Y^2 - 2Y\mu_Y + \mu_Y^2\} = E(Y^2) - 2\mu_Y E(Y) + E(\mu_Y^2) \\
 &= E(Y^2) - 2\mu_Y^2 + \mu_Y^2 = E(Y^2) - \mu_Y^2
 \end{aligned}$$

$$\begin{aligned}
 (b) \quad \text{Cov}(X, Y) &= E\{(X - \mu_X)(Y - \mu_Y)\} \\
 &= E\{XY - X\mu_Y - \mu_X Y + \mu_X \mu_Y\} \\
 &= E(XY) - \mu_Y E(X) - \mu_X E(Y) + E(\mu_X \mu_Y) \\
 &= E(XY) - 2\mu_X \mu_Y + \mu_X \mu_Y \\
 &= E(XY) - \mu_X \mu_Y
 \end{aligned}$$

If  $X$  and  $Y$  are independent, then  $E(XY) = E(X)E(Y)$  by Problem 4, and

$$\text{Cov}(X, Y) \stackrel{(b)}{=} E(XY) - E(X)E(Y) = 0$$

Note however by Problem 3 that zero covariance does not necessarily imply independence.

⑥ (a) Noting  $E(ax) = a\mu_x$ ,

$$\begin{aligned}\text{Var}(ax) &= E\{(ax - a\mu_x)^2\} \\ &= E\{a^2(x - \mu_x)^2\} = a^2 E\{(x - \mu_x)^2\} \\ &= a^2 \text{Var}(x)\end{aligned}$$

(b) Noting  $E(x+a) = \mu_x + a$ ,

$$\begin{aligned}\text{Var}(x+a) &= E\{(x+a - (\mu_x + a))^2\} \\ &= E\{(x+a - \mu_x - a)^2\} \\ &= E\{(x - \mu_x)^2\} = \text{Var}(x)\end{aligned}$$

⑦ Noting  $E(x+y) = \mu_x + \mu_y$ ,

$$\begin{aligned}\text{Var}(x+y) &= E\{(x+y - (\mu_x + \mu_y))^2\} \\ &= E\{((x - \mu_x) + (y - \mu_y))^2\} \\ &= E\{(x - \mu_x)^2 + 2(x - \mu_x)(y - \mu_y) + (y - \mu_y)^2\} \\ &= E\{(x - \mu_x)^2\} + 2E\{(x - \mu_x)(y - \mu_y)\} + E\{(y - \mu_y)^2\} \\ &= \text{Var}(x) + 2\text{Cov}(x, y) + \text{Var}(y)\end{aligned}$$



(8) Noting  $E(x+a) = \mu_x + a$  and  $E(y+b) = \mu_y + b$ ,

$$\begin{aligned} \text{Cov}(x+a, y+b) &= E\{(x+a - (\mu_x + a))(y+b - (\mu_y + b))\} \\ &= E\{(x+a - \mu_x - a)(y+b - \mu_y - b)\} \\ &= E\{(x - \mu_x)(y - \mu_y)\} = \text{Cov}(x, y) \end{aligned}$$

(9) (a)  $\text{Cov}(ax, by) = E\{(ax - a\mu_x)(by - b\mu_y)\}$

$$\begin{aligned} &= E\{a(x - \mu_x)b(y - \mu_y)\} \\ &= ab E\{(x - \mu_x)(y - \mu_y)\} = ab \text{Cov}(x, y) \end{aligned}$$

(b)  $\text{Corr}(ax, by) = \frac{\text{Cov}(ax, by)}{\sqrt{\text{Var}(ax) \text{Var}(by)}}$

$$= \frac{ab \text{Cov}(x, y)}{\sqrt{a^2 \text{Var}(x) b^2 \text{Var}(y)}}$$

$$= \frac{a}{|a|} \frac{b}{|b|} \frac{\text{Cov}(x, y)}{\sqrt{\text{Var}(x) \text{Var}(y)}}$$

$$= \text{sign}(a) \text{sign}(b) \text{Corr}(x, y),$$

provided  $a \neq 0$  and  $b \neq 0$

$$\begin{aligned}
 (10) \quad E(x_1 + x_2) &= \sum_{x_1} \sum_{x_2} (x_1 + x_2) p(x_1, x_2) \\
 &= \sum_{x_1} \sum_{x_2} (x_1 p(x_1, x_2) + x_2 p(x_1, x_2)) \\
 &= \sum_{x_1} \sum_{x_2} x_1 p(x_1, x_2) + \sum_{x_1} \sum_{x_2} x_2 p(x_1, x_2) \\
 &\quad \uparrow \qquad \qquad \qquad \uparrow \\
 &\quad g_1(x_1, x_2) \qquad \qquad g_2(x_1, x_2) \\
 &= E(x_1) + E(x_2)
 \end{aligned}$$

If  $x_1$  &  $x_2$  are continuous,

$$\begin{aligned}
 E(x_1 + x_2) &= \iint (x_1 + x_2) f(x_1, x_2) dx_1 dx_2 \\
 &= \iint x_1 f(x_1, x_2) dx_1 dx_2 \\
 &\quad + \iint x_2 f(x_1, x_2) dx_1 dx_2 \\
 &= E(x_1) + E(x_2)
 \end{aligned}$$

$$\textcircled{11} \text{ (a) } E\left(\sum_{i=1}^n y_i\right) = \sum_{i=1}^n E(y_i) = \sum_{i=1}^n \mu \\ = n\mu$$

$$\text{(b) } \text{Var}\left(\sum_{i=1}^n y_i\right) = \sum_{i=1}^n \text{Var}(y_i)$$

↑ Problems 7 and 5c

$$= n\sigma^2$$

$$\text{(c) } \text{Var}(\bar{y}) = \text{Var}\left(\frac{1}{n} \sum_{i=1}^n y_i\right) \\ = \frac{1}{n^2} \text{Var}\left(\sum_{i=1}^n y_i\right) = \frac{1}{n^2} n\sigma^2 = \frac{\sigma^2}{n}$$

$$\text{(d) } E(\bar{y}) = E\left(\frac{1}{n} \sum_{i=1}^n y_i\right) = \frac{1}{n} \sum_{i=1}^n E(y_i) \\ = \frac{1}{n} \cdot n\mu = \mu \quad \text{unbiased}$$

$$\text{(e) } \text{If } \sum_{i=1}^n a_i = 1, \quad E(L) = E\left(\sum_{i=1}^n a_i y_i\right) = \sum_{i=1}^n a_i E(y_i) \\ = \sum_{i=1}^n a_i \mu = \mu \sum_{i=1}^n a_i = \mu \cdot 1 = \mu \quad \text{unbiased}$$

$$\text{(f) } \text{Yes, } \bar{y} = \sum_{i=1}^n \left(\frac{1}{n}\right) y_i = \sum_{i=1}^n a_i y_i, \quad \text{with} \\ a_i = \frac{1}{n} \text{ for } i=1, \dots, n.$$

$$\text{(g) } \text{Var}(L) = \text{Var}\left(\sum_{i=1}^n a_i y_i\right) = \sum_{i=1}^n a_i^2 \text{Var}(y_i) \\ = \sum_{i=1}^n a_i^2 \sigma^2 = \sigma^2 \sum_{i=1}^n a_i^2$$

12

(a) No	(b) No	(c) Yes
(d) Yes	(e) No	(f) No
(g) Yes	(h) Yes	(i) No

13

$$A'B = \begin{pmatrix} 2 & 1 & 0 \\ 5 & -4 & 3 \end{pmatrix} \begin{pmatrix} 1 & 0 \\ 2 & 3 \\ -1 & 3 \end{pmatrix} = \begin{pmatrix} 2+2+0 & 0+3+0 \\ 5-8-3 & 0-12+9 \end{pmatrix}$$

$$(14) \quad c'd = 2 \cdot 1 + 1 \cdot 2 + 0 \cdot -1 = 4$$

$$cd' = \begin{pmatrix} 2 \\ 1 \\ 0 \end{pmatrix} (1 \ 2 \ -1) = \begin{pmatrix} 2 \cdot 1 & 2 \cdot 2 & 2 \cdot -1 \\ 1 \cdot 1 & 1 \cdot 2 & 1 \cdot -1 \\ 0 \cdot 1 & 0 \cdot 2 & 0 \cdot -1 \end{pmatrix}$$

$$(15) \quad AB = \left( \begin{array}{c|c} 4 & 4 \\ \hline 8 & 8 \end{array} \right) = AC$$

(a) Yes

(c) That was a joke.  $A^{-1}$  does not exist because the columns of  $A$  are linearly dependent.

16

$$\begin{pmatrix} 1 & 2 \\ 2 & 1 \end{pmatrix} \begin{pmatrix} 1 & 1 \\ 0 & 1 \end{pmatrix} = \begin{pmatrix} 1 & 3 \\ 2 & 3 \end{pmatrix}$$

but

$$\begin{pmatrix} 1 & 1 \\ 0 & 1 \end{pmatrix} \begin{pmatrix} 1 & 2 \\ 2 & 1 \end{pmatrix} = \begin{pmatrix} 3 & 3 \\ 2 & 1 \end{pmatrix}$$

17  $A^{-1}$  does not exist.

18  $(X'X)' = X'X'' = X'X$

19 The inverse is only defined for square matrices.