

Family (Last) Name _____

Given (First) Name Jerry

Student Number _____

STA 302s13 Quiz 5A

For reference, the general linear model in matrix form is $\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$, where \mathbf{X} is an $n \times (k+1)$ matrix of observable constants, the columns of \mathbf{X} are linearly independent, $\boldsymbol{\beta}$ is a $(k+1) \times 1$ vector of unknown constants (parameters), and $\boldsymbol{\epsilon}$ is an $n \times 1$ vector of unobservable random variables with $E(\boldsymbol{\epsilon}) = \mathbf{0}$ and $\text{cov}(\boldsymbol{\epsilon}) = \sigma^2 \mathbf{I}_n$, where $\sigma^2 > 0$ is an unknown constant parameter. The least squares estimator of $\boldsymbol{\beta}$ is $\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$.

1. (1 point) Suppose we want to estimate $\mathbf{a}'\boldsymbol{\beta}$ based on sample data. What estimator is the most natural choice?

$$\mathbf{a}'\hat{\boldsymbol{\beta}}$$

2. (2 points) Show that the estimator you have proposed is unbiased. Show the full calculation. Do not use $E(\hat{\boldsymbol{\beta}}) = \boldsymbol{\beta}$ directly. Show the full calculation.

$$\begin{aligned} E(\mathbf{a}'\hat{\boldsymbol{\beta}}) &= \mathbf{a}'E(\hat{\boldsymbol{\beta}}) = \mathbf{a}'E((\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}) \\ &= \mathbf{a}'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'E(\mathbf{Y}) = \mathbf{a}'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{X}\boldsymbol{\beta} = \mathbf{a}'\boldsymbol{\beta} \end{aligned}$$

3. (2 points) The natural estimator is a linear unbiased estimator of the form $\mathbf{c}_0'\mathbf{Y}$. What is the $n \times 1$ vector \mathbf{c}_0 ? Make sure your answer has the correct dimension. ~~You have more room than you need.~~

Circle your expression for \mathbf{c}_0 .

$$\mathbf{c}_0'\mathbf{Y} = \mathbf{a}'\hat{\boldsymbol{\beta}} = \mathbf{a}'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}, \text{ so}$$

$$\mathbf{c}_0' = \mathbf{a}'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}', \text{ and}$$

$$\mathbf{c}_0 = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{a}$$

4. (5 points) As you saw in homework and lecture, the *best* linear unbiased estimator is $\mathbf{c}'_0 \mathbf{Y}$. An important part of the proof is to show $(\mathbf{c} - \mathbf{c}_0)' \mathbf{c}_0 = 0$, using the constraint $\mathbf{X}'\mathbf{c} = \mathbf{a}$. Please carry out the calculation.

$$(\mathbf{c} - \mathbf{c}_0)' \hat{\mathbf{c}}_0 = \mathbf{c}' \hat{\mathbf{c}}_0 - \mathbf{c}_0' \hat{\mathbf{c}}_0$$

$$= \mathbf{c}' \mathbf{X} (\mathbf{X}' \mathbf{X})^{-1} \mathbf{a} - (\mathbf{X} (\mathbf{X}' \mathbf{X})^{-1} \mathbf{a})' \mathbf{X} (\mathbf{X}' \mathbf{X})^{-1} \mathbf{a}$$

using $\mathbf{X}'\hat{\mathbf{c}} = \mathbf{a} \Leftrightarrow \mathbf{c}'\mathbf{X} = \mathbf{a}'$,

$$= \mathbf{a}' \mathbf{X} (\mathbf{X}' \mathbf{X})^{-1} \mathbf{a} - \mathbf{a}' (\mathbf{X}' \mathbf{X})^{-1} \underbrace{\mathbf{X}' \mathbf{X} (\mathbf{X}' \mathbf{X})^{-1}}_{\mathbf{I}} \mathbf{a}$$

$$= 0$$

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STA 302s13 Quiz 4B

For reference, $\mathbf{Y} \sim N_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ means $M_{\mathbf{Y}}(\mathbf{t}) = e^{\mathbf{t}'\boldsymbol{\mu} + \frac{1}{2}\mathbf{t}'\boldsymbol{\Sigma}\mathbf{t}}$. You may use these facts without proof: $M_{\mathbf{A}\mathbf{Y}}(\mathbf{t}) = M_{\mathbf{Y}}(\mathbf{A}'\mathbf{t})$ and $M_{\mathbf{Y}+\mathbf{c}}(\mathbf{t}) = e^{\mathbf{t}'\mathbf{c}}M_{\mathbf{Y}}(\mathbf{t})$.

1. (4 points) Let $\mathbf{Z} = \boldsymbol{\Sigma}^{-1/2}(\mathbf{Y} - \boldsymbol{\mu})$. What is the distribution of \mathbf{Z} ? Show your work. Circle your final answer.

$$\begin{aligned} M_{\mathbf{Z}}(\mathbf{t}) &= M_{\mathbf{Y}-\boldsymbol{\mu}}(\boldsymbol{\Sigma}^{-1/2}\mathbf{t}) = M_{\mathbf{Y}-\boldsymbol{\mu}}(\boldsymbol{\Sigma}^{-1/2}\mathbf{t}) \\ &= e^{-\mathbf{t}'\boldsymbol{\mu}} M_{\mathbf{Y}}(\boldsymbol{\Sigma}^{-1/2}\mathbf{t}) \\ &= e^{-\mathbf{t}'\boldsymbol{\mu}} e^{\mathbf{t}'\boldsymbol{\mu} + \frac{1}{2}(\boldsymbol{\Sigma}^{-1/2}\mathbf{t})'\boldsymbol{\Sigma}(\boldsymbol{\Sigma}^{-1/2}\mathbf{t})} \\ &= e^{-\mathbf{t}'\boldsymbol{\mu}} e^{\mathbf{t}'\boldsymbol{\mu}} e^{\frac{1}{2}\mathbf{t}'\boldsymbol{\Sigma}^{-1/2}\boldsymbol{\Sigma}\boldsymbol{\Sigma}^{-1/2}\mathbf{t}} \\ &= e^{\frac{1}{2}\mathbf{t}'\underbrace{\boldsymbol{\Sigma}^{-1/2}\boldsymbol{\Sigma}^{1/2}}_{\mathbf{I}}\underbrace{\boldsymbol{\Sigma}^{1/2}\boldsymbol{\Sigma}^{-1/2}}_{\mathbf{I}}\mathbf{t}} = e^{\frac{1}{2}\mathbf{t}'\mathbf{t}} \end{aligned}$$

2. (4 points) What is the distribution of $\mathbf{Z}'\mathbf{Z}$? Give reasons for your answer. You may just cite well-known facts about functions of normal random variables.

So $\mathbf{Z} \sim N_p(\mathbf{0}, \mathbf{I})$

$\mathbf{Z}'\mathbf{Z} = \sum_{j=1}^p Z_j^2$, where $Z_j \sim N(0, 1)$,

So $Z_j^2 \sim \chi^2(1)$. The Z_j quantities are independent, so this is a sum of independent chi-squares, and $\mathbf{Z}'\mathbf{Z} \sim \chi^2(p)$