

Assignment 1

① For $x > 0$, $F_x(x) = \int_0^x 2t e^{-t^2} dt$

$$u = t^2 \\ du = 2t dt$$

$$\begin{array}{r} \times u \\ x \overline{) x^2} \\ \underline{0} \\ 0 \end{array}$$

$$= \int_0^{x^2} e^{-u} du = 1 - e^{-x^2}, \text{ so}$$

$$F_x(x) = (1 - e^{-x^2}) I(x > 0)$$

(b) $P(X > \frac{1}{2}) = e^{-\frac{1}{4}} = 0.7788$

②

<u>x</u>	<u>prob</u>	<u>x²</u>	<u>y = x² - 1</u>
-4	4/20	16	15
-3	3/20	9	8
-2	2/20	4	3
-1	1/20	1	0
0	0	0	-1
1	1/20	1	0
2	2/20	4	3
3	3/20	9	8
4	4/20	16	15

$$\begin{aligned}
 \textcircled{3} \quad f_Y(y) &= \frac{d}{dy} F_Y(y) = \frac{d}{dy} P(Y \leq y) = \frac{d}{dy} P(X^2 \leq y) \\
 &= \frac{d}{dy} P(|X| \leq y^{1/2}) = \frac{d}{dy} P(-y^{1/2} \leq X \leq y^{1/2}) \\
 &= \frac{d}{dy} (F_X(y^{1/2}) - F_X(-y^{1/2})) \\
 &= f_X(y^{1/2}) \cdot \frac{1}{2} y^{-1/2} - f_X(-y^{1/2}) \cdot (-1) \frac{1}{2} y^{-1/2} \\
 &= \frac{1}{2} I(-1 < y^{1/2} < 1) \cdot \frac{1}{2} y^{-1/2} + \frac{1}{2} I(-1 < -y^{1/2} < 1) \cdot \frac{1}{2} y^{-1/2} \\
 &= \frac{1}{4} y^{-1/2} \left[I(0 < y^{1/2} < 1) + I(-1 < -y^{1/2} < 0) \right] \\
 &= \frac{1}{4} y^{-1/2} \left[I(0 < y < 1) + I(1 > y^{1/2} > 0) \right] \\
 &= \frac{1}{4} y^{-1/2} \left[I(0 < y < 1) + I(0 < y < 1) \right] \\
 &= \frac{1}{2} y^{-1/2} I(0 < y < 1)
 \end{aligned}$$

$$\begin{aligned}
 (4) \quad f_z(z) &= \frac{d}{dz} F_z(z) = \frac{d}{dz} P(Z \leq z) = \frac{d}{dz} P\left(\frac{X-\mu}{\sigma} \leq z\right) \\
 &= \frac{d}{dz} P(X \leq \sigma z + \mu) = \frac{d}{dz} F_X(\sigma z + \mu) \\
 &= f_X(\sigma z + \mu) \cdot \sigma = \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2\sigma^2}(\sigma z + \mu - \mu)^2} \cdot \sigma \\
 &= \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}z^2}, \text{ pdf of } N(0, 1)
 \end{aligned}$$

$$\begin{aligned}
 (5) \quad M_Y(t) &= E(e^{Yt}) = E\left(e^{(a + \sum_{i=1}^n b_i X_i)t}\right) \\
 &= E\left(e^{at} e^{\sum_{i=1}^n b_i X_i t}\right) = e^{at} E\left(e^{\sum_{i=1}^n b_i X_i t}\right) \\
 &= e^{at} M_{\sum_{i=1}^n b_i X_i}(t) \stackrel{\text{ind}}{=} e^{at} \prod_{i=1}^n M_{b_i X_i}(t) \\
 &= e^{at} \prod_{i=1}^n M_{X_i}(b_i t) = e^{at} \prod_{i=1}^n e^{\mu b_i t + \frac{1}{2} \sigma^2 b_i^2 t^2} \\
 &= e^{(a + \mu \sum_{i=1}^n b_i)t + \frac{1}{2} (\sigma^2 \sum_{i=1}^n b_i^2)t^2}
 \end{aligned}$$

MGF of

$$N\left(a + \mu \sum_{i=1}^n b_i, \sigma^2 \sum_{i=1}^n b_i^2\right)$$

$$(6) f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} P(Y \leq y) = \frac{d}{dy} P(Z^2 \leq y)$$

(This is zero for $y < 0$, so let $y \geq 0$)

$$= \frac{d}{dy} P(|Z| \leq y^{1/2}) = \frac{d}{dy} P(-y^{1/2} \leq Z \leq y^{1/2})$$

$$= \frac{d}{dy} (F_Z(y^{1/2}) - F_Z(-y^{1/2}))$$

$$= f_Z(y^{1/2}) \cdot \frac{1}{2} y^{-1/2} - f_Z(-y^{1/2}) \cdot (-1) \cdot \frac{1}{2} y^{-1/2}$$

$$= \frac{1}{2} y^{-1/2} \left(\frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(y^{1/2})^2} + \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(-y^{1/2})^2} \right)$$

$$= \frac{1}{2} y^{-1/2} \left(\frac{2}{\sqrt{2\pi}} e^{-\frac{1}{2}y} \right), \text{ so}$$

$$f_Y(y) = \frac{\left(\frac{1}{2}\right)^{1/2}}{\Gamma\left(\frac{1}{2}\right)} e^{-\frac{1}{2}y} y^{\frac{1}{2}-1} \mathbb{I}(y \geq 0)$$

Gamma with $\alpha = \lambda = \frac{1}{2}$, which is $\chi^2(1)$.

$$\begin{aligned}
 \textcircled{7} \quad M_Y(t) &= M_{\sum Y_i}(t) \stackrel{\text{ind}}{=} \prod_{i=1}^n M_{Y_i}(t) \\
 &= \prod_{i=1}^n (1-2t)^{-Y_i} = (1-2t)^{-\sum_{i=1}^n Y_i}
 \end{aligned}$$

MGF of $\chi^2(\sum_{i=1}^n Y_i)$

$$\begin{aligned}
 \textcircled{8} \quad (a) \quad E(\bar{X}_n) &= E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} \sum_{i=1}^n E(X_i) \\
 &= \frac{1}{n} \sum_{i=1}^n \mu = \frac{n\mu}{n} = \mu
 \end{aligned}$$

$$\begin{aligned}
 (b) \quad \text{Var}(\bar{X}_n) &= \text{Var}\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n^2} \text{Var}\left(\sum_{i=1}^n X_i\right) \\
 &\stackrel{\text{ind}}{=} \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i) = \frac{1}{n^2} \sum_{i=1}^n \sigma^2 = \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n}
 \end{aligned}$$

9

	$x=1$	$x=2$	$x=3$	
$y=1$	$\frac{3}{12}$ (1)	$\frac{1}{12}$ (2)	$\frac{3}{12}$ (3)	$\frac{7}{12}$
$y=2$	$\frac{1}{12}$ (2)	$\frac{3}{12}$ (4)	$\frac{1}{12}$ (6)	$\frac{5}{12}$
	$\frac{4}{12}$	$\frac{4}{12}$	$\frac{4}{12}$	$\frac{12}{12}$

(a)

$$E(X) = 1 \cdot \frac{4}{12} + 2 \cdot \frac{4}{12} + 3 \cdot \frac{4}{12} = (1+2+3) \cdot \frac{1}{3} = \frac{6}{3} = 2$$

$$E(Y) = 1 \cdot \frac{7}{12} + 2 \cdot \frac{5}{12} = \frac{17}{12}$$

Put $Z = XY$ in the cells, in red

z	1	2	3	4	6
$P_z(z)$	$\frac{3}{12}$	$\frac{2}{12}$	$\frac{3}{12}$	$\frac{3}{12}$	$\frac{1}{12}$

$$E(Z) = \sum_{z} z P(z) = \frac{1}{12} (3+4+9+12+6) = \frac{34}{12}$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = \frac{34}{12} - 2 \cdot \frac{17}{12} = 0$$

(b) Not independent, because

$$P_{X,Y}(1,2) = \frac{1}{12} \neq P_X(1)P_Y(2) = \frac{4}{12} \cdot \frac{5}{12} = \frac{20}{144}$$

= $\frac{12}{144}$

$$\begin{aligned}
 \textcircled{10} \quad \text{Var}(X) &= E\left((X-\mu)^2\right) = E(X^2 - 2\mu X + \mu^2) \\
 &= E(X^2) - 2\mu E(X) + E(\mu^2) = E(X^2) - 2\mu^2 + \mu^2 \\
 &= E(X^2) - \mu^2 = E(X^2) - [E(X)]^2
 \end{aligned}$$

$$\begin{aligned}
 \textcircled{11} \quad \text{Cov}(X, Y) &= E\left((X-\mu_x)(Y-\mu_y)\right) \\
 &= E(XY - X\mu_y - \mu_x Y + \mu_x \mu_y) \\
 &= E(XY) - E(X)\mu_y - \mu_x E(Y) + E(\mu_x \mu_y) \\
 &= E(XY) - \mu_x \mu_y - \mu_x \mu_y + \mu_x \mu_y \\
 &= E(XY) - \mu_x \mu_y = E(XY) - E(X)E(Y)
 \end{aligned}$$

$$\begin{aligned}
 \textcircled{12} \quad \text{True: } E(XY) &= \sum_x \sum_y xy P_{XY}(x, y) \\
 &\stackrel{\text{ind}}{\downarrow} \sum_x \sum_y xy P_X(x) P_Y(y) \\
 &= \sum_x x P_X(x) \sum_y y P_Y(y) = E(X)E(Y)
 \end{aligned}$$

$$\text{And } \text{Cov}(X, Y) = E(XY) - E(X)E(Y) = 0$$

$\textcircled{13}$ False: Problem 9 is a counter-example

$$(14) \quad E(Y+Z) = E(Y) + E(Z) = \mu_y + \mu_z, \text{ so}$$

$$\begin{aligned} \text{Cov}(X, Y+Z) &= E\left((X-\mu_x)(Y+Z-(\mu_y+\mu_z))\right) \\ &= E\left[(X-\mu_x)(Y-\mu_y + Z-\mu_z)\right] \\ &= E\left[(X-\mu_x)(Y-\mu_y) + (X-\mu_x)(Z-\mu_z)\right] \\ &= E\left[(X-\mu_x)(Y-\mu_y)\right] + E\left[(X-\mu_x)(Z-\mu_z)\right] \\ &= \text{Cov}(X, Y) + \text{Cov}(X, Z) \end{aligned}$$

Could have used $\text{Cov}\left(\sum_{i=1}^n X_i, \sum_{j=1}^m Y_j\right) = \sum_{i=1}^n \sum_{j=1}^m \text{Cov}(X_i, Y_j)$

(15) Impossible because $\lim_{x \rightarrow -\infty} F(x) = 0$.

Or, if X is continuous then $f_X(x) = 0$ for all x and $\int_{-\infty}^{\infty} f_X(x) dx = 0 \neq 1$

If X is discrete, $F_X(x)$ has no jump points, so $P_X(x) = 0$ for all x and $\sum_x P_X(x) = 0 \neq 1$

$$(16) \quad F_X(x) = \int_{-\infty}^x f_X(t) dt = P(X \leq x)$$

$$\begin{aligned} (17) \quad F_{X|Y}(x|y) &= P(X \leq x | Y=y) = \frac{P(X \leq x, Y=y)}{P(Y=y)} \\ &\neq \frac{P(X \leq x, Y \leq y)}{P(Y \leq y)} = \frac{F_{XY}(x, y)}{F_Y(y)} \end{aligned}$$

(18) The statement is false. For example, let
 $P_X(1) = P_X(2) = P_X(3) = \frac{1}{3}$. $E(X) = 2$

EG $Y = \frac{1}{X}$, $P_Y(1) = P_Y(\frac{1}{2}) = P_Y(\frac{1}{3}) = \frac{1}{3}$, and

$$E(Y) = E\left(\frac{1}{X}\right) = \frac{1}{3} \left(\frac{6}{6} + \frac{3}{6} + \frac{2}{6} \right) = \frac{1}{3} \cdot \frac{11}{6}$$

$$= \frac{11}{18} \neq \frac{1}{2}$$

(19) Expected value is a linear operation. You can't just bring it inside the square.

This calculation says that every random variable has a variance of zero. All random variables are degenerate. In fact, there is no such thing as a random variable. All games of chance produce exactly the same result, every time.