

## Weibull Regression with R, Part Two\*

```

> rm(list=ls()); options(scipen=999)
> # install.packages("survival",dependencies=TRUE) # Only need to do this once
> library(survival) # Do this every time
> # install.packages("asaur",dependencies=TRUE) # Only need to do this once
> library(asaur)
> # help(pharmacoSmoking)
> head(pharmacoSmoking)
  id ttr relapse    grp age gender    race employment yearsSmoking
1  21 182      0  patchOnly 36  Male   white          ft           26
2 113  14      1  patchOnly 41  Male   white          other          27
3  39   5      1 combination 25 Female white          other          12
4  80  16      1 combination 54  Male   white          ft           39
5  87   0      1 combination 45  Male   white          other          30
6  29 182      0 combination 43  Male  hispanic        ft           30
  levelSmoking ageGroup2 ageGroup4 priorAttempts longestNoSmoke
1      heavy      21-49      35-49           0           0
2      heavy      21-49      35-49           3           90
3      heavy      21-49      21-34           3           21
4      heavy      50+       50-64           0           0
5      heavy      21-49      35-49           0           0
6      heavy      21-49      35-49           2          1825
> summary(pharmacoSmoking)
  id          ttr          relapse          grp
Min.   : 1.00   Min.   : 0.00   Min.   :0.000   combination:61
1st Qu.: 33.00  1st Qu.:  8.00   1st Qu.:0.000   patchOnly  :64
Median : 67.00  Median : 49.00   Median :1.000
Mean   : 66.15  Mean   : 77.44   Mean   :0.712
3rd Qu.: 99.00  3rd Qu.:182.00  3rd Qu.:1.000
Max.   :130.00  Max.   :182.00  Max.   :1.000

  age          gender          race          employment          yearsSmoking
Min.   :22.00   Female:81   black   :38   ft   :72   Min.   : 9.00
1st Qu.:41.00   Male  :44   hispanic: 8   other:39  1st Qu.:22.00
Median :49.00          other  : 2   pt   :14   Median :30.00
Mean   :48.84          white  :77          Mean   :30.88
3rd Qu.:56.00          Max.   :56.00
Max.   :86.00

  levelSmoking ageGroup2 ageGroup4 priorAttempts longestNoSmoke
heavy:89      21-49:66      21-34:16   Min.   : 0.00   Min.   : 0.0
light:36      50+ :59      35-49:50   1st Qu.: 1.00   1st Qu.: 7.0
          50-64:48   Median : 2.00   Median : 90.0
          65+ :11   Mean   : 12.68  Mean   : 539.7
          3rd Qu.: 5.00  3rd Qu.: 365.0
          Max.   :1000.00 Max.   :6205.0
> # Make fixed-up data frame called quit
> quit = within(pharmacoSmoking,{
+ DayOfRelapse = Surv(ttr+1,relapse)
+ contrasts(grp) = contr.treatment(2,base=2) # Patch only is reference category
+ colnames(contrasts(grp)) = c('Combo') # Names of dummy vars -- just one
+ # Collapse race categories
+ Race = as.character(race) # Small r race is a factor. This is easier to modify.
+ Race[Race!='white'] = 'blackOther'; Race=factor(Race)
+ }) # Finished making data frame quit
> with(quit, table(race,Race) )
      Race
race   blackOther white
black          38     0
hispanic         8     0
other            2     0
white            0    77

```

\* Copyright information is on the last page.

```
> full = survreg(DayOfRelapse ~ grp + age + gender + Race + employment
+               + yearsSmoking + levelSmoking + priorAttempts, dist='weibull', data=quit)
> summary(full)
```

Call:

```
survreg(formula = DayOfRelapse ~ grp + age + gender + Race +
employment + yearsSmoking + levelSmoking + priorAttempts,
data = quit, dist = "weibull")
```

	Value	Std. Error	z	p
(Intercept)	1.12177	0.97726	1.15	0.2510
grpCombo	1.09225	0.38186	2.86	0.0042
age	0.08432	0.03411	2.47	0.0134
genderMale	0.03631	0.41417	0.09	0.9301
Racewhite	0.25145	0.39143	0.64	0.5206
employmentother	-1.28799	0.46719	-2.76	0.0058
employmentpt	-1.28482	0.58631	-2.19	0.0284
yearsSmoking	-0.02351	0.03250	-0.72	0.4696
levelSmokinglight	-0.07347	0.43151	-0.17	0.8648
priorAttempts	-0.00105	0.00200	-0.52	0.6000
Log(scale)	0.54194	0.08917	6.08	0.0000000012

Scale= 1.72

Weibull distribution

Loglik(model)= -463.8    Loglik(intercept only)= -476.5

Chisq= 25.41 on 9 degrees of freedom, p= 0.0025

Number of Newton-Raphson Iterations: 5

n= 125

>

I am thinking about dropping Race, yearsSmoking, levelSmoking and priorAttempts. The last 3 variables all represent smoking history and could be correlated highly enough to wash out each other's effects. Test them simultaneously.

```
>
> # Fit the restricted model: Restricted by H0
> rest1 = survreg(DayOfRelapse ~ grp + age + gender + Race + employment ,
+               dist='weibull', data=quit)
> anova(rest1,full) # LR test
```

Terms

1 grp + age + gender + Race + employment

2 grp + age + gender + Race + employment + yearsSmoking + levelSmoking +

priorAttempts

	Resid.	Df	-2*LL	Test	Df	Deviance	Pr(>Chi)
1	117	928.3771		NA	NA	NA	NA
2	114	927.5513	=	3	0.8258271	0.8432801	

```
> # Is Race significant with those variables dropped?
```

```
> # Is Race significant with those variables dropped?
> summary(rest1)
```

```
Call:
survreg(formula = DayOfRelapse ~ grp + age + gender + Race +
  employment, data = quit, dist = "weibull")

```

	Value	Std. Error	z	p
(Intercept)	1.3905	0.8684	1.60	0.1093
grpCombo	1.1021	0.3794	2.91	0.0037
age	0.0637	0.0190	3.35	0.0008
genderMale	0.0561	0.4140	0.14	0.8921
Racewhite	0.1880	0.3788	0.50	0.6196
employmentother	-1.2821	0.4635	-2.77	0.0057
employmentpt	-1.2251	0.5837	-2.10	0.0358
Log(scale)	0.5444	0.0894	6.09	0.0000000011

Scale= 1.72

```
Weibull distribution
Loglik(model)= -464.2   Loglik(intercept only)= -476.5
  Chisq= 24.58 on 6 degrees of freedom, p= 0.00041
Number of Newton-Raphson Iterations: 5
n= 125
```

Decision: Drop race and gender.

```
> full2 = survreg(DayOfRelapse ~ grp + age + employment , dist='weibull',
data=quit)
> summary(full2)
```

```
Call:
survreg(formula = DayOfRelapse ~ grp + age + employment, data = quit,
  dist = "weibull")

```

	Value	Std. Error	z	p
(Intercept)	1.4957	0.8414	1.78	0.07545
grpCombo	1.1023	0.3793	2.91	0.00366
age	0.0643	0.0186	3.45	0.00055
employmentother	-1.2880	0.4617	-2.79	0.00528
employmentpt	-1.2123	0.5616	-2.16	0.03088
Log(scale)	0.5454	0.0894	6.10	0.000000001

Scale= 1.73

```
Weibull distribution
Loglik(model)= -464.3   Loglik(intercept only)= -476.5
  Chisq= 24.31 on 4 degrees of freedom, p= 0.000069
Number of Newton-Raphson Iterations: 5
n= 125
```

```
> # Test employment status controlling for age and experimental treatment.
> rest2 = survreg(DayOfRelapse ~ grp + age , dist='weibull', data=quit)
> anova(rest2,full2) # LR test
```

	Terms	Resid. Df	-2*LL	Test	Df	Deviance	Pr(>Chi)
1	grp + age	121	937.9007		NA	NA	NA
2	grp + age + employment	119	928.6554	=	2	9.245333	0.009826558

```

> # Test employment status with a Wald test.
> source("http://www.utstat.toronto.edu/~brunner/Rfunctions/Wtest.txt")
> # function(L,Tn,Vn,h=0) # H0: L theta = h
> # Tn is estimated theta, usually a vector.
> # Vn is the estimated asymptotic covariance matrix of Tn.
> # For Wald tests based on numerical MLEs, Tn = theta-hat,
> # and Vn is the inverse of the Hessian of the minus log likelihood.
>
> Vhat = vcov(full2); Vhat
      (Intercept)      grpCombo      age employmentother
(Intercept)  0.7079360800 -0.0320256900 -0.0147694486  0.111673731
grpCombo    -0.0320256900  0.1438698739 -0.0004703383 -0.013521493
age         -0.0147694486 -0.0004703383  0.0003472409 -0.003893727
employmentother 0.1116737308 -0.0135214927 -0.0038937268  0.213191081
employmentpt -0.0003554818 -0.0078279548 -0.0013434899  0.077138486
Log(scale)  -0.0098224903  0.0050290739  0.0002048412 -0.003182291
      employmentpt      Log(scale)
(Intercept) -0.0003554818 -0.0098224903
grpCombo    -0.0078279548  0.0050290739
age         -0.0013434899  0.0002048412
employmentother 0.0771384860 -0.0031822913
employmentpt  0.3153999894 -0.0035442716
Log(scale)  -0.0035442716  0.0079888732

> thetahat = full2$coefficients; thetahat
      (Intercept)      grpCombo      age employmentother
      1.4957374      1.1023048      0.0643414      -1.2880472
      employmentpt
      -1.2122529
>

```

Note that the asymptotic covariance matrix includes  $\log(\sigma)$ , but the "coefficients" vector does not.

```

> sigmahat = full2$scale; sigmahat
[1] 1.725305
> thetahat = c(thetahat,log(sigmahat))
>
> # H0: beta3=beta4=0. Express as H0: L theta = h
> eMat = rbind( c(0,0,0,1,0,0),
+             c(0,0,0,0,1,0) )
>
> Wtest(L=eMat, Tn=thetahat, Vn=Vhat)
      W      df      p-value
9.718885315 2.000000000 0.007754805
> anova(rest2,full2) # Repeating LR test for comparison
      Terms Resid. Df      -2*LL Test Df Deviance      Pr(>Chi)
1 grp + age      121 937.9007      NA      NA      NA
2 grp + age + employment 119 928.6554      = 2 9.245333 0.009826558
>
> # Test part time versus other
> pto = cbind(0,0,0,1,-1,0); pto
      [,1] [,2] [,3] [,4] [,5] [,6]
[1,] 0 0 0 1 -1 0
> Wtest(L=pto, Tn=thetahat, Vn=Vhat)
      W      df      p-value
0.01534747 1.00000000 0.90140640
>

```

Predict the day of relapse for a 50 year old patient who is employed full time and gets the patch-only treatment.

Weibull Regression:  $t_i = \exp\{\beta_0 + \beta_1 x_{i,1} + \dots + \beta_{p-1} x_{i,p-1}\} \cdot \epsilon_i^\sigma = e^{\mathbf{x}_i^\top \boldsymbol{\beta}} \epsilon_i^\sigma$ , where  $\epsilon_1 \sim \exp(1)$ .

- $t_i \sim \text{Weibull}$ , with  $\alpha = 1/\sigma$  and  $\lambda = e^{-\mathbf{x}_i^\top \boldsymbol{\beta}}$ .
- $E(t_i) = e^{\mathbf{x}_i^\top \boldsymbol{\beta}} \Gamma(\sigma+1)$ ,  $\text{Median}(t_i) = e^{\mathbf{x}_i^\top \boldsymbol{\beta}} \log(2)^\sigma$ ,  $h_i(t) = \frac{1}{\sigma} \exp\{-\frac{1}{\sigma} \mathbf{x}_i^\top \boldsymbol{\beta}\} t^{\frac{1}{\sigma}-1}$ .
- $S(t) = \exp\left\{-e^{-\frac{1}{\sigma} \mathbf{x}_i^\top \boldsymbol{\beta}} t^{\frac{1}{\sigma}}\right\}$

```
> thetahat
  (Intercept)      grpCombo      age employmentother
    1.4957374      1.1023048      0.0643414      -1.2880472
employmentpt
   -1.2122529      0.5454037

> x = c(1,0,50,0,0,0)
> xb = sum(x*tethahat)
>
> # a) The estimated mean
> exp(xb) * gamma(sigmahat+1)
[1] 175.5516
>
> # b) The estimated mean
> exp(xb) * log(2)^sigmahat
[1] 59.17273
>
> # I think the median is preferable to mean because the Weibull distribution
> # is skewed. Also, the predict function for Weibull regression works as expected
> # for medians (but not means).
>
> oldguy = data.frame(grp='patchOnly',age=50,employment='ft')
> predict(full2,newdata=oldguy,type='quantile',p=0.5,se=TRUE)
$fit
      1
59.17273

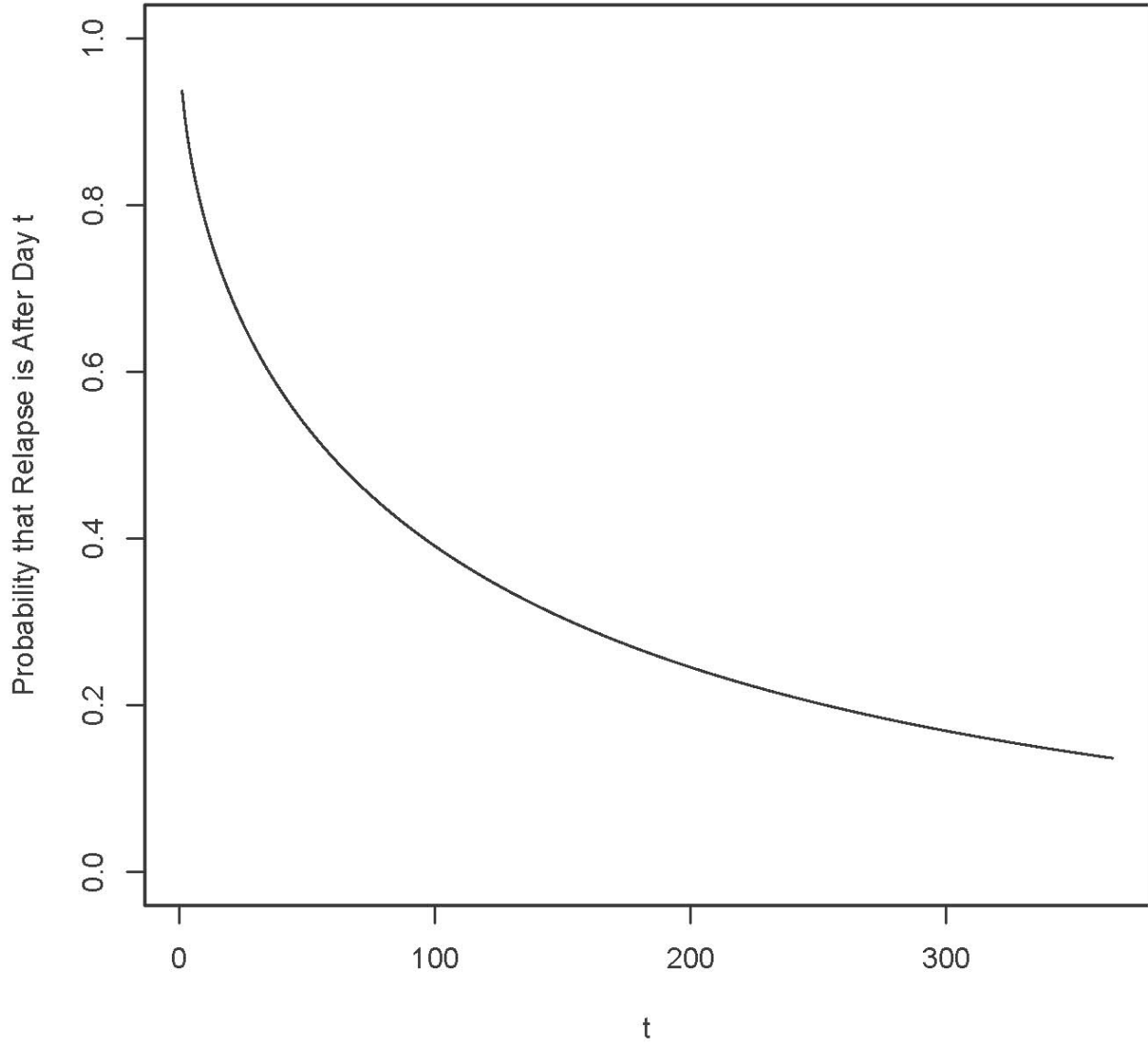
$se.fit
      1
18.87577

> # The 0.5 quantile is the median. se is from the delta method.
>
> # Estimate and plot S(t) for the old guy.
```

The se of  $S\text{-hat}(t)$  is straightforward in theory, but messy in practice. G-dot from Mathematica is very ugly. For example, in Wolfram Alpha, try  $D[\exp(-t^{1/s}) \exp(-(b_0 + 50 b_2)/s), b_0]$

```
> t = 1:365
> Shat = exp(-(exp(-xb/sigmahat)*t^(1/sigmahat)))
>
> plot(t,Shat,type='l',ylim=c(0,1),xlab='t',ylab='Probability that Relapse is After
Day t')
> tstring = expression(paste(hat(S)(t), " = Probability Relapsing After Day t"))
> title(tstring)
```

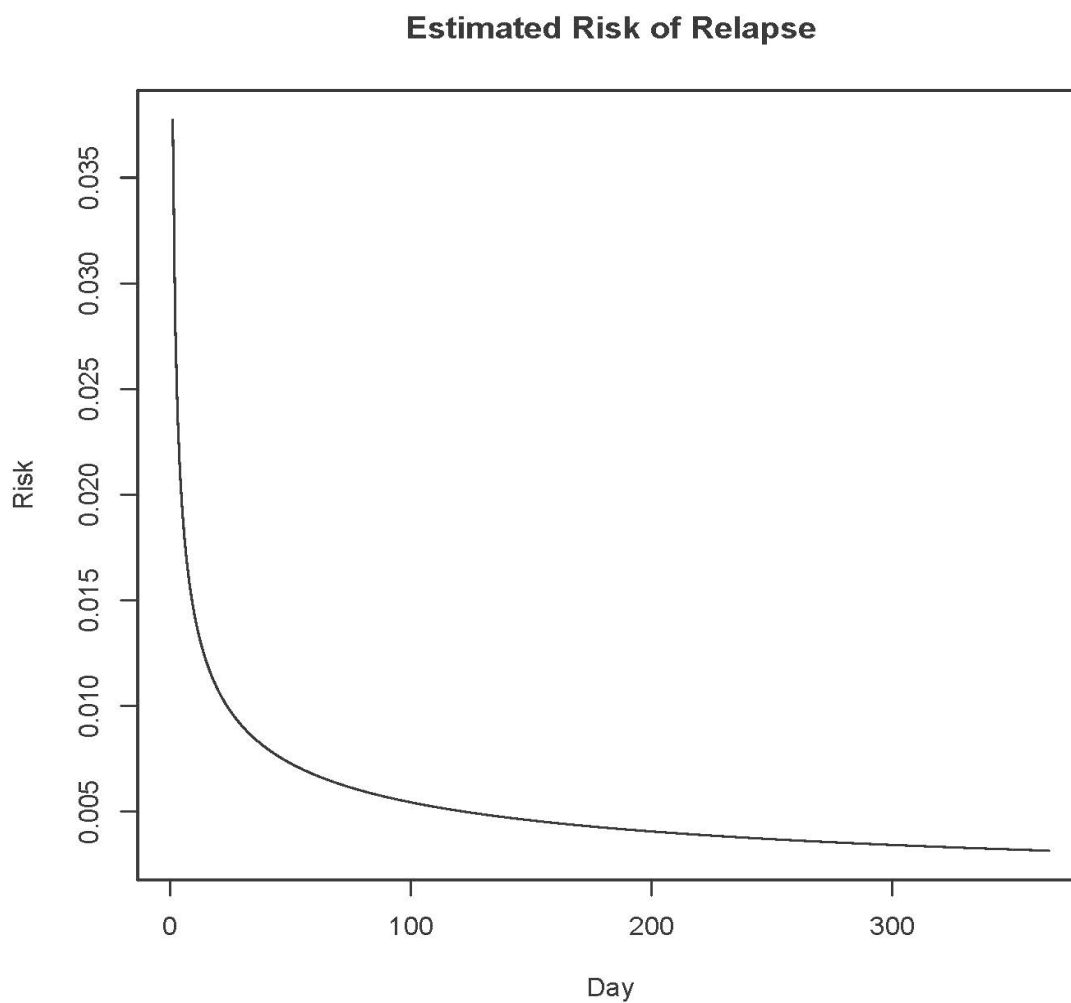
$$\hat{S}(t) = \text{Probability Relapsing After Day } t$$



Plot estimated hazard function for that 50 year old patient who is employed full time and gets the patch-only treatment.

$$\begin{aligned}h(t) &= \frac{f(t)}{S(t)} \\&= \frac{\alpha\lambda(\lambda t)^{\alpha-1} e^{-(\lambda t)^\alpha}}{e^{-(\lambda t)^\alpha}} \\&= \alpha\lambda^\alpha t^{\alpha-1} \\&= \frac{1}{\sigma} e^{-\frac{1}{\sigma} \mathbf{x}^\top \boldsymbol{\beta}} t^{\frac{1}{\sigma}-1}\end{aligned}$$

```
> h = 1/sigmahat * exp(-xb/sigmahat) * t^(1/sigmahat - 1)
> plot(t,h,type='l',xlab='Day',ylab='Risk',main='Estimated Risk of Relapse')
```



## LaTeX code for the record

```

\noindent
Weibull Regression:  $t_i = \exp\{\beta_0 + \beta_1 x_{i,1} + \dots + \beta_{p-1} x_{i,p-1}\} \cdot \epsilon_i^\sigma = e^{\mathbf{x}_i \boldsymbol{\beta}}$ 
 $\boldsymbol{\beta} \epsilon_i^\sigma$ ,
where  $\epsilon_1 \sim \exp(1)$ .
\begin{itemize}
\item  $t_i \sim$  Weibull, with  $\alpha = 1/\sigma$  and  $\lambda = e^{-\mathbf{x}_i \boldsymbol{\beta}}$ .
\item  $E(t_i) = e^{\mathbf{x}_i \boldsymbol{\beta}} \Gamma(\sigma + 1)$ ,
Median( $t_i$ ) =  $e^{\mathbf{x}_i \boldsymbol{\beta}} \log(2)^\sigma$ ,
 $h_i(t) = \frac{1}{\sigma} \exp\{-\frac{1}{\sigma} \mathbf{x}_i \boldsymbol{\beta} t^{\frac{1}{\sigma}-1}\}$ .
\item  $S(t) = e^{-\left( e^{-\frac{1}{\sigma} \mathbf{x}_i \boldsymbol{\beta} t^{\frac{1}{\sigma}}} \right)}$ 
\item  $S(t) = \exp\left\{ - e^{-\frac{1}{\sigma} \mathbf{x}_i \boldsymbol{\beta} t^{\frac{1}{\sigma}}} \right\}$ 
\end{itemize}

% Hazard calculation
\begin{eqnarray*}
h(t) & = & \frac{f(t)}{S(t)} \\
& = & \frac{\alpha \lambda (\lambda t)^{\alpha-1} e^{-(\lambda t)^\alpha}}{e^{-(\lambda t)^\alpha}} \\
& = & \alpha \lambda^\alpha t^{\alpha-1} \\
& = & \frac{1}{\sigma} e^{-\frac{1}{\sigma} \mathbf{x}_i \boldsymbol{\beta} t^{\frac{1}{\sigma}-1}}
\end{eqnarray*}

```

This document was prepared by [Jerry Brunner](#), University of Toronto. It is licensed under a Creative Commons Attribution - ShareAlike 3.0 Unported License:

[http://creativecommons.org/licenses/by-sa/3.0/deed.en\\_US](http://creativecommons.org/licenses/by-sa/3.0/deed.en_US). Use any part of it as you like and share the result freely. It is available in OpenOffice.org format from the course website:

<http://www.utstat.toronto.edu/~brunner/oldclass/312f23>