# STA 302 Quiz 2

Name	
Student No.	

- 1. For the MINITAB output below, n=10,  $H_0$  is  $\beta_1 \ge 0$  and  $H_a$  is  $\beta_1 < 0$ .
  - a) What is the critical value of the test statistic at  $\alpha = .01$ ?
  - b) What is the value of the test statistic t\*?
  - c) Do you reject  $H_0$ ? Yes or no.
- d) With this null and alternative hypothesis, what can you conclude about the presence of a linear relationship between X and Y?

Predictor	Coef	Stdev	t-ratio	P
Constant	6.6567	0.6445	10.33	0.000
C1	2.9740	0.1039	28.63	0.000
s = 0.9434	R-sq = :	99.0 <b>%</b>	R-sq(adj) =	98.9 <b>x</b>

- 2) For the MINITAB output above,
  - a) Construct a 99% confidence interval for  $\beta_1$ .
  - b) Are your results consistent with what you obtained in question 1? Explain.

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# Midterm Exam STA 302f 1990 Erindale College Aids Allowed: Calculator, Text book

- 1. Let  $Y_i = \beta_0 + X_i + \epsilon_i$ , i=1, ..., n where the  $X_i$  are known constants,  $\beta_0$  is an unknown constant, and the  $\epsilon_i$  are independent random variables with expected value zero and common variance  $\sigma^2$ .
  - a) (5 points) What is  $E(Y_i)$ ?
- b) (20 points) Find the least-squares estimate of  $\beta_0$ ; that is, find the value of  $\beta_0$  such that  $Q(\beta_0) = \sum\limits_{i=1}^n (Y_i E(Y_i))^2$  is minimized. Show that it is a minimum. Confine your answer to the space below.

For problems 2 through 18, write T for true or F for false. These problems are worth 3 points each.

- 2. \_\_\_\_\_ The least-squares line is chosen so as to minimize the sum of squared vertical distances of the points (on a scatterplot) from the line.
- 3. \_\_\_\_\_ If your data do not include any observations with  $X_i \le 0$ , it can be very misleading to interpret  $b_0$ .
- 4. \_\_\_\_\_ For the least-squares method to be valid, the error terms  $\epsilon_i$  must be distributed normally with mean zero and common variance  $\sigma^2$ .
- 5. \_\_\_\_\_ <u>SSE</u> represents the proportion of variation in the SSTO dependent variable that is explained by the independent variable.

- 6.  $\sum_{i=1}^{n} (Y_i \hat{Y}_i)^2 = 0$
- 7. \_\_\_\_\_ For the simple linear regression model (2.1) on p. 31,  $E(Y_i) = \beta_0 + \beta_1 X_i + \epsilon_i$ .
- 8. \_\_\_\_\_ Model (3.1) on p. 62 implies that if there is any relationship at all between X and Y, it must be linear.
- 9. \_\_\_\_\_ The confidence intervals and significance tests of chapter 3 are usually valid when n is large, even when the assumption of normality for the error terms is violated.
- 10. \_\_\_\_\_ We reject  $H_0$  at significance level  $\propto$  if and only if  $p>\propto$ .
- 11. \_\_\_\_\_\_ The confidence intervals and significance tests of chapter 3 are still valid when the  $X_i$  are random variables, provided that they are independent of  $\epsilon_i$ , their distribution does not involve  $\beta_0$ ,  $\beta_1$  or  $\sigma^2$ , and all probability statements are viewed as being conditional on the particular observed values of the  $X_i$ .
- 12. \_\_\_\_\_  $E(Y_i)=b_0+b_1X_i$ , where  $b_0$  and  $b_1$  are the least-squares estimates of  $\beta_0$  and  $\beta_1$  respectively.
- 13. \_\_\_\_\_ For simple linear regression, there is both a t-test and an F-test for  $H_0$ :  $\beta_1$ =0 versus  $H_a$ :  $\beta_1$ ≠0.
- 14. \_\_\_\_\_ For simple linear regression, if  $b_1=0$ , then  $b_0=\overline{Y}$ .
- 15. \_\_\_\_\_ Consider  $H_0$ :  $\beta_1 \ge 0$  versus  $H_a$ :  $\beta_1 < 0$ . Even if  $H_0$  is true, there could still be a negative correlation (in the population) between X and Y.
- 16. \_\_\_\_\_ Suppose you timed 1000 university students in the 100 meter dash, then waited a week and timed them again. Further, suppose that the mean and standard deviation of their times did not change from the first test to the second. Still, you would expect the very fastest students to run somewhat faster the second time.

- 17. \_\_\_\_\_ The model (3.1) on p. 64 implies that the parameter  $\sigma^2$  is normally distributed.
- 18. \_\_\_\_\_ We reject the null hypothesis that  $\beta_1$  equals a particular value (using a 2-tailed test) at significance level  $\propto$  if and only if the  $(1-\alpha)100\%$  confidence interval for  $\beta_1$  contains that particular value.
- 19. (10 points) For the ordinary simple linear regression model (2.1) on p. 31, show that  $\sum_{i=1}^{n} \widehat{Y}_i = \sum_{i=1}^{n} Y_i$ . This is a short proof.

Confine your answer to the space below.

Problems 20 through 24 refer to the following Minitab output.

These questions are worth 2 points each. Write your answers on the lines.

Worksheet retrieved from file: prob2.26 MTB > regress c2 1 c1

The regression equation is C2 = 183 + 0.262 C1

Predictor	Coef	Stdev	t-ratio	Р
Constant	182.97	12.72	14.38	0.000
C1	0.2616	0.1783	1.47	0.164

$$s = 10.29$$
 R-sq = 13.3% R-sq(adj) = 7.1%

Analysis of Variance

SOURCE	DF	SS	<b>M</b> S	F	Р
Regression	1	228.0	228.0	2.15	0.164
Error	14	1483.0	105.9		
Total	15	1711.0			

- 20. Is there a statistically significant relationship between X and Y at  $\propto$ =.20, two tailed?
- 21. What is the estimated standard deviation of  $b_0$ ? \_\_\_\_\_

22. Can you reject  $H_0$ :  $\beta_0$ =0 at  $\alpha$ =.01? \_\_\_\_\_\_

23. What is 
$$\sum_{i=1}^{n} (Y_i - \hat{Y}_i)^2$$
 ? \_\_\_\_\_\_

24. What proportion of the variation in the dependent variable is explained by the independent variable?

Questions 25 through 30 refer to question 2.26 (Robbery rate) on p. 58-59. Assume model (3.1) on p. 64. Some additional results are:

$$n = 16$$
  $\overline{X} = 69.875$   
 $b_0 = 182.97$   $b_1 = .2616$   
 $SSE = 1483.0$   $MSE = 105.9$   
 $\sum_{i=1}^{n} (X_i - \overline{X})^2 = 3331.75$   $\sum_{i=1}^{n} (Y_i - \overline{Y})^2 = 1711.0$ 

$$\sum_{i=1}^{n} (X_i - \overline{X})(Y_i - \overline{Y}) = 871.5 \qquad \sum_{i=1}^{n} X_i^2 = 81452.0$$

25. (3 points) The value of the correlation coefficient r is

- a) 0.187
- b) 0.365
- c) 0.262
- d) 0.133
- e) -0.133

- 26. (3 points) What is the predicted increase in the robbery rate for a population density increase of one person per unit area?
  - a) 182.97
  - b) 183.2316
  - c) .2616
  - d) 69.875
  - e) .0684
- 27. (3 points) A 95% confidence interval for  $\beta_1$  is given by
  - a) (.2253, 1.687)
  - b) (.1484, 3748)
  - c) (-.1208, .6440)
  - d) (.8977, 1.0564)
  - e) (-.1129, .8071)
- 28. (3 points)Is the t-test for  $H_0$ :  $\beta_1$ =0 statistically significant at  $\alpha$ =.05, two-tailed? Do you reject  $H_0$ ?
  - a) Yes, Yes
  - b) Yes, No
  - c) No, Yes
  - d) No, No
  - e) Sample size is too small for uniform convergence.
- 29. (3 points) A new city with a population density of 50 is to be observed. Give a 95% confidence interval for its predicted robbery rate.
  - a) (184.87, 207.23)
  - b) (189.29, 202.81)
  - c) (186.66, 205.44)
  - d) (185.702,206.398)
  - e) (.2253, 1.687)

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30. (4 points) Explain why  $b_0$  is meaningless here. Use the words "city" and "robbery" in your answer or you will receive no credit.

Erindale College - University of Toronto
Faculty of Arts and Science
December Examinations 1990
STA 302F

Duration - 3 hours

Aids allowed: Calculator, Textbook

Name (Please print	:)
Signature	
Student Number	

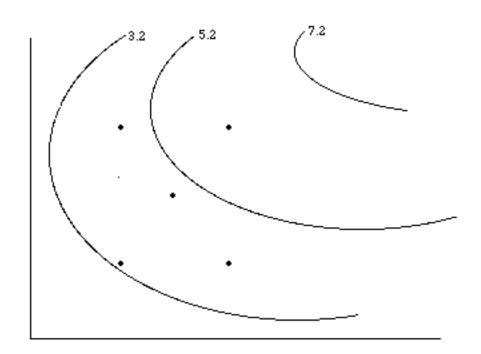
- 1. (5 pts) For model 7.18 on p. 237, derive the variance-covariance matrix of the vector  $\widehat{\mathbf{Y}}$ . Use only equations 6.45 through 6.47 on p. 203, the usual rules of matrix algebra and  $\mathbf{b} = (\mathbf{X'X})^{-1}\mathbf{X'Y}$ .
- 2. (7 pts) For model 7.18 on p. 237, show that  $\mathbf{X'}(\mathbf{Y} \widehat{\mathbf{Y}}) = \mathbf{0}_{p \times 1}$ . Use only equations 6.45 through 6.47 on p. 203, the usual rules of matrix algebra and  $\mathbf{b} = (\mathbf{X'X})^{-1}\mathbf{X'Y}$ .

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3. (3 pts) Refer to model 7.18 on p. 237. In addition,  $Var(\epsilon_i) = \frac{\sigma^2}{w_i}$ . Using only equations 6.45 through 6.47 on p. 203, the usual rules of matrix algebra and equation (11.61) on p. 419, show that for weighted least squares, **b** is unbiased for **\beta**.

4. (10 pts) Suppose we have a three-category independent variable and a single quantitative independent variable. The model is  $Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i3} + \beta_{12} X_{i1} X_{i2} + \beta_{13} X_{i1} X_{i3} + \epsilon_i, \text{ where } X_{i2} \text{ and } X_{i3} \text{ are indicator variables for the first two categories of the qualitative independent variable. Show that the test for equality of the three slopes is the same as testing <math>H_0$ :  $\beta_{12} = \beta_{13} = 0$ .

5. (5 pts) In a response surface drug study, the contour plot of the (predicted) response surface looks something like the picture below. The five dots • are the points at which sample data were collected. Your boss shows you the contour plot and asks what should be done next. How do you respond?



- 6. (7 pts) Consider the SMSA data set described on p. 1161-1162, using only variables 10, 11 and 12. It is claimed that even if we control for income, crime rate still varies by geographic region. Give the matrices  $\mathbf{C}$ ,  $\mathbf{\beta}$  and  $\mathbf{h}$  for  $\mathbf{H}_0$ :  $\mathbf{C}\mathbf{\beta} = \mathbf{h}$ . Assume that the rate at which crime rate changes as a function of income is the same in each geographic region.
- 7. (7 points) Refer again to the SMSA data set, this time restricting your attention to variables 4,6,7,10 and 11, and assuming that the independent variables do not interact. You want to simultaneously determine whether (a) controlling for all other variables, number of doctors and number of hospital beds are related to crime rate, and (b) the regression coefficient for percent of population in cities is equal to one. Give the matrices  $\mathbf{C}$ ,  $\mathbf{\beta}$  and  $\mathbf{h}$  for  $\mathbf{H}_0$ :  $\mathbf{C}\mathbf{\beta} = \mathbf{h}$ .

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Questions 8 through 11 refer to the following MINITAB output.

```
SUBC> dw.
The regression equation is
C4 = 1.63 + 0.380 C2 - 0.0231 C3 + 0.00042 C7 - 0.00362 C10 + 0.00506 C11
Predictor
                Coef
                           Stdev
                                    t-ratio
                                               0.212
Constant
               1.631
                           1.298
                                       1.26
C2
                                       5.58
                                               0.000
             0.37975
                         0.06808
С3
            -0.02308
                         0.02408
                                      -0.96
                                               0.340
C7
            0.000419
                        0.003034
                                       0.14
                                               0.890
C10
           -0.003622
                        0.003890
                                      -0.93
                                               0.354
            0.005057
                        0.001893
                                       2.67
C11
                                               0.009
s = 1.098
                R-sq = 35.9%
                                 R-sq(adj) = 32.9%
Analysis of Variance
SOURCE
             DF
                         SS
                                     MS
                                                F
             5
                     72.275
                                 14.455
                                            11.98
                                                     0.000
Regression
Error
            107
                    129.105
                                  1.207
                    201.380
Total
            112
```

Durbin-Watson statistic = 2.00

DF

1

1

1

1

1

SEQ SS

57.305

2.075

4.018

0.263

8.614

SOURCE

C2

C3

C7

C10

C11

MTB > regress c4 5 c2 c3 c7 c10 c11;

Continued on page 8

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## Questions 8-11 refer to the MINITAB output on page 7.

8. (	(8 pts)	) F	ill	in	the	blanks.
O. 1	$\mathcal{O}$ $\mathcal{P}$ $\mathcal{O}$	/ 1	1111	111	CLIC	DIGITING.

- 9. (2 pts) Once  $X_1$  has been taken into account, what proportion of the <u>remaining</u> variation in Y is explained by  $X_2$ ,  $X_3$ ,  $X_4$  and  $X_5$  together?
- 10. (1 pt) Is there evidence of autocorrelation in the error terms? (Yes or No)
- 11. (5 pts) Give a 95% Bonferroni joint confidence interval for  $\beta_2$  and  $\beta_3$ . Use the closest df in the table as an approximation.
- 12. (5 pts) Consider the model for which  $E(Y_i) = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_{12} X_{i1} \ X_{i2}. \ \ Demonstrate \ that \ \beta_{12} \neq 0 \ \ means$  that the relationship of Y to  $X_1$  depends upon the value of  $X_2$ .

Questions 13 through 55 are worth one point each. Answer each one T for true or F for false. Scores on this section will be corrected for guessing, so guess only if you think your chances of guessing right are better than 50-50.

- 13. \_\_\_\_It is impossible for a collection of error terms to be both independent and autocorrelated.
- 14. \_\_\_\_In a situation where we are comparing regression models with differing numbers of independent variables, it makes sense to use the adjusted  $R^2$  (adjusted for degrees of freedom) rather than the usual  $R^2$  = SSR/SSTO.
- 15. \_\_\_\_Suppose that  $\hat{Y} = 2.2 4X_1 + 11.2X_2 + 3.1X_1X_2$ . For a change of one unit in  $X_2$ , the predicted change in Y is 11.2 units.
- 16. \_\_\_\_ Weighted least squares is used in situations where the variance of the Y's does not appear to be equal for all levels of X.
- 17. \_\_\_\_ When there is multicolinearity, it is especially important to use two-tailed rather than one-tailed t-tests.
- 18. \_\_\_\_In stepwise regression, it is possible that an X variable brought in at an early stage will be subsequently dropped if it is no

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longer helpful in conjunction with X variables added at later stages.

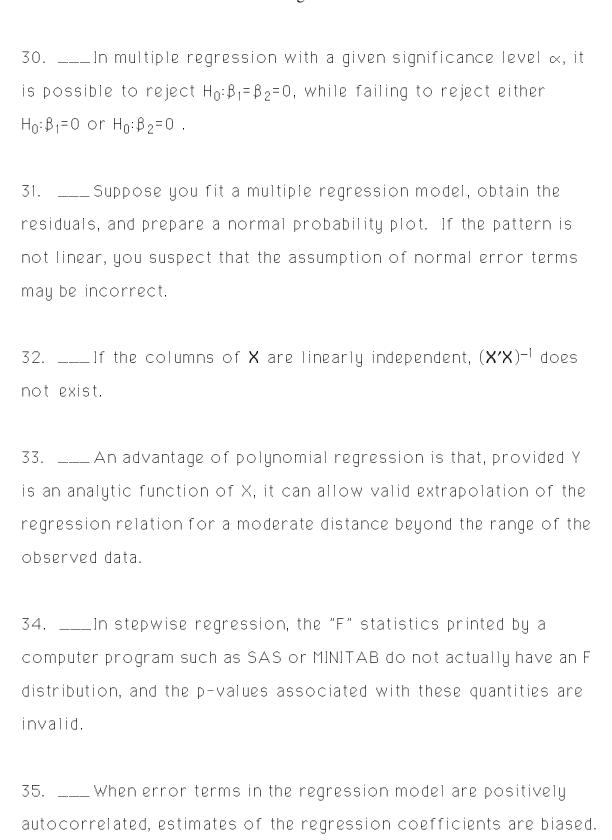
- 19. \_\_\_\_For testing the significance of the difference between two group means, suppose we use a simple linear regression model with normal error terms and  $X_i$  a dummy variable for group membership. The t-test for  $H_0$ :  $\beta$ =0 is identical to the usual independent (2-sample) t-test for  $H_0$ :  $\mu_1$ = $\mu_2$ .
- 20. \_\_\_ The concept of an interaction between two quantitative independent variables is undefined.
- 21. \_\_\_\_ Simple linear regression was used to test the difference in pain tolerance between males and females -- larger values of Y indicate more tolerance. The dummy variable X = 1 if the subject is female, and 0 if the subject is male. A positive value of  $b_1$  would indicate that in the sample, males had a higher pain tolerance than females.
- 22. \_\_\_For the model  $Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_{12} X_{i1} X_{i2} + \epsilon_i$ , the least-squares method fits a plane through the cloud of  $(X_1, X_2, Y)$  points so that the sum of squared vertical distances of the points from the plane is minimized.
- 23. \_\_\_\_ Suppose X is nationality (country of origin) and Y is freshman grade point average. It does not make sense to do simple

linear regression in this situation.

24. \_\_\_For the general regression model (7.18) on p. 237, the coefficient of multiple determination  $R^2$  is exactly equal to the squared correlation  $r^2$  between the  $Y_i$  and  $\hat{Y}_i$  variables.

# Continued on page 12

- 25. \_\_\_\_In stepwise regression, "tolerance" is the minimum proportion of explained variation that will allow an independent variable to be added to the model.
- 26. \_\_\_\_In the extra sum of squares approach to multiple regression, SSE(R)-SSE(F) = SSR(F)-SSR(R)
- 27. \_\_\_In multiple regression, the standardized regression coefficient for  $X_k$  equals the simple correlation between  $X_k$  and Y.
- 28. \_\_\_\_In well-designed experiments involving quantitative independent variables, a procedure for reducing the number of independent variables after the data are obtained is not necessary.
- 29. \_\_\_\_For the regression model 7.18 on p. 237, (Y-Xb)'(Y-Xb) is minimized if  $b = (X'X)^{-1}X'Y$ , provided that  $(X'X)^{-1}$  exists.



- 36. \_\_\_\_If two different dummy variables are used to represent sex of respondent in a survey, X'X will not be singular if the constant term  $\beta_0$  is omitted from the regression equation.
- 37. \_\_\_A market researcher would like to predict price of automobile purchased from the purchaser's age, income, sex and number of children. Multiple regression is a reasonable technique to apply.
- 38. \_\_\_\_ There can be problems with the numerical accuracy of  $(X'X)^{-1}$  when the determinant of X'X is close to zero.
- 39. \_\_\_\_In the presence of positively correlated error terms, the usual tests of statistical significance are still approximately valid for large samples.
- 40. \_\_\_\_If you use the extra sum of squares test  $H_0$ :  $\beta_k$ =0, the F\* statistic is exactly the square of t\* =  $\frac{b_k}{s\{b_k\}}$ .
- 41. \_\_\_\_If the independent variables are highly correlated among themselves, it is still possible to predict mean responses and new observations accurately, provided that R<sup>2</sup> is large and the predictions are made within the region of observations.

- 42. \_\_\_\_When severe multicolinearity exists, adding or deleting an independent variable from the model may substantially change the values of the other regression coefficients.
- 43. \_\_\_\_For the full model  $Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$  and the reduced model  $Y_i = \beta_0 + \epsilon_i$ , SSE(R) =  $\sum_{i=1}^{n} (Y_i \overline{Y})^2$
- 44. \_\_\_\_ When error terms are positively autocorrelated, MSE may seriously underestimate their variance.
- 45. \_\_\_\_It is possible that  $b_k = 0$ , and yet  $H_0$ :  $\beta_k = 0$  is rejected.
- 46. \_\_\_In a study with one independent variable that takes on 5 distinct values, the pure error lack of fit test is equivalent to fitting a 5th degree polynomial and then testing  $H_0$ :  $\beta_1 = \beta_2 = \beta_3 = \beta_4 = \beta_5 = 0.$
- 47. \_\_\_\_Let  $Y_i$  =sales,  $X_{i1}$ =advertising expenditures, and  $X_{i2}$  be a dummy variable for nationality (1 if the firm is Canadian, 0 if the firm is U.S.). The situation where Canadian and U.S. firms have the same slope but different intercepts is represented by the model:  $Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_{12} X_{i1} X_{i2} + \epsilon_i$ .

48When independent variables are strongly correlated with
each other, the residuals will not sum to exactly zero.
49In simple linear regression with autocorrelated error
terms, a good remedial measure is to take the natural log (In) of
both the X and the Y variables, adding a constant if necessary to
ensure that all the numbers are positive.
50 There can be problems with the numerical accuracy of
$(\mathbf{X'X})^{-1}$ when the magnitudes of the X variables differ greatly from
one another.
51 When the <u>dependent</u> variable is measured with error,
multiple regression can be extremely misleading.
52 When one or more <u>independent</u> variable are measured with
error, multiple regression can be extremely misleading.
53To represent a qualitative independent variable with k
categories in a multiple regression model with an intercept, only k-
1 dummy variables are needed.
54The estimated standard deviations of the regression
coefficients become small when the independent variables in the
model are highly correlated with one another.
55Suppose you weigh yourself every morning and record the
result. You would expect this time series to be negatively
autocorrelated.