

Name Jerry

Student Number \_\_\_\_\_

### STA 312f2012 Quiz 11

1. (2 points) For the Florida Death Penalty data, you were asked whether the model  $(PR, VR)(DP)$  fits adequately.

(a) Give the value of the Likelihood Ratio (not Pearson) test statistic. The answer is a number from your printout.

$$G^2 = 8.13$$

(b) What is the critical value at  $\alpha = 0.05$ ? The answer is a number.

$$7.815$$

(c) By this criterion, does the model fit adequately? Answer Yes or No.

No

2. (3 points) Planning another study of race and the death penalty, assume there are approximately equal numbers of Black and White prisoners, which is the case with the Florida data. In a  $2 \times 2$  table of race by death penalty, suppose the probability of a White prisoner getting the death penalty is 10%, while the probability of a Black prisoner getting the death penalty is 15%. What sample size is required for the power of the Pearson chi-square test to reach 0.80? The answer is a number that should be on your printout.

$$1374$$

3. (3 points) For the study of gender (G), being on the honour roll (H) and traffic accidents (A), you were asked for the smallest sample size that could reject the conditional independence model  $(AG)(GH)$  with probability 0.90 at  $\alpha = 0.05$ , given a set of cell probabilities. It does not matter whether you did it for a Pearson test or a likelihood ratio test; either will do. What is the required sample size? The answer is an integer.

$$n = 3613 \text{ for Likelihood ratio,} \\ 3982 \text{ for Pearson}$$

Please attach your printout or printouts. Make sure your name is written on each one.

# STA 312f2012 Formulas

$$Z_1 = \frac{\sqrt{n}(p-\pi_0)}{\sqrt{\pi_0(1-\pi_0)}} \quad Z_2 = \frac{\sqrt{n}(p-\pi_0)}{\sqrt{p(1-p)}} \quad p \pm z_{\alpha/2} \sqrt{\frac{p(1-p)}{n}}$$

> qnorm(0.975)  
[1] 1.959964  
> qnorm(0.995)  
[1] 2.575829

$$Z_2^2 \sim \chi^2(1, \lambda), \text{ with } \lambda = n \frac{(\pi - \pi_0)^2}{\pi(1-\pi)} \quad \theta = \frac{\pi_{11}\pi_{22}}{\pi_{12}\pi_{21}}$$

$$P(n_1, \dots, n_c) = \binom{n}{n_1 \dots n_c} \pi_1^{n_1} \dots \pi_c^{n_c} \quad \ell(\boldsymbol{\pi}) = \prod_{i=1}^n \pi_1^{y_{i,1}} \pi_2^{y_{i,2}} \dots \pi_c^{y_{i,c}} = \pi_1^{n_1} \pi_2^{n_2} \dots \pi_c^{n_c}$$

$$G^2 = -2 \log \left( \frac{\max_{\beta \in \mathcal{B}_0} \ell(\beta)}{\max_{\beta \in \mathcal{B}} \ell(\beta)} \right) = -2 \log \left( \frac{\ell_0}{\ell_1} \right) \quad \hat{\mu}_j = n \hat{\pi}_j \quad , \quad \hat{\mu}_{ij} = \frac{n_{i+j}}{n}$$

$$X^2 = \sum_{j=1}^c \frac{(n_j - \hat{\mu}_j)^2}{\hat{\mu}_j} = n \sum_{j=1}^c \frac{(p_j - \hat{\pi}_j)^2}{\hat{\pi}_j} \quad G^2 = 2 \sum_{j=1}^c n_j \log \left( \frac{n_j}{\hat{\mu}_j} \right) = 2n \sum_{j=1}^c p_j \log \left( \frac{p_j}{\hat{\pi}_j} \right)$$

$$\lambda = n \sum_{j=1}^c \frac{[\pi_j - \pi_j(M)]^2}{\pi_j(M)} \quad \lambda = 2n \sum_{j=1}^c \pi_j \log \left( \frac{\pi_j}{\pi_j(M)} \right)$$

$$X^2 = n \sum_{i=1}^I \sum_{j=1}^J \frac{(p_{ij} - p_{i+p+j})^2}{p_{i+p+j}} \quad G^2 = 2n \sum_{i=1}^I \sum_{j=1}^J p_{ij} \log \left( \frac{p_{ij}}{p_{i+p+j}} \right)$$

$$\lambda = n \sum_{i=1}^I \sum_{j=1}^J \frac{(\pi_{ij} - \pi_{i+\pi+j})^2}{\pi_{i+\pi+j}} \quad \lambda = 2n \sum_{i=1}^I \sum_{j=1}^J \pi_{ij} \log \left( \frac{\pi_{ij}}{\pi_{i+\pi+j}} \right)$$

$$\log \mu_{ijk} = \lambda + \lambda_i^X + \lambda_j^Y + \lambda_k^Z + \lambda_{ij}^{XY} + \lambda_{ik}^{XZ} + \lambda_{jk}^{YZ} + \lambda_{ijk}^{XYZ}$$

```
> df = 1:8
> CriticalValue = qchisq(0.95,df)
> round(rbind(df,CriticalValue),3)
      [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8]
df      1.000 2.000 3.000 4.000 5.00 6.000 7.000 8.000
CriticalValue 3.841 5.991 7.815 9.488 11.07 12.592 14.067 15.507
```

> # Table of non-centrality values needed for specified power

df	0.5	0.7	0.8	0.9
1	3.841022	6.172007	7.848883	10.50739
2	4.956736	7.701774	9.634685	12.65394
3	5.760482	8.792389	10.902570	14.17149
4	6.419476	9.682473	11.935286	15.40503
5	6.991270	10.452523	12.827607	16.46946
6	7.503313	11.140677	13.624286	17.41883
7	7.971192	11.768443	14.350527	18.28355
8	8.404641	12.349332	15.022138	19.08270

## Quiz 11 Printout

```
> # DeathPen (Q1)
>
> deathpen =
read.table("http://fisher.utstat.utoronto.ca/~brunner/312f12/code_n_data/de
athpen2.data",header=T)
> deathpen[1:4,]
  Prace Vrace Death
1 White White   Yes
2 White White   Yes
3 White White   Yes
4 White White   Yes
> attach(deathpen)
> deathrace = table(Prace, Vrace, Death)
> deathrace
, , Death = No

      Vrace
Prace   Black White
  Black    97    52
  White     9   132

, , Death = Yes

      Vrace
Prace   Black White
  Black     6    11
  White     0    19

> Q1 = loglin(deathrace, margin=list(c(1,2),3)); Q1
2 iterations: deviation 0
$lrt
[1] 8.131611

$pearson
[1] 6.977343

$df
[1] 3

$margin
$margin[[1]]
[1] "Prace" "Vrace"
```

```
$margin[[2]]  
[1] "Death"
```

```
>  
> # Q2 Power for another DP study.  
>  
> dummy2 = rbind(c(10,90),  
+               c(15,85))  
> eff2 = loglin(dummy2,margin=list(1,2))$pearson/200  
2 iterations: deviation 0  
> 7.848883/eff2  
[1] 1373.555
```

```
>  
>  
> # Q3 AHG test (AG)(GH). Want power of 0.90  
>  
> Male = rbind(c(0.005777638, 0.144224),  
+             c(0.024222362, 0.325776) )  
>  
> Female = rbind(c(0.004222372, 0.04577688),  
+              c(0.065777628, 0.38422312) )  
>  
> truth = numeric(8); dim(truth) = c(2,2,2)  
> truth[,1]=Male; truth[,2]=Female  
> truth  
, , 1
```

```
      [,1]      [,2]  
[1,] 0.005777638 0.144224  
[2,] 0.024222362 0.325776
```

```
, , 2
```

```
      [,1]      [,2]  
[1,] 0.004222372 0.04577688  
[2,] 0.065777628 0.38422312
```

```
> truth=truth*1000  
>  
> dummy = loglin(truth,margin=list(c(1,3),c(2,3)))  
2 iterations: deviation 5.684342e-14
```

```

> dummy
$lrt
[1] 3.501477

$pearson
[1] 3.177402

$df
[1] 2

$margin
$margin[[1]]
[1] 1 3

$margin[[2]]
[1] 2 3

> efflrt = dummy$lrt/1000; effpearson = dummy$pearson/1000
> nlrt = 12.65394/efflrt; nlrt
[1] 3613.886
> npearson = 12.65394/effpearson; npearson
[1] 3982.48

```